

CSE 312

# Foundations of Computing II

Lecture 21: Cont. Joint Distributions, Law of Total Expectation



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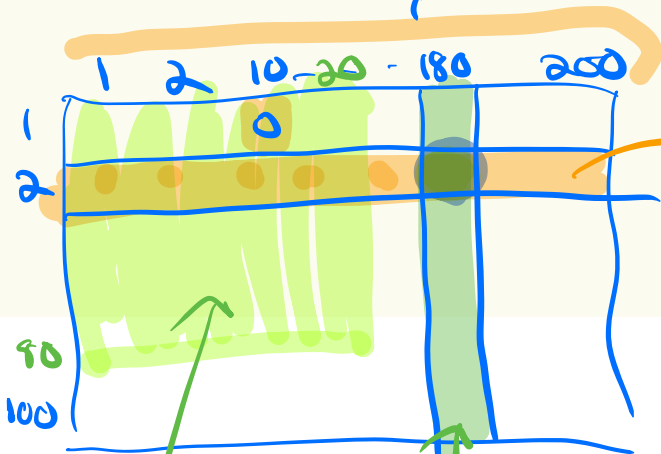
Slide Credit: Based on Stefano Tessaro's slides for 312 19au incorporating ideas from Alex Tsun, Rachel Lin, Hunter Schafer & myself 😊

## Agenda

- Continuous joint distributions 
- Conditional Expectation and Law of Total Expectation

	Discrete	Continuous $\rightarrow dx dy$
Joint PMF/PDF	$p_{X,Y}(x,y) = \mathbb{P}(X=x, Y=y)$	$f_{X,Y}(x,y) \neq \mathbb{P}(X=x, Y=y)$
Joint range/support $\Omega_{X,Y}$	$\{(x,y) \in \Omega_X \times \Omega_Y : p_{X,Y}(x,y) > 0\}$	$\{(x,y) \in \Omega_X \times \Omega_Y : f_{X,Y}(x,y) > 0\}$
Joint CDF	$F_{X,Y}(x,y) = \sum_{t \leq x, s \leq y} p_{X,Y}(t,s)$	$F_{X,Y}(x,y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(t,s) ds dt$
Normalization	$\sum_{x,y} p_{X,Y}(x,y) = 1$	$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = 1$ $\leftarrow$
Marginal PMF/PDF	$p_X(x) = \sum_y p_{X,Y}(x,y)$	$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$
Expectation	$\mathbb{E}[g(X,Y)] = \sum_{x,y} g(x,y) p_{X,Y}(x,y)$	$\mathbb{E}[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dx dy$

$$\sum_{x \in \Omega_X} \sum_{y \in \Omega_Y} p_{X,Y}(x,y) = 1$$



$\rightarrow p(X=2)$  summing row

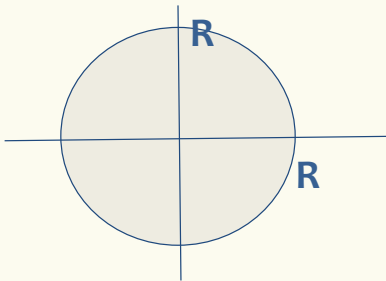
$\Omega_{X,Y} \neq \Omega_X \times \Omega_Y$   
 $X, Y$  not indep.

$$F_{XY}(80, 20)$$

$$\text{sim. } P(Y=180)$$

- Suppose that the surface of a disk is a circle with ~~area~~ <sup>radius</sup>  $R$  centered at the origin and that there is a single point imperfection at a location which is uniformly distributed across the surface of the disk. Let  $X$  and  $Y$  be the  $x$  and  $y$  coordinates of the imperfection (random variables) and let  $Z$  be the distance of the imperfection from the origin.
  - What is their joint density  $f(x,y)$ ?

$$f_{X,Y}(x,y) = \begin{cases} c & x^2 + y^2 \leq R^2 \\ 0 & \text{o.w.} \end{cases}$$



$$\iint_{x^2 + y^2 \leq R^2} c \, dx \, dy = 1$$

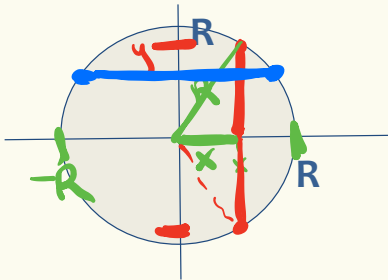
$$c \pi R^2 = 1$$

$$\Rightarrow c = \frac{1}{\pi R^2}$$

- Suppose that the surface of a disk is a circle with area  $R$  centered at the origin and that there is a single point imperfection at a location which is uniformly distributed across the surface of the disk. Let  $X$  and  $Y$  be the  $x$  and  $y$  coordinates of the imperfection (random variables) and let  $Z$  be the distance of the imperfection from the origin.

