CSE 312 Foundations of Computing II

Lecture 10: LOTUS, variance and independence among random variables.

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Agenda

- Recap 🖉
- LOTUS
- Variance
- Properties of Variance
- Independence of random variables

Review Random Variables

Definition. A random variable (RV) for a probability space (Ω, P) is a function $X: \Omega \to \mathbb{R}$.

The set of values that X can take on is its range/support: $X(\Omega)$

For a RV $X: \Omega \to \mathbb{R}$, the **probability mass function (pmf)** of X specifies, for any real number x, the probability that X = x

$$p_X(x) = P(X = x) = P(\{\omega \in \Omega \mid X(\omega) = x\}) \qquad \sum_{x \in \Omega_X} p_X(x) = 1$$

For a RV $X: \Omega \to \mathbb{R}$, the cumulative distribution function (cdf) of X specifies, for any real number x, the probability that $X \leq x$

$$F_X(x) = P(X \le x)$$

Review Expected Value of a Random Variable

Definition. Given a discrete $\mathbb{RV} X: \Omega \to \mathbb{R}$, the **expectation** or **expected value** or **mean** of *X* is

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot P(\omega)$$

or equivalently

$$\mathbb{E}[X] = \sum_{x \in \Omega_X} x \cdot P(X = x) = \sum_{x \in \Omega_X} x \cdot p_X(x)$$

Intuition: "Weighted average" of the possible outcomes (weighted by probability)

Recap Linearity of Expectation

Theorem. For any two random variables X and Y $\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y].$

Or, more generally: For any random variables X_1, \ldots, X_n ,

 $\mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n].$

Theorem. For any random variables *X*, and constants *a* and *b* $\mathbb{E}[aX + b] = a \cdot \mathbb{E}[X] + b.$

Using LOE to compute complicated expectations

Often boils down to the following three steps:

<u>Decompose</u>: Finding the right way to decompose the random variable into sum of simple random variables

 $X = X_1 + \dots + X_n$

• <u>LOE</u>: Apply linearity of expectation.

 $\mathbb{E}[X] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n].$

<u>Conquer</u>: Compute the expectation of each X_i

Often, X_i are indicator (0/1) random variables.

Indicator random variables – 0/1 valued

For any event A, can define the indicator random variable X_A for A $X_A = \begin{cases} 1 & \text{if event } A \text{ occurs} \\ 0 & \text{if event } A \text{ does not occur} \end{cases} \begin{cases} P(X_A = 1) = P(A) \\ P(X_A = 0) = 1 - P(A) \end{cases}$



 $\mathbb{E}[X_A] = P(A) = p$

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Linearity of Expectation – Even stronger

Theorem. For any random variables $X_1, ..., X_n$, and real numbers $a_1, ..., a_n \in \mathbb{R}$, $\mathbb{E}[a_1X_1 + \cdots + a_nX_n + b] = a_1\mathbb{E}[X_1] + \cdots + a_n\mathbb{E}[X_n] + b.$

Very important: In general, we do <u>not</u> have $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$

Linearity is special!

In general $\mathbb{E}[g(X)] \neq g(\mathbb{E}(X))$

E.g., $X = \begin{cases} +1 \text{ with prob } 1/2 \\ -1 \text{ with prob } 1/2 \end{cases}$

Then: $\mathbb{E}[X^2] \neq \mathbb{E}[X]^2$

How DO we compute $\mathbb{E}[g(X)]$?

Expected Value of g(X)

Definition. Given a discrete RV $X: \Omega \to \mathbb{R}$, the **expectation** or **expected** value or mean of g(X) is

$$\mathbb{E}[g(X)] = \sum_{\omega \in \Omega} g(X(\omega)) \cdot P(\omega)$$

or equivalently

$$\mathbb{E}[g(X)] = \sum_{x \in \Omega_X} g(x) \cdot P(X = x) = \sum_{x \in \Omega_X} g(x) \cdot p_X(x)$$

Also known as LOTUS: "Law of the unconscious statistician

(nothing special going on in the discrete case)

Example: from concept check

$$\mathbb{E}[g(X)] = \sum_{x \in \Omega_X} g(x) \cdot P(X = x)$$

- Toss a die; each side equally likely. *X* is the number showing
- $Y = X \mod 4$
- What is $\mathbb{E}[Y]$?

Pr(w)	ω	X
1/6	1	1
1/6	2	2
1/6	3	3
1/6	4	4
1/6	5	5
1/6	6	6

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Which game would you rather play?

Game 1: In every round, you win \$2 with probability 1/3, lose \$1 with probability 2/3.

 W_1 = payoff in a round of Game 1 $P(W_1 = 2) = \frac{1}{3}$, $P(W_1 = -1) = \frac{2}{3}$

Which game would you rather play?

