CSE 312
Foundations of Computing II
Lecture 18: CLT \& Polling

## Review The Normal Distribution

Definition. A Gaussian (or normal) random variable with parameters $\mu \in \mathbb{R}$ and $\sigma^{2} \geq 0$ has density

$$
f_{X}(x)=\frac{1}{\sqrt{2 \pi} \sigma} e^{-\frac{(x-\mu)^{2}}{2 \sigma^{2}}}
$$

We say that $X$ follows the Normal Distribution, and write $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$.

$\mathcal{N}(0,1)$. No closed form expression for CDF...

## Review The Normal Distribution. <br> $$
X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)
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Definition. A Gaussian (or normal) random variable $X$ with parameters $\mu \in \mathbb{R}$ and $\sigma \geq 0$ has density

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$$

Carl Friedrich Gauss

Fact. If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $\mathbb{E}[X]=\mu$, and $\operatorname{Var}(X)=\sigma^{2}$

Proof of expectation is easy because density curve is symmetric around $\mu$,

$$
f_{X}(\mu-x)=f_{X}(\mu+x), \text { but proof for variance requires integration of } e^{-x^{2} / 2}
$$

Review Standard (unit) normal $=\mathcal{N}(0,1)$

CDF. $\Phi(z)=P(Z \leq z)=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{Z} e^{-x^{2} / 2} \mathrm{~d} x$ for $Z \sim \mathcal{N}(0,1)$
Note: $\Phi(z)$ has no closed form - generally given via tables

