

**CSE 312**

# **Foundations of Computing II**

**Lecture 15: Normal Distribution & Central Limit Theorem**

## Announcements

- Midterm on Wed
- Review session by Rachel Lin tomorrow at 10:00am at Gates 271
- Practice session by Zhiyang Lim tomorrow at 4:30pm at Gates 271
- Concentrated office hour this week Monday and Tuesday. See schedule on Ed.
- Seat assignment on Ed.

# Review Continuous RVs 13

## Probability Density Function (PDF).

$f: \mathbb{R} \rightarrow \mathbb{R}$  s.t.

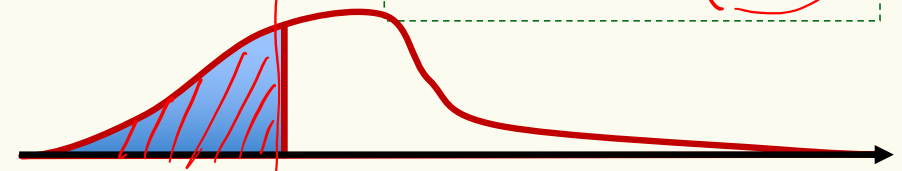
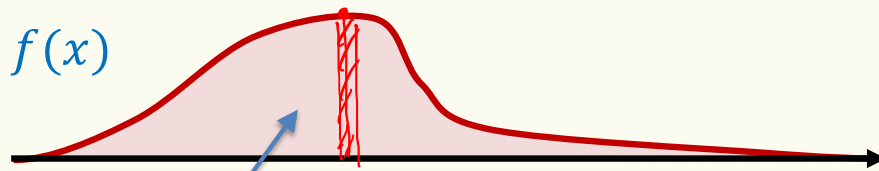
- $f(x) \geq 0$  for all  $x \in \mathbb{R}$
- $\int_{-\infty}^{+\infty} f(x) dx = 1$

## Cumulative Distribution Function (CDF).

$$F(y) = \int_{-\infty}^y f(x) dx$$

$$P(X \leq y) = P(X \leq y)$$

Theorem.  $f(x) = \frac{dF(x)}{dx}$



Density  $\neq$  Probability!

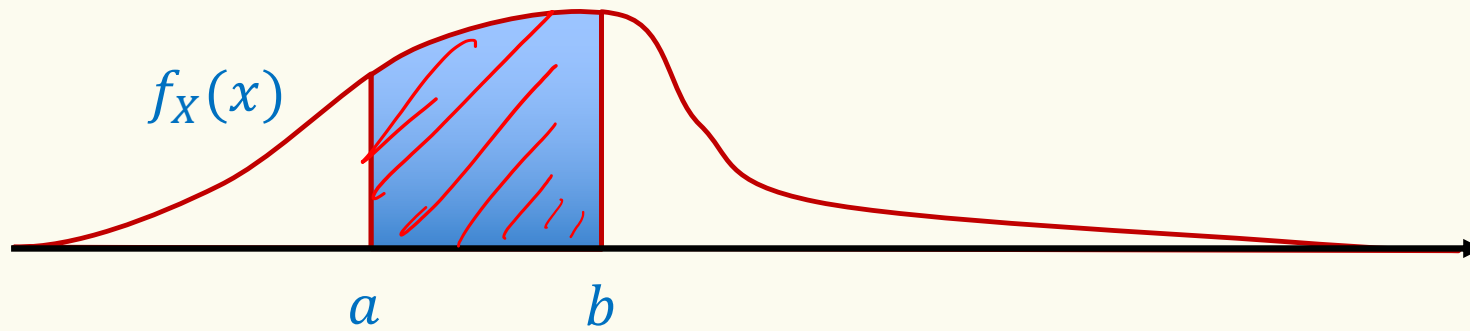
$$P[X \approx x] = \epsilon \cdot f(x)$$

$$\in [-\frac{\epsilon}{2} + x, x + \frac{\epsilon}{2}]$$

$$F_X(y) = P(X \leq y)$$

$$\frac{P[X \approx x]}{P[X \approx y]} = \frac{\epsilon f(x)}{\epsilon f(y)}$$

## Review Continuous RVs



$$P(X \in [a, b]) = \int_a^b f_X(x) dx = F_X(b) - F_X(a)$$

## Review Exponential Distribution

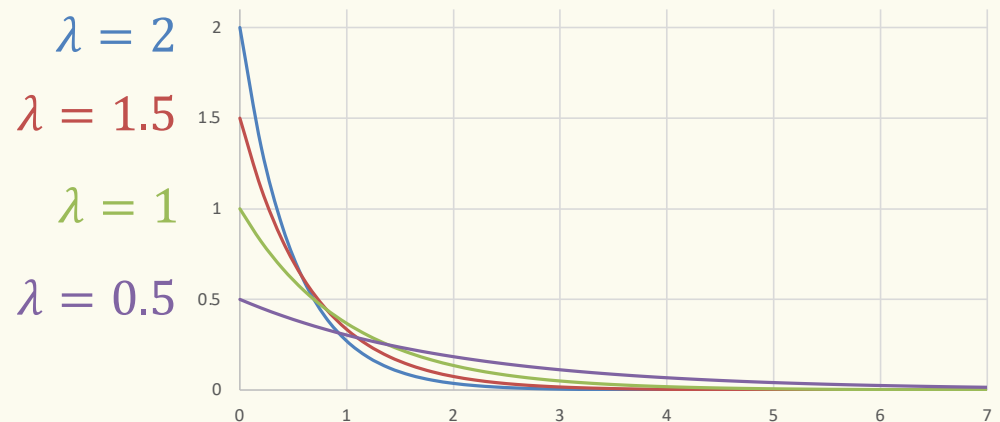
**Definition.** An **exponential random variable**  $X$  with parameter  $\lambda \geq 0$  is follows the exponential density

$$\lim_{n \rightarrow \infty} \text{Geo}(n, \lambda) \rightarrow \text{Exp}(\lambda) \quad f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & x < 0 \end{cases}$$

We write  $X \sim \text{Exp}(\lambda)$  and say  $X$  that follows the exponential distribution.

