Announcements

- Don't Cheat
- Proposals due tonight

Logistic, SVM, and Perceptron

Machine Learning – CSE546 Kevin Jamieson University of Washington

October 25, 2018



How do we solve for *w*? The last two lectures!

Perceptron is optimizing what?

Perceptron update rule:

$$\begin{bmatrix} w_{k+1} \\ b_{k+1} \end{bmatrix} = \begin{bmatrix} w_k \\ b_k \end{bmatrix} + y_k \begin{bmatrix} x_k \\ 1 \end{bmatrix} \mathbf{1} \{ y_i(b + x_i^T w) < 0 \}$$

SVM objective:

$$\sum_{i=1}^{n} \max\{0, 1 - y_i(b + x_i^T w)\} + \lambda ||w||_2^2 = \sum_{i=1}^{n} \ell_i(w, b)$$

$$\nabla_{w}\ell_{i}(w,b) = \begin{cases} -x_{i}y_{i} + \frac{2\lambda}{n}w & \text{if } y_{i}(b + x_{i}^{T}w) < 1\\ \frac{2\lambda}{n} & \text{otherwise} \end{cases}$$

$$\nabla_{b}\ell_{i}(w,b) = \begin{cases} -y_{i} & \text{if } y_{i}(b + x_{i}^{T}w) < 1\\ 0 & \text{otherwise} \end{cases}$$
Perceptron is almost SGD on SVM with $\lambda = 0, \eta = 1!$

SVMs vs logistic regression

• We often want probabilities/confidences, logistic wins here?

SVMs vs logistic regression

- We often want probabilities/confidences, logistic wins here?
- No! Perform isotonic regression or non-parametric bootstrap for probability calibration. Predictor gives some score, how do we transform that score to a probability?

Bootstrap

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Limitations of CV

- An 80/20 split throws out a relatively large amount of data if only have, say, 20 examples.
- Test error is informative, but how accurate is this number? (e.g., 3/5 heads vs. 30/50)
- How do I get confidence intervals on statistics like the median or variance of a distribution?
- Instead of the error for the entire dataset, what if I want to study the error for a *particular example* x?

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The Bootstrap: Developed by Efron in 1979.

"The most important innovation in statistics of the last 40 years" — famous ML researcher and statistician, 2015

Given dataset drawn iid samples with CDF F_Z :

$$\mathcal{D} = \{z_1, \dots, z_n\} \stackrel{i.i.d.}{\sim} F_Z$$

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For b=1,...,B define the *b*th *bootstrapped* dataset as drawing *n* samples **with replacement** from *D*

 $\mathcal{D}^{*b} = \{z_1^{*b}, \dots, z_n^{*b}\} \stackrel{i.i.d.}{\sim} \widehat{F}_{Z,n}$ and the *b*th bootstrapped statistic as: $\theta^{*b} = t(\mathcal{D}^{*b})$

Given dataset drawn iid samples with CDF F_Z :



Given dataset drawn iid samples with CDF F_Z :



Applications

Common applications of the bootstrap:

- Estimate parameters that escape simple analysis like the variance or median of an estimate
- Confidence intervals
- Estimates of error for a particular example:



Figures from Hastie et al

Takeaways

Advantages:

- Bootstrap is very generally applicable. Build a confidence interval around anything
- Very simple to use
- Appears to give meaningful results even when the amount of data is very small
- Very strong asymptotic theory (as num. examples goes to infinity)

Takeaways

Advantages:

- Bootstrap is very generally applicable. Build a confidence interval around anything
- Very simple to use
- Appears to give meaningful results even when the amount of data is very small
- Very strong **asymptotic theory** (as num. examples goes to infinity)

Disadvantages

- Very few meaningful finite-sample guarantees
- Potentially computationally intensive
- Reliability relies on test statistic and rate of convergence of empirical CDF to true CDF, which is unknown
- Poor performance on "extreme statistics" (e.g., the max)

Not perfect, but better than nothing.