## Review Table of $\boldsymbol{\Phi ( z ) C D F ~ o f ~ S t a n d a r d ~ N o r m a l ~}$

$$
P(Z \leq 0.98)=\Phi(0.98) \approx 0.8365
$$

## For what $a$ is

$$
P(Z>a) \leq 0.01 ?
$$

| $\Phi$ Table: $\mathbb{P}(Z \leq z)$ when $Z \sim \mathcal{N}(0,1)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| 0.0 | 0.5 | 0.50399 | 0.50798 | 0.51197 | 0.51595 | 0.51994 | 0.52392 | 0.5279 | 0.53188 | 0.53586 |
| 0.1 | 0.53983 | 0.5438 | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 0.57535 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
| 0.3 | 0.61791 | 0.62172 | 0.62552 | 0.6293 | 0.63307 | 0.63683 | 0.64058 | 0.64431 | 0.64803 | 0.65173 |
| 0.4 | 0.65542 | 0.6591 | 0.66276 | 0.6664 | 0.67003 | 0.67364 | 0.67724 | 0.68082 | 0.68439 | 0.68793 |
| 0.5 | 0.69146 | 0.69497 | 0.69847 | 0.70194 | 0.7054 | 0.70884 | 0.71226 | 0.71566 | 0.71904 | 0.7224 |
| 0.6 | 0.72575 | 0.72907 | 0.73237 | 0.73565 | 0.73891 | 0.74215 | 0.74537 | 0.74857 | 0.75175 | 0.7549 |
| 0.7 | 0.75804 | 0.76115 | 0.76424 | 0.7673 | 0.77035 | 0.77337 | 0.77637 | 0.77935 | 0.7823 | 0.78524 |
| 0.8 | 0.78814 | 0.79103 | 0.79389 | 0.79673 | 0.79955 | 0.80234 | 0.80511 | 0.80785 | 0.81057 | 0.81327 |
| 0.9 | 0.81594 | 0.81859 | 0.82121 | 0.82381 | 0.82639 | 0.82894 | 0.83147 | 0.83398 | 0.83646 | 0.83891 |
| 1.0 | 0.84134 | 0.84375 | 0.84614 | 0.84849 | 0.85083 | 0.85314 | 0.85543 | 0.85769 | 0.85993 | 0.86214 |
| 1.1 | 0.86433 | 0.8665 | 0.86864 | 0.87076 | 0.87286 | 0.87493 | 0.87698 | 0.879 | 0.881 | 0.88298 |
| 1.2 | 0.88493 | 0.88686 | 0.88877 | 0.89065 | 0.89251 | 0.89435 | 0.89617 | 0.89796 | 0.89973 | 0.90147 |
| 1.3 | 0.9032 | 0.9049 | 0.90658 | 0.90824 | 0.90988 | 0.91149 | 0.91309 | 0.91466 | 0.91621 | 0.91774 |
| 1.4 | 0.91924 | 0.92073 | 0.9222 | 0.92364 | 0.92507 | 0.92647 | 0.92785 | 0.92922 | 0.93056 | 0.93189 |
| 1.5 | 0.93319 | 0.93448 | 0.93574 | 0.93699 | 0.93822 | 0.93943 | 0.94062 | 0.94179 | 0.94295 | 0.94408 |
| 1.6 | 0.9452 | 0.9463 | 0.94738 | 0.94845 | 0.9495 | 0.95053 | 0.95154 | 0.95254 | 0.95352 | 0.95449 |
| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608 | 0.96164 | 0.96246 | 0.96327 |
| 1.8 | 0.96407 | 0.96485 | 0.96562 | 0.96638 | 0.96712 | 0.96784 | 0.96856 | 0.96926 | 0.96995 | 0.97062 |
| 1.9 | 0.97128 | 0.97193 | 0.97257 | 0.9732 | 0.97381 | 0.97441 | 0.975 | 0.97558 | 0.97615 | 0.9767 |
| 2.0 | 0.97725 | 0.97778 | 0.97831 | 0.97882 | 0.97932 | 0.97982 | 0.9803 | 0.98077 | 0.98124 | 0.98169 |
| 2.1 | 0.98214 | 0.98257 | 0.983 | 0.98341 | 0.98382 | 0.98422 | 0.98461 | 0.985 | 0.98537 | 0.98574 |
| 2.2 | 0.9861 | 0.98645 | 0.98679 | 0.98713 | 0.98745 | 0.98778 | 0.98809 | 0.9884 | 0.9887 | 0.98899 |
| 2.3 | 0.98928 | 0.98956 | 0.98983 | 0.9901 | 0.99036 | 0.99061 | 0.99086 | 0.99111 | 0.99134 | 0.99158 |
| 2.4 | 0.9918 | 0.99202 | 0.99224 | 0.99245 | 0.99266 | 0.99286 | 0.99305 | 0.99324 | 0.99343 | 0.99361 |
| 2.5 | 0.99379 | 0.99396 | 0.99413 | 0.9943 | 0.99446 | 0.99461 | 0.99477 | 0.99492 | 0.99506 | 0.9952 |
| 2.6 | 0.99534 | 0.99547 | 0.9956 | 0.99573 | 0.99585 | 0.99598 | 0.99609 | 0.99621 | 0.99632 | 0.99643 |
| 2.7 | 0.99653 | 0.99664 | 0.99674 | 0.99683 | 0.99693 | 0.99702 | 0.99711 | 0.9972 | 0.99728 | 0.99736 |
| 2.8 | 0.99744 | 0.99752 | 0.9976 | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999 |

## Review Table of $\boldsymbol{\Phi}(\mathbf{z})$ CDF of Standard Normal

$$
P(Z \leq 0.98)=\Phi(0.98) \approx 0.8365
$$

## For what $a$ is

$$
P(Z \geq a) \leq 0.01 ?
$$

For any $a \geq 2.33$
$P(Z>a) \leq 0.01$.

