## CSE 312 Foundations of Computing II

Lecture 10: LOTUS, variance and independence among random variables.

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#### Agenda

- Recap 🖉
- LOTUS
- Variance
- Properties of Variance
- Independence of random variables

#### **Review Random Variables**

**Definition.** A random variable (RV) for a probability space  $(\Omega, P)$  is a function  $X: \Omega \to \mathbb{R}$ .

The set of values that X can take on is its range/support:  $\mathbb{Z} \longrightarrow \mathbb{N}_X$ 

For a RV  $X: \Omega \to \mathbb{R}$ , the **probability mass function (pmf)** of X specifies, for any real number x, the probability that X = x

$$p_X(x) = P(X = x) = P(\{\omega \in \Omega \mid X(\omega) = x\}) \qquad \sum_{x \in \Omega_X} p_X(x) = 1$$

For a RV  $X: \Omega \to \mathbb{R}$ , the cumulative distribution function (cdf) of X specifies, for any real number x, the probability that  $X \leq x$ 

$$F_X(x) = P(X \le x)$$

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#### **Review Expected Value of a Random Variable**

**Definition.** Given a discrete  $\mathbb{RV} X: \Omega \to \mathbb{R}$ , the **expectation** or **expected value** or **mean** of *X* is

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot P(\omega)$$

or equivalently

$$\mathbb{E}[X] = \sum_{x \in \Omega_X} x \cdot P(X = x) = \sum_{x \in \Omega_X} x \cdot p_X(x)$$

Intuition: "Weighted average" of the possible outcomes (weighted by probability)

#### **Recap Linearity of Expectation**

**Theorem.** For any two random variables *X* and *Y*  $\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y].$ 

Or, more generally: For any random variables  $X_1, \ldots, X_n$ ,

 $\mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n].$ 

**Theorem.** For any random variables *X*, and constants *a* and *b*  $\mathbb{E}[aX + b] = a \cdot \mathbb{E}[X] + b.$  Using LOE to compute complicated expectations



Often boils down to the following three steps:

<u>Decompose</u>: Finding the right way to decompose the random variable into sum of simple random variables

 $X = X_1 + \dots + X_n$ 

• <u>LOE</u>: Apply linearity of expectation.

 $\mathbb{E}[X] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n].$ 

<u>Conquer</u>: Compute the expectation of each X<sub>i</sub>

Often,  $X_i$  are indicator (0/1) random variables.

#### Indicator random variables – 0/1 valued

For any event *A*, can define the indicator random variable  $X_A$  for *A*  $X_A = \begin{cases} 1 & \text{if event } A \text{ occurs} \\ 0 & \text{if event } A \text{ does not occur} \end{cases} \begin{cases} P(X_A = 1) = P(A) \\ P(X_A = 0) = 1 - P(A) \end{cases}$ 



 $\mathbb{E}[X_A] = P(A) = p$ 

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#### Linearity of Expectation – Even stronger

**Theorem.** For any random variables  $X_1, ..., X_n$ , and real numbers  $a_1, ..., a_n \models \mathbb{R}$ ,  $\mathbb{E}[a_1X_1 + \cdots + a_nX_n + b] = a_1\mathbb{E}[X_1] + \cdots + a_n\mathbb{E}[X_n] + b.$ 

#### Very important: In general, we do <u>not</u> have $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$



#### Linearity is special!

In general  $\mathbb{E}[g(X)] \neq g(\mathbb{E}(X))$ 

E.g.,  $X = \begin{cases} +1 \text{ with prob } 1/2 \\ -1 \text{ with prob } 1/2 \end{cases}$ 

Then:  $\mathbb{E}[X^2] \neq \mathbb{E}[X]^2$ How DO we compute  $\mathbb{E}[g(X)]$ ?

$$E(x) = \sum_{\omega \in \mathcal{X}} x(\omega) r(\omega)$$

$$E(x) = \sum_{x \in \Omega_X} x P(X=x)$$
Expected Value of  $g(X)$ 

$$Y = g(X)$$
Definition. Given a discrete RV  $X: \Omega \to \mathbb{R}$ , the expectation or expected value or mean of  $g(X)$  is
$$E[g(X)] = \sum_{\omega \in \Omega} g(X(\omega)) \cdot P(\omega)$$
or equivalently
$$E[g(X)] = \sum_{x \in \Omega_X} g(x) \cdot P(X=x) = \sum_{x \in \Omega_X} g(x) \cdot p_X(x)$$

Also known as LOTUS: "Law of the unconscious statistician

(nothing special going on in the discrete case)

#### **Example: from concept check**

$$\mathbb{E}[g(X)] = \sum_{x \in \Omega_X} g(x) \cdot P(X = x)$$

• Toss a die; each side equally likely. *X* is the number showing



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#### Which game would you rather play?