Warm up: risk prediction with logistic regression

Boss gives you a bunch of data on loans defaulting or not:

$$\{(x_i, y_i)\}_{i=1}^n \ x_i \in \mathbb{R}^d, \ y_i \in \{-1, 1\}$$

- You model the data as: $P(Y = y | x, w) = \frac{1}{1 + \exp(-y w^T x)}$
- And compute the maximum likelihood estimator:

$$\widehat{w}_{MLE} = \arg\max_{w} \prod_{i=1}^{n} P(y_i | x_i, w)$$

For a new loan application x, boss recommends to give loan if your model says they will repay it with probability at least .95 (i.e. low risk):

Give loan to x if
$$\frac{1}{1 + \exp(-\widehat{w}_{MLE}^T x)} \ge .95$$

 One year later only half of loans are paid back and the bank folds. What might have happened?

How would you use the bootstrap to do this differently?

Decision Theory

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- Learn: f:X —>Y
 - □ X features
 - □ Y target classes $Y \in \{-1, 1\}$
- Expected loss of f:

Loss function:

 $\ell(f(x),y) = \mathbf{1}\{f(x) \neq y\}$

$$\mathbb{E}_{XY}[\mathbf{1}\{f(X) \neq Y\}] = \mathbb{E}_X[\mathbb{E}_{Y|X}[\mathbf{1}\{f(x) \neq Y\}|X = x]]$$
$$\mathbb{E}_{Y|X}[\mathbf{1}\{f(x) \neq Y\}|X = x] = 1 - P(Y = f(x)|X = x)$$

Bayes optimal classifier:

$$f(x) = \arg\max_{y} \mathbb{P}(Y = y | X = x)$$

Model of logistic regression:

$$P(Y = y | x, w) = \frac{1}{1 + \exp(-y \, w^T x)}$$

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Bayes rule:
$$\mathbb{P}(Y = y | X = x) = \frac{\mathbb{P}(X = x | Y = y)\mathbb{P}(Y = y)}{P(X = x)}$$

- Learn: f:X —>Y
 - □ **X** features

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 $\mathbb{P}(y=1) = 1/3 \ \mathbb{P}(y=0) = 2/3$



$$\mathbb{P}(X = x) = \mathbb{P}(X = x | Y = 0) \mathbb{P}(Y = 0) + \mathbb{P}(X = x | Y = 1) \mathbb{P}(Y = 1)$$

=: $(1 - \pi) P_0(x) + \pi P_1(x)$

Suppose
$$P_0(x) = \mathcal{N}(x; \mu_0, \sigma^2)$$
 $P_1(x) = \mathcal{N}(x; \mu_1, \sigma^2)$

$$f(x) = \arg \max_{y} \mathbb{P}(Y = y | X = x)$$
$$= \arg \max_{y} \mathbb{P}(X = x | Y = y) \mathbb{P}(Y = y)$$

$$f(x) = 1$$
 if $\frac{P_1(x)\pi}{P_0(x)(1-\pi)} \ge 1$

Let

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$$f(x) = 1 \text{ if } \frac{\mu_1 - \mu_0}{\sigma^2} \left(x - \frac{\mu_1 + \mu_0}{2} \right) \ge -\log\left(\frac{\pi}{1 - \pi}\right)$$
$$f(x) = 1 \text{ if } x \ge \frac{\mu_1 + \mu_0}{2} - \frac{\sigma^2}{\mu_1 - \mu_0} \log\left(\frac{\pi}{1 - \pi}\right)$$

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Let
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$$=: (1 - \pi) P_0(x) + \pi P_1(x)$$

