

CSE 312

# Foundations of Computing II

Lecture 8: More on random variables; expectation

Anonymous Questions: [www.slido.com/1891306](http://www.slido.com/1891306)

## Last Class:

- Random Variables
- Probability Mass Function (PMF)
- Cumulative Distribution Function (CDF)

## Today:

- Recap
- Expectation
- Linearity of Expectation
- Indicator Random Variables

**Kandinsky**

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## Review Random Variables

**Definition.** A **random variable (RV)** for a probability space  $(\Omega, P)$  is a function  $X: \Omega \rightarrow \mathbb{R}$ .

The set of values that  $X$  can take on is its *range/support*:  $X(\Omega)$  or  $\Omega_X$

## Example: Returning Homeworks

$s_1, s_2, s_3$

- Class with 3 students, randomly hand back homeworks. All permutations equally likely.
- Let  $X$  be the number of students who get their own HW

$s_1, s_2, s_3$

$\Pr(\omega)$	$\omega$	$X(\omega)$
1/6	(1, 2, 3)	3
1/6	(1, 3, 2)	1
1/6	(2, 1, 3)	1
1/6	(2, 3, 1)	0
1/6	(3, 1, 2)	0
1/6	(3, 2, 1)	1

$$\Omega_X = \{0, 1, 3\}$$



## Example: Returning Homeworks

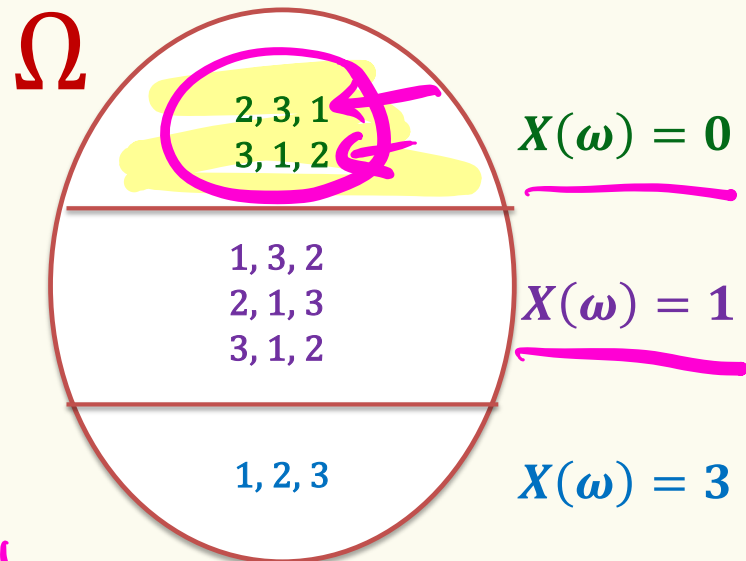
- Class with 3 students, randomly hand back homeworks. All permutations equally likely.
- Let  $X$  be the number of students who get their own HW

$\Pr(\omega)$	$\omega$	$X(\omega)$
1/6	1, 2, 3	3
1/6	1, 3, 2	1
1/6	2, 1, 3	1
1/6	2, 3, 1	0
1/6	3, 1, 2	0
1/6	3, 2, 1	1

$$X(\omega) = i$$

$$= \{\omega \mid X(\omega) = i\}$$

$$P(X(\omega) = 0) = \frac{1}{3}$$



## Review Random Variables

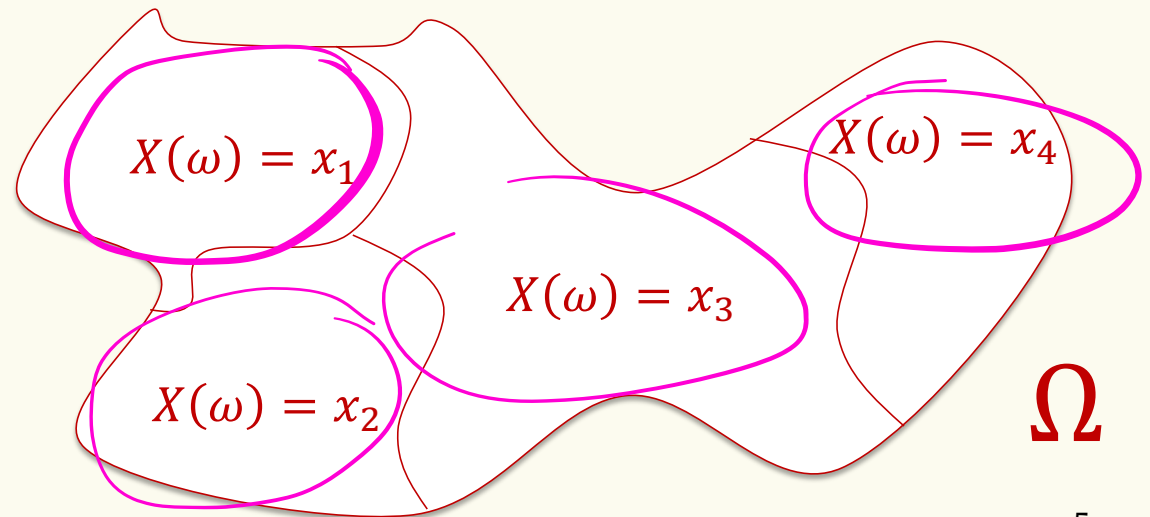
**Definition.** A **random variable (RV)** for a probability space  $(\Omega, P)$  is a function  $X: \Omega \rightarrow \mathbb{R}$ .

