

Today:

- + signal detect. theory
- + decoding from multi. neurons
- + inf. theory

Review of Bayes' Law:

$$P(\vec{F}, \vec{S}) = P(\vec{F}|\vec{S})P(\vec{S}) = P(\vec{S}|\vec{F})P(\vec{F})$$

posterior likelihood joint conditional marginal
 $P(\vec{S}|\vec{F}) = \frac{P(\vec{F}|\vec{S})P(\vec{S})}{P(\vec{F})}$ prior

Signal Detection Theory

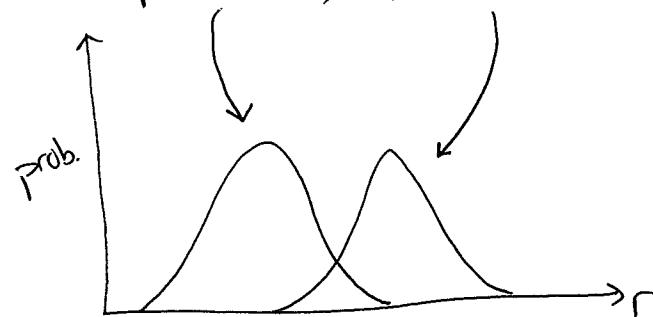
Goal: given a ~~brain~~ neural response to a binary stim,
determine if stim is 0 or 1

E.g., $S = 0$ no tiger
 $= 1$ tiger

$r \in [0, r_{\max}]$ firing rate

Relevant quantities: prior: $p(S=0)$, $p(S=1) = 1 - p(S=0)$

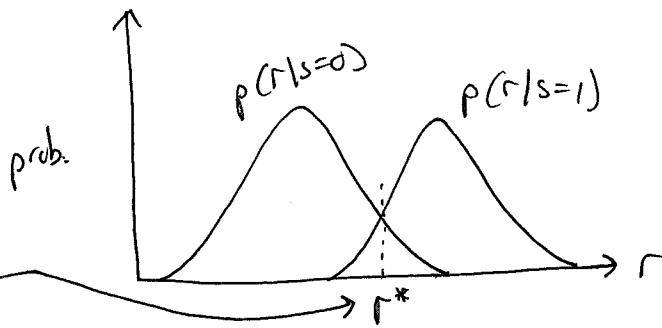
likelihood: $p(r|S=0)$, $p(r|S=1)$



Specific goal: figure out rule for determining if $S=0$ or 1 given r ,

i.e., choose threshold r^* s.t. if $r > r^*$ guess 1
 $\quad \quad \quad$ if $r < r^*$ guess 0

Simplest $s|b$:



Max. likelihood (ML): choose 1 when $p(r|s=1) > p(r|s=0)$
0 otherwise $\rightarrow r^*$ is $s|b$ to $p(r|s=1) = p(r|s=0)$

false pos. rate = ~~$p(s=0)p(\text{choose } 1|s=0)$~~

$$p(s=0)p(\text{choose } 1|s=0) = p(s=0)p(r > r^* | s=0)$$
$$= p(s=0) \int_{r^*}^{\infty} p(r|s=0) dr$$

false neg. rate = $p(s=1)p(\text{choose } 0|s=1) = p(s=1)p(r < r^* | s=1)$

$$= p(s=1) \int_{-\infty}^{r^*} p(r|s=1) dr$$

problem: consider when ~~r = r*~~ $r = r^* \rightarrow$ most of the time $s=0$ (not trigger)

ML neglects prior

better $s|b$: choose 1 when $p(s=1|r) > p(s=0|r)$

Maximum a posteriori (MAP) i.e., $\frac{p(s=1)p(r|s=1)}{p(r)} > \frac{p(s=0)p(r|s=0)}{p(r)}$

this incorporates prior probability \rightarrow ~~p(s=1)~~ if $p(s=1)$ very small,

then $p(r|s=1)$ must be very large (rel. to $p(r|s=0)$) to choose 1

$\hookrightarrow r^*$ is $s|b$ to:

$$p(s=1)p(r|s=1) = p(s=0)p(r|s=0)$$

$$\text{i.e., } \frac{p(r|s=1)}{p(r|s=0)} = \frac{p(s=1)}{p(s=0)}$$

②

alternative sl^b: consider cost of guessing incorrectly

$C(\text{choose } 1, s=0)$ ← guess tiger but no tiger (low cost)

