CSE 312 Foundations of Computing II

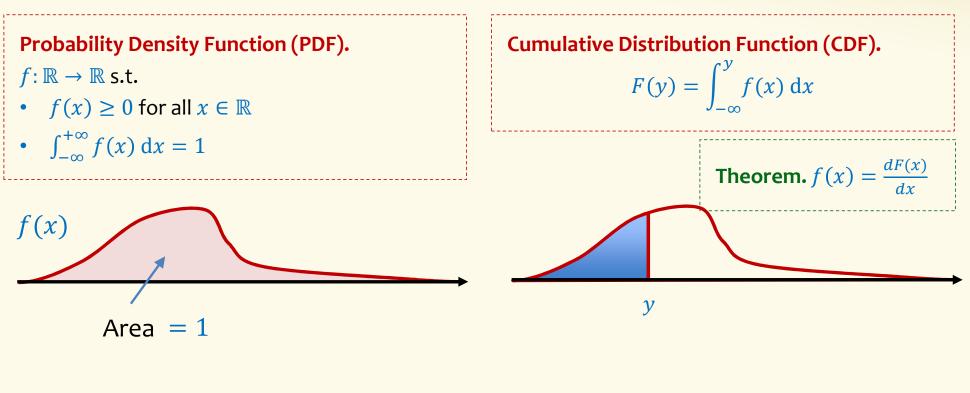
Lecture 16: Normal Distribution & Central Limit Theorem

1

Announcements

- Midterm on <u>Wednesday</u>
 - Read instructions on edstem carefully
 - Look at the sample midterm
- Review session is tomorrow 4pm (zoom link will be posted)
- Feedback form: https://forms.gle/NLvU4Pt6HiHZd1Zz7

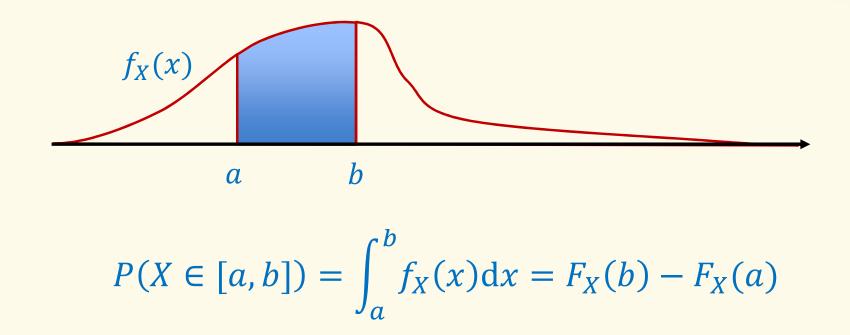
Review Continuous RVs



Density \neq Probability !

 $F_X(y) = P(X \le y)$

Review Continuous RVs

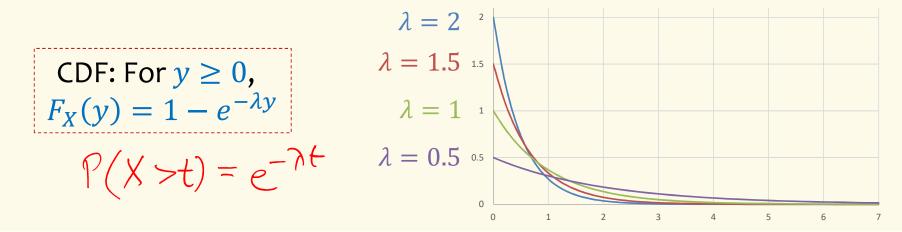


Review Exponential Distribution

Definition. An **exponential random variable** *X* with parameter $\lambda \ge 0$ follows the exponential density

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x \ge 0\\ 0 & x < 0 \end{cases}$$

We write $X \sim \text{Exp}(\lambda)$ and say X that follows the exponential distribution.



