## CSE 312 <br> Foundations of Computing II

Lecture 6: Chain Rule and Independence

TWT PAUL G. ALLEN SCHOOL OF COMPUTER SCIENCE \& ENGINEERING
slido.com/4171468
for polls and anonymous questions

## Thank you for your feedback!!!

- Several people mentioned that I was going too fast.
- Slow me down! Ask questions!!! That's your job!!
- Watch Summer 2020 videos before class (at half speed)
- Do the reading before class.
- If you want more practice
- Do all the section problems!
- Problems in all three readings.
- MIT "Mathematics for Computer Science" 6.042J (sections on counting \& probability)
- Get the book "A First Course in Probability" by Sheldon Ross


## Agenda

- Recap
- Chain Rule
- Independence
- Conditional independence
- Infinite process


## Review Conditional \& Total Probabilities

- Conditional Probability

$$
P(B \mid A)=\frac{P(A \cap B)}{P(A)}
$$

- Bayes Theorem

$$
P(A \mid B)=\frac{P(B \mid A) P(A)}{P(B)} \quad \text { if } P(A) \neq 0, P(B) \neq 0
$$

- Law of Total Probability

$$
E_{1}, \ldots, E_{n} \text { partition } \Omega
$$



$$
P(F)=\sum_{i=1}^{n} P\left(F \cap E_{i}\right)=\sum_{i=1}^{n} P\left(F \mid E_{i}\right) P\left(E_{i}\right)
$$

## Bayes Theorem with Law of Total Probability

Bayes Theorem with LTP: Let $E_{1}, E_{2}, \ldots, E_{n}$ be a partition of the sample space, and $F$ and event. Then,

$$
P\left(E_{1} \mid F\right)=\frac{P\left(F \mid E_{1}\right) P\left(E_{1}\right)}{P(F)}=\frac{P\left(F \mid E_{1}\right) P\left(E_{1}\right)}{\sum_{i=1}^{n} P\left(F \mid E_{i}\right) P\left(E_{i}\right)}
$$

Simple Partition: In particular, if $E$ is an event with non-zero probability, then

$$
P(E \mid F)=\frac{P(F \mid E) P(E)}{P(F \mid E) P(E)+P\left(F \mid E^{C}\right) P\left(E^{C}\right)}
$$

## Example - Zika Testing



A disease caused by Zika virus that's spread through mosquito bites.

Usually no or mild symptoms (rash); sometimes severe symptoms (paralysis).

During pregnancy: may cause birth defects.

Suppose you took a Zika test, and it returns "positive", what is the likelihood that you actually have the disease?

- Tests for diseases are rarely $100 \%$ accurate.


## Example - Zika Testing

Z have Zika
$T$ test positive

Suppose we know the following Zika stats

- A test is $98 \%$ effective at detecting Zika ("true positive") $\quad P(T \mid Z)=0.98$
- However, the test may yield a "false positive" $1 \%$ of the time $P\left(T \mid Z^{c}\right)=0.01$
- $0.5 \%$ of the US population has Zika. $P(Z)=0.005$

What is the probability you have Zika (event $Z$ ) given that you test positive (event $T$ )?

By Bayes Rule, $P(Z \mid T)=\frac{P(T \mid Z) P(Z)}{P(T)}$

By the Law of Total Probability, $P(T)=P(T \mid Z) P(Z)+P\left(T \mid Z^{c}\right) P\left(Z^{c}\right)$

## Philosophy - Updating Beliefs

While it's not $98 \%$ that you have the disease, your beliefs changed significantly

Z = you have Zika
T = you test positive for Zika


Prior: $\mathrm{P}(\mathrm{Z})$


Posterior: $\mathrm{P}(\mathrm{Z} \mid \mathrm{T})$

## Example - Zika Testing

Have zika blue, don't pink
What is the probability you have Zika (event $Z$ ) given that you test positive (event $T$ ).
$P(T \mid Z)=$
$P\left(T \mid Z^{C}\right)=$
$P(Z)=$


Suppose we had 1000 people:

- 5 have Zika and all test positive
- 985 do not have Zika and test negative
- 10 do not have Zika and test positive


## Example - Zika Testing

Have zika blue, don't pink
Picture below gives us the following Zika stats

- A test is $100 \%$ effective at detecting Zika ("true positive"). $\quad P(T \mid Z)=5 / 5=1$
- However, the test may yield a "false positive" $1 \%$ of the time $\quad P\left(T \mid Z^{C}\right)=10 / 995$
- $0.5 \%$ of the US population has Zika.

$$
P(Z)=\frac{5}{1000}=0.005
$$

What is the probability you have Zika (event $Z$ ) given that you test positive (event $T$ )?


Suppose we had 1000 people:

- 5 have Zika and all test positive
- 985 do not have Zika and test negative
- 10 do not have Zika and test positive


## Example - Zika Testing

Have zika blue, don't pink
Picture below gives us the following Zika stats

- A test is $100 \%$ effective at detecting Zika ("true positive"). $\quad P(T \mid Z)=5 / 5=1$
- However, the test may yield a "false positive" $1 \%$ of the time $\quad P\left(T \mid Z^{C}\right)=10 / 995$
- $0.5 \%$ of the US population has Zika. $5 \%$ have it.

$$
P(Z)=\frac{995}{1000}=0.005
$$

What is the probability you have Zika (event Z) given that you test positive (event $T$ )?


Suppose we had 1000 people:

- 5 have Zika and and test positive
- 985 do not have Zika and test negative
- 10 do not have Zika and test positive

$$
\frac{5}{5+10}=\frac{1}{3} \approx 0.33
$$

Suppose we know the following Zika stats

- A test is $98 \%$ effective at detecting Zika ("true positive") $\quad P(T \mid Z)=0.98$
- However, the test may yield a "false positive" $1 \%$ of the time $P\left(T \mid Z^{C}\right)=0.01$
- $0.5 \%$ of the US population has Zika. $\quad P(Z)=0.005$

What is the probability you test negative (event $T^{c}$ ) give you have Zika (event $Z$ )?

## Conditional Probability Define a Probability Space

The probability conditioned on $A$ satisfies the required axioms.

Example. $\mathbb{P}\left(\mathcal{B}^{c} \mid \mathcal{A}\right)=1-\mathbb{P}(\mathcal{B} \mid \mathcal{A})$

## Conditional Probability Define a Probability Space

The probability conditioned on $A$ follows the same properties as (unconditional) probability.

Example. $\mathbb{P}\left(\mathcal{B}^{c} \mid \mathcal{A}\right)=1-\mathbb{P}(\mathcal{B} \mid \mathcal{A})$

Formally. $(\Omega, \mathbb{P})$ is a probability space $+\mathbb{P}(\mathcal{A})>0$


Axiom 1 (Non-negativity): $P(E) \geq 0$.
Axiom 2 (Normalization): $P(\Omega)=1$
Axiom 3 (Countable Additivity): If $E$ and $F$ are mutually exclusive, then $P(E \cup F)=P(E)+P(F)$

## Agenda

- Recap
- Chain Rule
- Independence
- Conditional independence
- Infinite process


## Chain Rule

$$
\mathbb{P}(\mathcal{B} \mid \mathcal{A})=\frac{\mathbb{P}(\mathcal{A} \cap \mathcal{B})}{\mathbb{P}(\mathcal{A})} \quad \square \mathbb{P}(\mathcal{A}) \mathbb{P}(\mathcal{B} \mid \mathcal{A})=\mathbb{P}(\mathcal{A} \cap \mathcal{B})
$$

## Chain Rule

$$
\mathbb{P}(\mathcal{B} \mid \mathcal{A})=\frac{\mathbb{P}(\mathcal{A} \cap \mathcal{B})}{\mathbb{P}(\mathcal{A})} \quad \square \mathbb{P}(\mathcal{A}) \mathbb{P}(\mathcal{B} \mid \mathcal{A})=\mathbb{P}(\mathcal{A} \cap \mathcal{B})
$$