CDF: For  $y \geq 0$ ,

$$F_X(y) = 1 - e^{-\lambda y}$$



# Agenda

- Normal Distribution ◀
- Practice with Normals
- Central Limit Theorem (CLT)

# The Normal Distribution



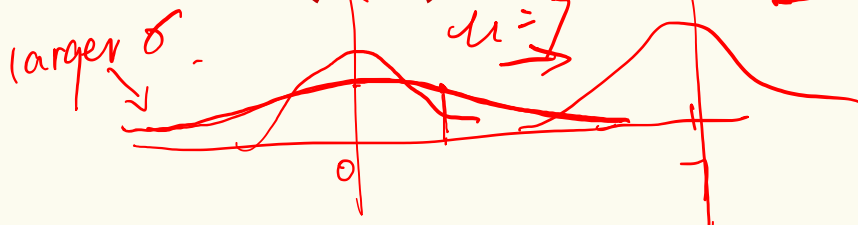
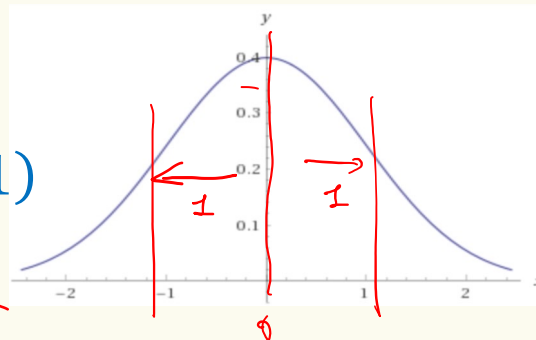
Carl Friedrich Gauss  
x=7

**Definition.** A **Gaussian (or normal) random variable** with parameters  $\mu \in \mathbb{R}$  and  $\sigma \geq 0$  has density

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

We say that  $X$  follows the Normal Distribution, and write  $X \sim \mathcal{N}(\mu, \sigma^2)$ .

**Standard (unit) normal =  $\mathcal{N}(0, 1)$**



$$f_Y(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

$$\int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} dx = \sqrt{2\pi}$$

$$\frac{1}{\sqrt{2\pi}} e^{-\frac{(x-7)^2}{2}}$$

x=0

# The Normal Distribution

$(x=)$

**Definition.** A **Gaussian (or normal) random variable** with parameters  $\mu \in \mathbb{R}$  and  $\sigma \geq 0$  has density

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$x = \mu + \epsilon$   
 $x' = \mu - \epsilon$   
 $f(x) = f(x')$



Carl Friedrich Gauss

We say that  $X$  follows the Normal Distribution, and write  $X \sim \mathcal{N}(\mu, \sigma^2)$ .

**Fact.** If  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then  $\mathbb{E}[X] = \mu$ , and  $\text{Var}(X) = \sigma^2$

Proof of expectation is easy because density curve is symmetric around  $\mu$ ,

$f_X(\mu - x) = f_X(\mu + x)$ , but proof for variance requires integration of  $e^{-x^2/2}$



# The Normal Distribution

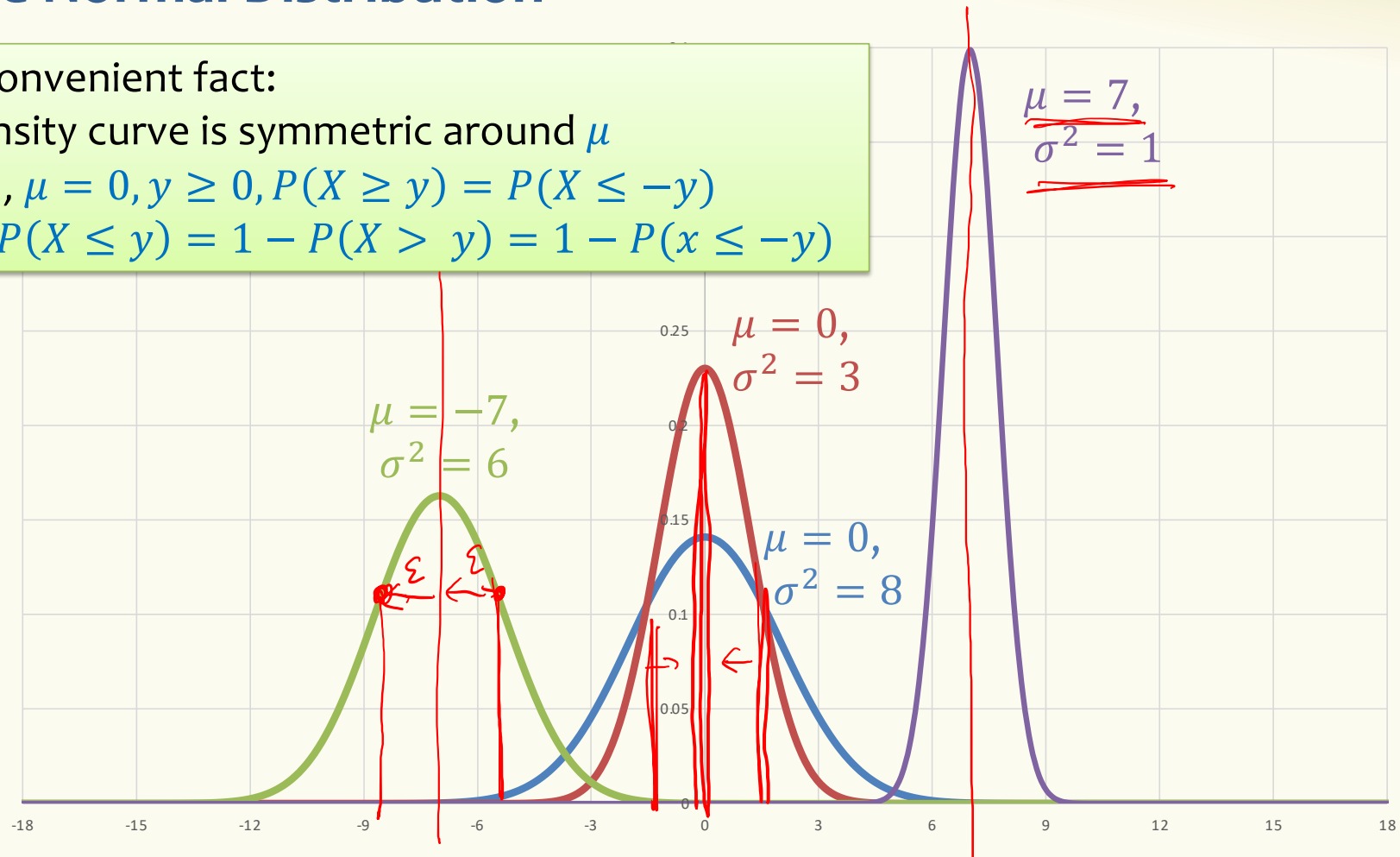
Aka a “Bell Curve” (imprecise name)