| $\Phi$ Table: $\mathbb{P}(Z \leq z)$ when $Z \sim \mathcal{N}(0,1)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
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| 0.1 | 0.53983 | 0.5438 | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 0.57535 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
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| 0.4 | 0.65542 | 0.6591 | 0.66276 | 0.6664 | 0.67003 | 0.67364 | 0.67724 | 0.68082 | 0.68439 | 0.68793 |
| 0.5 | 0.69146 | 0.69497 | 0.69847 | 0.70194 | 0.7054 | 0.70884 | 0.71226 | 0.71566 | 0.71904 | 0.7224 |
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| 0.7 | 0.75804 | 0.76115 | 0.76424 | 0.7673 | 0.77035 | 0.77337 | 0.77637 | 0.77935 | 0.7823 | 0.78524 |
| 0.8 | 0.78814 | 0.79103 | 0.79389 | 0.79673 | 0.79955 | 0.80234 | 0.80511 | 0.80785 | 0.81057 | 0.81327 |
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| 1.0 | 0.84134 | 0.84375 | 0.84614 | 0.84849 | 0.85083 | 0.85314 | 0.85543 | 0.85769 | 0.05090 | 0.86214 |
| 1.1 | 0.86433 | 0.8665 | 0.86864 | 0.87076 | 0.87286 | 0.87493 | 0.87698 | 0.879 | 0.881 | 0.88298 |
| 1.2 | 0.88493 | 0.88686 | 0.88877 | 0.89065 | 0.89251 | 0.89435 | 0.89617 | 0.89796 | 0.89973 | 0.90147 |
| 1.3 | 0.9032 | 0.9049 | 0.90658 | 0.90824 | 0.90988 | 0.91149 | 0.91309 | 0.91466 | 0.91621 | 0.91774 |
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| 1.5 | 0.93319 | 0.93448 | 0.93574 | 0.93699 | 0.93822 | 0.93943 | 0.94062 | 0.94179 | 0.94295 | 0.94408 |
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| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608 | 0.96164 | 0.96246 | 0.96327 |
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| 2.8 | 0.99744 | 0.99752 | 0.9976 | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999 |

## Review Properties of normal distributions

- Normal distributions stay normal under shifting and scaling.
- To "standardize" a normal random variable $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, you subtract the mean and divide by the standard deviation, i.e.,

$$
\frac{x-\mu}{\sigma} \sim \mathcal{N}(0,1)
$$

- This allows you to use the standard normal tables (showing $\Phi(z)=$ $P(Z \leq z)$ for $Z \sim \mathcal{N}(0,1)$ ) to do calculations for any normal distribution.

Review Analyzing non-standard normal in terms of $\mathcal{N}(0,1)$

If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $\frac{X-\mu}{\sigma} \sim \mathcal{N}(0,1)$

Therefore,

$$
F_{X}(z)=P(X \leq z)=P\left(\frac{X-\mu}{\sigma} \leq \frac{z-\mu}{\sigma}\right)=\Phi\left(\frac{z-\mu}{\sigma}\right)
$$

## Example

Let $X \sim \mathcal{N}(0,1)$.

$$
\begin{aligned}
P(|X-\mu|<k \sigma) & =P\left(\frac{|X-\mu|}{\sigma}<k\right)= \\
& =P\left(-k<\frac{X-\mu}{\sigma}<k\right)=\Phi(k)-\Phi(-k)
\end{aligned}
$$

e.g. $k=1: \quad 68 \%$
$k=2: 95 \%$
$k=3: 99 \%$

## Review closure under addition

Fact. If $X \sim \mathcal{N}\left(\mu_{X}, \sigma_{X}^{2}\right), \mathrm{Y} \sim \mathcal{N}\left(\mu_{Y}, \sigma_{Y}^{2}\right)$ (both independent normal RV) then $\mathrm{a} X+b Y+c \sim \mathcal{N}\left(a \mu_{X}+b \mu_{Y}+c, a^{2} \sigma_{X}^{2}+b^{2} \sigma_{Y}^{2}\right)$

## Agenda

- Central Limit Theorem (CLT)
- Polling


## Normal Distributions EVERYWHERE - why?



## S\&P 500 Returns after Elections




## Sums of i.i.d. RVs look normal!

$X_{1}, \ldots, X_{n}$ i.i.d. with expectation $\mu$ and variance $\sigma^{2}$
Consider

$$
S_{n}=X_{1}+\cdots+X_{n}
$$

## Empirical observation:

$S_{n}$ looks like a normal RV as $n$ grows.