$C(\text{choose } 0, s=1)$ ← guess no tiger but tiger! (high cost)

in this case:

given r , choose 1 when $E_s[C(\text{choose } 1, s)|r] < E_s[C(\text{choose } 0, s)|r]$

0 otherwise

$$\begin{aligned} E_s[C(\text{choose } 1, s)|r] &= p(s=0|r) C(\text{choose } 1|s=0) + p(s=1|r) C(\text{choose } 1|s=1) \\ &= p(s=0|r) C(\text{choose } 1|s=0) \\ &= \underbrace{p(s=0)p(r|s=0)}_{p(r)} C(\text{choose } 1|s=0) \end{aligned}$$

0 (assuming no cost for correct guess)

$$E_s[C(\text{choose } 0, s)|r] = \underbrace{p(s=1)p(r|s=1)}_{p(r)} C(\text{choose } 0|s=1)$$

i.e., choose 1 when $\frac{p(s=0)p(r|s=0)}{p(r)} C(\text{choose } 1|s=0) <$

$$\frac{p(s=1)p(r|s=1)}{p(r)} C(\text{choose } 0|s=1)$$

if $C(\text{choose } 0|s=1)$ very high, ~~more likely~~ more likely to choose 1
(and avoid tiger)

summary: 3 main things to consider:

likelihood, prior, cost

Decoding from multiple neurons

More general case: decode continuous stim from multiple neurons

↳ same general idea: $p(\vec{s})$ prior

$p(\vec{r}|\vec{s})$ likelihood

$p(\vec{s}|\vec{r})$ ~~posterior~~ posterior

$$\hookrightarrow = \frac{p(\vec{r}|\vec{s})p(\vec{s})}{p(\vec{r})}$$

Max likelihood:

$$\text{guess } \vec{s}^* = \underset{s}{\operatorname{argmax}} p(\vec{r}|\vec{s})$$

Max a posteriori:

$$\begin{aligned} \text{guess } \vec{s}^* &= \underset{s}{\operatorname{argmax}} p(\vec{s}|\vec{r}) = \underset{s}{\operatorname{argmax}} \frac{p(\vec{r}|\vec{s})p(\vec{s})}{p(\vec{r})} \\ &= \underset{s}{\operatorname{argmax}} p(\vec{r}|\vec{s})p(\vec{s}) \end{aligned}$$

Example: see slides 1-6

Slide 1 notes: "Gaussian" refers only to shape (this is tuning curve, not distribution)

$f_a(s)$ gives mean firing rate
as function of stim.

Slide 3 notes: need $p(\vec{r}|\vec{s})$ for both ML + MAP

$$\text{mean #spikes in } T = f_a(s)T$$

$$\text{observed #spikes in } T = r_a T$$

④

slide 4 notes: $\operatorname{argmax}_s p(\vec{r}|s) = \operatorname{argmax} \ln(p(\vec{r}|s))$

~~RECALL~~

$$\ln(p(\vec{r}|s)) = \ln\left(\prod_{a=1}^N \frac{(f_a(s)T)^{r_a T}}{(r_a T)!} \exp(-f_a(s)T)\right)$$

$$= \sum_{a=1}^N \ln((f_a(s)T)^{r_a T}) - \ln((r_a T)!) - f_a(s)T$$

$$= \sum_{a=1}^N r_a T \ln(f_a(s)T) - \underbrace{\sum_{a=1}^N \ln((r_a T)!)}$$

no dep. on s

$$- \underbrace{\sum_{a=1}^N f_a(s)T}_{\text{"const."}}$$

$$= \sum_{a=1}^N r_a + \ln(f_a(s)T) + C$$

const. w.r.t s

to find max likelihood:

$$\begin{aligned} 0 &= \frac{d}{ds} \ln(p(\vec{r}|s)) = \sum_{a=1}^N r_a T \frac{d}{ds} \ln(f_a(s)T) \\ &= \sum_{a=1}^N r_a T \frac{\frac{d}{ds} f_a(s)T}{f_a(s)T} \quad \longrightarrow \quad \begin{aligned} f_a(s) &= r_{\max} e^{-\frac{1}{2\sigma_a^2}(s-s_a)^2} \\ \frac{d}{ds} f_a(s) &= r_{\max} e^{-\frac{1}{2\sigma_a^2}(s-s_a)^2} \times \\ &\quad \left(-\frac{1}{\sigma_a^2} 2(s-s_a) \right) \\ &= -(s-s_a) \cancel{f_a(s)} \end{aligned} \\ &= -\sum_{a=1}^N \frac{r_a T (s-s_a)}{\sigma_a^2} \end{aligned}$$

