CSE 312

## Foundations of Computing II

Lecture 16: CLT \& Polling


## Review The Normal Distribution

Aka a "Bell Curve" (imprecise name)

A convenient fact: Symmetry around $\mu$ E.g., for standard normal

$$
P(X \leq-y)=1-P(x \leq y)
$$



Review Closure of normal distribution - Under Shifting and Scaling
Fact. If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $Y=\underline{a} X+\underline{b} \sim \mathcal{N}\left(a \mu+b, \underline{a^{2} \sigma^{2}}\right)$


## Review How Many Standard Deviations Away?



## Review

Table of $\Phi(\mathbf{z})$ CDF of Standard Normal Distribution
$\Phi$ Table: $\mathbb{P}(Z \leq z)$ when $Z \sim \mathcal{N}(0,1)$

| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.0 | 0.5 | 0.50399 | 0.50798 | 0.51197 | 0.51595 | 0.51994 | 0.52392 | 0.5279 | 0.53188 | 0.53586 |
| 0.1 | 0.53983 | 0.5438 | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 5 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
| 0.3 | 0.61791 | 0.62172 | 0.62552 | 0.6293 | 0.63307 | 0.63683 | 0.64058 | 0.64431 | 0.64803 | 0.65173 |
| 0.4 | 0.65542 | 0.6591 | 0.66276 | 0.6664 | 0.67003 | 0.67364 | 0.67724 | 0.68082 | 0.68439 | 0.68793 |
| 0.5 | 0.69146 | 0.69497 | 0.69847 | 0.70194 | 0.7054 | 0.7088 | 0.7122 | 0.71566 | 0.71904 | . 7224 |
| 0.6 | 0.72575 | 0.72907 | 0.73237 | 0.7356 | 0.73891 | 0.7421 | 0.7453 | 0.74857 | 0.751 | 0.7549 |
| 0.7 | 0.75804 | 0.76115 | 0.76424 | 0.7673 | 0.77035 | 0.77337 | 0.77637 | 0.77935 | 0.7823 | 0.78524 |
| 0.8 | 0.78814 | 0.79103 | 0.79389 | 0.79673 | 0.79955 | 0.80234 | 0.80511 | 0.80785 | 0.81057 | 0.81327 |
| 0.9 | 0.81594 | 0.81859 | 0.82121 | 0.82381 | 0.82639 | 0.82894 | 0.83147 | 0.83398 | 0.83646 | 0.83891 |
| 1.0 | 0.84 | 0.84375 | , 84614 | 0.8484 | 0.85083 | 0.8531 | 0.85543 | 0.8576 | 0.859 | . 86214 |
| 1.1 | 0.86433 | 0.8665 | 0.86864 | 0.87076 | 0.87286 | 0.87493 | 0.87698 | 0.879 | 0.881 | 0.88298 |
| 1.2 | 0.88493 | 0.88686 | 0.88877 | 0.89065 | 0.89251 | 0.89435 | 0.89617 | 0.89796 | 0.89973 | 0.90147 |
| 1.3 | 0.9032 | 0.9049 | 0.90658 | 0.90824 | 0.90988 | 0.91149 | 0.91309 | 0.91466 | 0.91621 | 0.91774 |
| 1.4 | 0.91924 | 0.92073 | 0.9222 | 0.92364 | 0.92507 | 0.92647 | 0.92785 | 0.92922 | 0.93056 | 0.93189 |
| 1.5 | 0.93319 | 0.93448 | 0.93574 | 0.93699 | 0.93822 | 0.93943 | 0.94062 | 0.94179 | 0.94295 | 0.94408 |
| 1.6 | 0.9452 | 0.9463 | 0.94738 | 0.94 | 0.9495 | 0.95053 | 0.95154 | 0.95254 | 0.95352 | 0.95449 |
| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608 | 0.96164 | 0.96246 | 0.96327 |
| 1.8 | 0.96407 | 0.96485 | 0.96562 | 0.96638 | 0.96712 | 0.96784 | 0.9685 | 0.96926 | 0.96995 | 0.97062 |
| 1.9 | 0.97128 | 0.97193 | 0.97257 | 0.9732 | 0.97381 | 0.97441 | 0.975 | 0.97558 | 0.97615 | 0.9767 |
| 2.0 | 0.97725 | 0.97778 | 0.97831 | 0.97882 | 0.97932 | 0.97982 | 0.9803 | 0.98077 | 0.98124 | 0.98169 |
| 2.1 | 0.98214 | 0.98257 | 0.983 | 0.98341 | 0.98382 | 0.98422 | 0.98461 | 0.985 | 0.98537 | 0.98574 |
| 2.2 | 0.9861 | 0.98645 | 0.98679 | 0.98713 | 0.98745 | 0.98778 | 0.98809 | 0.9884 | 0.9887 | 0.98899 |
| 2.3 | 0.98928 | 0.98956 | 0.98983 | 0.9901 | 0.99036 | 0.99061 | 0.99086 | 0.99111 | 0.99134 | 0.99158 |
| 2.4 | 0.9918 | 0.99202 | 0.99224 | 0.99245 | 0.99266 | 0.99286 | 0.99305 | 0.99324 | 0.99343 | 0.99361 |
| 2.5 | 0.99379 | 0.9939 | 0.99413 | 0.9943 | 0.99446 | 0.99461 | 0.99477 | 0.99492 | 0.99506 | 0.9952 |
| 2.6 | 0.99534 | 0.99547 | 0.9956 | 0.9 | 0.99585 | 0.99598 | 0.99609 | 0.99621 | 0.99632 | 0.99643 |
| 2.7 | 0.99653 | 0.99664 | 0.99674 | 0.99683 | 0.99693 | 0.99702 | 0.99711 | 0.9972 | 0.99728 | 0.99736 |
| 2.8 | 0.99744 | 0.99752 | 0.9976 | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.09865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999 |

