# Relative Abundance Estimation with Microbiome Data 

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5/17/2021

## High-Throughput Sequencing of Microbial Communities

- Motivation: how can we determine what microbes live in a given environment?
- Isolation/culturing
- Useful but limited: most microbes aren’t culturable!
- Sequencing
- Attempt to detect and identify microbes by sequencing genetic material in samples
- Feasible with high-throughput sequencing techniques developed and refined over ~ the past two decades


## High-Throughput Sequencing of Microbial Communities

- High-throughput sequencing (whole-genome or marker gene)
- Complex measurement process with many steps
- Sample collection and storage
- DNA extraction
- DNA amplification
- Sequencing
- Taxonomic assignment


## High-Throughput Sequencing of Microbial Communities

- Measurement output: table $W_{n \times J}$ of taxon counts
- $W_{i j}$ : count of reads assigned to taxon $j$ in sample $i$

| Sample $\hat{\sim}$ | Atopobium.vaginae $\stackrel{\text { c }}{ }$ | Prevotella.bivia $\hat{}$ | Sneathia.amnii $\hat{\sim}$ | Streptococcus.agalactiae |
| :---: | :---: | :---: | :---: | :---: |
| 1 | 1028 | 1 | 14947 | 2 |
| 2 | 0 | 6 | 2 | 0 |
| 3 | 1424 | 21708 | 7 | 0 |
| 4 | 0 | 1854 | 6501 | 0 |

## What Do Read Counts (Allegedly) Tell Us?

According to microbiome folk wisdom, $W_{i j}$ (\# of reads assigned to taxon $j$ and sample $i$ )

- Does not to reflect "absolute abundance"
- i.e., $W_{i^{\prime} j}>W_{i j}$ does not imply that taxon $j$ is present in higher concentration in sample $i^{\prime}$ than in sample $i$

| Sample | Atopobium.vaginae $\stackrel{\text { a }}{ }$ | Prevotella.bivia | Sneathia.amnii | Streptococcus.agalactiae |
| :---: | :---: | :---: | :---: | :---: |
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## What Do Read Counts (Allegedly) Tell Us?

Also according to microbiome folk wisdom, $W_{i j}$ (\# of reads assigned to taxon $j$ and sample $i$ )

- reflects "relative abundance" in sense that $W_{i j} \propto p_{i j}$, where $p_{i j}$ is the true proportion of detectable microbes in sample $i$ belonging to taxon $j$

| Sample $\hat{*}$ | Atopobium.vaginae $\hat{*}$ | Prevotella.bivia $\hat{\sim}$ | Sneathia.amnii $\hat{\square}$ | Streptococcus.agalactiae |
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## What Do Read Counts (Allegedly) Tell Us?

- "Relative abundance" interpretation motivates estimator for $p_{i j}$ (true prop. of microbes in sample $i$ belonging to taxon $j$ )

$$
\hat{p}_{i j}=\frac{W_{i j}}{\sum_{j=1}^{J} W_{i j}}
$$

| Sample $\hat{\sim}$ | Atopobium.vaginae $\stackrel{\text { c }}{ }$ | Prevotella.bivia $\hat{}$ | Sneathia.amnii $\hat{\sim}$ | Streptococcus.agalactiae |
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- "Relative abundance" interpretation motivates estimator for $p_{i j}$ (true prop. of microbes in sample $i$ belonging to taxon $j$ )

$$
\hat{p}_{i j}=\frac{W_{i j}}{\sum_{j=1}^{J} W_{i j}}
$$

- Focus of this talk: $\hat{p}_{i j}$ a reasonable estimator of $p_{i j}$ ?
- How can we evaluate performance?
- Can we do better?


## Some Statistical Framing



## Some Statistical Framing

Goal: reason about states of nature / data-generating mechanism using observations + what we know about how they were generated

## States of Nature

True microbial composition(s)
$\left\{p_{i j}\right\}$ of communities of interest


Data Generating Mechanism
Sample collection, preparation, sequencing, taxonomic assignment, etc.