A convenient fact:

Density curve is symmetric around  $\mu$

E.g.,  $\mu = 0, y \geq 0, P(X \geq y) = P(X \leq -y)$

→  $P(X \leq y) = 1 - P(X > y) = 1 - P(x \leq -y)$



## Closure of normal distribution – Under Shifting and Scaling

**Fact.** If  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then  $Y = aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$

**Proof.**  $\mathbb{E}[Y] = a \mathbb{E}[X] + b = a\mu + b$   
 $\text{Var}(Y) = a^2 \text{Var}(X) = a^2\sigma^2$

Can show with algebra that the PDF of  $Y = aX + b$  is still normal.

A very useful fact:  $\frac{X-\mu}{\sigma} \sim \mathcal{N}(0, 1)$

## CDF of normal distribution

$$Y = \frac{X - \mu}{\sigma} = \frac{1}{\sigma} X - \left(\frac{\mu}{\sigma}\right)$$

$\frac{1}{\sigma} E[X] - \frac{\mu}{\sigma} = 0$   
 $\frac{1}{\sigma^2} \text{Var}[X] = 1$

**Fact.** If  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then  $Y = aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$

**Standard (unit) normal =  $\mathcal{N}(0, 1)$**

**CDF.**  $\Phi(y) = P(Y \leq y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^y e^{-x^2/2} dx$  for  $Y \sim \mathcal{N}(0, 1)$

Note:  $\Phi(y)$  has no closed form – generally given via tables

If  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then  $F_X(z) = P(X \leq z) = P\left(\frac{X - \mu}{\sigma} \leq \frac{z - \mu}{\sigma}\right) = \Phi\left(\frac{z - \mu}{\sigma}\right)$

Useful fact:  $\frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1)$

$Y = \frac{X - \mu}{\sigma}$        $Y \sim \mathcal{N}(0, 1)$

# Table of Standard Cumulative Normal Density

$$\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

$x - x$

$\forall y \geq 0$   $P(X \leq -y) = 1 - P(X \leq y)$   $z = 1.09$

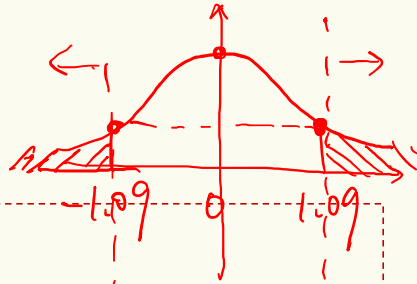
$P(Z \leq 1.09) = \Phi(1.09) \approx 0.8621$