The set of values that  $X$  can take on is its *range/support*:  $X(\Omega)$  or  $\Omega_X$

$$\{X = x_i\} = \{\omega \in \Omega \mid X(\omega) = x_i\}$$

Random variables **partition** the sample space.

$$\sum_{x \in X(\Omega)} P(X = x) = 1$$



$$\Omega_X = \{x_1, x_2, x_3, x_4\}$$

## Review PMF and CDF

### Definitions:

For a RV  $X: \Omega \rightarrow \mathbb{R}$ , the **probability mass function (pmf)** of  $X$  specifies, for any real number  $x$ , the probability that  $X = x$

$$p_X(x) = P(X = x) = P(\{\omega \in \Omega \mid X(\omega) = x\})$$

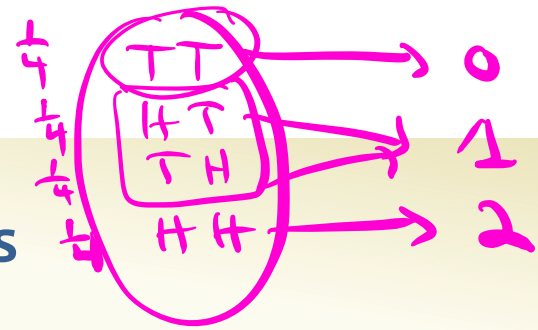
$$\sum_{x \in \Omega_X} p_X(x) = 1$$

For a RV  $X: \Omega \rightarrow \mathbb{R}$ , the **cumulative distribution function (cdf)** of  $X$  specifies, for any real number  $x$ , the probability that  $X \leq x$

$$F_X(x) = P(X \leq x)$$

r.v.