Agenda

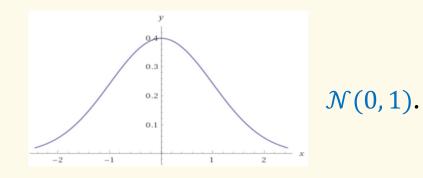
- Normal Distribution
- Practice with Normals
- Central Limit Theorem (CLT)

The Normal Distribution

Definition. A Gaussian (or normal) random variable with parameters $\mu \in \mathbb{R}$ and $\sigma \ge 0$ has density

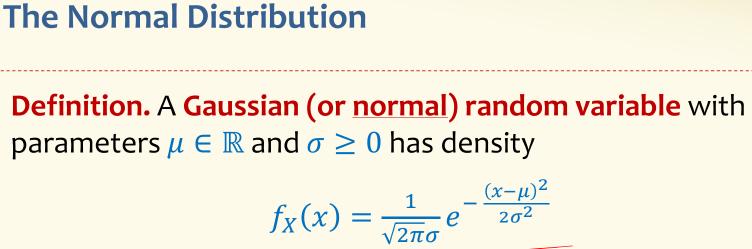
$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

We say that X follows the Normal Distribution, and write $X \sim \mathcal{N}(\mu, \sigma^2)$.





Carl Friedrich Gauss





Carl Friedrich Gauss

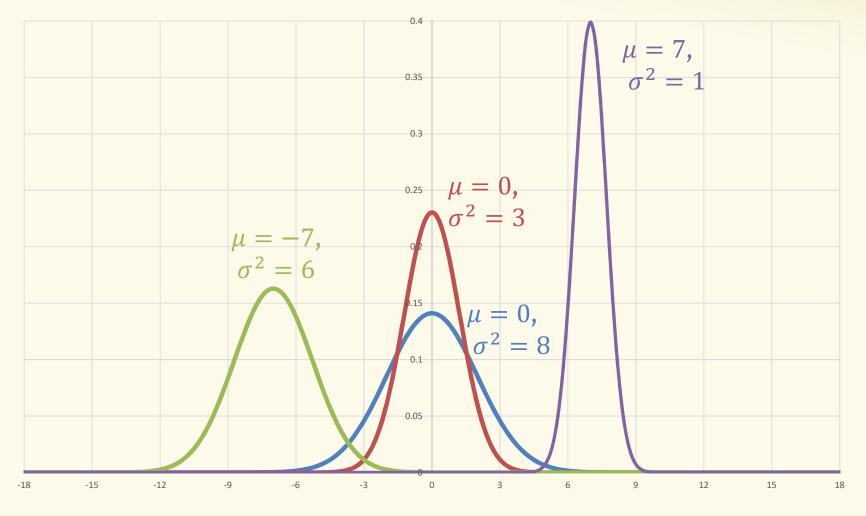
We say that X follows the Normal Distribution, and write $X \sim \mathcal{N}(\mu, \sigma^2)$.

Fact. If $X \sim \mathcal{N}(\mu, \sigma^2)$, then $\mathbb{E}[X] = \mu$, and $Var(X) = \sigma^2$

Proof of expectation is easy because density curve is symmetric around μ , $f_X(\mu - x) = f_X(\mu + x)$, but proof for variance requires integration of $e^{-x^2/2}$

The Normal Distribution

Aka a "Bell Curve" (imprecise name)



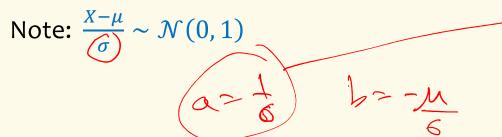
Closure of normal distribution – Under Shifting and Scaling

Fact. If $X \sim \mathcal{N}(\mu, \sigma^2)$, then $Y = aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$

Proof.
$$\mathbb{E}[Y] = a \mathbb{E}[X] + b \neq a\mu + b$$

 $Var(Y) = a^2 Var(X) = a^2 \sigma^2$

Can show with algebra that the PDF of Y = aX + b is still normal.



 $-\underline{M} \sim 0$

CDF of normal distribution

Fact. If $X \sim \mathcal{N}(\mu, \sigma^2)$, then $Y = aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$

Standard (unit) normal = $\mathcal{N}(0, 1)$ CDF. $\Phi(z) = P(Z \le z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{z} e^{-x^{2}/2} dx$ for $Z \sim \mathcal{N}(0, 1)$

Note: $\Phi(z)$ has no closed form – generally given via tables

If
$$X \sim \mathcal{N}(\mu, \sigma^2)$$
, then $F_X(z) = P(X \le z) = P\left(\frac{X-\mu}{\sigma} \le \frac{z-\mu}{\sigma}\right) = \Phi\left(\frac{z-\mu}{\sigma}\right)$