Theorem. (Chain Rule) For events $\mathcal{A}_{1}, \mathcal{A}_{2}, \ldots, \mathcal{A}_{n}$,

$$
\mathbb{P}\left(\mathcal{A}_{1} \cap \cdots \cap \mathcal{A}_{n}\right)=\mathbb{P}\left(\mathcal{A}_{1}\right) \cdot \mathbb{P}\left(\mathcal{A}_{2} \mid \mathcal{A}_{1}\right) \cdot \mathbb{P}\left(\mathcal{A}_{3} \mid \mathcal{A}_{1} \cap \mathcal{A}_{2}\right)
$$

$$
\cdots \mathbb{P}\left(\mathcal{A}_{n} \mid \mathcal{A}_{1} \cap \mathcal{A}_{2} \cap \cdots \cap \mathcal{A}_{n-1}\right)
$$

An easy way to remember: We have $n$ tasks and we can do them sequentially, conditioning on the outcome of previous tasks

## Chain Rule Example

Have a Standard 52-Card Deck. Shuffle It, and draw the top 3 cards in order. (uniform probability space).


$$
)=P(A \cap B \cap C) ?
$$

A: Ace of Spades First
B: 10 of Clubs Second
C: 4 of Diamonds Third

## Chain Rule Example

Have a Standard 52-Card Deck. Shuffle It, and draw the top 3 cards in order. (uniform probability space).


## Agenda

- Recap
- Chain Rule
- Independence
- Conditional independence
- Infinite process


## Independence

Definition. Two events $\mathcal{A}$ and $\mathcal{B}$ are (statistically) independent if

$$
\mathbb{P}(\mathcal{A} \cap \mathcal{B})=\mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B})
$$

Alternatively,

- If $\mathbb{P}(\mathcal{A}) \neq 0$, equivalent to $\mathbb{P}(\mathcal{B} \mid \mathcal{A})=\mathbb{P}(B)$
- If $\mathbb{P}(\mathcal{B}) \neq 0$, equivalent to $\mathbb{P}(\mathcal{A} \mid \mathcal{B})=\mathbb{P}(\mathcal{A})$
"The probability that $\mathcal{B}$ occurs after observing $\mathcal{A}$ " -- Posterior
= "The probability that $\mathcal{B}$ occurs" -- Prior


## Example -- Independence

Toss a coin 3 times. Each of 8 outcomes equally likely.

- $A=\{$ at most one $T\}=\{\mathrm{HHH}, \mathrm{HHT}, \mathrm{HTH}, \mathrm{THH}\}$
- $B=\{$ at most 2 Heads $\}=\{H H H\}^{c}$ Independent?

$$
\mathbb{P}(\mathcal{A} \cap \mathcal{B}) \stackrel{?}{=} \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B})
$$

## Multiple Events - Mutual Independence

Definition. Events $A_{1}, \ldots, A_{n}$ are mutually independent if for every non-empty subset $I \subseteq\{1, \ldots, n\}$, we have

$$
P\left(\bigcap_{i \in I} A_{i}\right)=\prod_{i \in I} P\left(A_{i}\right) .
$$

## Example - Network Communication

Each link works with the probability given, independently
i.e., mutually independent events $A, B, C, D$ with

$$
\begin{aligned}
& P(A)=p \\
& P(B)=q \\
& P(C)=r \\
& P(D)=s
\end{aligned}
$$



## Example - Network Communication

Each link works with the probability given, independently

$$
\begin{aligned}
& P(A)=p \\
& P(B)=q \\
& P(C)=r \\
& P(D)=s
\end{aligned}
$$

i.e., mutually independent events $A, B, C, D$

What is $P(1-4$ connected $)$ ?