## Example – Two fair independent coin flips

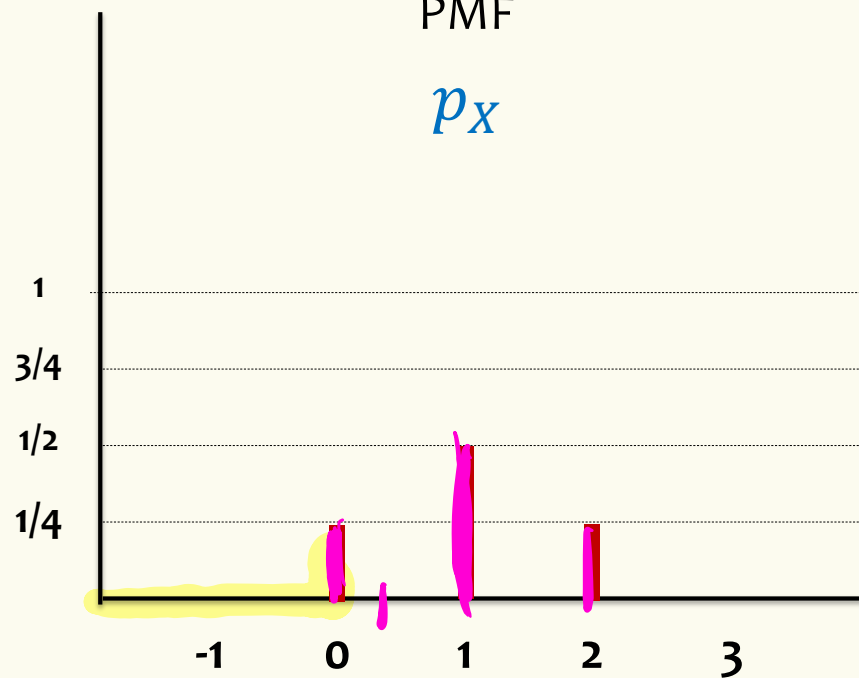


$X =$  number of heads

Probability Mass Function

PMF

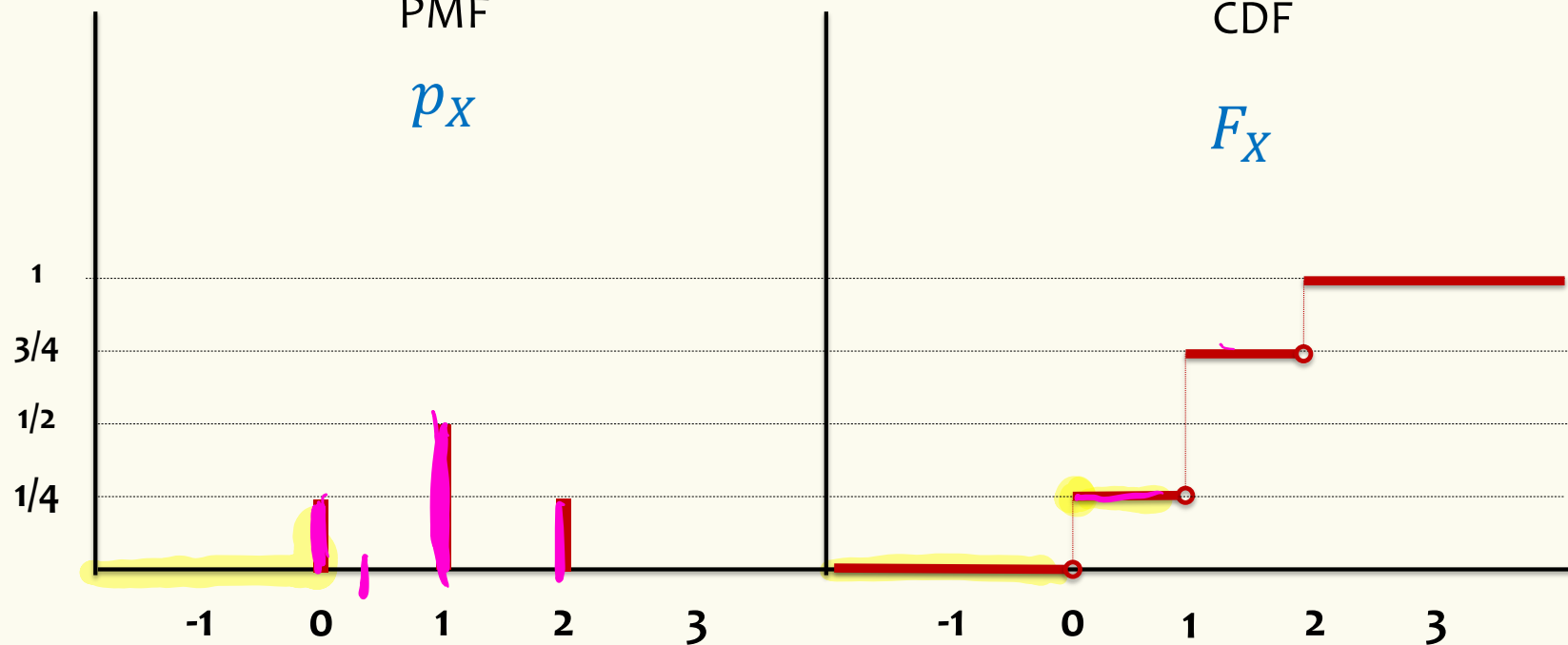
$p_X$



Cumulative Distribution Function

CDF

$F_X$





## Example: Returning Homeworks

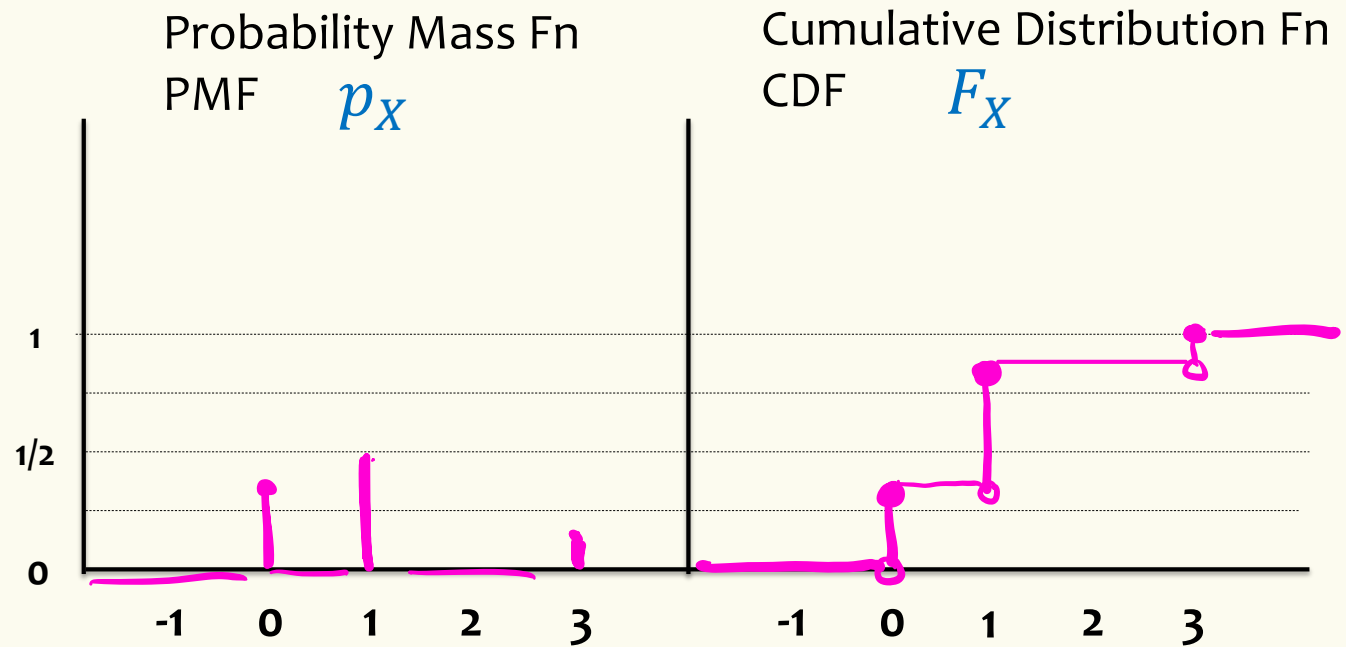
- Class with 3 students, randomly hand back homeworks. All permutations equally likely.
- Let  $X$  be the number of students who get their own HW