**Game 1:** In every round, you win \$2 with probability 1/3, lose \$1 with probability 2/3.

 $W_{1} = \text{payoff in a round of Game 1}$   $P(W_{1} = 2) = \frac{1}{3}, P(W_{1} = -1) = \frac{2}{3}$   $E(W_{1}) = 2 \cdot \frac{1}{3} + (-1)^{2} = 0$ 

#### Which game would you rather play?

**Game 1:** In every round, you win \$2 with probability 1/3, lose \$1 with probability 2/3.

$$W_1 = \text{payoff in a round of Game 1}$$
  
 $P(W_1 = 2) = \frac{1}{3}, P(W_1 = -1) = \frac{2}{3}$   
 $\mathbb{E}[W_1] = 0$ 

**Game 2:** In every round, you win \$10 with probability 1/3, lose \$5 with probability 2/3.

$$W_{2} = \text{payoff in a round of Game 2}$$

$$P(W_{2} = 10) = \frac{1}{3}, P(W_{2} = -5) = \frac{2}{3}$$

$$\mathbb{E}[W_{2}] = 0$$



Same expectation, but clearly a very different distribution. We want to capture the difference – New concept: Variance

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New quantity (random variable): How far from the expectation?

 $W_{1} - \mathbb{E}[W_{1}]$  $\mathbb{E}[W_{1} - \mathbb{E}[W_{1}]]$  $= \mathbb{E}[W_{1}] - \mathbb{E}[\mathbb{E}[W_{1}]]$  $= \mathbb{E}[W_{1}] - \mathbb{E}[W_{1}]]$ = 0

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A better quantity (random variable): How far from the expectation?

$$\mathbb{E}[(W_2 - \mathbb{E}[W_2])^2] = \frac{2}{3}(25) + \frac{1}{3}(100)$$

$$= 50$$



g(x) = (x - E(X))

#### Variance

$$\mathbb{E}[g(X)] = \sum_{x \in \Omega_X} g(x) \cdot P(X = x)$$



Intuition: Variance (or standard deviation) is a quantity that measures, in expectation, how "far" the random variable is from its expectation.

## $\mathbb{E}[g(X)] = \sum_{x \in \Omega_X} g(x) \cdot P(X = x)$ Variance – Example 1 $\operatorname{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \sum p_X(x) \cdot (x - \mathbb{E}[X])^2$ X fair die • $P(X = 1) = \dots = P(X = 6) = 1/6$ • $\mathbb{E}[X] = 3.5$ $\operatorname{Var}(X) = \sum_{x} P(X = x) \cdot (x - \mathbb{E}[X])^{2}$ $= \frac{1}{6} \left( 1 - 3.5 \right)^{2} + \frac{1}{6} \left( 2 - 3.5 \right)^{2} + \dots + \frac{1}{6} \left( 6 - 3.5 \right)^{2}$

#### Variance – Example 1

#### X fair die

- $P(X = 1) = \dots = P(X = 6) = 1/6$
- $\mathbb{E}[X] = 3.5$

 $Var(X) = \sum_{x} P(X = x) \cdot (x - \mathbb{E}[X])^{2}$ =  $\frac{1}{6} [(1 - 3.5)^{2} + (2 - 3.5)^{2} + (3 - 3.5)^{2} + (4 - 3.5)^{2} + (5 - 3.5)^{2} + (6 - 3.5)^{2}]$ =  $\frac{2}{6} [2.5^{2} + 1.5^{2} + 0.5^{2}] = \frac{2}{6} [\frac{25}{4} + \frac{9}{4} + \frac{1}{4}] = \frac{35}{12} \approx 2.91677 \dots$ 

#### **Variance in Pictures**

Captures how much "spread' there is in a pmf

All pmfs have same expectation

= frall of truse F(X)



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 $= \sum_{x} (x)^{2} (x - E(x))^{2} = a^{2} Vor(X)$ 

#### Variance – Properties



 $= E(X^{*}) - [E(X)]^{*}$ 



#### Variance – Example 1

# X fair die • $\mathbb{P}(X = 1) = \dots = \mathbb{P}(X = 6) = 1/6$ • $\mathbb{E}[X] = \frac{21}{6}$ • $\mathbb{E}[X^2] = \frac{91}{6} - \frac{1}{2} + \frac{1}{6} +$ $Var(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{91}{6} - \left(\frac{21}{6}\right)^2 = \frac{105}{36} \approx 2.91677$ (X-E(x))2]

#### Variance of Indicator Random Variables

Suppose that  $X_A$  is an indicator RV for event A with P(A) = p so XA=) G.W.