## Example - Network Communication

If each link works with the probability given, independently: What's the probability that nodes 1 and $\mathbf{4}$ can communicate?
$P(1-4$ connected $)=P((A \cap B) \cup(C \cap D))$

$$
=P(A \cap B)+P(C \cap D)-P(A \cap B \cap C \cap D)
$$

$P(A \cap B)=P(A) \cdot P(B)=p q$
$P(C \cap D)=P(C) \cdot P(D)=r s$
$P(A \cap B \cap C \cap D)$
$=P(A) \cdot P(B) \cdot P(C) \cdot P(D)=$ pqrs

$$
P(1-4 \text { connected })=p q+r s-p q r s
$$

## Independence - Another Look

Definition. Two events $A$ and $B$ are (statistically) independent if

$$
P(A \cap B)=P(A) \cdot P(B) .
$$

"Equivalently." $P(A \mid B)=P(A)$.

Events generated independently $\boldsymbol{\rightarrow}$ their probabilities satisfy independence
But events can be independent without being generated by independent processes.


This can be counterintuitive!

## Sequential Process


$P(\mathbf{R} \mid 3 \mathrm{R} 3 \mathrm{~B})=$

Setting: An urn contains:

- 3 red and 3 blue balls w/ probability $3 / 5$
- 3 red and 1 blue balls $w /$ probability $1 / 10$
- 5 red and 7 blue balls $w /$ probability 3/10 We draw a ball at random from the urn.

Are $R$ and $3 R 3 B$ independent?
$P(\mathbf{R})=$

Sequential Process


Are $R$ and $3 R 3 B$ independent?

$$
P(\mathbf{R} \mid 3 \mathrm{R} 3 \mathrm{~B})=\frac{1}{2}
$$

Sequential Process


Are R and 3R3B independent?

## Ball drawn

Setting: An urn contains:

- 3 red and 3 blue balls w/ probability 3/5
- 3 red and 1 blue balls $w /$ probability $1 / 10$
- 5 red and 7 blue balls $w /$ probability 3/10 We draw a ball at random from the urn.

$$
P(\mathbf{R})=\frac{3}{5} \times \frac{1}{2}+\frac{1}{10} \times \frac{3}{4}+\frac{3}{10} \times \frac{5}{12}=\frac{1}{2}
$$

$$
P(\mathbf{R} \mid 3 \mathrm{R} 3 \mathrm{~B})=\frac{1}{2}
$$

$$
\text { Independent! } P(\mathbf{R})=P(\mathbf{R} \mid 3 \mathbf{R} 3 \mathbf{B})
$$



Often probability space $(\Omega, \mathbb{P})$ is defined using independence

## Example - Biased coin

We have a biased coin comes up Heads with probability 2/3; Each flip is independent of all other fips. Suppose it is tossed 3 times.
$\mathbb{P}(H H H)=$
$\mathbb{P}(T T T)=$
$\mathbb{P}(H T T)=$

## Example - Biased coin

We have a biased coin comes up Heads with probability 2/3, independently of other flips. Suppose it is tossed 3 times.
$\mathbb{P}(2$ heads in 3 tosses $)=$
A) $(2 / 3)^{2} 1 / 3$
B) $2 / 3$
C) $3(2 / 3)^{2} 1 / 3$
D) $(1 / 3)^{2}$

## Independence as an assumption

- People often assume it without justification
- Example: A skydiver has two chutes
$A$ : event that the main chute doesn't open $\quad P(A)=0.02$
$B$ : event that the back-up doesn't open $\quad P(B)=0.1$
- What is the chance that at least one opens assuming independence?


## Independence as an assumption

- People often assume it without justification
- Both chutes could fail because of the same rare event e.g., freezing rain.


## Agenda

- Recap
- Chain Rule
- Independence
- Conditional independence
- Infinite process


## Conditional Independence

Definition. Two events $A$ and $B$ are independent conditioned on $C$ if

$$
P(C) \neq 0 \text { and } P(A \cap B \mid C)=P(A \mid C) \cdot P(B \mid C) .
$$

- If $P(A \cap C) \neq 0$, equivalent to $P(B \mid A \cap C)=P(B \mid C)$
- If $P(B \cap C) \neq 0$, equivalent to $P(A \mid B \cap C)=P(A \mid C)$