$\Pr(\omega)$	$\omega$	$X(\omega)$
1/6	1, 2, 3	3
1/6	1, 3, 2	1
1/6	2, 1, 3	1
1/6	2, 3, 1	0
1/6	3, 1, 2	0
1/6	3, 2, 1	1

$$P(X = 0) = 1/3$$

$$P(X = 1) = 1/2$$

$$P(X = 3) = 1/6$$



## Example – Number of Heads

We flip  $n$  coins, independently, each heads with probability  $p$

$$\Omega = \{HH \cdots HH, HH \cdots HT, HH \cdots TH, \dots, TT \cdots TT\}$$

$$|\Omega| = 2^n$$

$X = \#$  of heads

$$\Omega_X = \{0, 1, 2, \dots, n\}$$

$$p_X(k) = P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$$

$$P(\underbrace{HH \cdots H}_k \underbrace{T \cdots T}_{n-k}) = p^k (1-p)^{n-k}$$

$$P(\underbrace{T \cdots T}_{n-k-1} \underbrace{H \cdots H}_k \underbrace{T}_{1}) = p^k (1-p)^{n-k}$$

## Example – Number of Heads

We flip  $n$  coins, independently, each heads with probability  $p$

$$\Omega = \{\text{HH} \cdots \text{HH}, \text{HH} \cdots \text{HT}, \text{HH} \cdots \text{TH}, \dots, \text{TT} \cdots \text{TT}\}$$

$X = \#$  of heads

$$p_X(k) = P(X = k) = \binom{n}{k} \cdot p^k \cdot (1 - p)^{n-k}$$

# of sequences with  $k$  heads

Prob of sequence w/  $k$  heads

# Agenda

- Random Variables
- Probability Mass Function (PMF)
- Cumulative Distribution Function (CDF)
- Expectation ◀

## Expectation (Idea)

$$\Omega_X = \{0, 1, \dots, 20\}$$

**Example.** Toss a coin 20 times independently with probability  $\frac{1}{4}$  of coming up heads on each toss.

$$\underline{20} \cdot \frac{1}{4} =$$

$X$  = number of heads

How many heads do you *expect* to see?

5

What if you toss it independently  $n$  times and it comes up heads with probability  $p$  each time?

$np.$

## Review Expected Value of a Random Variable

**Definition.** Given a discrete RV  $X: \Omega \rightarrow \mathbb{R}$ , the **expectation** or **expected value** or **mean** of  $X$  is

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot P(\omega)$$

or equivalently

$$\mathbb{E}[X] = \sum_{x \in X(\Omega)} x \cdot P(X = x) = \sum_{x \in \Omega_X} \underbrace{x}_{\text{outcome}} \underbrace{p_X(x)}_{\text{probability}}$$