Sum of independent normal is still normal

Fact. If $X \sim \mathcal{N}\left(\mu_{X}, \sigma_{X}^{2}\right), \mathrm{Y} \sim \mathcal{N}\left(\mu_{Y}, \sigma_{Y}^{2}\right)$ (both independent normal RV) then $\mathrm{a} X+b Y+c \sim \mathcal{N}\left(a \mu_{X}+b \mu_{Y}+c, a^{2} \sigma_{X}^{2}+b^{2} \sigma_{Y}^{2}\right)$


## Agenda

- Central Limit Theorem (CLT)
- Polling



## Gaussian in Nature

Empirical distribution of collected data often resembles a Gaussian ...

e.g. Height distribution resembles Gaussian.
R.A.Fisher (1918) observed that the height is likely the outcome of the sum of many independent random parameters, i.e., can written as

$$
X=X_{1}+\cdots+X_{n}
$$

## Sum of Independent RVs

i.i.d. = independent and identically distributed
$X_{1}, \ldots, X$ Xi.i.d. with expectation $\mu$ and variance $\sigma^{2}$
Define

$$
S_{n}=X_{2}+\cdots+X_{n}
$$

$\mathbb{E}\left[S_{n}\right]=\mathbb{E}\left[X_{1}\right]+\cdots+\mathbb{E}\left[X_{n}\right]=\underline{n \mu}$
$\operatorname{Var}\left(S_{n}\right)=\operatorname{Var}\left(X_{1}\right)+\cdots+\operatorname{Var}\left(X_{n}\right)=n \sigma^{2} \rightarrow \quad \underline{\sqrt{n} \cdot \sigma}$

Empirical observation: $S_{n}$ looks like a normal RV as $n$ grows.


## Example: Sum of $n$ i.i.d. Exp(1) random variables


(a) $n=1$

(e) $n=12$
-

(b) $n=2$

(f) $n=25$



## CLT (Idea)



## CLT (Idea)



## Central Limit Theorem

$X_{1}, \ldots, X_{n}$ i.i.d., each with expectation $\mu$ and variance $\sigma^{2}$
$\longrightarrow N\left(n \mu, n \sigma^{2}\right)$
Define $S_{n}=X_{1}+\cdots+X_{n}$ and

$$
Y_{n}=\frac{S_{n}-n \mu}{(\sigma \sqrt{n})} \rightarrow N(0,1)
$$

$\mathbb{E}\left[Y_{n}\right]=\frac{1}{\sigma \sqrt{n}}\left(\mathbb{E}\left[S_{n}\right]-n \mu\right)=\frac{1}{\sigma \sqrt{n}}(n \mu-n \mu)=0$
$\operatorname{Var}\left(Y_{n}\right)=\frac{1}{\sigma^{2} n}\left(\operatorname{Var}\left(S_{n}-n \mu\right)\right)=\frac{\operatorname{Var}\left(S_{n}\right)}{\sigma^{2} n}=\frac{\sigma^{2} n}{\sigma^{2} n}=1$