Table of Standard Cumulative Normal Density



	Φ Table: $\mathbb{P}(Z \leq z)$ when $Z \sim \mathcal{N}(0, 1)$										
	z (0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
$P(Z \le 1.09) = \Phi(1.09) \approx 0.8621$	0.0 (0.5	0.50399	0.50798	0.51197	0.51595	0.51994	0.52392	0.5279	0.53188	0.53586
		0.53983	0.5438	0.54776	0.55172	0.55567	0.55962	0.56356	0.56749	0.57142	0.57535
		0.57926	0.58317	0.58706	0.59095	0.59483	0.59871	0.60257	0.60642	0.61026	0.61409
()(+)(09)		0.61791	0.62172	0.62552	0.6293	0.63307	0.63683	0.64058	0.64431	0.64803	0.65173
What is $P(Z \le -1.09)? \bigcirc -1?(Z \le 1.09)$		0.65542	0.6591	0.66276	0.6664	0.67003	0.67364	0.67724	0.68082	0.68439	0.68793
What is		0.69146	0.69497	0.69847	0.70194	0.7054	0.70884	0.71226	0.71566	0.71904	0.7224
$\mathcal{D}(\mathbf{z} \rightarrow \mathbf{z}) = \mathbf{z} - \mathbf$		0.72575	0.72907	0.73237	0.73565	0.73891	0.74215	0.74537	0.74857	0.75175	0.7549
P(Z < -1.09)? (2) $P(Z < -1.09)$		$0.75804 \\ 0.78814$	0.76115 0.79103	0.76424 0.79389	0.7673	0.77035	0.77337	0.77637	0.77935 0.80785	0.7823	0.78524
		0.78814 0.81594	0.79103 0.81859	0.79389 0.82121	0.79673 0.82381	0.79955 0.82639	0.80234 0.82894	0.80511 0.83147	0.80785	0.81057 0.83646	0.81327 0.83891
		$0.81394 \\ 0.84134$	0.81859 0.84375	0.82121 0.84614	0.82381 0.84849	0.82039	0.82894 0.85314	0.85147 0.85543	0.85398 0.85769	0.85993	
Poll:		0.86433	0.8665	0.86864	0.87076	0.87286	0.87493	0.87698	0.879	0.881	0.88298
FUII.		0.88493	0.88686	0.88877	0.89065	0.89251	0.89435	0.89617	0.89796	0.89973	0.90147
n allas com/naulhanna an 9	10000	0.9032	0.9049	0.90658	0.90824	0.90988	0.91149	0.91309	0.91466	0.91621	0.91774
pollev.com/paulbeame028	1.4 (0.91924	0.92073	0.9222	0.92364	0.92507	0.92647	0.92785	0.92922	0.93056	0.93189
	1.5 (0.93319	0.93448	0.93574	0.93699	0.93822	0.93943	0.94062	0.94179	0.94295	0.94408
a. 0.1379	1.6 (0.9452	0.9463	0.94738	0.94845	0.9495	0.95053	0.95154	0.95254	0.95352	0.95449
	1.7 (0.95543	0.95637	0.95728	0.95818	0.95907	0.95994	0.9608	0.96164	0.96246	0.96327
b. 0.8621		0.96407	0.96485	0.96562	0.96638	0.96712	0.96784	0.96856	0.96926	0.96995	0.97062
51 010021		0.97128	0.97193	0.97257	0.9732	0.97381	0.97441	0.975	0.97558	0.97615	0.9767
C. 0		0.97725	0.97778	0.97831	0.97882	0.97932	0.97982	0.9803	0.98077	0.98124	0.98169
		0.98214	0.98257	0.983	0.98341	0.98382	0.98422	0.98461	0.985	0.98537	0.98574
d. Not able to compute		0.9861	0.98645	0.98679	0.98713	0.98745	0.98778	0.98809	0.9884	0.9887	0.98899
d. Not able to compute		$\begin{array}{c c} 0.98928 \\ \hline 0.9918 \end{array}$	0.98956 0.99202	$\frac{0.98983}{0.99224}$	0.9901 0.99245	0.99036 0.99266	$\frac{0.99061}{0.99286}$	0.99086 0.99305	$\begin{array}{c} 0.99111 \\ 0.99324 \end{array}$	$\begin{array}{r} 0.99134 \\ 0.99343 \end{array}$	$\begin{array}{r} 0.99158 \\ 0.99361 \end{array}$
·		0.9918	0.99202	0.99224 0.99413	0.99245	0.99266 0.99446	0.99280 0.99461	0.99305 0.99477	0.99324 0.99492	0.99545	0.99501 0.9952
		0.99579	0.99590 0.99547	0.99413 0.9956	0.9943 0.99573	0.99440 0.99585	0.99401 0.99598	0.99477	0.99492 0.99621	0.99500	0.9952 0.99643
same e		0.99653	0.99664	0.99574	0.99683	0.99693	0.99598	0.99009	0.99021	0.99032	0.99736
1.10 7		0.99744	0.99752	0.9976	0.99767	0.99774	0.99781	0.99788	0.99795	0.99801	0.99807
		0.99813	0.99819	0.99825	0.99831	0.99836	0.99841	0.99846	0.99851	0.99856	0.99861
		0.99865	0.99869	0.99874	0.99878	0.99882	0.99886	0.99889	0.99893	0.99896	0.999
		I						1	1	1]