$\sum_{a=1}^N \frac{r_a T}{\sigma_a^2} = \sum_{a=1}^N \frac{s_a r_a T}{\sigma_a^2} \rightarrow s^* = \frac{\sum_{a=1}^N s_a r_a / \sigma_a^2}{\sum_{a=1}^N r_a / \sigma_a^2}$

$$MAP: s^* = \underset{s}{\operatorname{argmax}} \ln(p(s|\vec{r}))$$

$$\ln(p(s|\vec{r})) = \cancel{\ln(p(\vec{r}))} + \ln(p(s))$$

$$\ln(p(\vec{r}|s) p(s)/p(r)) = \ln(p(\vec{r}|s)) + \ln(p(s)) - \ln(p(r))$$

$$= \sum_{a=1}^N r_a T \ln(f_a(s)^T) + C + \ln p(s) + D$$

can solve if $p(s)$ is Gaussian w/ mean s_{prior} & var Σ_{prior}
 (see slide 5 for $s^{(n)}$)

* often can't find analytical $s^{(n)}$ so do numerical optimization

Slide 5 notes: ML & MAP are equivalent when prior is constant

can also use Bayesian inference:

instead of $s^* = \underset{s}{\operatorname{argmax}} p(s|\vec{r})$

use $s^* = \int s dp(s|\vec{r}) s = E_{p(s|\vec{r})}[s]$

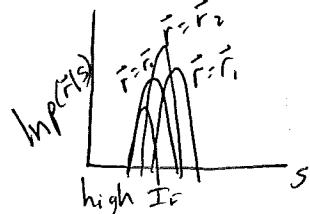
↳ minimizes least squares loss

usually hard to compute

$$E_s [(s - s^*)^2 | \vec{r}]$$

Slide 7 notes: averages are over $\vec{r}|s$

(6)



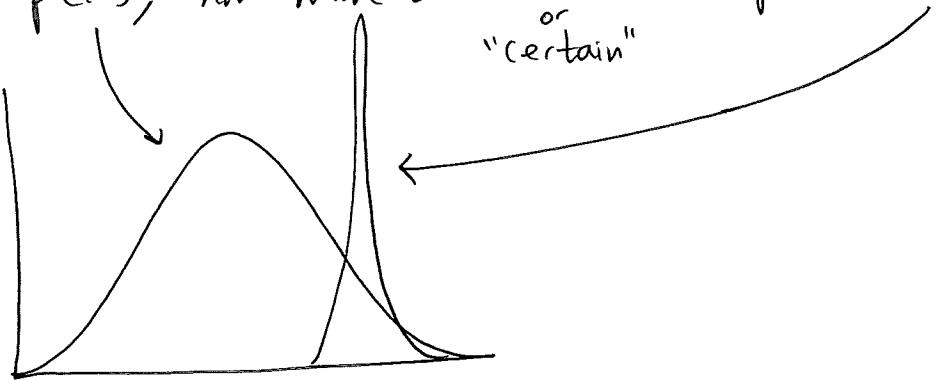
Information Theory

before: talked about estimating stim. from response

now: how much does response tell about stim

→ info theory provides more general way of quantifying this

Q: given prior $p(x)$, how much ~~is~~ "narrower" is posterior $p(x|y=y_i)$?



Entropy quantifies "uncertainty" of distribution

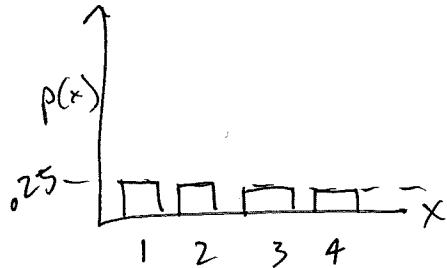
(more general than variance + has nice math. properties)

$$H(X) = - \sum_x p(x) \log(p(x))$$

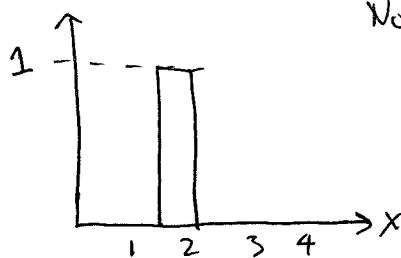
see "A mathematical theory of communication"