Φ Table:  $P(Z \leq z)$  when  $Z \sim \mathcal{N}(0, 1)$

z	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
0.0	0.5	0.50399	0.50798	0.51197	0.51595	0.51994	0.52392	0.5279	0.53188	0.53586
0.1	0.53983	0.5438	0.54776	0.55172	0.55567	0.55962	0.56356	0.56749	0.57142	0.57535
0.2	0.57926	0.58317	0.58706	0.59095	0.59483	0.59871	0.60257	0.60642	0.61026	0.61409
0.3	0.61791	0.62172	0.62552	0.6293	0.63307	0.63683	0.64058	0.64431	0.64803	0.65173
0.4	0.65542	0.6591	0.66276	0.6664	0.67003	0.67364	0.67724	0.68082	0.68439	0.68793
0.5	0.69146	0.69497	0.69847	0.70194	0.7054	0.70884	0.71226	0.71566	0.71904	0.7224
0.6	0.72575	0.72907	0.73237	0.73565	0.73891	0.74215	0.74537	0.74857	0.75175	0.7549
0.7	0.75804	0.76115	0.76424	0.7673	0.77035	0.77337	0.77637	0.77935	0.7823	0.78524
0.8	0.78814	0.79103	0.79389	0.79673	0.79955	0.80234	0.80511	0.80785	0.81057	0.81327
0.9	0.81594	0.81859	0.82121	0.82381	0.82639	0.82894	0.83147	0.83398	0.83646	0.83891
1.0	0.84134	0.84375	0.84614	0.84849	0.85083	0.85314	0.85543	0.85769	0.85993	0.86214
1.1	0.86433	0.8665	0.86864	0.87076	0.87286	0.87493	0.87698	0.879	0.881	0.88298
1.2	0.88493	0.88686	0.88877	0.89065	0.89251	0.89435	0.89617	0.89796	0.89973	0.90147
1.3	0.9032	0.9049	0.90658	0.90824	0.90988	0.91149	0.91309	0.91466	0.91621	0.91774
1.4	0.91924	0.92073	0.9222	0.92364	0.92507	0.92647	0.92785	0.92922	0.93056	0.93189
1.5	0.93319	0.93448	0.93574	0.93699	0.93822	0.93943	0.94062	0.94179	0.94295	0.94408
1.6	0.9452	0.9463	0.94738	0.94845	0.9495	0.95053	0.95154	0.95254	0.95352	0.95449
1.7	0.95543	0.95637	0.95728	0.95818	0.95907	0.95994	0.9608	0.96164	0.96246	0.96327
1.8	0.96415	0.96485	0.96562	0.96638	0.96712	0.96784	0.96856	0.96926	0.96995	0.97062
1.9	0.9713	0.97193	0.97257	0.9732	0.97381	0.97441	0.975	0.97558	0.97615	0.9767
2.0	0.97725	0.97778	0.97831	0.97882	0.97932	0.97982	0.9803	0.98077	0.98124	0.98169
2.1	0.98214	0.98257	0.983	0.98341	0.98382	0.98422	0.98461	0.985	0.98537	0.98574
2.2	0.98611	0.98645	0.98679	0.98713	0.98745	0.98778	0.98809	0.9884	0.9887	0.98899
2.3	0.98938	0.98956	0.98983	0.9901	0.99036	0.99061	0.99086	0.99111	0.99134	0.99158
2.4	0.99182	0.99202	0.99224	0.99245	0.99266	0.99286	0.99305	0.99324	0.99343	0.99361
2.5	0.99379	0.99396	0.99413	0.9943	0.99446	0.99461	0.99477	0.99492	0.99506	0.9952
2.6	0.99533	0.99543	0.99552	0.9956	0.99568	0.99575	0.99582	0.99589	0.99596	0.99603
2.7	0.99609	0.99616	0.99623	0.9963	0.99637	0.99643	0.99649	0.99655	0.99661	0.99667
2.8	0.99673	0.99678	0.99684	0.99689	0.99694	0.99699	0.99704	0.99709	0.99714	0.99719
2.9	0.99724	0.99729	0.99734	0.99739	0.99744	0.99749	0.99754	0.99759	0.99764	0.99769
3.0	0.99774	0.99779	0.99784	0.99789	0.99794	0.99799	0.99804	0.99809	0.99814	0.99819
3.1	0.99824	0.99829	0.99834	0.99839	0.99844	0.99849	0.99854	0.99859	0.99864	0.99869
3.2	0.99874	0.99879	0.99884	0.99889	0.99894	0.99899	0.99904	0.99909	0.99914	0.99919
3.3	0.99924	0.99929	0.99934	0.99939	0.99944	0.99949	0.99954	0.99959	0.99964	0.99969
3.4	0.99974	0.99979	0.99984	0.99989	0.99994	0.99999	1.00004	1.00009	1.00014	1.00019

What is

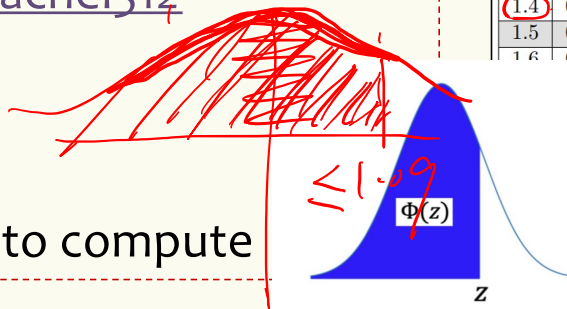
$P(Z \leq -1.09)$ ?



Poll:

[pollev.com/rachel312](http://pollev.com/rachel312)

- a. 0.1379
- b. 0.8621
- c. 0
- d. Not able to compute



Density curve is symmetric around  $\mu = 0$   
 $\rightarrow P(X \leq -y) = P(X \geq y) = 1 - P(X \leq y)$   
 $= 1 - P(Z \leq 1.09)$

## Closure of the normal -- under addition

$X + Y$

**Fact.** If  $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$ ,  $Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$  (both independent normal RV) then  $aX + bY + c \sim \mathcal{N}(a\mu_X + b\mu_Y + c, a^2\sigma_X^2 + b^2\sigma_Y^2)$

Note: The special thing is that **the sum of normal RVs is still a normal RV.**

The values of the expectation and variance are **not** surprising.

$X, Y$  are independent

### Why not surprising?

- Linearity of expectation (always true)
- When  $X$  and  $Y$  are independent,  $\text{Var}(aX + bY) = a^2\text{Var}(X) + b^2\text{Var}(Y)$

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

# Agenda

- Normal Distribution
- Practice with Normals ◀
- Central Limit Theorem (CLT)

## What about Non-standard normal?

If  $X \sim \mathcal{N}(\underline{\mu}, \underline{\sigma^2})$ , then  $\overset{Y}{=} \frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1)$

Therefore,

$$\underline{F_X(z)} = \underline{P(X \leq z)} = P\left(\frac{X - \mu}{\sigma} \leq \frac{z - \mu}{\sigma}\right) = \Phi\left(\frac{z - \mu}{\sigma}\right)$$