12

 $\Phi(z)$

7

Closure of the normal -- under addition

Fact. If $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$, $Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$ (both independent normal RV) then $aX + bY + c \sim \mathcal{N}(a\mu_X + b\mu_Y + c, a^2\sigma_X^2 + b^2\sigma_Y^2)$

Note: The special thing is that **the sum of normal RVs is still a normal RV**. The values of the expectation and variance are **not** surprising.

Why not?

- Linearity of expectation (always true)
- When X and Y are independent, $Var(aX + bY) = a^2Var(X) + b^2Var(Y)$

Agenda

- Normal Distribution
- Practice with Normals
- Central Limit Theorem (CLT)

What about Non-standard normal?

If
$$X \sim \mathcal{N}(\mu, \sigma^2)$$
, then $\frac{x - \mu}{\sigma} \sim \mathcal{N}(0, 1)$
Therefore,
 $F_X(z) = P(X \le z) = P\left(\frac{X - \mu}{\sigma} \le \frac{z - \mu}{\sigma}\right) = \Phi\left(\frac{z - \mu}{\sigma}\right)$

Example	
	V-11 1.2-11 1.2-0.4
Let $X \sim \mathcal{N}(0.4, 4 = 2^2)$.	$X - M = \frac{1 \cdot 2 - M}{6} = \frac{1 \cdot 2 - 0 \cdot 4}{2}$
6-2	Y10-780.
$P(X \le 1.2) = P\left(\frac{X - 0.4}{2} \le \frac{1.2}{2}\right)$	$\frac{2-0.4}{2}$
	2
N(0) $(X - 0.4)$	
$P\left(\frac{n-6\pi}{2}\right) \le 0.4$	$\left(\Phi \right) = \Phi(0.4) \approx 0.6554$
	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$
$\sim \mathcal{N}(0,1)$	0.3 0.6179 0.6217
	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$
	0.6 0.7257 0.7291
	0.7 0.7580 0.7611

Example
Let
$$X \sim \mathcal{N}(3, 16)$$
.

$$P(2 < X < 5) = P\left(\frac{2-3}{4} < \frac{X-3}{\sqrt{4}} < \frac{5-3}{4}\right)$$

$$= P\left(-\frac{1}{4} < Z < \frac{1}{2}\right) = \Phi\left(\frac{1}{2}\right) - \Phi\left(-\frac{1}{4}\right)$$

$$= \Phi\left(\frac{1}{2}\right) - \Phi\left(-\frac{1}{4}\right) = \Phi\left(\frac{1}{2}\right) - \left(1 - \Phi\left(\frac{1}{4}\right)\right)$$

$$= \Phi\left(\frac{1}{2}\right) - \left(1 - \Phi\left(\frac{1}{4}\right)\right) \approx 0.29017$$

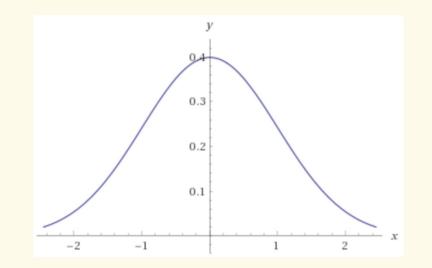
Example – How Many Standard Deviations Away?