#### Variance of Indicator Random Variables

Suppose that  $X_A$  is an indicator RV for event A with P(A) = p so  $\mathbb{E}[X_A] = P(A) = p$ 

Since  $X_A$  only takes on values 0 and 1, we always have  $X_A^2 = X_A$  so





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#### **Random Variables and Independence**

Comma is shorthand for AND

**Definition.** Two random variables *X*, *Y* are **(mutually) independent** if for all *x*, *y*,

$$P(X = y, Y = y) = P(X = x) \cdot P(Y = y)$$

Intuition: Knowing X doesn't help you guess Y and vice versa

**Definition.** The random variables  $X_1, ..., X_n$  are **(mutually) independent** if for all  $x_1, ..., x_n$ ,

$$P(X_1 = x_1, \dots, X_n = x_n) = P(X_1 = x_1) \cdots P(X_n = x_n)$$

Note: No need to check for all subsets, but need to check for all values!

#### Example

Let X be the number of heads in n independent coin flips of the same coin with probability p of coming up heads.

Let  $Y = X \mod 2$  be the parity (even/odd) of X.

Are X and Y independent?

$$P(X = 3, Y = 0) = 0$$

$$Poll: slido.com/3680281$$

$$A. Yes$$

$$B. No$$

$$P(X=3) = 3 (1-p)^{3}$$

$$P(X=0) = 3$$

$$P(X=0) = 3$$

#### Example

Make 2n independent coin flips of the same coin with probability p of coming up heads. .

Let X be the number of heads in the first n flips and Y be the number of heads in the last n flips.

Are X and Y independent?

Poll: slido.com/3680281

A. YesB. No

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- Properties of Independent Random Variables

#### Important Facts about Independent Random Variables

**Theorem.** If *X*, *Y* independent,  $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$ 

**Theorem.** If *X*, *Y* independent, Var(X + Y) = Var(X) + Var(Y)

**Corollary.** If  $X_1, X_2, ..., X_n$  mutually independent,  $\operatorname{Var}\left(\sum_{i=1}^n X_i\right) = \sum_i^n \operatorname{Var}(X_i)$ 

#### (Not Covered) Proof of $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$

**Theorem.** If *X*, *Y* independent,  $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$ 

Proof

Let 
$$x_i, y_i, i = 1, 2, ...$$
 be the possible values of  $X, Y$ .  

$$\mathbb{E}[X \cdot Y] = \sum_i \sum_j x_i \cdot y_j \cdot P(X = x_i \land Y = y_j)$$
independence
$$= \sum_i \sum_j x_i \cdot y_i \cdot P(X = x_i) \cdot P(Y = y_j)$$

$$= \sum_i x_i \cdot P(X = x_i) \cdot \left(\sum_j y_j \cdot P(Y = y_j)\right)$$

$$= \mathbb{E}[X] \cdot \mathbb{E}[Y]$$
Note: NOT true in general; see earlier example  $\mathbb{E}[X^2] \neq \mathbb{E}[X]^2$ 

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#### (Not Covered) Proof of Var(X + Y) = Var(X) + Var(Y)

**Theorem.** If *X*, *Y* independent, Var(X + Y) = Var(X) + Var(Y)

Proof

$$Var(X + Y) = \mathbb{E}[(X + Y)^{2}] - (\mathbb{E}[X + Y])^{2}$$
  

$$= \mathbb{E}[X^{2} + 2XY + Y^{2}] - (\mathbb{E}[X] + \mathbb{E}[Y])^{2}$$
  

$$= \mathbb{E}[X^{2}] + 2 \mathbb{E}[XY] + \mathbb{E}[Y^{2}] - (\mathbb{E}[X]^{2} + 2 \mathbb{E}[X] \mathbb{E}[Y] + \mathbb{E}[Y]^{2})$$
  

$$= \mathbb{E}[X^{2}] - \mathbb{E}[X]^{2} + \mathbb{E}[Y^{2}] - \mathbb{E}[Y]^{2} + 2 \mathbb{E}[XY] - 2 \mathbb{E}[X] \mathbb{E}[Y]$$
  

$$= Var(X) + Var(Y) + 2 \mathbb{E}[XY] - 2 \mathbb{E}[X] \mathbb{E}[Y]$$
  

$$= Var(X) + Var(Y)$$
  
equal by independence

### **Example – Coin Tosses**

We flip n independent coins, each one heads with probability p

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$$X_i = \begin{cases} 1, \ i^{\text{th}} \text{ outcome is heads} \\ 0, \ i^{\text{th}} \text{ outcome is tails.} \end{cases}$$
  
-  $Z = \text{number of heads}$   
What is  $\mathbb{E}[Z]$ ? What is  $\text{Var}(Z)$ ?  
P( $Z = k$ ) =  $\binom{n}{k}p^k(1-p)^{n-k}$   
Note:  $X_1, \dots, X_n$  are mutually independent! [Verify it formally!]  
Var( $Z$ ) =  $\sum_{i=1}^{n} \text{Var}(X_i) = n \cdot p(1-p)$   
Note  $\text{Var}(X_i) = p(1-p)$ 