*Handwritten notes: A red arrow points from the fraction  $\frac{X - \mu}{\sigma}$  in the probability expression to the  $\frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1)$  statement above. The fraction  $\frac{z - \mu}{\sigma}$  in the final expression is underlined.*

## Example

$\mu =$        $\sigma = 2$

Let  $X \sim \mathcal{N}(\underline{0.4}, 4 = 2^2)$ .

$$\begin{aligned} P(\underline{X \leq 1.2}) &= P\left(\frac{X - 0.4}{2} \leq \frac{1.2 - 0.4}{2}\right) \\ &= P\left(\frac{X - 0.4}{2} \leq \underline{0.4}\right) = \Phi(0.4) \approx 0.6554 \\ &\quad \sim \mathcal{N}(0, 1) \end{aligned}$$

0.1	0.5398	0.5438
0.2	0.5793	0.5832
0.3	0.6179	0.6217
0.4	0.6554	0.6591
0.5	0.6915	0.6950
0.6	0.7257	0.7291
0.7	0.7580	0.7611



## Example

Let  $X \sim \mathcal{N}(3, 16)$ .  $\sigma^2 = 16$   $\sigma = 4$   $N \sim (0, 1)$

$$P(2 < X < 5) = P\left(\frac{2-3}{4} < \frac{X-3}{4} < \frac{5-3}{4}\right)$$

$$= P\left(-\frac{1}{4} < Z < \frac{1}{2}\right)$$

$$= \Phi\left(\frac{1}{2}\right) - \Phi\left(-\frac{1}{4}\right)$$

Density curve is symmetric around  $\mu = 0$   
 $\rightarrow P(X \leq -y) = P(X \geq y) = 1 - P(X \leq y)$

$$= \Phi\left(\frac{1}{2}\right) - \left(1 - \Phi\left(\frac{1}{4}\right)\right) \approx 0.29017$$

## Example – How Many Standard Deviations Away?

Let  $X \sim \mathcal{N}(\mu, \sigma^2)$ .

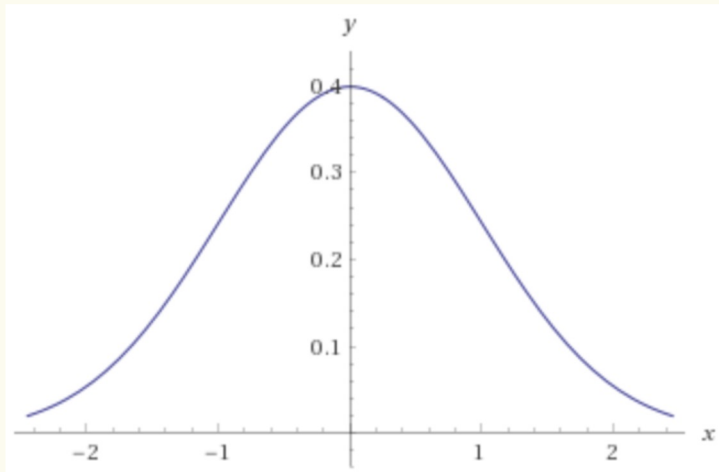
$$\begin{aligned} P(|X - \mu| < k\sigma) &= P\left(\frac{|X - \mu|}{\sigma} < k\right) = \\ &= P\left(-k < \frac{X - \mu}{\sigma} < k\right) = \Phi(k) - \Phi(-k) \end{aligned}$$

e.g.  $k = 1$ : 68%

$k = 2$ : 95%

$k = 3$ : 99%

# Halloween Brain Break



**Normal Distribution**



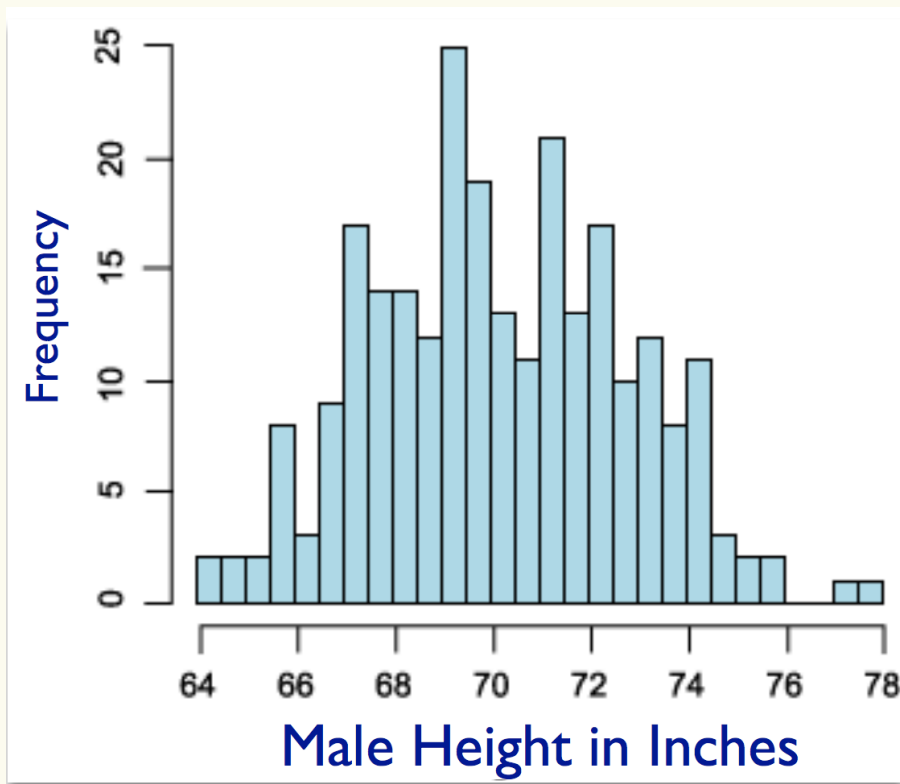
**Paranormal Distribution**

# Agenda

- Normal Distribution
- Practice with Normals
- Central Limit Theorem (CLT) ◀

## Gaussian in Nature

Empirical distribution of collected data often resembles a Gaussian ...



e.g. Height distribution resembles Gaussian.