## Central Limit Theorem

$X_{1}, \ldots, X_{n}$ i.i.d., each with expectation $\mu$ and variance $\sigma^{2}$

Define $S_{n}=X_{1}+\cdots+X_{n}$ and

$$
z_{n}=\left(\frac{S_{n}}{n}\right)
$$

$\mathbb{E}\left[Z_{n}\right]=\underline{\underline{n \mu}}=\mu$
$\operatorname{Var}\left(Z_{n}\right)=\frac{\operatorname{Var}\left(S_{n}\right)}{\mathrm{n}^{2}}=\frac{\sigma^{2} n}{\underline{n^{2}}}=\frac{\sigma^{2}}{n} \downarrow$

Central Limit Theorem

$$
N(0,1) \propto Y_{n}=\frac{X_{1}+\cdots+X_{n}-n \mu}{\sigma \sqrt{n}}
$$

Theorem. (Central Limit Theorem) The CDF of $Y_{n}$ converges to the CDF of the standard normal $\mathcal{N}(0,1)$, i.e.,

$$
\lim _{n \rightarrow \infty} P\left(Y_{n} \leq y\right)=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{y} e^{-x^{2} / 2} \mathrm{~d} x=\Phi(y)
$$

## Central Limit Theorem

$$
Y_{n}=\frac{X_{1}+\cdots+X_{n}-n \mu}{\sigma \sqrt{n}}
$$

Theorem. (Central Limit Theorem) The CDF of $Y_{n}$ converges to the CDF of the standard normal $\mathcal{N}(0,1)$, i.e.,

$$
\left.\lim _{n \rightarrow \infty} P\left(Y_{n}\right) \leq y\right)=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{y} e^{-x^{2} / 2} \mathrm{~d} x
$$

Also stated as:

- $\lim _{n \rightarrow \infty} Y_{n} \rightarrow \rightarrow X_{i} \sim N\left(\mu, \sigma^{2}\right)$
- $\underline{\lim }_{n \rightarrow \infty} \frac{\left(\frac{1}{n} \sum_{i=1}^{n} X_{i}\right.}{z_{n}} \rightarrow \mathcal{N}\left(\mu_{,} \frac{\sigma^{2}}{n}\right)$ for $\mu=\mathbb{E}\left[X_{i}\right]$ and $\sigma^{2}=\operatorname{Var}\left(X_{i}\right)$


## CLT $\rightarrow$ Normal Distribution EVERYWHERE



S\&P 500 Returns after Elections



## Examples from:

https://galtonboard.com/probabilityexamplesinlife

## Agenda

- Central Limit Theorem (CLT) Review
- Polling


## Magic Mushrooms

In Fall 2020, Oregonians voted on whether to legalize the therapeutic use of "magic mushirooms".

Poll to determine the fraction $p$ of the population expected to vote in favor.

- Call up a random sample of $n$ people to ask their opinion $P=E\left[\bar{P}_{1}\right]$
- Report the empirical fraction


1. Is this agood estimate?
traek 40 w to choose $n$ ?

- in populetió


## Polling Accuracy

## Often see claims that say

$$
\begin{aligned}
& \text { "Our poll found } 80 \% \text { support. This poll is accurate to within } \\
& 5 \% \text { with } 98 \% \text { probability** } \\
& \bar{p} \in[p[\bar{P} \in[p 5 \%, \quad p+5 \%]] \geqslant 98 \%
\end{aligned}
$$

Will unpack what this and how they sample enough people to know this is true.