Suppose $P_0(x) = \mathcal{N}(x; \mu_0, \sigma^2)$ $P_1(x) = \mathcal{N}(x; \mu_1, \sigma^2)$

$$f(x) = 1 \text{ if } \frac{P_1(x)\pi}{P_0(x)(1-\pi)} \ge 1 \qquad f(x) = 1 \text{ if } x \ge \frac{\mu_1 + \mu_0}{2} - \frac{\sigma^2}{\mu_1 - \mu_0} \log(\frac{\pi}{1-\pi})$$

$$\pi = 1/2$$
 $\pi \in (1/2, 1)$ $\pi \in (0, 1/2)$



Same ideas extend to higher dimensions:

$$P_1(x) = \mathcal{N}(x; \mu_1, \Sigma_1)$$

$$P_0(x) = \mathcal{N}(x; \mu_0, \Sigma_0)$$

$$f(x) = 1$$
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Cases:

$$\Sigma_0 = \Sigma_1 :$$

 $\Sigma_0 \neq \Sigma_1$:

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$$f(x) = 1$$
 if $\frac{P_1(x)\pi}{P_0(x)(1-\pi)} \ge 1$

In practice we observe $\{(x_i, y_i)\}_{i=1}^n$

$$\widehat{\mu}_{k} = \frac{1}{|\{i : y_{i} = k\}|} \sum_{i:y_{i} = k} x_{i}$$
$$\widehat{\Sigma}_{k} = \frac{1}{|\{i : y_{i} = k\}| - 1} \sum_{i:y_{i} = k} (x_{i} - \widehat{\mu}_{k})(x_{i} - \widehat{\mu}_{k})^{T}$$

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Discriminative learning directly models $\mathbb{P}(Y = y | X = x)$

Example:

Generative learning models $\mathbb{P}(X=x,Y=y)=\mathbb{P}(X=x|Y=y)\mathbb{P}(Y=y)$

Example:

Hypothesis testing

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You are Amazon and wish to detect transactions with stolen credit cards.

For each transaction we observe a **feature vector X**: { email-address, age of account, anonymous PO box, price of items, copies of purchased item, etc. } and the transaction is either **real (Y=0)** or **fraudulent (Y=1)**

Hypothesis testing:

H0:
$$X \sim P_0$$
 $P_k = \mathbb{P}(X = x | Y = k)$
H1: $X \sim P_1$

Your job is to build a (possibly randomized) decision function $\delta(x) \in \{0, 1\}$

$$\mathbb{P}(X = x) = \pi \mathbb{P}_1(x) + (1 - \pi) \mathbb{P}_0(x)$$

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Bayesian Hypothesis Testing:

Assume
$$\mathbb{P}(Y=1) = \pi$$

 $\mathbb{P}(X=x) = \pi P_1(x) + (1-\pi)P_0(x)$

$$\arg\min_{\delta} \mathbb{P}_{XY}(Y \neq \delta(X))$$

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Minimax Hypothesis Testing:

$$\arg\min_{\delta} \max\{\mathbb{P}(\delta(X) = 0 | Y = 1), \mathbb{P}(\delta(X) = 1 | Y = 0)\}$$

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Neyman-Pearson Hypothesis Testing:

$$\arg\max_{\delta} \mathbb{P}(\delta(X) = 1 | Y = 1), \text{ subject to } \mathbb{P}(\delta(X) = 1 | Y = 0) \le \alpha \}$$

Neyman-Pearson Testing

Hypothesis testing:

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Neyman-Pearson Hypothesis Testing:

$$\arg\max_{\delta} \mathbb{P}(\delta(X) = 1 | Y = 1), \text{ subject to } \mathbb{P}(\delta(X) = 1 | Y = 0) \le \alpha \}$$

Theorem: The optimal test δ^* has the form $\mathbb{P}(\delta^*(X) = 1) = \begin{cases} 1 & \text{if } \frac{P_1(x)}{P_0(x)} > \eta \\ \gamma & \text{if } \frac{P_1(x)}{P_0(x)} = \eta \\ 0 & \text{if } \frac{P_1(x)}{P_0(x)} < \eta \end{cases}$ and satisfies $\mathbb{P}(\delta^*(X) = 1 | Y = 0) = \alpha$