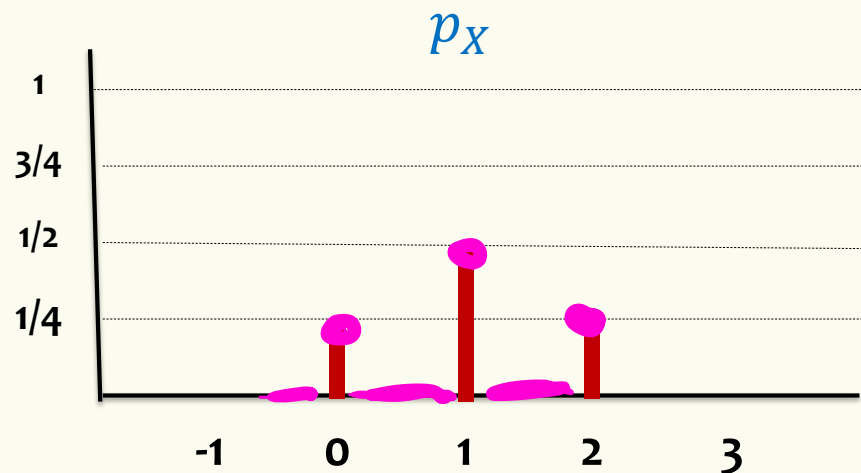
Intuition: “Weighted average” of the possible outcomes (weighted by probability)

# Expectation

np

**Example.** Two fair coin flips  
 $\Omega = \{TT, HT, TH, HH\}$   
 $X =$  number of heads

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot P(\omega)$$
$$\Rightarrow \mathbb{E}[X] = \sum_{x \in X(\Omega)} x \cdot P(X = x)$$



What is  $\mathbb{E}[X]$ ?

$$\mathbb{E}[X] = 0 \cdot p_X(0) + 1 \cdot p_X(1) + 2 \cdot p_X(2)$$
$$= 0 \cdot \frac{1}{4} + 1 \cdot \frac{1}{2} + 2 \cdot \frac{1}{4} = \frac{1}{2} + \frac{1}{2} = 1$$

0 17

## Example: Returning Homeworks

$$\Rightarrow \mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot P(\omega)$$

$$\Rightarrow \mathbb{E}[X] = \sum_{x \in X(\Omega)} x \cdot P(X = x)$$

- Class with 3 students, randomly hand back homeworks. All permutations equally likely.
- Let  $X$  be the number of students who get their own HW
- What is  $\mathbb{E}[X]$ ?

$\Pr(\omega)$	$\omega$	$X(\omega)$
1/6	2, 3, 1	0
1/6	3, 1, 2	0
1/6	1, 3, 2	1
1/6	3, 2, 1	1
1/6	2, 1, 3	1
1/6	1, 2, 3	3

$$\mathbb{E}(X) = 0 \cdot \frac{1}{6} + 0 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6}$$

$$= 0 \cdot P(X=0) + 1 \cdot P(X=1) + 3 \cdot P(X=3)$$





## Example: Returning Homeworks

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot P(\omega)$$

$$\mathbb{E}[X] = \sum_{x \in X(\Omega)} x \cdot P(X = x)$$

- Class with 3 students, randomly hand back homeworks. All permutations equally likely.
- Let  $X$  be the number of students who get their own HW

$\Pr(\omega)$	$\omega$	$X(\omega)$
1/6	1, 2, 3	3
1/6	1, 3, 2	1
1/6	2, 1, 3	1
1/6	2, 3, 1	0
1/6	3, 1, 2	0
1/6	3, 2, 1	1

$$\begin{aligned}\mathbb{E}[X] &= 3 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} + 0 \cdot \frac{1}{6} + 0 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} \\ &= 6 \cdot \frac{1}{6} = 1\end{aligned}$$

$$E(Z) = \sum_{x \in \mathcal{Z}} x \cdot P(Z=x)$$

$$\mathcal{Z}_x = \{1, 2, 3, \dots\}$$

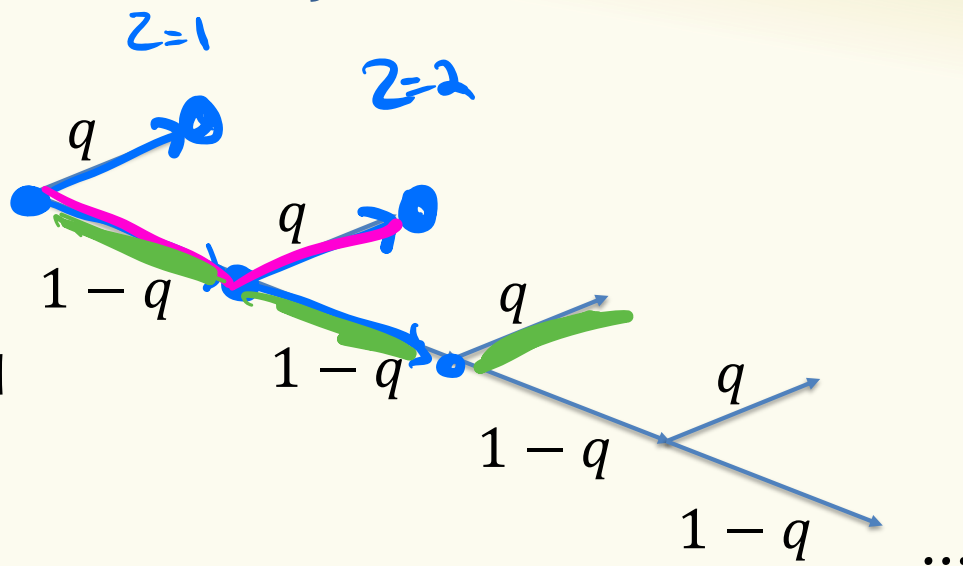
### Example – Flipping a biased coin until you see heads