## Some Statistical Framing



## How Well Does the Naive Estimator Perform?

Sample vs. True Compositions of 40 Samples Sequenced by Brooks et al. (2015)


Taxon

- Atopobium vaginae
- Gardnerella vaginalis
- Lactobacillus crispatus_cluster
- Lactobacillus iners
- Prevotella bivia
- Sneathia amnii
- Streptococcus agalactiae


## How Well Does the Naive Estimator Perform?

Sample vs. True Compositions of 40 Samples Sequenced by Brooks et al. (2015)


## How Well Does the Naive Estimator Perform?

Sample Read Proportions in Two Even Mixtures


Taxon
Gardnerella vaginalis

- Lactobacillus iners
- Sneathia amnii


## How Well Does the Naive Estimator Perform?

Sample Read Proportions in Two Even Mixtures


## What's Going On?

- McLaren et al. (2019)
- Observe $\hat{p}_{i j}$ does not perform well as an estimator of $p_{i j}$
- Suggesting that $W_{i j} \propto p_{i j}$ does not hold in general
- i.e., read counts across taxa in a sample are not approximately proportional to true relative abundances
- Hypothesis: any given sequencing protocol will be better at detecting some microbial taxa than others


## Multiplicative Distortion: "Efficiencies"

- McLaren et al. (2019)
- A given sample handling/sequencing/postprocessing protocol will preferentially detect some microbes over others
- Formalize this idea in terms of a detection "efficiency" $e_{j}$ of taxon j
- Instead of $W_{i j} \propto p_{i j}$, posit $W_{i j} \propto e_{j} p_{i j}$ (at least approximately)

$$
\Rightarrow \hat{p}_{i j}=\frac{W_{i j}}{\sum_{j=1}^{J} W_{i j}} \approx \frac{e_{j} p_{i j}}{\sum_{j=1}^{J} e_{j} p_{i j}} \neq p_{i j} \text { (in general) }
$$

## Efficiencies: An Example

$$
\underbrace{\mathbb{E}\left[W_{i .}\right]}_{\begin{array}{c}
\text { Expected } \\
\text { counts in } \\
\text { sample } i
\end{array}}=\underbrace{\left(\rho_{i .}\right.}_{\begin{array}{c}
\text { True relative } \\
\text { abundance } \\
\text { profile of } \\
\text { sample } i
\end{array}} \circ \exp \underbrace{(\boldsymbol{\beta})}_{\begin{array}{c}
\text { Taxon- } \\
\text { specific } \\
\text { "efficiencies" }
\end{array}}) \cdot \underbrace{k_{i}}_{\begin{array}{c}
\text { Proportionality } \\
\text { term }
\end{array}}
$$

## Simulated Sequencing Data on a Simple Community



Consider a specimen

- containing taxa $A, B$, and $C$
- in relative abundances 0.5 , 0.25 , and 0.25 , respectively


## Efficiencies: An Example

## Simulated Sequencing Data on a Simple Community



Consider a specimen

- containing taxa $A, B$, and $C$
- in relative abundances 0.5, 0.25 , and 0.25 , respectively

We can simulate* sequencing data under

- assumption $\mathbb{E}\left[W_{i}\right] \propto \rho_{i}$.
* First setting (with same efficiencies across taxa): each count simulated as a negative binomial with mean $\mu_{j}=500 * \rho_{j}$ and size parameter $s=5$ (s.t. $\left.\operatorname{Var}\left(W_{i j}\right)=\mu_{j}+\mu_{j}^{2} / s\right)$


## Efficiencies: An Example



Simulated Sequencing Data on a Simple Community


Consider a specimen

- containing taxa $A, B$, and $C$
- in relative abundances 0.5 , 0.25 , and 0.25 , respectively

We can simulate* sequencing data under

- assumption $\mathbb{E}\left[W_{i}\right] \propto \rho_{i}$.
- assumption $\mathbb{E}\left[W_{i}\right] \propto \exp (\boldsymbol{\beta}) \circ \rho_{i}$.
- with $\exp (\boldsymbol{\beta})=(2,8,1)$
* First setting (with same efficiencies across taxa): each count simulated as a negative binomial with mean $\mu_{j}=500 * \rho_{j}$ and size parameter $s=5$ (s.t.
$\left.\operatorname{Var}\left(W_{i j}\right)=\mu_{j}+\mu_{j}^{2} / s\right)$
* Second setting (differing efficiencies): each count simulated as a negative binomial with $\mu_{j}=[500 / \exp (\overline{\boldsymbol{\beta}})] * \rho_{i j}$ and size parameter $s=5$ (s.t. $\left.\operatorname{Var}\left(W_{i j}\right)=\mu_{j}+\mu_{j}^{2} / s\right)$


## Efficiencies: An Example



Detection efficiency of taxon $B$ relative to taxon $C$ $\exp \left(\beta_{2}\right)=8$