R.A.Fisher (1918) observed that the height is likely the outcome of the sum of many independent random parameters, i.e., can be written as

$$X = X_1 + \dots + X_n$$

## Sum of Independent RVs

i.i.d. = independent and identically distributed

$X_1, \dots, X_n$  i.i.d. with expectation  $\mu$  and variance  $\sigma^2$

Define

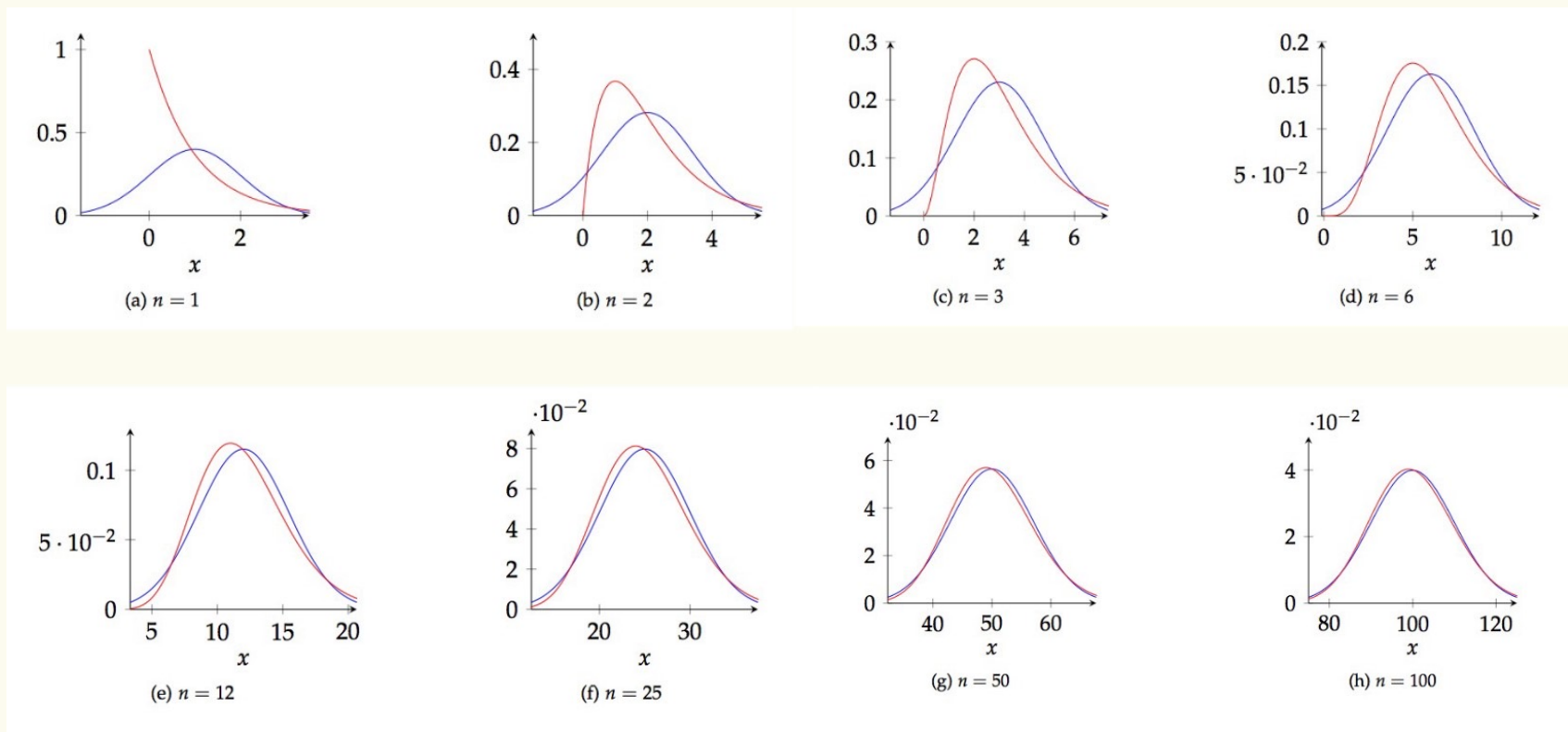
$$S_n = X_1 + \dots + X_n$$

$$\mathbb{E}[S_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n] = n\mu$$

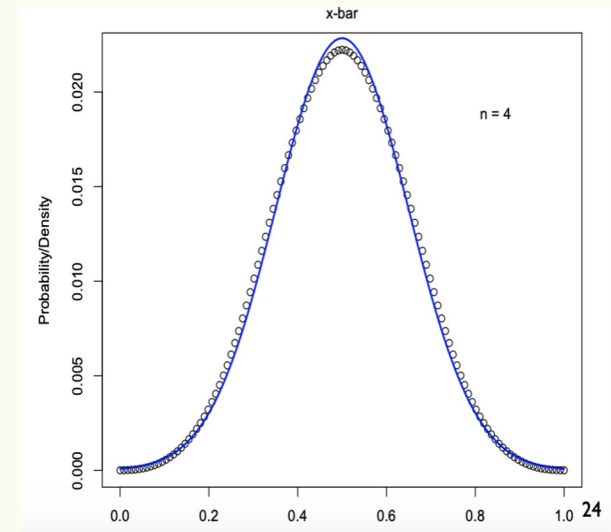