Game 1: In every round, you win \$2 with probability 1/3, lose \$1 with probability 2/3.

$$W_1 = \text{payoff in a round of Game 1}$$

 $P(W_1 = 2) = \frac{1}{3}, P(W_1 = -1) = \frac{2}{3}$
 $\mathbb{E}[W_1] = 0$

Game 2: In every round, you win \$10 with probability 1/3, lose \$5 with probability 2/3.

$$W_2 = \text{payoff in a round of Game 2}$$

 $P(W_2 = 10) = \frac{1}{3}, P(W_2 = -5) = \frac{2}{3}$
 $\mathbb{E}[W_2] = 0$



Same expectation, but clearly a very different distribution. We want to capture the difference – New concept: Variance



New quantity (random variable): How far from the expectation? $W_1 - \mathbb{E}[W_1]$



New quantity (random variable): How far from the expectation?

 $W_{1} - \mathbb{E}[W_{1}]$ $\mathbb{E}[W_{1} - \mathbb{E}[W_{1}]]$ $= \mathbb{E}[W_{1}] - \mathbb{E}[\mathbb{E}[W_{1}]]$ $= \mathbb{E}[W_{1}] - \mathbb{E}[W_{1}]$ = 0







A better quantity (random variable): How far from the expectation?

$$\mathbb{E}[(W_2 - \mathbb{E}[W_2])^2] = \frac{2}{3} \cdot 25 + \frac{1}{3} \cdot 100 = 50$$



We say that W_2 has "higher variance" than W_1 .

 $Var(W) = \mathbb{E}[(W - \mathbb{E}[W])^2]$

Variance

$$\mathbb{E}[g(X)] = \sum_{x \in \Omega_X} g(x) \cdot P(X = x)$$



Intuition: Variance (or standard deviation) is a quantity that measures, in expectation, how "far" the random variable is from its expectation.

Variance – Example 1

X fair die

- $P(X = 1) = \dots = P(X = 6) = 1/6$
- $\mathbb{E}[X] = 3.5$

$$Var(X) = \sum_{x} P(X = x) \cdot (x - \mathbb{E}[X])^2$$

 $\mathbb{E}[g(X)] = \sum_{x \in \Omega_X} g(x) \cdot P(X = x)$ $Var(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \sum_x p_X(x) \cdot (x - \mathbb{E}[X])^2$

Variance – Example 1

X fair die

- $P(X = 1) = \dots = P(X = 6) = 1/6$
- $\mathbb{E}[X] = 3.5$

 $Var(X) = \sum_{x} P(X = x) \cdot (x - \mathbb{E}[X])^{2}$ = $\frac{1}{6} [(1 - 3.5)^{2} + (2 - 3.5)^{2} + (3 - 3.5)^{2} + (4 - 3.5)^{2} + (5 - 3.5)^{2} + (6 - 3.5)^{2}]$ = $\frac{2}{6} [2.5^{2} + 1.5^{2} + 0.5^{2}] = \frac{2}{6} [\frac{25}{4} + \frac{9}{4} + \frac{1}{4}] = \frac{35}{12} \approx 2.91677 \dots$

Variance in Pictures

Captures how much "spread' there is in a pmf

All pmfs have same expectation



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Variance – Properties

Definition. The variance of a (discrete) RV X is $Var(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \sum_x p_X(x) \cdot (x - \mathbb{E}[X])^2$

Theorem. For any $a, b \in \mathbb{R}$, $Var(a \cdot X + b) = a^2 \cdot Var(X)$

Variance – Properties

Definition. The variance of a (discrete) RV *X* is $Var(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \sum_x p_X(x) \cdot (x - \mathbb{E}[X])^2$ The same $Var(X) = \mathbb{E}[Y^2] = \mathbb{E}[Y^2]^2$

Theorem. $Var(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$

Variance

Theorem.
$$Var(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

Proof: $\operatorname{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2]$ $= \mathbb{E}[X^2 - 2\mathbb{E}[X] \cdot X + \mathbb{E}[X]^2]$ $= \mathbb{E}(X^2) - 2\mathbb{E}[X]\mathbb{E}[X] + \mathbb{E}[X]^2$ $= \mathbb{E}[X^2] - \mathbb{E}[X]^2$ (linearity of expectation!) $\mathbb{E}[X^2] \text{ and } \mathbb{E}[X]^2$ are different !