Ratio of counts in taxon $B$ to counts in taxon $\mathrm{C} \approx 8$ times too large

## Simulated Sequencing Data on a Simple Community

Consider a specimen

- containing taxa $A, B$, and $C$
- in relative abundances 0.5, 0.25 , and 0.25 , respectively

We can simulate* sequencing data under

- assumption $\mathbb{E}\left[W_{i .}\right] \propto \rho_{i}$.
- assumption $\mathbb{E}\left[W_{i}\right] \propto \exp (\boldsymbol{\beta}) \circ \rho_{i}$.
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$\left.\operatorname{Var}\left(W_{i j}\right)=\mu_{j}+\mu_{j}^{2} / s\right)$


## Estimating Relative Abundance in Presence of Efficiencies

- McLaren et al. (2019)
- Method for estimating $p_{i j}$ and $e_{j}$ via a centered log-ratio transformation of counts $W_{i j}$
- Need to know presence/absence in advance
- Zero counts, spurious counts an issue

| Sample $\hat{\sim}$ | Atopobium.vaginae $\hat{\square}$ | Prevotella.bivia $\hat{\square}$ | Sneathia.amnii $\hat{\sim}$ | Streptococcus.agalactiae |
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- McLaren et al. (2019)
- Method for estimating $p_{i j}$ and $e_{j}$ via a centered log-ratio transformation of counts $W_{i j}$
Log(0)
undefined • Need to know presence/absence in advance
- Zero counts, spurious counts an issue

| Sample | Atopobium. vaginae $\hat{\square}$ | Prevotella.bivia Sneathia.amnii $\stackrel{\text { ar }}{ }$ |  | Streptococcus.agalactiae |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | - 1028 | 1 | 4947 |  | 2 |
| 2 | $0$ | 6 | 2 |  | 0 |
| 3 | 1424 | 21708 | 7 |  | 0 |
| 4 | 0 | 1854 | 6501 |  | 0 |

## Estimating Relative Abundance in Presence of Efficiencies

- McLaren et al. (2019)
- Method for estimating $p_{i j}$ and $e_{j}$ via a centered log-ratio transformation of counts $W_{i j}$

$$
W_{i j}>0 \text { when } p_{i j}=0
$$

undefined - Need to know presence/absence in adyance

- zero counts, spurious counts an issue

| Sample | Atopobium.vaginae $\stackrel{\text { 人 }}{ }$ | Prevotella.bivia Sneat/ra.amnii $\overbrace{\text { a }}$ \$treptococcus.agalactia ${ }^{\text {a }}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | $028$ |  | 11947 |  | (2) |
| 2 | $0$ | $6$ | (2) |  | 0 |
| 3 | 1424 | 21708 | 7 | $\rightarrow$ | 0 |
| 4 | 0 | 1854 | 6501 |  | 0 |

## Generalizing McLaren et al.

- Clausen-Willis approach: model counts $W$ directly
- Attempt to model spurious reads (due to, e.g., contamination) in addition to detection efficiencies
- Mean model for a count $W_{i j}$

$$
\begin{aligned}
& \qquad \mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \gamma, \tilde{\mathbf{p}}, \tilde{\gamma}\right] \\
& =\quad \underbrace{p_{i j} \exp \left(\gamma_{i}+\beta_{j}\right)}_{\text {contribution of sample }}+\quad{ }^{\tilde{p}_{i j} \exp (\tilde{\gamma})} \\
& \text { spurious read sources }^{\tilde{p}^{2}}
\end{aligned}
$$

## Mean Model Details

$$
\mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \gamma, \tilde{\mathbf{p}}, \tilde{\boldsymbol{\gamma}}\right]
$$



## Mean Model Details

True relative abundance of taxon jin sample I

$$
\mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \boldsymbol{\gamma}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]
$$



## Mean Model Details

## True relative abundance $\mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \boldsymbol{\gamma}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]$



Proportionality constant

## Mean Model Details

## True relative abundance $\mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \boldsymbol{\gamma}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]$



Proportionality constant

Log efficiency $e_{j}$

## Mean Model Details

## True relative abundance of taxon j in sample I <br> $\mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \boldsymbol{\gamma}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]$

(Unknown) relative abundance profile of spurious read source


## Mean Model Details

$$
\begin{aligned}
& \text { True relative abundance } \\
& \text { of taxon j in sample \| }
\end{aligned} \quad \mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \gamma, \tilde{\mathbf{p}}, \tilde{\gamma}\right]
$$

