## CSE 312: Foundations of Computing II

## Section 6: Joint Distributions

## 1. Random Stick

You hold a stick of unit length (1). Someone comes along and breaks off a random piece at some point $Y \sim$ Unif $(0,1)$. Now you hold a stick of length $Y$. Another person comes along and breaks off another piece from the remaining part of the stick that you hold at point $X \sim \operatorname{Unif}(0, Y)$. You are left with a stick of length $X$. Find the PDF $f_{X}(x)$, mean $\mathbb{E}[X]$ using LTE and variance $\operatorname{Var}(X)$ using LTE as well.
Solution:
(a) First, let's solve for the PDF $f_{X}(x)$. First notice that:

$$
f_{Y}(y)= \begin{cases}1, & y \in(0,1) \\ 0, & \text { otherwise }\end{cases}
$$

Further:

$$
f_{X \mid Y}(x \mid y)= \begin{cases}\frac{1}{y}, & 0<x<y<1 \\ 0, & \text { otherwise }\end{cases}
$$

This means that by the law of total probability and the definition of marginal distributions we have:

$$
\begin{aligned}
f_{X}(x) & =\int_{-\infty}^{\infty} f_{X \mid Y}(x \mid y) f_{Y}(y) d y \\
& =\int_{x}^{1} \frac{1}{y} d y \\
& =-\ln (x), \text { for } x \in(0,1)
\end{aligned}
$$

(b) To solve for the expected value of $X$ we will use conditional expectation. First note that:

$$
(X \mid Y=y) \sim U n i f(0, y)
$$

which tells us that

$$
\mathbb{E}[X \mid Y=y]=\frac{1}{2}(0+y)=\frac{1}{2} y
$$

Thus:

$$
\begin{aligned}
\mathbb{E}[X] & =\int_{0}^{1} \mathbb{E}[X \mid Y=y] f_{Y}(y) d y \\
& =\int_{0}^{1} \frac{1}{2} y \cdot 1 d y \\
& =\frac{1}{4}
\end{aligned}
$$

(c) We can similarly solve for the variance of $X$. First we note that:

$$
\mathbb{E}\left[X^{2}\right]=\operatorname{Var}(X)+\mathbb{E}[X]^{2}
$$

from the definition of variance. If we add conditioning, we get:

$$
\begin{aligned}
\mathbb{E}\left[X^{2} \mid Y=y\right] & =\operatorname{Var}(X \mid Y=y)+\mathbb{E}[X \mid Y=y]^{2} \\
& =\frac{1}{12}(y-0)^{2}+\left(\frac{1}{2} y\right)^{2} \\
& =\frac{1}{3} y^{2}
\end{aligned}
$$

Which allows us to calculate that:

$$
\begin{aligned}
\mathbb{E}\left[X^{2}\right] & =\int_{0}^{1} \mathbb{E}\left[X^{2} \mid Y=y\right] f_{Y}(y) d y \\
& =\int_{0}^{1} \frac{1}{3} y^{2} \cdot 1 d y \\
& =\frac{1}{9}
\end{aligned}
$$

So we finally have:

$$
\operatorname{Var}(X)=\mathbb{E}\left[X^{2}\right]-\mathbb{E}[X]^{2}=\frac{1}{9}-\left(\frac{1}{4}\right)^{2}=\frac{7}{144}
$$

## 2. Another Urn Question

An urn has 12 balls, 5 red ones and 7 green ones. Draw 3 balls. Let $X$ denote the number of red balls in the sample. Compute $\operatorname{Var}(X)$ when sampling is done:
(a) With replacement
(b) Without replacement

## Solution:

(a) We start by introducing the indicator variables:

$$
\begin{aligned}
X_{1} & = \begin{cases}1, & \text { first ball is red } \\
0, & \text { first ball is green }\end{cases} \\
X_{2} & = \begin{cases}1, & \text { second ball is red } \\
0, & \text { second ball is green }\end{cases} \\
X_{3} & = \begin{cases}1, & \text { third ball is red } \\
0, & \text { third ball is green }\end{cases}
\end{aligned}
$$