## Central Limit Theorem

$$
\begin{aligned}
& \mathbb{E}\left[S_{n}\right]=\mathbb{E}\left[X_{1}\right]+\cdots+\mathbb{E}\left[X_{n}\right]=n \mu \\
& \operatorname{Var}\left(S_{n}\right)=\operatorname{Var}\left(X_{1}\right)+\cdots+\operatorname{Var}\left(X_{n}\right)=n \sigma^{2}
\end{aligned}
$$

$X_{1}, \ldots, X_{n}$ i.i.d., each with expectation $\mu$ and variance $\sigma^{2}$

Define $S_{n}=X_{1}+\cdots+X_{n}$,

$$
Y_{n}=\frac{S_{n}-n \mu}{\sigma \sqrt{n}}
$$

Then distribution of $Y_{n}=\frac{S_{n}-n \mu}{\sigma \sqrt{n}}$ converges to that of a normal distribution with mean 0 and variance 1 as $n \rightarrow \infty$.

## Central Limit Theorem

$$
Y_{n}=\frac{X_{1}+\cdots+X_{n}-n \mu}{\sigma \sqrt{n}}
$$

Theorem. (Central Limit Theorem) The CDF of $Y_{n}$ converges to the CDF of the standard normal $\mathcal{N}(0,1)$, i.e.,

$$
\lim _{n \rightarrow \infty} P\left(Y_{n} \leq y\right)=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{y} e^{-x^{2} / 2} \mathrm{~d} x
$$

## Summary Central Limit Theorem

$X_{1}, \ldots, X_{n}$ i.i.d., each with expectation $\mu$ and variance $\sigma^{2}$
Define $S_{n}=X_{1}+\cdots+X_{n}$ and $\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}$. and $Y_{n}=\frac{S_{n}-n \mu}{\sigma \sqrt{n}}$
mean
variance

CLT:

## CLT application

- You buy lightbulbs that burn out according to an exponential distribution with parameter $\lambda=1.8$ lightbulbs per year.
- You buy a pack of 10 (independent) light bulbs. What is the probability that your 10-pack lasts at least 5 years?


## Table of Standard Cumulative Normal Density

| $\Phi$ Table: $\mathbb{P}(Z \leq z)$ when $Z \sim \mathcal{N}(0,1)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| 0.0 | 0.5 | 0.50399 | 0.50798 | 0.51197 | 0.51595 | 0.51994 | 0.52392 | 0.5279 | 0.53188 | 0.53586 |
| 0.1 | 0.53983 | 0.5438 | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 0.57535 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
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| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608 | 0.96164 | 0.96246 | 0.96327 |
| 1.8 | 0.96407 | 0.96485 | 0.96562 | 0.96638 | 0.96712 | 0.96784 | 0.96856 | 0.96926 | 0.96995 | 0.97062 |
| 1.9 | 0.97128 | 0.97193 | 0.97257 | 0.9732 | 0.97381 | 0.97441 | 0.975 | 0.97558 | 0.97615 | 0.9767 |
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| 2.8 | 0.99744 | 0.99752 | 0.9976 | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999 |

## Outline of CLT steps

- Write the event you are interested in, in terms of a sum of i.i.d. random variables.
- Normalize RV to have mean o and standard deviation 1.
- Write event in terms of $\Phi$, the CDF of a $\mathcal{N}(0,1)$.
- Look up in table.

Outline of CLT steps - extra step if random variables are discrete.

- Write the event you are interested in, in terms of a sum of i.i.d. random variables.
- Apply continuity correction if RVs are discrete (see yesterday's section and Tsun Section 5.7.4)
- Normalize RV to have mean 0 and standard deviation 1.
- Write event in terms of $\Phi$, the CDF of a $\mathcal{N}(0,1)$.
- Look up in table.


## Agenda

- Central Limit Theorem (CLT)
- Polling


## Magic Mushrooms

Suppose there will soon be a vote on whether to legalize the therapeutic use of "magic mushrooms" and we want to conduct a poll to try to predict what will happen.

Poll to determine the fraction $p$ of the population expected to vote in favor.

- Call up a random sample of $n$ people to ask their opinion
- Report the empirical fraction

Questions

- Is this a good estimate?
- How to choose $n$ ?



## Polling Accuracy

Often see claims that say
"Our poll found $80 \%$ support. This poll is accurate to within 5\% with 98\% probability*"

Will unpack what this means and how they sample enough people to know this is true.