Contrast to Plain Independence. Two events $A$ and $B$ are independent if

$$
P(A \cap B)=P(A) \cdot P(B) .
$$

- If $P(A) \neq 0$, equivalent to $P(B \mid A)=P(B)$
- If $P(B) \neq 0$, equivalent to $P(A \mid B)=P(A)$


## Example - Throwing Dice

Suppose that Coin 1 has probability of heads 0.3
and Coin 2 has probability of head 0.9 .
We choose one of the coins randomly with equal probability and flip that coin 2 times independently. What is the probability we get all heads?

$$
\begin{array}{cc}
P(H H)=P\left(H H \mid C_{1}\right) \cdot P\left(C_{1}\right)+P\left(H H \mid C_{2}\right) \cdot P\left(C_{2}\right) & \text { Law of Total Probability } \\
=P\left(H \mid C_{1}\right)^{2} P\left(C_{1}\right)+P\left(H \mid C_{2}\right)^{2} P\left(C_{2}\right) & \text { Conditional Independence } \\
=0.3^{2} \cdot 0.5+0.9^{2} \cdot 0.5=0.45 & C_{i}=\operatorname{coin} i \text { was selected }
\end{array}
$$

## Example - Throwing Dice

Suppose that Coin 1 has probability of heads 0.3 and Coin 2 has probability of head 0.9.
We choose one coin randomly with equal probability and flip that coin 2 times independently. What is the probability we get all heads?

$$
\begin{aligned}
& P\left(H H \mid C_{1}\right)=P\left(H \mid C_{1}\right)^{2} \\
& P(H H) \neq P(H)^{2}
\end{aligned}
$$

$$
\begin{aligned}
& P(H H)=P\left(H \mid C_{1}\right)^{2} P\left(C_{1}\right)+P\left(H \mid C_{2}\right)^{2} P\left(C_{2}\right)=0.3^{2} \cdot 0.5+0.9^{2} \cdot 0.5=0.45 \\
& P(H)=P\left(H \mid C_{1}\right) \cdot P\left(C_{1}\right)+P\left(H \mid C_{2}\right) \cdot P\left(C_{2}\right) \doteq 0.6 \quad P(H H)=0.36
\end{aligned}
$$

## Agenda

- Recap
- Chain Rule
- Independence
- Conditional independence
- Infinite process

Often probability space $(\Omega, P)$ is given implicitly via sequential process

- Experiment proceeds in $n$ sequential steps, each step follows some local rules defined by the chain rule and independence
- Natural extension: Allows for easy definition of experiments where $|\Omega|=\infty$


## Example - Throwing A Die Repeatedly

Alice and Bob are playing the following game.
A 6-sided die is thrown, and each time it's thrown, regardless of the history, it is equally likely to show any of the six numbers

If it shows $1,2 \rightarrow$ Alice wins.
If it shows $3 \rightarrow$ Bob wins.
Otherwise, play another round
What is $\operatorname{Pr}\left(\right.$ Alice wins on $1^{\text {st }}$ round $)=$ $\operatorname{Pr}\left(\right.$ Alice wins on $2^{\text {st }}$ round $)=$ $\operatorname{Pr}\left(\right.$ Alice wins on $i^{\text {th }}$ round $)=$ ? $\operatorname{Pr}($ Alice wins $)=$ ?

## Sequential Process - defined in terms of independence

A 6-sided die is thrown, and each time it's thrown, regardless of the history, it is equally likely to show any of the six numbers

Local Rules: In each round, toss a die

- If it shows $1,2 \rightarrow$ Alice wins
- If it shows $3 \rightarrow$ Bob wins
- Else, play another round
$\operatorname{Pr}($ Alice wins on $i$ th round $\mid$ nobody won in rounds $1 . . i-1)=1 / 3$


## Sequential Process - Example



$\mathcal{A}_{2}$

Local Rules: In each round

- If it shows $1,2 \rightarrow$ Alice wins
- If it shows $3 \rightarrow$ Bob wins
- Else, play another round

Events:

- $\mathcal{A}_{i}=$ Alice wins in round $i$
- $\mathcal{N}_{i}=$ nobody wins in round $i$