Then, note that $X_{1}, X_{2}, X_{3}$ are all $\operatorname{Ber}\left(\frac{5}{12}\right.$ and are independent.
$\operatorname{Var}\left(X_{1}\right)=\operatorname{Var}\left(X_{2}\right)=\operatorname{Var}\left(X_{3}\right)=p(1-p)=\frac{5}{12}\left(1-\frac{5}{12}\right)=\frac{35}{144}$
So:

$$
\begin{array}{rlr}
\operatorname{Var}(X) & =\operatorname{Var}\left(X_{1}+X_{2}+X_{3}\right) & \\
& =\operatorname{Var}\left(X_{1}\right)+\operatorname{Var}\left(X_{2}\right)+\operatorname{Var}\left(X_{3}\right)[\text { since they are all independent }] \\
& =\frac{35}{144}+\frac{35}{144}+\frac{35}{144} & =\frac{35}{48}
\end{array}
$$

(b) We can consider this as taking out 3 balls. We will define $X_{1}$ for the first ball being red, $X_{2}$ for the second ball being red, and $X_{3}$ for the third ball being red. The independence of $X_{1}, X_{2}$, and $X_{3}$ is no longer true. So, we will want to solve for the covariance matrix and find the sum of the entries. However, the marginal distributions of $X_{1}, X_{2}$, and $X_{3}$ are $\operatorname{Ber}\left(\frac{5}{12}\right)$, since each has a probability $\frac{5}{12}$ of being red.

We can prove this because there are 5 red balls and 12 total balls. We want to calculate the probability that the ith ball is red, after we choose 3 balls from the urn. There are a total of $\mathbb{P}(12,3)$ ways to order
the 3 balls we picked from the urn. There are 5 ways to fix the red ball at the ith position. There are $\mathbb{P}(12-1,3-1)$ ways to order the remaining 11 balls for the other 2 positions. This leaves us with:

$$
\begin{aligned}
\mathbb{P}\left(X_{i}=1\right) & =\frac{5 P(11,2)}{P(12,3} \\
& =\frac{5 \frac{11!}{9!}}{\frac{12!}{9!}} \\
& =\frac{5}{12}
\end{aligned}
$$

Which means that $X_{i} \sim \operatorname{Ber}\left(\frac{5}{12}\right)$.
We calculated the variance for this above, and we have:

$$
\operatorname{Cov}\left(X_{i}, X_{i}\right)=\operatorname{Var}\left(X_{i}\right)=\frac{35}{144}
$$

Now, we have $\operatorname{Cov}\left(X_{1}, X_{2}\right)=\mathbb{E}\left[X_{1} X_{2}\right]-\mathbb{E}\left[X_{1}\right] \mathbb{E}\left[X_{2}\right]$.
Note that $X_{1} \cdot X_{2}$ is only 1 when both are 1 . We start with 5 red balls of the 12 in the first choice, and then 4 of the remaining 11 in the second choice. So:

$$
\mathbb{E}\left[X_{1} X_{2}\right]=1 \cdot \frac{5}{12} \cdot \frac{4}{11}=\frac{5}{33}
$$

Then, since $X_{1}$ is the first choice, with 5 red balls of the 12 :

$$
\mathbb{E}\left[X_{1}\right]=\frac{5}{12}
$$

For the second choice we have:

$$
\begin{aligned}
\mathbb{E}\left[X_{2}\right] & =\mathbb{E}\left[X_{2} \mid X_{1}=1\right] \mathbb{P}\left(X_{1}=1\right)+\mathbb{E}\left[X_{2} \mid X_{1}=0\right] \mathbb{P}\left(X_{1}=0\right) \\
& =\frac{4}{11} \cdot \frac{5}{12}+\frac{5}{11} \cdot \frac{7}{12} \\
& =\frac{5}{12}
\end{aligned}
$$

All together this means:

$$
\begin{aligned}
\operatorname{Cov}\left(X_{1}, X_{2}\right) & =\mathbb{E}\left[X_{1} X_{2}\right]-\mathbb{E}\left[X_{1}\right] \mathbb{E}\left[X_{2}\right] \\
& =-\frac{35}{1584}
\end{aligned}
$$

In fact, you will find that for any $i, \mathbb{E}\left[X_{i}\right]=\mathbb{P}\left(X_{1}=1\right)=\frac{5}{12}$, when we consider each of these variables marginally. Further, for any $i \neq j, \mathbb{E}\left[X_{i} X_{j}\right]=\frac{5}{33}$, we can solve for these manually considering all cases, or consider the similarity to the hat check or cat and mitten problem. We have:

$$
\begin{aligned}
\mathbb{E}\left[X_{i} X_{j}\right] & =\mathbb{P}\left(X_{i}=1\right) \cdot \mathbb{P}\left(X_{j}=1 \mid X_{i}=1\right) \\
& =\frac{5}{12} \cdot \frac{4}{11} \\
& =\frac{5}{33}
\end{aligned}
$$