Let $X \sim \mathcal{N}(\mu, \sigma^2)$.

$$P(|X - \mu| < k\sigma) = P\left(\frac{|X - \mu|}{\sigma} < k\right) =$$
$$= P\left(-k < \frac{X - \mu}{\sigma} < k\right) = \Phi(k) - \Phi(-k)$$

e.g. k = 1: 68% k = 2: 95% k = 3: 99%

Brain Break





Normal Distribution

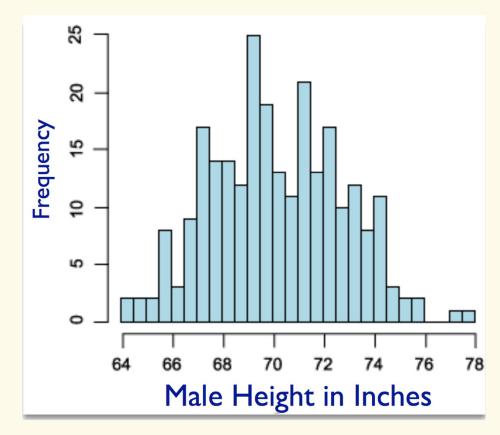
Paranormal Distribution

Agenda

- Normal Distribution
- Practice with Normals
- Central Limit Theorem (CLT)

Gaussian in Nature

Empirical distribution of collected data often resembles a Gaussian ...



e.g. Height distribution resembles Gaussian.

R.A.Fisher (1918) observed that the height is likely the outcome of the sum of many independent random parameters, i.e., it can written as $X = X_1 + \dots + X_n$

Sum of Independent RVs

i.i.d. = independent and identically distributed

 X_1, \ldots, X_n i.i.d. with expectation μ and variance σ^2

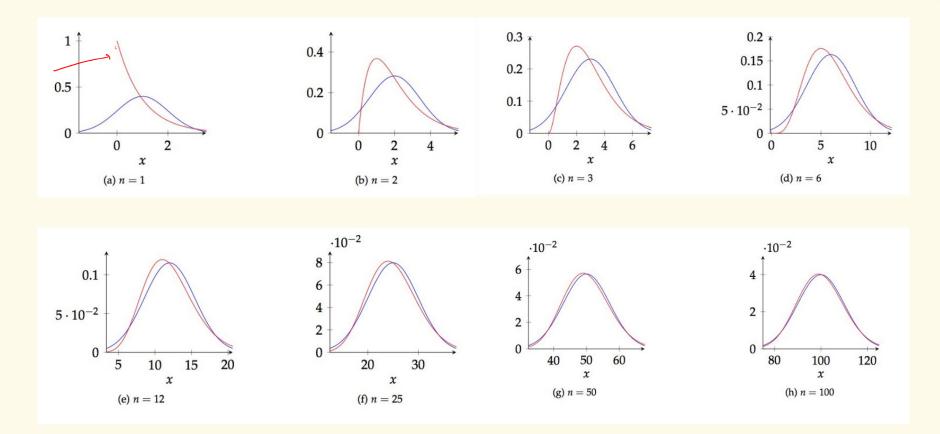
Define

$$S_n = X_1 + \dots + X_n$$

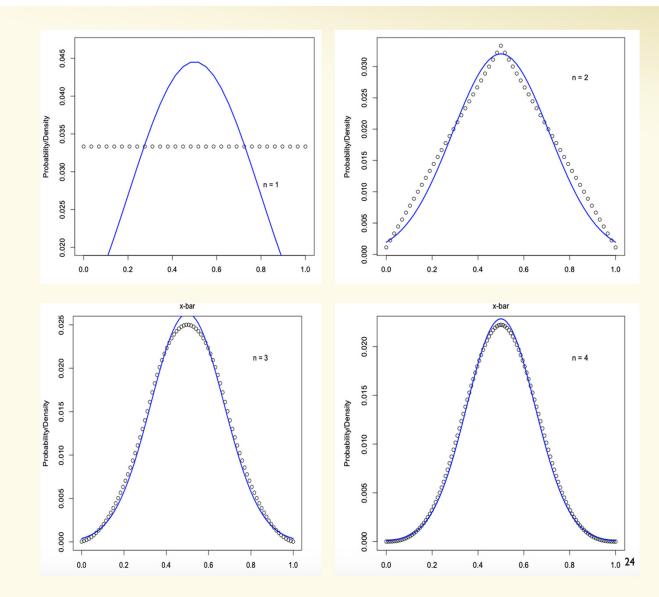
 $\mathbb{E}[S_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n] = n\mu$ $Var(S_n) = Var(X_1) + \dots + Var(X_n) = n\sigma^2$

Empirical observation: S_n looks like a normal RV as n grows.