* When it is $95 \%$ this is sometimes written as " 19 times out of 20 "


## Formalizing Polls

Population size $N$, true fraction of voting in favor $p$, sample size $n$.
Problem: We don't know $p$, want to estimate it

## Polling Procedure

for $i=1, \ldots, n$ :

1. Pick uniformly random person to call (prob: $1 / N$ )
2. Ask them how they will vote

$$
X_{i}=\left\{\begin{array}{lr}
\underline{1}, & \text { voting in favor } \\
0, & \text { otherwise }
\end{array}\right.
$$

Report our estimate of $p$ :

$$
\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}
$$

## Formalizing Polls

Population size $N$, true fraction of voting in favor $p$, sample size $n$.

Problem: We don't know $p$
Polling Procedure
for $i=1, \ldots, n$ :

What type of rv. is $X_{i}$ ?
Poll: pollev.com/rachel312

|  | Type | $\mathbb{E}\left[X_{i}\right]$ | $\operatorname{Var}\left(X_{i}\right)$ |
| :---: | :---: | :---: | :--- |
| a. Bernoulli | $p$ | $p(1-p)$ |  |
| b. | Bernoulli | $p$ | $p^{2}$ |
| c. Geometric | $p$ | $\frac{1-p}{p^{2}}$ |  |
| d. | Binomial | $n p$ | $n p(1-p)$ |

1. Pick uniformly random person to call (prob: $1 / N$ )
2. Ask them how they will vote

$$
X_{i}=\left\{\begin{array}{lr}
1 \\
0
\end{array}\right) \quad \begin{array}{r}
\text { voting in favor } \\
\text { otherwise }
\end{array}
$$

Report our estimate of $p$ :
$X=\frac{1}{n} \sum_{i=1}^{n} X_{i}$

## Random Variables

What type of r.v. is $X_{i}$ ?

|  | Type | $\mathbb{E}\left[X_{i}\right]$ | $\operatorname{Var}\left(X_{i}\right)$ |
| :---: | :---: | :---: | :---: |
| a. | Bernoulli | $p$ | $p(1-p)$ |
| b. | Bernoulli | $p$ | $p^{2}$ |
| C. | Geometric | $p$ | $\frac{1-p}{p^{2}}$ |
| d. | Binomial | $\mathrm{n} p$ | $n p(1-p)$ |

What about $\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}$ ?

|  | Poll: pollev.com/rachel312 |  |
| :--- | :--- | :--- |
|  | $\mathbb{E}[\bar{X}]$ | $\operatorname{Var}(\bar{X})$ |
| a. | $n p$ | $n p(1-p)$ |
| b. | $p$ | $p(1-p)$ |
| c. | $p$ | $p(1-p) / n$ |
| d. | $p / n$ | $p(1-p) / n$ |

## Roadmap: Bounding Error

Goal: Find the value of $n$ such that $98 \%$ of the time, the estimate $\bar{X}$ is within $5 \%$ of the true $p$

Get good estimate if $\bar{X}$ lands in this region


Want $P(|\bar{X}-p|>0.05) \leq 0.02$

## Central Limit Theorem

pollev.com/rachel312
Poll: In the limit $\bar{X}$ is...?
a. $\quad \mathcal{N}(0,1)$
b. $\mathcal{N}(p, p(1-p))$
c. $\quad \mathcal{N}(p, p(1-p) / n)$
d. I don't know

As $n \rightarrow \infty$,

$$
\frac{X_{1}+X_{2}+\cdots X_{n}-n \mu}{\sigma \sqrt{n}} \rightarrow \mathcal{N}(0,1)
$$

As $n \rightarrow \infty$,

$$
\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i} \rightarrow \mathcal{N}\left(\mu, \frac{\sigma^{2}}{n}\right)
$$

## Roadmap: Bounding Error



Want $P(|\bar{X}-p|>0.05) \leq 0.02$

## Roadmap: Bounding Error

Goal: Find the value of $n$ such that $98 \%$ of the time, the estimate $\bar{X}$ is within $5 \%$ of the true $p$

1. Define probability of a "bad event" $P(|\bar{X}-p|>0.05) \leq 0.02$
2. Apply CLT
3. Convert to a standard normal
4. Solve for $n$