So, for any $i \neq j, \operatorname{Cov}\left(X_{i}, X_{j}\right)=-\frac{35}{1584}$.
This gives us the following covariance matrix:

|  | $X_{1}$ | $X_{2}$ | $X_{3}$ |
| :---: | :---: | :---: | :---: |
| $X_{1}$ | $\frac{35}{144}$ | $-\frac{35}{1554}$ | $-\frac{35}{1584}$ |
| $X_{2}$ | $-\frac{35}{1584}$ | $\frac{35}{144}$ | $-\frac{35}{1584}$ |
| $X_{3}$ | $-\frac{35}{1584}$ | $-\frac{35}{1584}$ | $\frac{35}{144}$ |

So, we have the following:

$$
\begin{aligned}
\operatorname{Cov}(X, X) & =\operatorname{Cov}\left(\sum_{i=1}^{3} X_{i}, \sum_{j=1}^{3} X_{j}\right) \\
& =\sum_{i=1}^{3} \sum_{j=1}^{3} \operatorname{Cov}\left(X_{i}, X_{j}\right) \\
& =3 \cdot \frac{35}{144}+2\binom{3}{2} \cdot\left(-\frac{35}{1584}\right) \\
& =\frac{105}{176}
\end{aligned}
$$

## 3. Continuous Joint Density

The joint probability density function of $X$ and $Y$ is given by

$$
f_{X, Y}(x, y)= \begin{cases}\frac{6}{7}\left(x^{2}+\frac{x y}{2}\right) & 0<x<1,0<y<2 \\ 0 & \text { otherwise } .\end{cases}
$$

(a) Verify that this is indeed a joint density function.
(b) Compute the marginal density function of $X$.
(c) Find $P\left(\left.Y>\frac{1}{2} \right\rvert\, X<\frac{1}{2}\right)$.
(d) Find $E(X)$.
(e) Find $E(Y)$

## Solution:

(a) A joint density function will integrate to 1 over all possible values. Thus, we integrate over the joint range range using Wolfram Alpha, and see that:

$$
\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X, Y}(x, y) d x d y=\int_{0}^{2} \int_{0}^{1} \frac{6}{7}\left(x^{2}+\frac{x y}{2}\right) d x d y=1
$$

We also need to check that the density is nonnegative, but that is easily seen to be true.
(b) We apply the definition of the marginal density function of $X$, using the fact that we only need to integrate over the values where the joint density is positive:

$$
f_{X}(x)= \begin{cases}\int_{0}^{2} \frac{6}{7}\left(x^{2}+\frac{x y}{2}\right) d y=\frac{6}{7} x(2 x+1) & 0<x<1 \\ 0 & \text { otherwise }\end{cases}
$$

(c) By the definition of conditional probability:

$$
\mathbb{P}\left(\left.Y>\frac{1}{2} \right\rvert\, X<\frac{1}{2}\right)=\frac{\mathbb{P}\left(Y>\frac{1}{2}, X<\frac{1}{2}\right)}{\mathbb{P}\left(X<\frac{1}{2}\right)}
$$

For the numerator, we have

$$
\begin{aligned}
\mathbb{P}(Y & \left.>\frac{1}{2}, X<\frac{1}{2}\right)=\int_{1 / 2}^{\infty} \int_{-\infty}^{1 / 2} f_{X, Y}(x, y) d x d y \\
& =\int_{1 / 2}^{2} \int_{0}^{1 / 2} \frac{6}{7}\left(x^{2}+\frac{x y}{2}\right) d x d y=\frac{69}{448}
\end{aligned}
$$

For the denominator, we can integrate using the marginal distribution that we found before:

$$
\int_{0}^{1 / 2} \frac{6}{7} x(2 x+1) d x=\frac{5}{28}
$$

Putting these together, we get:

$$
\mathbb{P}\left(\left.Y>\frac{1}{2} \right\rvert\, X<\frac{1}{2}\right)=\frac{\frac{69}{448}}{\frac{5}{28}}=0.8625
$$

(d) By definition, and using $\Omega_{X}=(0,1)$ :

$$
\mathbb{E}[X]=\int_{0}^{1} f_{X}(x) x d x=\int_{0}^{1} \frac{6}{7} x(2 x+1) x d x=\frac{5}{7}
$$

(e) By definition, and using $\Omega_{Y}=(0,2)$ :

$$
\mathbb{E}[Y]=\int_{0}^{2} f_{Y}(y) y d y=\int_{0}^{2} \frac{1}{14}(3 y+4) y d y=\frac{8}{7}
$$