Shannon 1948

E.g.



$$H(X) = -4 \times 0.25 \times \log 0.25 = \log 4$$



Note: entropy is f^n of entire distribution

$$H(X) = -1 \times \log 1 = 0 \quad (\text{note: } 0 \log 0 \approx 0)$$

⑦

$$\text{Mutual Info : } MI(R,S) = H[S] - E_r[H[S|R=r]] = H[R] - E_s[H[R|S=s]]$$

Example:

$$S = 0 \text{ or } 1 \quad , \quad P(S=0) = .9 \quad , \quad P(S=1) = .1$$

(no tiger) (tiger)

response entropy

noise
entropy

$r = 0$ or 1 , $p(r=0) = .8$, $p(r=1) = .2$

$$P(r=1|s=0) = 1, P(r=1|s=1) = 0$$

$$p(r=0|s=0) = .9, \quad p(r=0|s=1) = .1$$

How much info does spike/nonspike contain about s ? no tiger tiger

$$H[R] = - \sum_r p(r) \log p(r) = - [.8 \log .8 + .2 \log .2] \approx .5$$

$$E_s[H(R|S=s)] = p(s=0)H(R|S=0) + p(s=1)H(R|S=1)$$

~~$\log \left(\frac{1}{\epsilon} \right)$~~

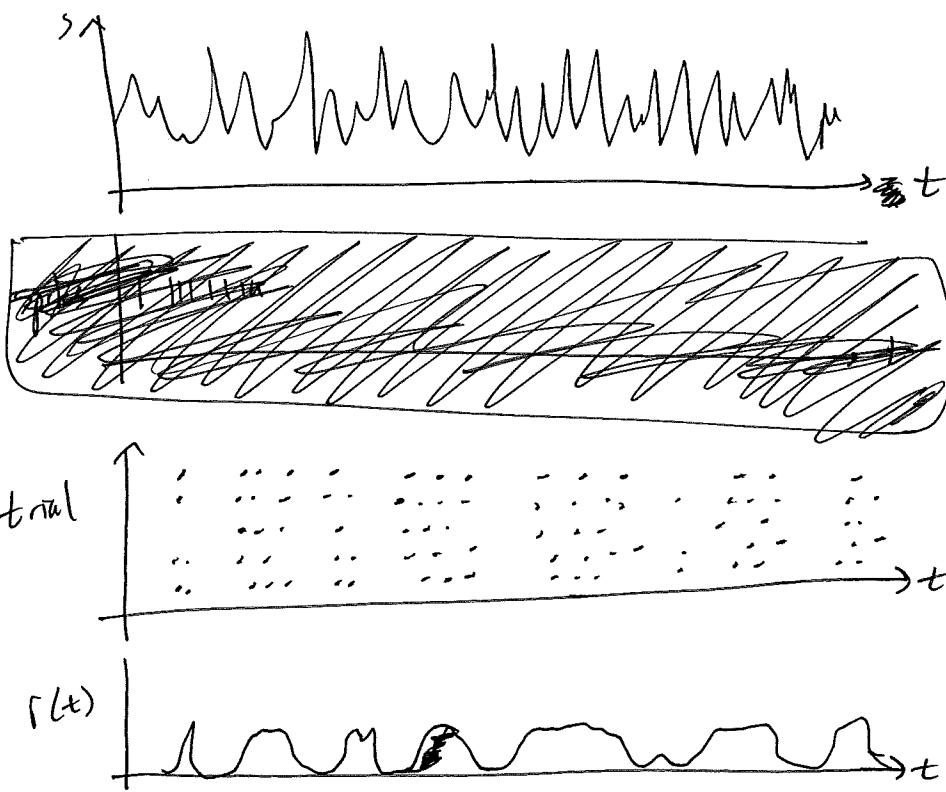
$$\begin{aligned}
 &= .9 \left[- (p(r=0|s=0) \log p(r=0|s=0) + p(r=1|s=0) \log p(r=1|s=0)) \right] \\
 &\quad + .1 \left[- (p(r=0|s=1) \log p(r=0|s=1) + p(r=1|s=1) \log p(r=1|s=1)) \right] \\
 &= -.9 [.9 \log .9 + .1 \log .1] - .1 [.1 \log .1 + .9 \log .9] \\
 &= (.9 \log .9 + .1 \log .1) \approx .325
 \end{aligned}$$

$$MI(R,S) = H(R) - E_S[H(R|S=S)] \approx .5 - .325 = .175 \text{ bits}$$

What about arbitrary stimulus?