- Biased coin, each flip indep:

$$P(H) = q > 0$$

$$P(T) = 1 - q$$

- $Z = \#$  of coin flips until first head



pmf

$$P(Z = i) = (1-q)^{i-1} q$$

$$P(Z=1) = q$$

$$P(Z=2) = (1-q)q$$

$$P(Z=3) = (1-q)^2 q$$

$$E[Z] = \sum_{i=1}^{\infty} i P(Z=i) = \sum_{i=1}^{\infty} i (1-q)^{i-1} q = \frac{1}{q}$$

## Example – Flipping a biased coin until you see heads

- Biased coin, each flip indep:

$$P(H) = q > 0$$

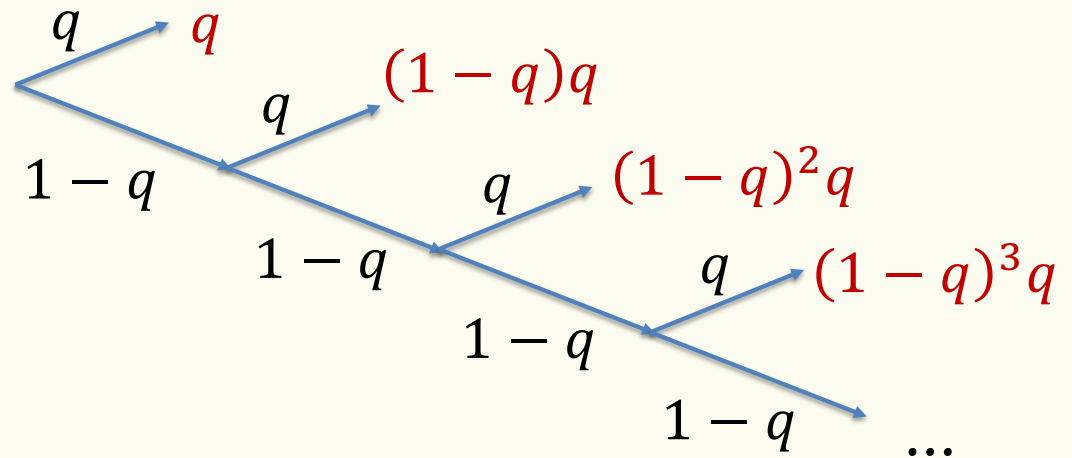
$$P(T) = 1 - q$$

- $Z = \#$  of coin flips until first head

$$P(Z = i) = q (1 - q)^{i-1}$$

$$\mathbb{E}[Z] = \sum_{i=1}^{\infty} i \cdot P(Z = i) = \sum_{i=1}^{\infty} i \cdot q(1 - q)^{i-1}$$

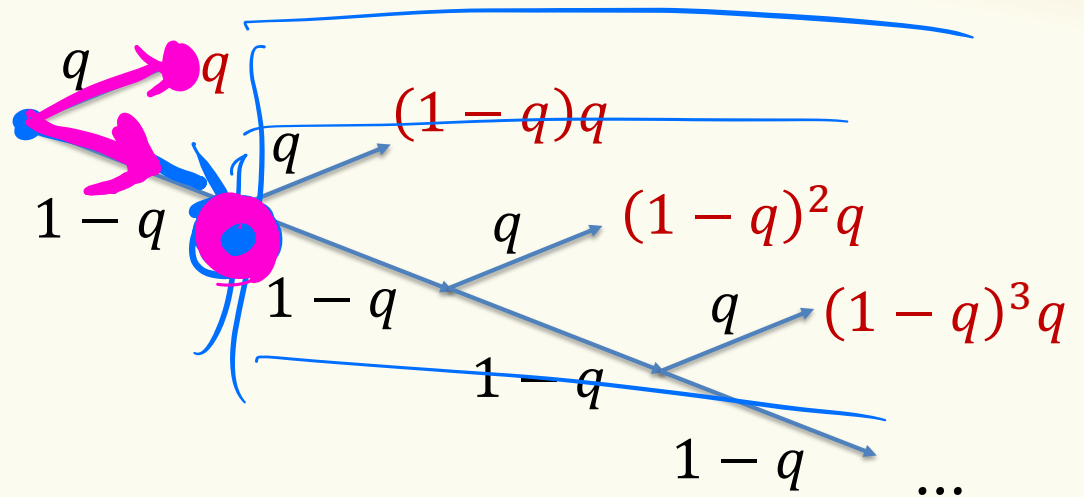
Converges, so  $\mathbb{E}[Z]$  is finite



Can calculate this directly but...