– What is the range of  $X$  &  $Y$  and the marginal density of  $X$  and of  $Y$ ?

$$\mathcal{L}_X = [-R, R], \quad \mathcal{L}_Y = [-R, R]$$



$$f_X(x) = \int_{-\sqrt{R^2-x^2}}^{\sqrt{R^2-x^2}} \frac{1}{\pi R^2} dy = \frac{2\sqrt{R^2-x^2}}{\pi R^2}$$

$$f_X(x) = \begin{cases} \frac{2\sqrt{R^2-x^2}}{\pi R^2} & x^2 \leq R^2 \\ 0 & \text{o.w.} \end{cases}$$

$$f_Y(y) = \begin{cases} \frac{2\sqrt{R^2 - y^2}}{\pi R^2} & |y| \leq R \\ 0 & \text{o.w.} \end{cases}$$

- Suppose that the surface of a disk is a circle with area  $R$  centered at the origin and that there is a single point imperfection at a location with is uniformly distributed across the surface of the disk. Let  $X$  and  $Y$  be the  $x$  and  $y$  coordinates of the imperfection (random variables) and let  $Z$  be the distance of the imperfection from the origin.

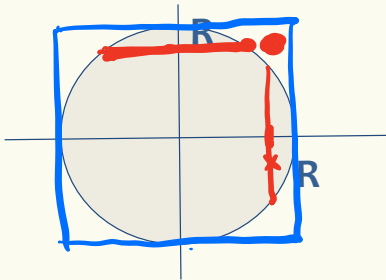
– Are  $X$  and  $Y$  independent?

$$\Delta_x \times \Delta_y$$

$$\Delta_{x,y} = \Delta_x \times \Delta_y$$

$$f_X(x) > 0$$

$$\underline{f_{X,Y}(x,y) = 0} \quad \text{by } \underline{f_X(x), f_Y(y) > 0}$$

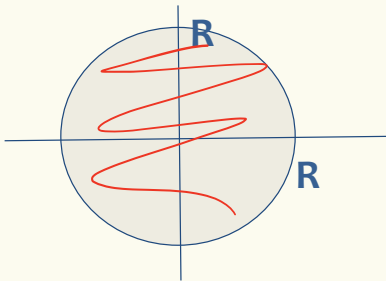


- Suppose that the surface of a disk is a circle with area  $R$  centered at the origin and that there is a single point imperfection at a location which is uniformly distributed across the surface of the disk. Let  $X$  and  $Y$  be the  $x$  and  $y$  coordinates of the imperfection (random variables) and let  $Z$  be the distance of the imperfection from the origin.

– What is  $E(Z)$ ?

$$Z = \sqrt{x^2 + y^2}$$

$$Z = g(x, y)$$



$$\iint_{x^2 + y^2 \leq R^2} \frac{\sqrt{x^2 + y^2}}{\pi R^2} dx dy$$

$$\iint_{x, y} g(x, y) f(x, y) dx dy$$



## All of this generalizes to more than 2 random variables

	Discrete	Continuous
<b>Joint PMF/PDF</b>	$p_{X,Y}(x, y) = \mathbb{P}(X = x, Y = y)$	$f_{X,Y}(x, y) \neq \mathbb{P}(X = x, Y = y)$
<b>Joint range/support</b> $\Omega_{X,Y}$	$\{(x, y) \in \Omega_X \times \Omega_Y : p_{X,Y}(x, y) > 0\}$	$\{(x, y) \in \Omega_X \times \Omega_Y : f_{X,Y}(x, y) > 0\}$
<b>Joint CDF</b>	$F_{X,Y}(x, y) = \sum_{t < x, s < y} p_{X,Y}(t, s)$	$F_{X,Y}(x, y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(t, s) ds dt$
<b>Normalization</b>	$\sum_{x,y} p_{X,Y}(x, y) = 1$	$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx dy = 1$
<b>Marginal PMF/PDF</b>	$p_X(x) = \sum_y p_{X,Y}(x, y)$	$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$
<b>Expectation</b>	$\mathbb{E}[g(X, Y)] = \sum_{x,y} g(x, y) p_{X,Y}(x, y)$	$\mathbb{E}[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X,Y}(x, y) dx dy$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y,Z}(x, y, z) dx dy dz = 1$$

$$f_{X,Y}(x, y) = \int_{-\infty}^{\infty} f(x, y, z) dz$$

$$f_X(x) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y, z) dy dz$$

## Agenda

- Continuous joint distributions
- Conditional Expectation and Law of Total Expectation 