(Unknown) relative
abundance profile of
spurious read source


Proportionality constant
" " spurious read sources

Intensity of spurious reads

## Mean Model Details: A Bit More Generality

$$
\begin{gathered}
\qquad \mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \boldsymbol{\gamma}, \tilde{\mathbf{p}}, \tilde{\gamma}\right] \\
=\quad \underbrace{p_{i j} \exp \left(\gamma_{i}+\beta_{j}\right)}_{\text {contribution of sample }}+\quad \underbrace{\tilde{p}_{i j} \exp (\tilde{\gamma})}
\end{gathered}
$$

$=\underbrace{Z_{i} p^{j} \exp \left(\gamma_{i}+X_{i} \beta^{j}\right)}_{\text {contribution of sample }}+\underbrace{\tilde{Z}_{i}\left(\tilde{p}^{j} \circ \exp (\tilde{\gamma})\right)}_{\text {" spurious read sources }}$
$=\underbrace{\left[\left(\mathbf{Z} \mathbf{p} \circ \exp \left(\boldsymbol{\gamma} \mathbf{1}_{J}^{T}+\mathbf{X} \boldsymbol{\beta}\right)\right.\right.}+\quad \tilde{\mathbf{Z}}\left[\tilde{\mathbf{p}} \circ \exp \left(\tilde{\boldsymbol{\gamma}} \mathbf{1}_{J}^{T}+\tilde{\mathbf{X}} \boldsymbol{\beta}\right)\right]]_{i j}$
contribution of samples
contribution of spurious read sources

## Mean Model Details: A Bit More Generality

$$
\begin{gathered}
\mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \boldsymbol{\gamma}, \tilde{\mathbf{p}}, \tilde{\gamma}\right] \\
=\underbrace{\left[\left(\mathbf{Z} \mathbf{p} \circ \exp \left(\gamma \mathbf{1}_{J}^{T}+\mathbf{X} \boldsymbol{\beta}\right)+\tilde{\mathbf{Z}}\left[\tilde{\mathbf{p}} \circ \exp \left(\tilde{\boldsymbol{\gamma}} 1_{J}^{T}+\tilde{\mathbf{X}} \boldsymbol{\beta}\right)\right]\right]_{i j}\right.}
\end{gathered}
$$

contribution of samples contribution of spurious read sources
More general form allows us to

- Easily incorporate technical replicates
- Model differing efficiencies across samples
- E.g., due to different protocols in different batches
- Model multiple sources of spurious reads
- And more

More details in supplemental slides if you're interested

## Defining an Estimator

We estimate unknown parameters in mean model

$$
\mu_{i j}:=\mathbb{E}\left[W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \boldsymbol{\gamma}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]
$$

by modeling

$$
W_{i j} \mid \boldsymbol{\beta}, \mathbf{p}, \boldsymbol{\gamma}, \tilde{\mathbf{p}}, \tilde{\gamma} \sim \operatorname{Poisson}\left(\mu_{i j}\right)
$$

and estimate parameters via maximum likelihood

## An Applied Example

- Data from Brooks et al. (2015)
- 40 whole-cell samples of known composition prepped and sequenced (via 16S) together
- All specimens composed of some combination of 7 common bacterial species in the vaginal microbiome
- We observe some spurious reads (nonzero number of reads in taxa known to be absent in a particular sample)
- Probably reasonable to model a single detection efficiency for each taxon


## An Applied Example (cont.)

Proposed mean model:

$$
\begin{gathered}
\mathbf{E}\left[\mathbf{W}_{n \times J} \mid \mathbf{p}, \boldsymbol{\gamma}, \boldsymbol{\beta}, \tilde{\mathbf{p}}, \tilde{\boldsymbol{\gamma}}\right]= \\
\left(\mathbf{p} \circ \exp \left(\boldsymbol{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)+\mathbf{1}_{n}\left[\tilde{\mathbf{p}} \circ \exp \left(\tilde{\boldsymbol{\gamma}} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)\right]\right.
\end{gathered}
$$

- $n=40$ samples sequenced
- $K=40$ unique specimens
- $J=7$ taxa considered
- $p=1(1 \times J)$ efficiency effect
- $\tilde{n}=1$ spurious read source


## An Applied Example (cont.)

Proposed mean model:

## $\mathbf{E}\left[\mathbf{W}_{n \times J} \mid \mathbf{p}, \boldsymbol{\gamma}, \boldsymbol{\beta}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]=$



## An Applied Example (cont.)

Proposed mean model:

## $\mathbf{E}\left[\mathbf{W}_{n \times J} \mid \mathbf{p}, \boldsymbol{\gamma}, \boldsymbol{\beta}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]=$

$$
\left(\mathbf{p} \circ \exp \left(\gamma \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)+\mathbf{1}_{n}\left[\tilde{\mathbf{p}} \circ \exp \left(\tilde{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)\right]\right.
$$


$n \times 1=40 \times 1$ matrix of read depth values for samples
$i=1, \ldots, 40$

- $n=40$ samples sequenced
- $K=40$ unique specimens
- $J=7$ taxa considered
- $p=1(1 \times J)$ efficiency effect
- $\tilde{n}=1$ spurious read source
$\boldsymbol{\beta}=\left[\beta_{1}, \ldots, \beta_{7}\right]^{T}$, constrain $\beta_{7}=0$; $\beta_{j}$ has interp.
$\log$ relative eff. of taxon $j$ rel. to taxon 7
(for $j=1, \ldots, 6$ )


## An Applied Example (cont.)

Proposed mean model:

$$
\left(\mathbf{p} \circ \exp \left(\gamma \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)+\mathbf{1}_{n}\left[\tilde{\mathbf{p}} \circ \exp \left(\tilde{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)\right]\right.
$$

$n \times J=40 \times 7$ matrix of
true relative abundances
$p_{i j}$ for $i=1, \ldots, 40$ and
$j=1, \ldots, 7$
$n \times 1=40 \times 1$
matrix of read depth values for samples
$i=1, \ldots, 40$

- $n=40$ samples sequenced
- $K=40$ unique specimens
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- $p=1(1 \times J)$ efficiency effect
- $\tilde{n}=1$ spurious read source
$\boldsymbol{\beta}=\left[\beta_{1}, \ldots, \beta_{7}\right]^{T}$, constrain $\beta_{7}=0$; $\beta_{j}$ has interp. log relative eff. of taxon $j$ rel. to taxon 7 (for $j=1, \ldots, 6$ )


## An Applied Example (cont.)

Proposed mean model:

$\left(\mathbf{p} \circ \exp \left(\boldsymbol{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)+\mathbf{1}_{n}\left[\tilde{\mathbf{p}} \circ \exp \left(\tilde{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)\right]\right.$


- $n=40$ samples sequenced
- $K=40$ unique specimens
- $J=7$ taxa considered
- $p=1(1 \times J)$ efficiency effect
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$\boldsymbol{\beta}=\left[\beta_{1}, \ldots, \beta_{7}\right]^{T}$, constrain $\beta_{7}=0$; $\beta_{j}$ has interp. log relative eff. of taxon $j$ rel. to taxon 7
(for $j=1, \ldots, 6$ )


## Mean Model Example (cont.)

## Model all samples as having on average same abundance of spurious reads

Proposed mean model:

$$
\mathbf{E}\left[\mathbf{W}_{n \times J} \mid \mathbf{p}, \gamma, \boldsymbol{\beta}, \tilde{\mathbf{p}}, \tilde{\boldsymbol{\gamma}}\right]=
$$

$$
\left(\mathbf{p} \circ \exp \left(\boldsymbol{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)+\mathbf{1}_{n}\left[\tilde{\mathbf{p}}^{\prime} \circ \exp \left(\tilde{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)\right]\right.
$$

Single source of spurious reads $\left[\tilde{p}_{11}, \ldots, \tilde{p}_{1 J}\right]$
$n \times J=40 \times 7$ matrix of
true relative abundances $p_{i j}$ for $i=1, \ldots, 40$ and $j=1, \ldots, 7$


- $n=40$ samples sequenced
- $K=40$ unique specimens
- $J=7$ taxa considered
- $p=1(1 \times J)$ efficiency effect
- $\tilde{n}=1$ spurious read source
$\boldsymbol{\beta}=\left[\beta_{1}, \ldots, \beta_{7}\right]^{T}$, constrain $\beta_{7}=0$; $\beta_{j}$ has interp. log relative eff. of taxon $j$ rel. to taxon 7
(for $j=1, \ldots, 6$ )


## Mean Model Example (cont.)

Proposed mean model:

## Model all samples as having on average same abundance of spurious reads <br> Single source of spurious reads $\left[\tilde{p}_{11}, \ldots, \tilde{p}_{1 J}\right]$

$$
\mathbf{E}\left[\mathbf{W}_{n \times J} \mid \mathbf{p}, \gamma, \boldsymbol{\beta}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]=
$$

$$
\left(\mathbf{p} \circ \exp \left(\gamma \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)+\mathbf{1}_{n}\left[\tilde{\mathbf{p}} \circ \exp \left(\tilde{\boldsymbol{\gamma}} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)\right]\right.
$$



Single intensity of spurious reads $\tilde{\gamma}$
$n \times J=40 \times 7$ matrix of
true relative abundances $p_{i j}$ for $i=1, \ldots, 40$ and $j=1, \ldots, 7$

## Performance on Brooks (2015) Data

Fit mean model

$$
\mathbf{E}\left[\mathbf{W}_{n \times J} \mid \mathbf{p}, \boldsymbol{\gamma}, \boldsymbol{\beta}, \tilde{\mathbf{p}}, \tilde{\gamma}\right]=\left(\mathbf{p} \circ \exp \left(\boldsymbol{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)+\mathbf{1}_{n}\left[\tilde{\mathbf{p}} \circ \exp \left(\tilde{\gamma} \mathbf{1}_{J}^{T}+\boldsymbol{\beta}\right)\right]\right.
$$

- One sample per unique specimen
- One set of log efficiencies $\boldsymbol{\beta}$
- One source of spurious reads $\tilde{\mathbf{p}}$ modeled as having same read abundance across samples

To data from 40 samples sequenced together

- Use known true compositions of first 10 samples
- All other compositions estimated from data


## Performance on Brooks (2015) Data

Compositions of 30 Samples Sequenced by Brooks et al. (2015)


## Performance on Brooks (2015) Data

- Additionally, this model fit estimated 6 relative abundances with spurious counts to be zero
- $\approx 10 \%$ of taxon-sample pairs with spurious reads (with true relative abundance zero)
- Better choice of $\tilde{Z}$ might perform better


## Future Work

- Inference via a modified bootstrap
- Predicting efficiencies in taxa not present in specimens of known composition
- Investigating use of covariates for spurious reads (e.g., DNA concentration)


## Acknowledgments

This work is joint with my fabulous advisor, Dr.
Amy D. Willis, Assistant Professor, UW
Department of Biostatistics, and it is the National Institute Of General Medical Sciences of the National Institutes of Health under Award Number R35GM133420

## statistical diversity lab



I am grateful as well for the invaluable feedback my colleagues in the Statistical Diversity Lab have provided in the development of this model and presentation


Bryan Martin
(UW Statistics)


Cecilia Shi
(UW Statistics)


Pauline Trinh
(UW DEOHS)


Maria Valdez Cabrera
(UW Biostatistics)


Sarah Teichman
(UW Statistics)

## Thank You!

## References

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McLaren, Michael R., Amy D. Willis, and Benjamin J. Callahan. "Consistent and correctable bias in metagenomic sequencing experiments." elife 8 (2019): e46923.

## Supplemental slides

| $K \times J$ matrix of true |
| :--- |
| relative abundances |
| of taxa $j=1, \ldots, J$ in |
| specimens $k=1, \ldots, K$ |

## Mean Model Details

| $p \times J$ matrix of |
| :--- |
| (log) efficiency |
| parameters |

$n \times \tilde{K}$ matrix linking samples to sources of spurious reads: columns may depend on/ $\exp (\gamma)$
$\tilde{\mathbf{Z}}\left[\tilde{\mathbf{p}} 。 \exp \left(\tilde{\gamma} \mathbf{1}_{J}^{T}+\tilde{\mathbf{X}} \boldsymbol{\beta}\right)\right]$
$\left(\mathbf{Z p} \circ \exp \left(\gamma \mathbf{1}_{J}^{T}+\mathbf{X} \boldsymbol{\beta}\right) \quad+\right.$
contribution of samples
$n \times K$ matrix
linking samples to originating specimens;
$Z_{i k}=1$ if
sample $i$ was taken from
specimen $k$

contribution of spúrious read sources

$\tilde{n} \times 1$ matrix of spurious read intensities

- $n$ : \# samples sequenced
- K: \# unique specimens
- J: \# taxa considered
- $p: \#(1 \times J)$ efficiency effects
- $\tilde{n}$ : \# spurious read sources