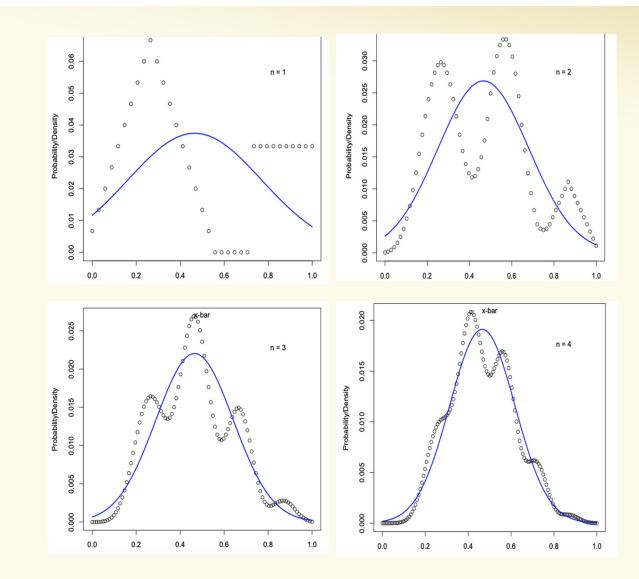
Example: Sum of n i.i.d. Exp(1) random variables











Central Limit Theorem

 X_1, \ldots, X_n i.i.d., each with expectation μ and variance σ^2

Define
$$S_n = X_1 + \dots + X_n$$
 and
 $Y_n = \frac{S_n - n\mu}{\sigma\sqrt{n}}$ $\int A \sigma^2$
 $\mathbb{E}[Y_n] = \frac{1}{\sigma\sqrt{n}} (\mathbb{E}[S_n] - n\mu) = \frac{1}{\sigma\sqrt{n}} (n\mu - n\mu) = 0$
 $\operatorname{Var}(Y_n) = \frac{1}{\sigma^2 n} (\operatorname{Var}(S_n - n\mu)) = \frac{\operatorname{Var}(S_n)}{\sigma^2 n} = \frac{\sigma^2 n}{\sigma^2 n} = 1$

Central Limit Theorem

$$X_n = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

Theorem. (Central Limit Theorem) The CDF of Y_n converges to the CDF of the standard normal $\mathcal{N}(0,1)$, i.e.,

$$\lim_{n \to \infty} P(Y_n \le y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{y} e^{-x^2/2} \mathrm{d}x$$

Central Limit Theorem

$$X_n = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

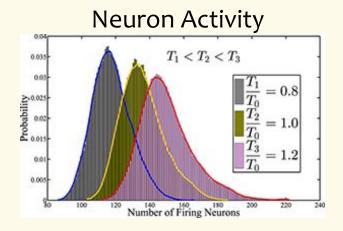
Theorem. (Central Limit Theorem) The CDF of Y_n converges to the CDF of the standard normal $\mathcal{N}(0,1)$, i.e.,

$$\lim_{n \to \infty} P(Y_n \le y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{y} e^{-x^2/2} \mathrm{d}x$$

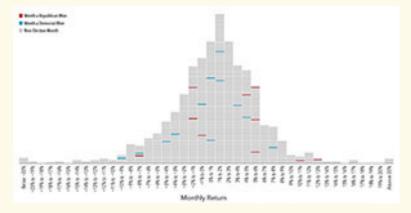
Also stated as:

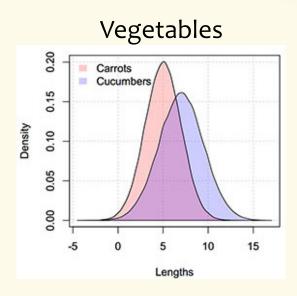
- $\lim_{n\to\infty} Y_n \to \mathcal{N}(0,1)$
- $\lim_{n \to \infty} \frac{1}{n} \sum_{i=1}^{n} X_i \to \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$ for $\mu = \mathbb{E}[X_i]$ and $\sigma^2 = \operatorname{Var}(X_i)$

CLT → **Normal Distribution EVERYWHERE**



S&P 500 Returns after Elections





Examples from: https://galtonboard.com/probabilityexamplesinlife