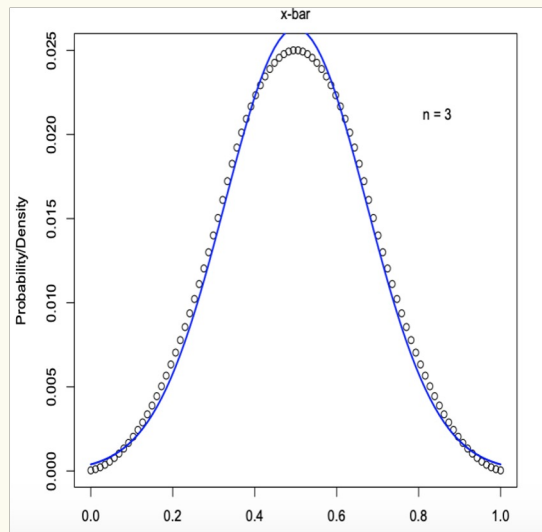
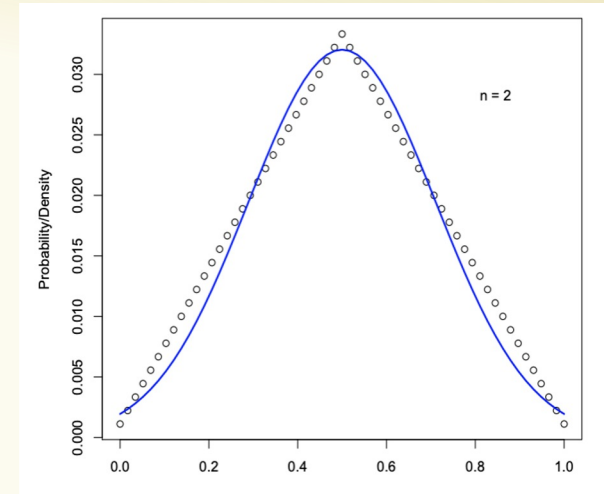
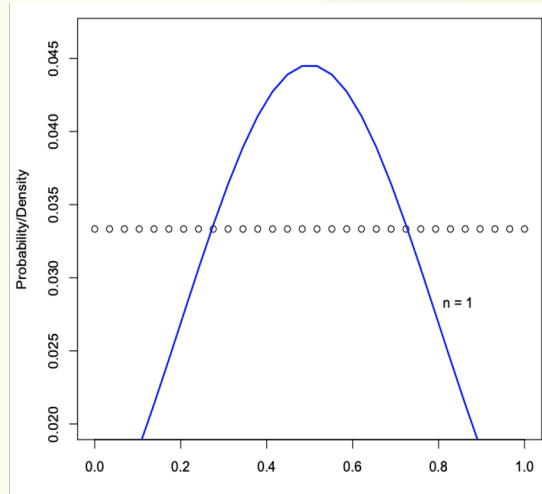
$$\text{Var}(S_n) = \text{Var}(X_1) + \dots + \text{Var}(X_n) = n\sigma^2$$

**Empirical observation:**  $S_n$  looks like a normal RV as  $n$  grows.

## Example: Sum of $n$ i.i.d. $\text{Exp}(1)$ random variables

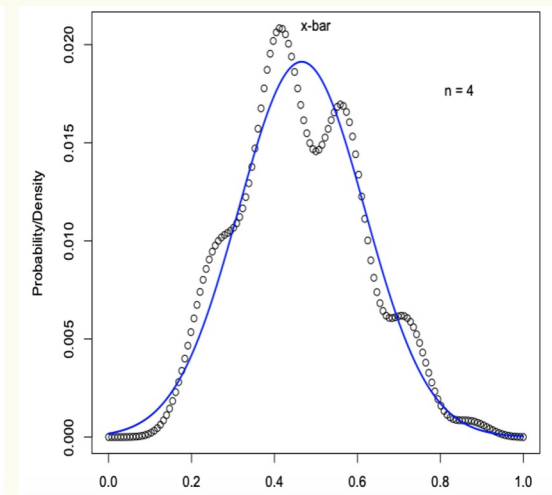
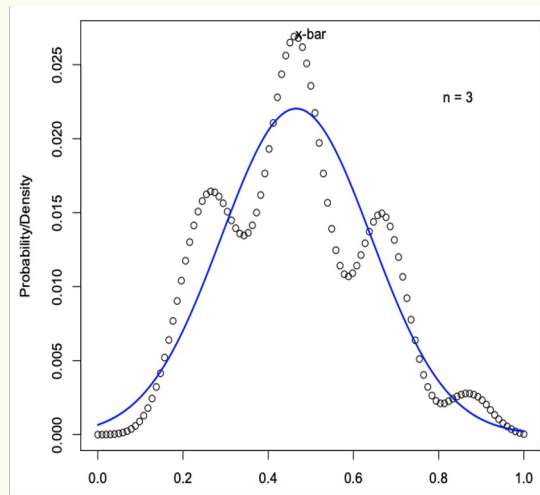
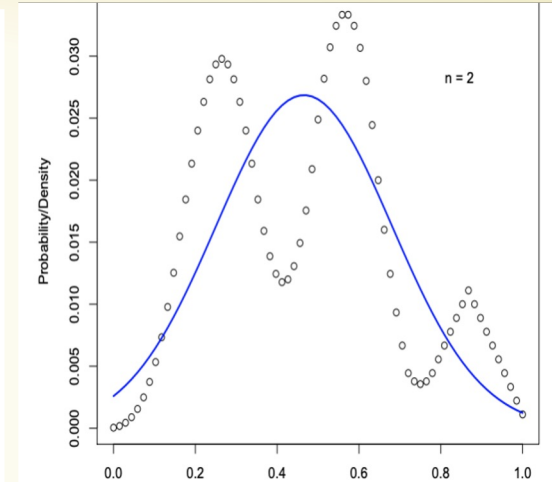
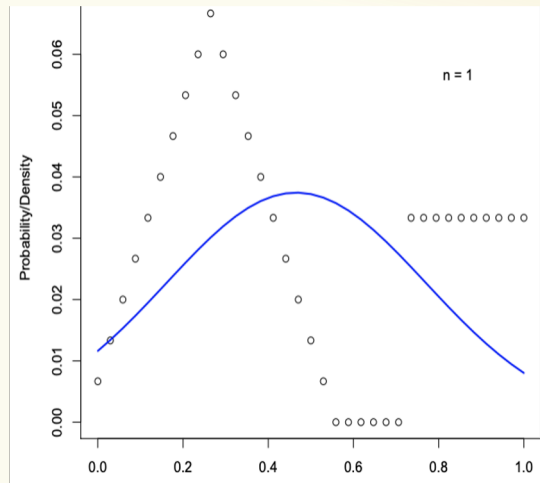