See slides 11-12

Context:



$$MI(R, S) = H(R) - E_S[H(R|S=S)]$$

$$= -p \log p - (1-p) \log(1-p)$$

$$\text{avg. rate} = \bar{r}$$

$$P(r=1) = \bar{r} \Delta t$$

$$P(r=0) = 1 - \bar{r} \Delta t$$

$$P(r=1|s) = r(t) \Delta t = p(t)$$

$$P(r=0|s) = 1 - r(t) \Delta t = 1 - p(t)$$

$$\hookrightarrow = \sum_S p(S) H(R|S=S) \quad (\text{hard to calculate!})$$

$$\text{But... law of large #s says: } E_S[H(R|S=S)] \approx \frac{1}{n} \sum_{i=1}^n H(R|S=\vec{S}_i)$$

where \vec{S}_i are sampled from $p(\vec{S})$

$$\text{But } \vec{S}(t) \text{ are samples from } p(\vec{S}), \text{ so } \rightarrow \frac{1}{n_t} \sum_{t=1}^{n_t} H(R|S=\vec{S}(t))$$

$$= \frac{1}{n_t} \sum_{t=1}^{n_t} -p(r=0|\vec{S}(t)) \underbrace{\log p(r=0|\vec{S}(t))}_{-p(r=1|\vec{S}(t)) \log p(r=1|\vec{S}(t))} - p(r=1|\vec{S}(t)) \log p(r=1|\vec{S}(t))$$

$$= \frac{1}{n_t} \sum_{t=1}^{n_t} -p(t) \log p(t) - (1-p(t)) \log(1-p(t)) \xrightarrow{(9)} \frac{1}{T} \int_0^T dt [p(t) \log p(t) + (1-p(t)) \log(1-p(t))]$$

$$\begin{aligned} \therefore MI(R, S) &= H(R) - E_{\vec{S}}(H(R|S=\vec{s})) \\ &= -p \log p - (1-p) \log (1-p) \\ &\quad + \frac{1}{T} \int_0^T dt [p(t) \log p(t) + (1-p(t)) \log (1-p(t))] \end{aligned}$$

(note: \vec{r}_{st} is expected spikes in 1 time bin)

Important: doesn't require explicit reference to stimulus, just assumes that $\vec{S}(t)$ was sampled from its prior distribution $p(\vec{S})$

\downarrow
more math

\downarrow

$$MI(R, S) \approx \frac{1}{T} \int_0^T dt \frac{r(t)}{\bar{r}} \log \frac{r(t)}{\bar{r}}$$

slide 18 notes: dim reduction via STA,
covariance analysis, etc.

slide 23 notes: spike trains might contain more info than single spikes treated independently

slide 18 notes: replace $\int_0^T dt$ w/ $\int d^k s P(s_1, s_2, \dots)$

slide 24 notes:

E.g., $\vec{s}(t_1) \rightarrow$

010 trial 4
010 trial 2
011 trial 3
010 trial 4
001 trial 5
01. trial 6

length-3 words:

$\vec{s}(t_2) \rightarrow$
000 trial 1
100 trial 2
000 :
200 :
000 trial 6
000

:

For each stim $s(t)$
get distr. of words
 $p(w|s(t))$

$$\text{To calc. } MI(w, s) = H[w] - E_{\vec{s}}[H(w|s=\vec{s})]$$

need $p(w)$ — estimate from all words of given length

need $p(w|\vec{s}(t))$ for all t — estimate across trials at time t

use law of large numbers again to go from

$$E_{\vec{s}}[H(w|s=\vec{s})] = \sum_s p(\vec{s}) H(w|s=\vec{s}) \approx \frac{1}{N_t} \sum_t H(w|s=\vec{s}(t))$$

But as length of words gets longer, harder to estimate ~~$p(w|\vec{s}(t))$~~
~~length~~ ~~MI(w,s)~~

Therefore, extrapolate:

length	1	2	3	4	...	10
MI(w,s)	85	84	84	83	...	80