## Sequential Process - Example

## Events:

- $\mathcal{A}_{i}=$ Alice wins in round $i$
- $\mathcal{N}_{i}=$ nobody wins in rounds $1 . . i$
$\mathbb{P}\left(\mathcal{A}_{2}\right)=\mathbb{P}\left(\mathcal{N}_{1} \cap \mathcal{A}_{2}\right)$


$$
\mathcal{N}_{2}
$$

$2^{\text {nd }}$ roll indep of $1^{\text {st }}$ roll

## Sequential Process - Example

## Events:

- $\mathcal{A}_{i}=$ Alice wins in round $i$
- $\mathcal{N}_{i}=$ nobody wins in rounds $1 . . i$

$$
\begin{aligned}
\mathbb{P}\left(\mathcal{A}_{2}\right) & =\mathcal{P}\left(\mathcal{N}_{1} \cap \mathcal{A}_{2}\right) \\
& =\mathcal{P}\left(\mathcal{N}_{1}\right) \times \mathcal{P}\left(\mathcal{A}_{2} \mid \mathcal{N}_{1}\right) \\
& =\frac{1}{2} \times \frac{1}{3}=\frac{1}{6}
\end{aligned}
$$



The event $\mathcal{A}_{2}$ implies $\mathcal{N}_{1}$, and this means that $\mathcal{A}_{2} \cap \mathcal{N}_{1}=\mathcal{A}_{2}$ $2^{\text {nd }}$ roll indep of $1^{\text {st }}$ roll


## Sequential Process - Example



$$
\begin{aligned}
& \mathbb{P}\left(\mathcal{A}_{i}\right)=\mathcal{P}\left(\mathcal{N}_{1} \cap \mathcal{N}_{2} \cap \cdots \cap \mathcal{N}_{i-1} \cap \mathcal{A}_{i}\right) \\
& =\mathcal{P}\left(\mathcal{N}_{1}\right) \times \mathcal{P}\left(\mathcal{N}_{2} \mid \mathcal{N}_{1}\right) \times \mathcal{P}\left(\mathcal{N}_{3} \mid \mathcal{N}_{1} \cap \mathcal{N}_{2}\right) \\
& \quad \cdots \times \mathcal{P}\left(\mathcal{N}_{i-1} \mid \mathcal{N}_{1} \cap \mathcal{N}_{2} \cap \cdots \cap \mathcal{N}_{i-1}\right) \times \mathcal{P}\left(\mathcal{A}_{i} \mid \mathcal{N}_{1} \cap \mathcal{N}_{2} \cap \cdots \cap \mathcal{N}_{i-1}\right) \\
& \\
& \quad=\left(\frac{1}{2}\right)^{i-1} \times \frac{1}{3}
\end{aligned}
$$

## Sequential Process -- Example

$\mathcal{A}_{i}=$ Alice wins in round $i \quad \mathbb{P}\left(\mathcal{A}_{i}\right)=\left(\frac{1}{2}\right)^{i-1} \times \frac{1}{3}$
What is the probability that Alice wins?

## Sequential Process -- Example

$\mathcal{A}_{i}=$ Alice wins in round $i \quad \mathbb{P}\left(\mathcal{A}_{i}\right)=\left(\frac{1}{2}\right)^{i-1} \times \frac{1}{3}$
What is the probability that Alice wins?

$$
\begin{aligned}
& \mathbb{P}\left(\mathcal{A}_{1} \cup \mathcal{A}_{2} \cup \cdots\right)=\sum_{i=1}^{\infty} \mathbb{P}\left(\mathcal{A}_{i}\right) \\
& \sum_{i=1}^{\infty}\left(\frac{1}{2}\right)^{i-1} \times \frac{1}{3}=\frac{1}{3} \times 2=\frac{2}{3}
\end{aligned}
$$

$$
\text { All } \mathcal{A}_{i} \text { 's are disjoint. }
$$

$$
\text { Fact. If }|x|<1 \text {, then } \sum_{i=0}^{\infty} x^{i}=\frac{1}{1-x} .
$$