## Conditional Expectation

**Definition.** Let  $X$  be a discrete random variable then the **conditional expectation** of  $X$  given event  $A$  is

$$E[X | A] = \sum_{x \in \Omega(X)} x \Pr(X = x | A)$$

- Linearity of expectation still applies here

$$E[aX + bY + c | A] = aE[X | A] + bE[Y | A] + c$$

## Conditional Expectation

$$A \circ Y=y$$

**Definition.** Let  $X$  be a discrete random variable then the **conditional expectation** of  $X$  given event  $Y = y$  is

$$E[X | Y = y] = \sum_{x \in \Omega(X)} x \Pr(X = x | Y = y)$$

conditional prob mass fn,  
 $f_X | Y=y$

$$\sum_x \underbrace{P(X=x | Y=y)}_{\text{check}} = 1$$

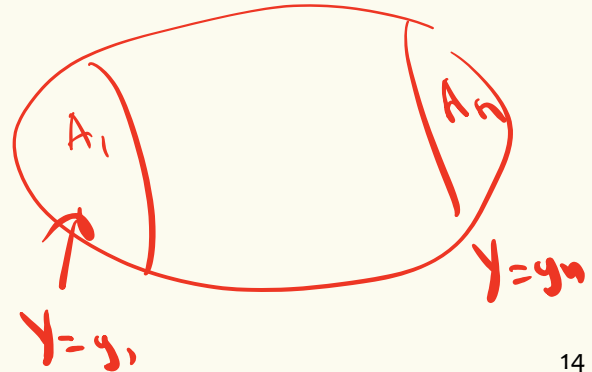
- Linearity of expectation still applies here

$$E[aX + bY + c | Y = y] = aE[X | Y = y] + bE[Y | Y = y] + c$$

## Law of Total Expectation

**Law of Total Expectation (event version).** Let  $X$  be a random variable and let events  $A_1, \dots, A_n$  partition the sample space. Then,

$$E[X] = \sum_{i=1}^n E[X|A_i] \Pr(A_i)$$



## Proof of Law of Total Expectation

Follows from Law of Total Probability and manipulating sums

$$\begin{aligned} E[X] &= \sum_{x \in \Omega(X)} x \Pr(X = x) \\ &= \sum_{x \in \Omega(X)} x \sum_{i=1}^n \Pr(X = x | A_i) \Pr(A_i) && \text{(by LTP)} \\ &= \sum_{i=1}^n \Pr(A_i) \left[ \sum_{x \in \Omega(X)} x \Pr(X = x | A_i) \right] && \text{(change order of sums)} \\ &= \sum_{i=1}^n \Pr(A_i) E[X | A_i] && \text{(def of cond. expect.)} \end{aligned}$$

## Law of Total Expectation

Law of iterated exps.

**Law of Total Expectation (random variable version).** Let  $X$  be a random variable and  $Y$  be a discrete random variable. Then,

$$E[X] = \sum_{y \in \Omega(Y)} E[X|Y=y] \Pr(Y=y)$$

$$E[E(X|Y)]$$

random variable takes

$y \in \Omega_Y$

$E(X|Y=y)$  w.p.  $\Pr(Y=y)$

$$\Omega_Y = \{1, 2, 3\}$$

$$E(X|Y) = \begin{cases} E(X|Y=1) \\ E(X|Y=2) \\ E(X|Y=3) \end{cases}$$

$$\begin{cases} \Pr(Y=1) \\ \Pr(Y=2) \\ \Pr(Y=3) \end{cases}$$

## Example: Flipping Coins

Suppose wanted to analyze flipping a random number of coins. Suppose someone gave us  $Y \sim Poi(5)$  fair coins and we wanted to compute the expected number of heads  $X$  from flipping those coins.