* When it is $95 \%$ this is sometimes written as " 19 times out of 20 "


## Formalizing Polls

Population size $N$, true fraction of voting in favor $p$, sample size $n$.
Problem: We don't know $p$, want to estimate it

## Polling Procedure

for $i=1, \ldots, n$ :

1. Pick uniformly random person to call (prob: $1 / N$ )
2. Ask them how they will vote

$$
X_{i}=\left\{\begin{array}{lr}
1, & \text { voting in favor } \\
0, & \text { otherwise }
\end{array}\right.
$$

Report our estimate of $p: \quad \bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}$

## Formalizing Polls

Population size $N$, true fraction of voting in favor $p$, sample size $n$.

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## Polling Procedure

 for $i=1, \ldots, n$ :
## What type of r.v. is $X_{i}$ ?

Poll:

|  | Type | $\mathbb{E}\left[X_{i}\right]$ | $\operatorname{Var}\left(X_{i}\right)$ |
| :--- | :--- | :--- | :--- |
| a. Bernoulli | $p$ | $p(1-p)$ |  |
| b. Bernoulli | $p$ | $p^{2}$ |  |
| c. Geometric | $p$ | $\frac{1-p}{p^{2}}$ |  |
| d. Binomial | $\mathrm{n} p$ | $n p(1-p)$ |  |

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2. Ask them how they will vote

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X_{i}=\left\{\begin{array}{lr}
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0, & \text { otherwise }
\end{array}\right.
$$

Report our estimate of $p$ :

$$
\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}
$$

## Random Variables

What type of r.v. is $X_{i}$ ?

| Type | $\mathbb{E}\left[X_{i}\right]$ | $\operatorname{Var}\left(X_{i}\right)$ |
| :---: | :---: | :---: |
| Bernoulli | $p$ | $p(1-p)$ |

What about $\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}$ ?

| Poll: |  |  |
| :--- | :--- | :--- |
|  | $\mathbb{E}[\bar{X}]$ | $\operatorname{Var}(\bar{X})$ |
| a. | $n p$ | $n p(1-p)$ |
| b. $p$ | $p(1-p)$ |  |
| c. $p$ | $p(1-p) / n$ |  |
| d. $p / n$ | $p(1-p) / n$ |  |

## Central Limit Theorem

With i.i.d random variables $X_{1}, X_{2}, \ldots, X_{n}$ where $\mathbb{E}\left[X_{i}\right]=\mu$ and $\operatorname{Var}\left(X_{i}\right)=\sigma^{2}$

As $n \rightarrow \infty$,

$$
\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i} \rightarrow \mathcal{N}\left(p, \frac{p(1-p)}{n}\right)
$$

## Roadmap: Bounding Error

Goal: Find the value of $n$ such that $98 \%$ of the time, the estimate $\bar{X}$ is within $5 \%$ of the true $p$

Get good estimate if $\bar{X}$ lands in this region


How does the value of $n$ play a role?

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Question: for what $n$ is $P(|\bar{X}-p|>0.05) \leq 0.02$

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## Recall!!

## For what $a$ is

$$
P(Z>a) \leq 0.01 ?
$$

## Equivalently

$P(Z \leq a) \geq 0.99$ ?
For any $a \geq 2.33$
$P(Z>a) \leq 0.01$.