## Following the Road Map

1. Want $P(|\bar{X}-p|>0.05) \leq 0.02$

2. By CLT $\bar{X} \rightarrow \mathcal{N}\left(\mu, \sigma^{2}\right)$ where $\mu=p$ and $\sigma^{2}=p(1-p) / n$
3. Define $Z=\frac{\bar{X}-\mu}{\sigma}=\frac{\bar{X}-p}{\sigma}$. Then, by the $\operatorname{CLT} Z \rightarrow \mathcal{N}(0,1)$

$$
P(|\bar{X}-p|>0.05)=P(|Z| \cdot \sigma>0.05)
$$

$$
\frac{1}{\sqrt{p(1-p)}} \text { is always } \geq 2
$$

$$
\begin{aligned}
& =P(|Z|>0.05 / \sigma)=P(|Z|>0.05 \\
& \leq P(|Z|>0.1 \sqrt{n})
\end{aligned}
$$

## Following the Road Map

1. Want $P(|\bar{X}-p|>0.05) \leq 0.02$

2. By CLT $\bar{X} \rightarrow \mathcal{N}\left(\mu, \sigma^{2}\right)$ where $\mu=p$ and $\sigma^{2}=p(1-p) / n$
3. Define $Z=\frac{\bar{X}-\mu}{\sigma}=\frac{\bar{X}-p}{\sigma}$. Then, by the CLT $Z \rightarrow \mathcal{N}(0,1)$
$P(|\bar{X}-p|>0.05)=P(|Z| \cdot \sigma>0.05)$ $\frac{1}{\sqrt{p(1-p)}}$ is always $\geq 2$

$$
\left.\frac{-P(|7|>005 / \sigma)-P(|7|}{\frac{-0 \text { ose } n \text { so that this is at most } 0.02}{\leq P(|Z|>0.1 \sqrt{n})}}>0.05 \xrightarrow{\sqrt{p(1-p)}}\right)
$$

## 4. Solve for $n$

We want $P(|Z|>0.1 \sqrt{n}) \leq 0.02$ where $Z \rightarrow \mathcal{N}(0,1)$

- If we actually had $Z \sim \mathcal{N}(0,1)$ then enough to show that $P(Z>0.1 \sqrt{n}) \leq 0.01$ since $\mathcal{N}(0,1)$ is symmetric about 0
- Now $P(Z>z)=1-\Phi(z)$ where $\Phi(z)$ is the CDF of the Standard Normal Distribution
- So, want to choose $n$ so that $0.1 \sqrt{n} \geq z$ where $\Phi(z) \geq 0.99$


## Table of $\Phi(\mathbf{z})$ CDF of

 Standard Normal DistributionChoose $n$ so
$0.1 \sqrt{n} \geq z$ where $\Phi(z) \geq 0.99$