## Example: Computer Failures

Suppose your computer operates in a sequence of steps, and that at each step  $i$  your computer will fail with probability  $p$  (independently of other steps). Let  $X$  be the number of steps it takes your computer to fail. What is  $E[X]$ ?

$$Y = \begin{cases} 1 & \text{if first trial failure} \\ 0 & \text{o.w.} \end{cases}$$

$$E(X) = E(X | Y=1) P(Y=1) + E(X | Y=0) P(Y=0)$$

$\quad \quad \quad \uparrow \quad \quad \quad \cdot p$

$\uparrow + \# \text{ additional trials/steps to failure,}$

$$E(1 + X | Y=0) (1-p)$$

$\quad \quad \quad \uparrow \quad \quad \quad \cdot (1-p)$   
 $\quad \quad \quad 1 + E(X)$

$$E(X) = p + (1-p)(1+E(X))$$

$$= p + \underbrace{1-p}_{1} + \underbrace{(1-p)E(X)}$$



$$pE(X) = 1$$

$$E(X) = \frac{1}{p}$$

## Elevator rides

The number of people who enter an elevator on the ground floor is a Poisson random variable with mean 10. If there are  $N$  floors above the ground floor, and if each person is equally likely to get off at any one of the  $N$  floors, independently of where others get off, compute the expected number of stops the elevator will make before discharging all the passengers.

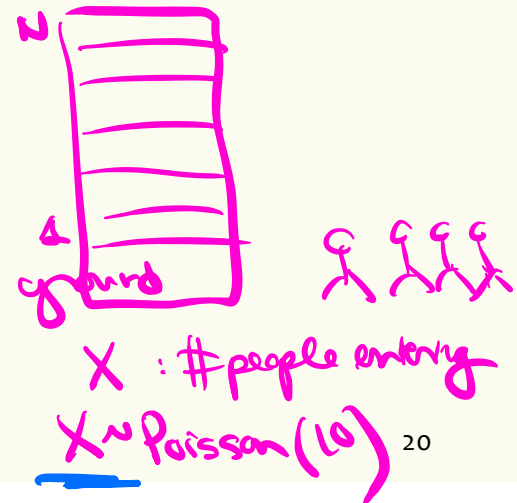
$Y$ : # stops elevator makes.

$$E(Y) = \sum_{k=0}^{\infty} E(Y | X=k) P(X=k)$$

$e^{-10} \frac{10^k}{k!}$

$$Y = Y_1 + Y_2 + \dots + Y_n$$

$$Y_i = \begin{cases} 1 & \text{stops on } i^{\text{th}} \text{ floor} \\ 0 & \text{o.w.} \end{cases}$$



$$E(Y|X=k) = E(Y_1 + Y_2 + \dots + Y_N | X=k)$$

LOE  
for card effects

$$\stackrel{\uparrow}{=} \sum_{i=1}^N E(Y_i | X=k)$$

$P(\text{step on } i^{\text{th}} \text{ floor} | X=k)$

$Z_i$ : # of people who step on  $i^{\text{th}}$  floor |  $X=k$

$\text{Bin}(k, \frac{1}{N})$

$$\Pr(Z_i > 0) = 1 - \Pr(Z_i = 0)$$
$$\left( \frac{N-1}{N} \right)^k$$

## Reference Sheet (with continuous RVs)

	Discrete	Continuous
<b>Joint PMF/PDF</b>	$p_{X,Y}(x, y) = P(X = x, Y = y)$	$f_{X,Y}(x, y) \neq P(X = x, Y = y)$
<b>Joint CDF</b>	$F_{X,Y}(x, y) = \sum_{t \leq x} \sum_{s \leq y} p_{X,Y}(t, s)$	$F_{X,Y}(x, y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(t, s) ds dt$
<b>Normalization</b>	$\sum_x \sum_y p_{X,Y}(x, y) = 1$	$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx dy = 1$
<b>Marginal PMF/PDF</b>	$p_X(x) = \sum_y p_{X,Y}(x, y)$	$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$
<b>Expectation</b>	$E[g(X, Y)] = \sum_x \sum_y g(x, y) p_{X,Y}(x, y)$	$E[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X,Y}(x, y) dx dy$
<b>Conditional PMF/PDF</b>	$p_{X Y}(x   y) = \frac{p_{X,Y}(x, y)}{p_Y(y)}$	$f_{X Y}(x   y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}$
<b>Conditional Expectation</b>	$E[X   Y = y] = \sum_x x p_{X Y}(x   y)$	$E[X   Y = y] = \int_{-\infty}^{\infty} x f_{X Y}(x   y) dx$
<b>Independence</b>	$\forall x, y, p_{X,Y}(x, y) = p_X(x) p_Y(y)$	$\forall x, y, f_{X,Y}(x, y) = f_X(x) f_Y(y)$