| $\Phi$ Table: $\mathbb{P}(Z \leq z)$ when $Z \sim \mathcal{N}(0,1)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| 0.0 | 0.5 | 0.50399 | 0.50798 | 0.51197 | 0.51595 | 0.51994 | 0.52392 | 0.5279 | 0.53188 | 0.53586 |
| 0.1 | 0.53983 | 0.5438 | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 0.57535 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
| 0.3 | 0.61791 | 0.62172 | 0.62552 | 0.6293 | 0.63307 | 0.63683 | 0.64058 | 0.64431 | 0.64803 | 0.65173 |
| 0.4 | 0.65542 | 0.6591 | 0.66276 | 0.6664 | 0.67003 | 0.67364 | 0.67724 | 0.68082 | 0.68439 | 0.68793 |
| 0.5 | 0.69146 | 0.69497 | 0.69847 | 0.70194 | 0.7054 | 0.70884 | 0.71226 | 0.71566 | 0.71904 | 0.7224 |
| 0.6 | 0.72575 | 0.72907 | 0.73237 | 0.73565 | 0.73891 | 0.74215 | 0.74537 | 0.74857 | 0.75175 | 0.7549 |
| 0.7 | 0.75804 | 0.76115 | 0.76424 | 0.7673 | 0.77035 | 0.77337 | 0.77637 | 0.77935 | 0.7823 | 0.78524 |
| 0.8 | 0.78814 | 0.79103 | 0.79389 | 0.79673 | 0.79955 | 0.80234 | 0.80511 | 0.80785 | 0.81057 | 0.81327 |
| 0.9 | 0.81594 | 0.81859 | 0.82121 | 0.82381 | 0.82639 | 0.82894 | 0.83147 | 0.83398 | 0.83646 | 0.83891 |
| 1.0 | 0.84134 | 0.84375 | 0.84614 | 0.84849 | 0.85083 | 0.85314 | 0.85543 | 0.85769 | 0.85993 | 0.86214 |
| 1.1 | 0.86433 | 0.8665 | 0.86864 | 0.87076 | 0.87286 | 0.87493 | 0.87698 | 0.879 | 0.881 | 0.88298 |
| 1.2 | 0.88493 | 0.88686 | 0.88877 | 0.89065 | 0.89251 | 0.89435 | 0.89617 | 0.89796 | 0.89973 | 0.90147 |
| 1.3 | 0.9032 | 0.9049 | 0.90658 | 0.90824 | 0.90988 | 0.91149 | 0.91309 | 0.91466 | 0.91621 | 0.91774 |
| 1.4 | 0.91924 | 0.92073 | 0.9222 | 0.92364 | 0.92507 | 0.92647 | 0.92785 | 0.92922 | 0.93056 | 0.93189 |
| 1.5 | 0.93319 | 0.93448 | 0.93574 | 0.93699 | 0.93822 | 0.93943 | 0.94062 | 0.94179 | 0.94295 | 0.94408 |
| 1.6 | 0.9452 | 0.9463 | 0.94738 | 0.94845 | 0.9495 | 0.95053 | 0.95154 | 0.95254 | 0.95352 | 0.95449 |
| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608 | 0.96164 | 0.96246 | 0.96327 |
| 1.8 | 0.96407 | 0.96485 | 0.96562 | 0.96638 | 0.96712 | 0.96784 | 0.96856 | 0.96926 | 0.96995 | 0.97062 |
| 1.9 | 0.97128 | 0.97193 | 0.97257 | 0.9732 | 0.97381 | 0.97441 | 0.975 | 0.97558 | 0.97615 | 0.9767 |
| 2.0 | 0.97725 | 0.97778 | 0.97831 | 0.97882 | 0.97932 | 0.97982 | 0.9803 | 0.98077 | 0.98124 | 0.98169 |
| 2.1 | 0.98214 | 0.98257 | 0.983 | 0.98341 | 0.98382 | 0.98422 | 0.98461 | 0.985 | 0.98537 | 0.98574 |
| 2.2 | 0.9861 | 0.98645 | 0.98679 | 008713 | 0.98745 | 0.98778 | 0.98809 | 0.9884 | 0.9887 | 0.98899 |
| 2.3 | 0.98928 | 0.98956 | 0.98983 | 0.9901 | 0.99036 | 0.99061 | 0.99086 | 0.99111 | 0.99134 | 0.99158 |
| 2.4 | 0.9918 | 0.99202 | 0.99224 | 0.99245 | 0.99266 | 0.99286 | 0.99305 | 0.99324 | 0.99343 | 0.99361 |
| 2.5 | 0.99379 | 0.99396 | 0.99413 | 0.9943 | 0.99446 | 0.99461 | 0.99477 | 0.99492 | 0.99506 | 0.9952 |
| 2.6 | 0.99534 | 0.99547 | 0.9956 | 0.99573 | 0.99585 | 0.99598 | 0.99609 | 0.99621 | 0.99632 | 0.99643 |
| 2.7 | 0.99653 | 0.99664 | 0.99674 | 0.99683 | 0.99693 | 0.99702 | 0.99711 | 0.9972 | 0.99728 | 0.99736 |
| 2.8 | 0.99744 | 0.99752 | 0.9976 | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999 |