## Example – Flipping a biased coin until you see heads

- Biased coin, each flip indep:
  - $P(H) = q > 0$
  - $P(T) = 1 - q$
- $Z = \#$  of coin flips until first head



**Another view:** If you get heads first try you get  $Z = 1$ ;  
 If you get tails you have used one try and have the same experiment left

$$\begin{aligned} \mathbb{E}[Z] &= q \cdot 1 + (1 - q) [1 + \mathbb{E}(Z)] \\ &= \underbrace{q + (1 - q)}_1 + \underbrace{(1 - q) \mathbb{E}(Z)} \end{aligned}$$

$$q \mathbb{E}(Z) = 1 \quad \Rightarrow \quad \mathbb{E}(Z) = \frac{1}{q}$$

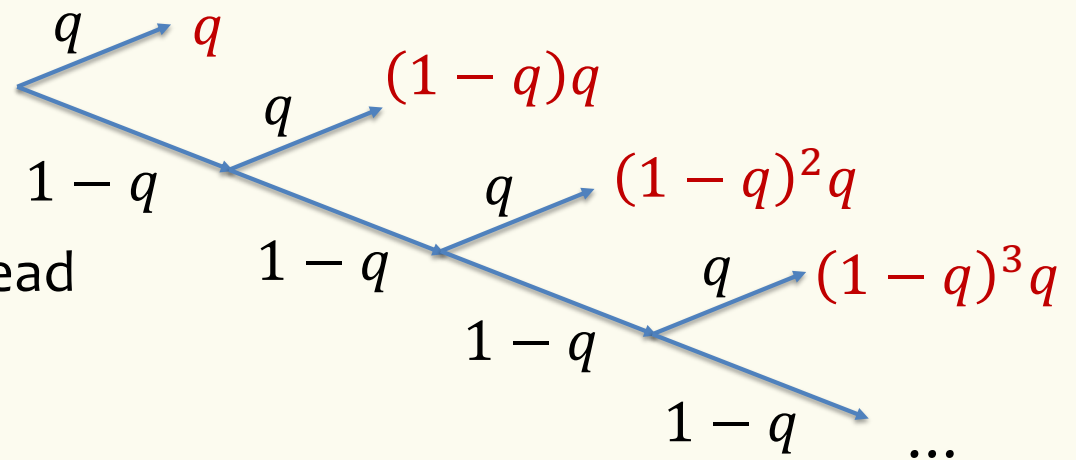
## Example – Flipping a biased coin until you see heads

- Biased coin:

$$P(H) = q > 0$$

$$P(T) = 1 - q$$

- $Z = \#$  of coin flips until first head



**Another view:** If you get heads first try you get  $Z = 1$ ;  
If you get tails you have used one try and have the same experiment left

$$\mathbb{E}[Z] = q \cdot 1 + (1 - q)(1 + \mathbb{E}(Z))$$

Solving gives  $q \cdot \mathbb{E}[Z] = q + (1 - q) = 1$  Implies  $\mathbb{E}[Z] = 1/q$

## Example – Coin Tosses

We flip  $n$  coins, each toss independent, probability  $p$  of coming up heads.

$Z$  is the number of heads, what is  $\mathbb{E}(Z)$ ?

$$P(Z=k) = \begin{cases} \binom{n}{k} p^k (1-p)^{n-k} & k \in \{0, 1, \dots, n\} \\ 0 & \text{otherwise.} \end{cases}$$

$$\mathbb{E}(Z) = \sum_{k=0}^n k P(Z=k)$$

$$\approx \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k}$$

## Example – Coin Tosses

We flip  $n$  coins, each toss independent; heads with probability  $p$ ,  $Z$  is the number of heads, what is  $\mathbb{E}[Z]$ ?

$$\begin{aligned}\mathbb{E}[Z] &= \sum_{k=0}^n k \cdot P(Z = k) = \sum_{k=0}^n k \cdot \binom{n}{k} p^k (1-p)^{n-k} \\ &= \sum_{k=0}^n k \cdot \frac{n!}{k! (n-k)!} p^k (1-p)^{n-k} = \sum_{k=1}^n \frac{n!}{(k-1)! (n-k)!} p^k (1-p)^{n-k} \\ &= np \sum_{k=1}^n \frac{(n-1)!}{(k-1)! (n-k)!} p^{k-1} (1-p)^{n-k} \\ &= np \sum_{k=0}^{n-1} \frac{(n-1)!}{k! (n-1-k)!} p^k (1-p)^{(n-1)-k} \\ &= np \sum_{k=0}^{n-1} \binom{n-1}{k} p^k (1-p)^{(n-1)-k} = np(p + (1-p))^{n-1} = np \cdot 1 = np\end{aligned}$$



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Can we solve it more elegantly, please?