From table $z=2.33$ works

$\Phi$ Table: $\mathbb{P}(Z \leq z)$ when $Z \sim \mathcal{N}(0,1)$

| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.0 | 0.5 | 0.50399 | 0.50798 | 0.51197 | 0.51595 | 0.51994 | 0.52392 | 0.5279 | 0.53188 | 0.53586 |
| 0.1 | 0.53983 | 0.5438 | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 0.57535 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
| 0.3 | 0.61791 | 0.62172 | 0.62552 | 0.6293 | 0.63307 | 0.63683 | 0.64058 | 0.64431 | 0.64803 | 0.65173 |
| 0.4 | 0.65542 | 0.6591 | 0.66276 | 0.6664 | 0.67003 | 0.67364 | 0.67724 | 0.68082 | 0.68439 | 0.68793 |
| 0.5 | 0.69146 | 0.69497 | 0.69847 | 0.70194 | 0.7054 | 0.70884 | 0.71226 | 0.71566 | 0.71904 | 0.7224 |
| 0.6 | 0.72575 | 0.72907 | 0.73237 | 0.73565 | 0.73891 | 0.74215 | 0.74537 | 0.74857 | 0.75175 | 0.7549 |
| 0.7 | 0.75804 | 0.76115 | 0.76424 | 0.7673 | 0.77035 | 0.77337 | 0.77637 | 0.77935 | 0.7823 | 0.78524 |
| 0.8 | 0.78814 | 0.79103 | 0.79389 | 0.79673 | 0.79955 | 0.80234 | 0.80511 | 0.80785 | 0.81057 | 0.81327 |
| 0.9 | 0.81594 | 0.81859 | 0.82121 | 0.82381 | 0.82639 | 0.82894 | 0.83147 | 0.83398 | 0.83646 | 0.83891 |
| 1.0 | 0.84134 | 0.84375 | 0.84614 | 0.84849 | 0.85083 | 0.85314 | 0.85543 | 0.85769 | 0.85993 | 0.86214 |
| 1.1 | 0.86433 | 0.8665 | 0.86864 | 0.87076 | 0.87286 | 0.87493 | 0.87698 | 0.879 | 0.881 | 0.88298 |
| 1.2 | 0.88493 | 0.88686 | 0.88877 | 0.89065 | 0.89251 | 0.89435 | 0.89617 | 0.89796 | 0.89973 | 0.90147 |
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| 1.4 | 0.91924 | 0.92073 | 0.9222 | 0.92364 | 0.92507 | 0.92647 | 0.92785 | 0.92922 | 0.93056 | 0.93189 |
| 1.5 | 0.93319 | 0.93448 | 0.93574 | 0.93699 | 0.93822 | 0.93943 | 0.94062 | 0.94179 | 0.94295 | 0.94408 |
| 1.6 | 0.9452 | 0.9463 | 0.94738 | 0.94845 | 0.9495 | 0.95053 | 0.95154 | 0.95254 | 0.95352 | 0.95449 |
| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608 | 0.96164 | 0.96246 | 0.96327 |
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| 1.9 | 0.97128 | 0.97193 | 0.97257 | 0.9732 | 0.97381 | 0.97441 | 0.975 | 0.97558 | 0.97615 | 0.9767 |
| 2.0 | 0.97725 | 0.97778 | 0.97831 | 0.97882 | 0.97932 | 0.97982 | 0.9803 | 0.9807 | 0.98124 | 0.98169 |
| 2.1 | 0.98214 | 0.98257 | 0.983 | 0.98341 | 0.98382 | 0.98422 | 0.98461 | 0.985 | 0.98537 | 0.98574 |
| 2.2 | 0.9861 | 0.98645 | 0.9867 | 9.981 | 0.9874 | 0.98778 | 0.98809 | 0.9884 | 0.9887 | 0.98899 |
| 2.3 | 0.98928 | 0.98956 | 0.9 | 0.9901 | 0.99036 | 0.99061 | 0.99086 | 0.99111 | 0.99134 | 0.99158 |
| 2.4 | 0.9918 | 0.99202 | 0.99224 | 0.900 | 0.99266 | 0.99286 | 0.99305 | 0.99324 | 0.99343 | 0.99361 |
| 2.5 | 0.99379 | 0.99396 | 0.99413 | 0.9943 | 0.99446 | 0.99461 | 0.99477 | 0.99492 | 0.99506 | 0.9952 |
| 2.6 | 0.99534 | 0.99547 | 0.9956 | 0.99573 | 0.9958 | 0.9959 | 0.99609 | 0.99621 | 0.99632 | 0.99643 |
| 2.7 | 0.99653 | 0.99664 | 0.99674 | 0.99683 | 0.99693 | 0.99702 | 0.99711 | 0.9972 | 0.99728 | 0.99736 |
| 2.8 | 0.99744 | 0.99752 | 0.9976 | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999 |

## 4. Solve for $n$

Choose $n$ so
$0.1 \sqrt{n} \geq z$ where $\Phi(z) \geq 0.99$

From table $z=2.33$ works

- So we can choose $0.1 \sqrt{n} \geq 2.33$ or $\sqrt{n} \geq 23.3$
- Then $n \geq 543 \geq(23.3)^{2}$ would be good enough ... if we had $Z \sim \mathcal{N}(0,1)$
- We only have $Z \rightarrow \mathcal{N}(0,1)$ so there is some loss due to approximation error.
- Maybe instead consider $z=3.0$ with $\Phi(z) \geq 0.99865$ and $n \geq 30^{2}=900$ to cover any loss.


## Idealized Polling

So far, we have been discussing "idealized polling". Real life is normally not so nice :

Assumed we can sample people uniformly at random, not really possible in practice

- Not everyone responds
- Response rates might differ in different groups
- Will people respond truthfully?

Makes polling in real life much more complex than this idealized mode!!