## Recap



Goal: Find the value of $n$ such that $98 \%$ of the time, the estimate $\bar{X}$ is within $5 \%$ of the true $p$

1. Define question. For what $n$ is $P(|\bar{X}-p|>0.05) \leq 0.02$
2. Apply CLT: By CLT $\bar{X} \rightarrow \mathcal{N}\left(\mu, \sigma^{2}\right)$ where $\mu=p$ and $\sigma^{2}=$ $p(1-p) / n$
3. Convert to a standard normal. Specifically, define $Z=$ $\frac{\bar{x}-\mu}{\sigma}=\frac{\bar{x}-p}{\sigma}$. Then, by the CLT $Z \rightarrow \mathcal{N}(0,1)$
4. Solve for $n$

## In more detail

1. The question: for what $n$ is $P(|\bar{X}-p|>0.05) \leq 0.02$
2. By CLT $\bar{X} \rightarrow \mathcal{N}\left(\mu, \sigma^{2}\right)$ where $\mu=p$ and $\sigma^{2}=p(1-p) / n$
3. $P(|\bar{X}-p|>0.05)=P(\bar{X}-p>0.05)+P(\bar{X}-p<-0.05)$
4. $P(|\bar{X}-p|>0.05) \leq 0.02$ equivalent to $P(\bar{X}-p>0.05) \leq 0.01$
5. Equivalent to $P\left(\begin{array}{c}\bar{x}-p\end{array} \frac{\sqrt{p(1-p)}}{\sqrt{n}}>0.05 / \frac{\sqrt{p(1-p)}}{\sqrt{n}}\right) \leq 0.01$
6. Equivalent to $P\left(Z>\frac{0.05}{\sqrt{p(1-p)} / \sqrt{n}}\right) \leq 0.01$

## Recall!!

## For what $a$ is

$$
P(Z>a) \leq 0.01 ?
$$

For any $a \geq 2.33$
$P(Z>a) \leq 0.01$.

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| 2.7 | 0.99653 | 0.99664 | 0.99674 | 0.99683 | 0.99693 | 0.99702 | 0.99711 | 0.9972 | 0.99728 | 0.99736 |
| 2.8 | 0.99744 | 0.99752 | 0.9976 | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999 |

So we know that if $\frac{0.05}{\sqrt{p(1-p)} / \sqrt{n}} \geq 2.33$


Then $P\left(Z>\frac{0.05}{\sqrt{p(1-p)} / \sqrt{n}}\right) \leq 0.01$.
So just need to solve this: $\sqrt{n} \geq \frac{2.33 \sqrt{p(1-p)}}{0.05}$
Since we don't know $p$, we will just take $n$ so it works for all $p$. RHS is largest with $p=0.5$. So we take
$\sqrt{n} \geq \frac{2.33 \cdot 0.5}{0.05}$ or $\sqrt{n} \geq 23.3$
Then $n \geq 543 \geq(23.3)^{2}$ would be good enough.

- Since only have $Z \rightarrow \mathcal{N}(0,1)$ so there is some loss due to approximation error.
- So should increase $n$ a bit more to be safe.

Zooming back out: we found an approximate"confidence interval"

We are trying to estimate some parameter (e.g. p). We output an estimator $\bar{X}$ such that $P(|\bar{X}-p|>\epsilon) \leq \delta$ for some $(\epsilon, \delta)$.

- Often found using CLT
- We say that we are $(1-\delta) * 100 \%$ confident that the result of our poll $(\bar{X})$ is an accurate estimate of $p$ to within $\epsilon^{*} 100 \%$ percent.
- In our example, $(\epsilon=0.05, \delta=0.02)$.


## Idealized Polling

So far, we have been discussing "idealized polling". Real life is normally not so nice $:$ :

Assumed we can sample people uniformly at random, not really possible in practice

- Not everyone responds
- Response rates might differ in different groups
- Will people respond truthfully?

Makes polling in real life much more complex than this idealized model!