# CLT (Idea)





# CLT (Idea)



## Central Limit Theorem

$X_1, \dots, X_n$  i.i.d., each with expectation  $\mu$  and variance  $\sigma^2$

Define  $S_n = X_1 + \dots + X_n$  and

$$Y_n = \frac{S_n - n\mu}{\sigma\sqrt{n}}$$

$$\mathbb{E}[Y_n] = \frac{1}{\sigma\sqrt{n}} (\mathbb{E}[S_n] - n\mu) = \frac{1}{\sigma\sqrt{n}} (n\mu - n\mu) = 0$$

$$\text{Var}(Y_n) = \frac{1}{\sigma^2 n} (\text{Var}(S_n - n\mu)) = \frac{\text{Var}(S_n)}{\sigma^2 n} = \frac{\sigma^2 n}{\sigma^2 n} = 1$$

## Central Limit Theorem

$$Y_n = \frac{X_1 + \cdots + X_n - n\mu}{\sigma\sqrt{n}}$$

**Theorem. (Central Limit Theorem)** The CDF of  $Y_n$  converges to the CDF of the standard normal  $\mathcal{N}(0,1)$ , i.e.,

$$\lim_{n \rightarrow \infty} P(Y_n \leq y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^y e^{-x^2/2} dx$$

## Central Limit Theorem

$$Y_n = \frac{X_1 + \cdots + X_n - n\mu}{\sigma\sqrt{n}}$$

**Theorem. (Central Limit Theorem)** The CDF of  $Y_n$  converges to the CDF of the standard normal  $\mathcal{N}(0,1)$ , i.e.,

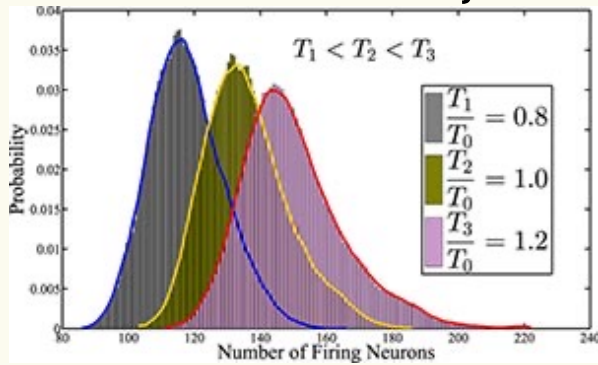
$$\lim_{n \rightarrow \infty} P(Y_n \leq y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^y e^{-x^2/2} dx$$

Also stated as:

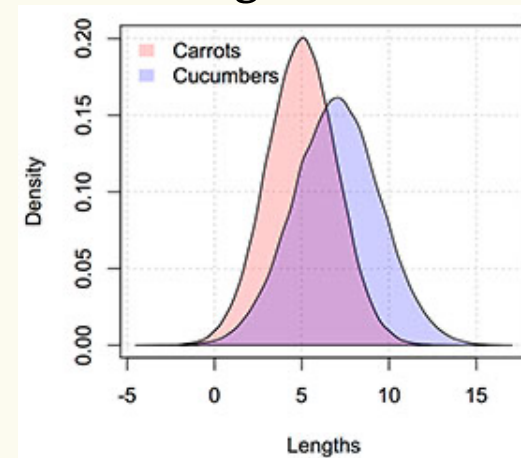
- $\lim_{n \rightarrow \infty} Y_n \rightarrow \mathcal{N}(0,1)$
- $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i \rightarrow \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$  for  $\mu = \mathbb{E}[X_i]$  and  $\sigma^2 = \text{Var}(X_i)$

# CLT → Normal Distribution EVERYWHERE

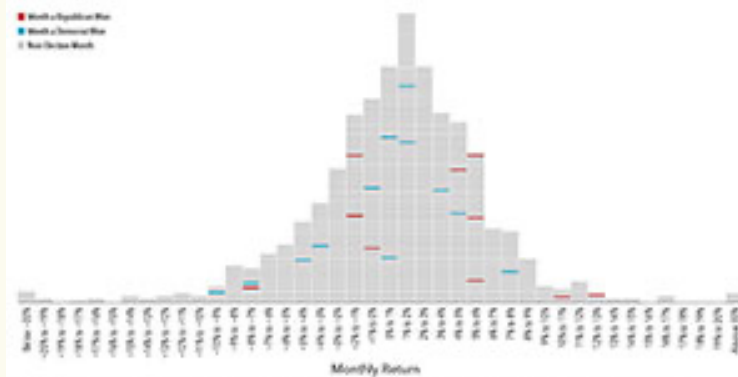
## Neuron Activity



## Vegetables



## S&P 500 Returns after Elections



Examples from:  
<https://galtonboard.com/probabilityexamplesinlife>