Variance – Example 1

X fair die

- $\mathbb{P}(X = 1) = \dots = \mathbb{P}(X = 6) = 1/6$
- $\mathbb{E}[X] = \frac{21}{6}$
- $\mathbb{E}[X^2] = \frac{91}{6}$

$$Var(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{91}{6} - \left(\frac{21}{6}\right)^2 = \frac{105}{36} \approx 2.91677$$

Variance of Indicator Random Variables

Suppose that X_A is an indicator RV for event A with P(A) = p so $\mathbb{E}[X_A] = P(A) = p$

$\operatorname{Var}(X_A) = \mathbb{E}[X_A^2] - \mathbb{E}[X_A]^2 =$

Variance of Indicator Random Variables

Suppose that X_A is an indicator RV for event A with P(A) = p so $\mathbb{E}[X_A] = P(A) = p$

Since X_A only takes on values 0 and 1, we always have $X_A^2 = X_A$ so

 $Var(X_A) = \mathbb{E}[X_A^2] - \mathbb{E}[X_A]^2 = \mathbb{E}[X_A] - \mathbb{E}[X_A]^2 = p - p^2 = p(1 - p)$

In General, $Var(X + Y) \neq Var(X) + Var(Y)$

Proof by counter-example:

- Let *X* be a r.v. with pmf *P*(*X* = 1) = *P*(*X* = −1) = 1/2 – What is E[*X*] and Var(*X*)?
- Let Y = -X
 - What is $\mathbb{E}[Y]$ and Var(Y)?

What is Var(X + Y)?

In General, $Var(X + Y) \neq Var(X) + Var(Y)$

Proof by counter-example:

Recall glued coins



- Let X_1 be a r.v. that indicates if the first coin comes up heads.
- Let X_2 be a r.v. that indicates if the second coin comes up heads.



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Random Variables and Independence

Comma is shorthand for AND

Definition. Two random variables *X*, *Y* are **(mutually) independent** if for all *x*, *y*,

$$P(X = x, Y = y) = P(X = x) \cdot P(Y = y)$$

Intuition: Knowing X doesn't help you guess Y and vice versa

Definition. The random variables $X_1, ..., X_n$ are **(mutually) independent** if for all $x_1, ..., x_n$,

$$P(X_1 = x_1, \dots, X_n = x_n) = P(X_1 = x_1) \cdots P(X_n = x_n)$$

Note: No need to check for all subsets, but need to check for all values!

Let *X* be the number of heads in *n* independent coin flips of the same coin with probability *p* of coming up heads. Let $Y = X \mod 2$ be the parity (even/odd) of *X*. Are *X* and *Y* independent?

Let *X* be the number of heads in *n* independent coin flips of the same coin with probability *p* of coming up heads. Let $Y = X \mod 2$ be the parity (even/odd) of *X*. Are *X* and *Y* independent?

> Poll: slido.com/3680281

A. YesB. No

Make 2n independent coin flips of the same coin with probability p of coming up heads. .

Let X be the number of heads in the first n flips and Y be the number of heads in the last n flips.

Are *X* and *Y* independent?

Make 2n independent coin flips of the same coin with probability p of coming up heads. .

Let X be the number of heads in the first n flips and Y be the number of heads in the last n flips.

Are X and Y independent?

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A. YesB. No

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- Properties of Independent Random Variables

Important Facts about Independent Random Variables

Theorem. If *X*, *Y* independent, $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$

Theorem. If *X*, *Y* independent, Var(X + Y) = Var(X) + Var(Y)

Corollary. If $X_1, X_2, ..., X_n$ mutually independent, $\operatorname{Var}\left(\sum_{i=1}^n X_i\right) = \sum_i^n \operatorname{Var}(X_i)$

(Not Covered) Proof of $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$

Theorem. If X, Y independent, $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$ Proof Let x_i , y_i , i = 1, 2, ... be the possible values of X, Y. $\mathbb{E}[X \cdot Y] = \sum_{i} \sum_{j} x_i \cdot y_j \cdot P(X = x_i \wedge Y = y_j)$ independence $=\sum_{i}\sum_{i}x_{i}\cdot y_{i}\cdot P(X=x_{i})\cdot P(Y=y_{j})$ $=\sum_{i} x_{i} \cdot P(X = x_{i}) \cdot \left(\sum_{i} y_{j} \cdot P(Y = y_{j})\right)$ $= \mathbb{E}[X] \cdot \mathbb{E}[Y]$ Note: NOT true in general; see earlier example $\mathbb{E}[X^2] \neq \mathbb{E}[X]^2$

(Not Covered) Proof of Var(X + Y) = Var(X) + Var(Y)

Theorem. If *X*, *Y* independent, Var(X + Y) = Var(X) + Var(Y)

Proof

Example – Coin Tosses

We flip n independent coins, each one heads with probability p

-
$$X_i = \begin{cases} 1, \ i^{\text{th}} \text{ outcome is heads} \\ 0, \ i^{\text{th}} \text{ outcome is tails.} \end{cases}$$

- $Z = \text{number of heads}$
What is $\mathbb{E}[Z]$? What is $\text{Var}(Z)$?
P($Z = k$) = $\binom{n}{k}p^k(1-p)^{n-k}$
Note: X_1, \dots, X_n are mutually independent! [Verify it formally!]
Var(Z) = $\sum_{i=1}^{n} \text{Var}(X_i) = n \cdot p(1-p)$
Note $\text{Var}(X_i) = p(1-p)$