## Agenda

- Random Variables
- Probability Mass Function (PMF)
- Cumulative Distribution Function (CDF)
- Expectation
- Linearity of Expectation ◀





$X(\omega)$   
 $Y(\omega)$

$$Z(\omega) = X(\omega) + Y(\omega)$$

## Linearity of Expectation

**Theorem.** For **any** two random variables  $X$  and  $Y$   
(no conditions whatsoever on the random variables)

(defined on  
common prob space)

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y].$$

Or, more generally: For any random variables  $X_1, \dots, X_n$ ,

$$\mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n].$$

**Because:**  $\mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[(X_1 + \dots + X_{n-1}) + X_n]$   
 $= \mathbb{E}[X_1 + \dots + X_{n-1}] + \mathbb{E}[X_n] = \dots$

## Linearity of Expectation – Proof



$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot P(\omega)$$
$$\mathbb{E}[X] = \sum_{x \in X(\Omega)} x \cdot P(X = x)$$

**Theorem.** For **any** two random variables  $X$  and  $Y$   
( $X, Y$  do not need to be independent)

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y].$$

$$\begin{aligned} \mathbb{E}[X + Y] &= \sum_{\omega} P(\omega) \overbrace{(X(\omega) + Y(\omega))}^{Z(\omega)} \\ &= \underbrace{\sum_{\omega} P(\omega) X(\omega)} + \underbrace{\sum_{\omega} P(\omega) Y(\omega)} \\ &= \underline{\mathbb{E}[X]} + \underline{\mathbb{E}[Y]} \end{aligned}$$

$$E(X)$$

## Using LOE to compute complicated expectations

Often boils down to the following three steps:

- Decompose: Finding the right way to decompose the random variable into sum of simple random variables

$$X = X_1 + \dots + X_n$$

- LOE: Apply linearity of expectation.

$$E[X] = E[X_1] + \dots + E[X_n].$$

- Conquer: Compute the expectation of each  $X_i$

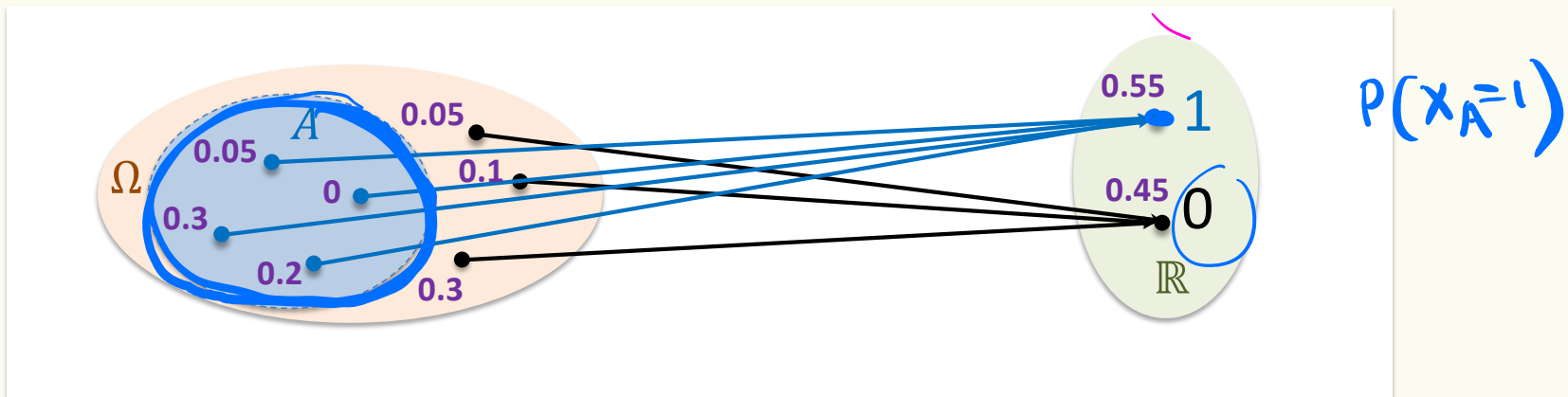
Often,  $X_i$  are **indicator** (0/1) random variables.

## Indicator random variables – 0/1 valued

For any event  $A$ , can define the **indicator** random variable  $X_A$  for  $A$

$$X_A = \begin{cases} 1 & \text{if event } A \text{ occurs} \\ 0 & \text{if event } A \text{ does not occur} \end{cases}$$

$$\begin{aligned} P(X_A = 1) &= P(A) \\ P(X_A = 0) &= 1 - P(A) \end{aligned}$$



$$\underline{E(X_A)} = \underline{0 \cdot P(X_A=0)} + \underline{1 \cdot P(X_A=1)}$$

$$= P A$$