## Recognition Part I

## CSE 576

## What we have seen so far: Vision as Measurement Device



Real-time stereo on Mars


Physics-based Vision



Structure from Motion


Virtualized Reality Slide Credit: Alyosha

## Visual Recognition

- What does it mean to "see"?
- "What" is "where", Marr 1982
- Get computers to "see"


## Visual Recognition

## Verification



## Visual Recognition

## Classification:

Is there a car in this picture?


## Visual Recognition

## Detection:

Where is the car in this picture?


## Visual Recognition

## Pose Estimation:



## Visual Recognition

## Activity Recognition:



## Visual Recognition

## Object Categorization:



## Visual Recognition

## Segmentation

## Sky

## Car

## Object recognition Is it really so hard?

This is a chair


Find the chair in this image


Output of normalized correlation


## Object recognition Is it really so hard?

Find the chair in this image


Pretty much garbage
Simple template matching is not going to make it

## Challenges 1: view point variation



## Challenges 2: illumination



## Challenges 3: occusion

## Challenges 4: scale



## Challenges 5: deformation



Xu, Beihong 1943

## Challenges 6: background clutter

Klimt, 1913


## Challenges 7: object intra-class variation


slide by Fei-Fei, Fergus \& Torralba

## Let's start with finding Faces



How to tell if a face is present?

## One simple method: skin detection



Skin pixels have a distinctive range of colors

- Corresponds to region(s) in RGB color space
- for visualization, only $R$ and $G$ components are shown above

Skin classifier

- A pixel $X=(R, G, B)$ is skin if it is in the skin region
- But how to find this region?


## Skin detection



Learn the skin region from examples

- Manually label pixels in one or more "training images" as skin or not skin
- Plot the training data in RGB space
- skin pixels shown in orange, non-skin pixels shown in blue
- some skin pixels may be outside the region, non-skin pixels inside. Why?


## Skin classifier

- Given $X=(R, G, B)$ : how to determine if it is skin or not?


## Skin classification techniques

Skin classifier


- Given $X=(R, G, B)$ : how to determine if it is skin or not?
- Nearest neighbor
- find labeled pixel closest to $X$
- choose the label for that pixel
- Data modeling
- Model the distribution that generates the data (Generative)
- Model the boundary (Discriminative)


## Classification

- Probabilistic
- Supervised Learning
- Discriminative vs. Generative
- Ensemble methods
- Linear models
- Non-linear models


## Let's play with probability for a bit

Remembering simple stuff

## Probability

## Basic probability

- X is a random variable
- $P(X)$ is the probability that $X$ achieves a certain value

- $0 \leq P(X) \leq 1$
- $\begin{gathered}\int_{-\infty}^{\infty} P(X) d X=1 \\ \text { continuous } \mathrm{X}\end{gathered} \quad$ or $\quad \begin{gathered}\sum P(X)=1 \\ \text { discrete } \mathrm{X}\end{gathered}$
- Conditional probability: $\mathrm{P}(\mathrm{X} \mid \mathrm{Y})$
- probability of $X$ given that we already know $Y$


## Thumbtack \& Probabilities

$P($ Heads $)=\theta, P($ Tails $)=1-\theta$


Flips are i.i.d.:

- Independent events $D=\left\{x_{i} \mid i=1 \ldots n\right\}, P(D \mid \theta)=\prod_{i} P\left(x_{i} \mid \theta\right)$
- Identically distributed according to Binomial distribution

Sequence $D$ of $\alpha_{H}$ Heads and $\alpha_{T}$ Tails

$$
P(\mathcal{D} \mid \theta)=\theta^{\alpha_{H}}(1-\theta)^{\alpha_{T}}
$$

## Maximum Likelihood Estimation

Data: Observed set $D$ of $\alpha_{H}$ Heads and $\alpha_{T}$ Tails Hypothesis: Binomial distribution
Learning: finding $\theta$ is an optimization problem

- What's the objective function?

MLE: Choose $\theta$ to maximize probability of $D$

$$
\begin{aligned}
& P(\mathcal{D} \mid \theta)=\theta^{\alpha_{H}}(1-\theta)^{\alpha_{T}} \\
\hat{\theta} & =\arg \max _{\theta} P(\mathcal{D} \mid \theta) \\
& =\arg \max _{\theta} \ln P(\mathcal{D} \mid \theta)
\end{aligned}
$$

## Parameter learning

$\hat{\theta}=\arg \max _{\theta} \ln P(\mathcal{D} \mid \theta)$
$=\arg \max _{\theta} \ln \theta^{\alpha_{H}}(1-\theta)^{\alpha_{T}}$
Set derivative to zero, and solve!

$$
\begin{aligned}
& \frac{d}{d \theta} \ln P(\mathcal{D} \mid \theta)=\frac{d}{d \theta}\left[\ln \theta^{\alpha_{H}}(1-\theta)^{\alpha_{T}}\right] \\
& \quad=\frac{d}{d \theta}\left[\alpha_{H} \ln \theta+\alpha_{T} \ln (1-\theta)\right] \\
& \quad=\alpha_{H} \frac{d}{d \theta} \ln \theta+\alpha_{T} \frac{d}{d \theta} \ln (1-\theta) \\
& \quad=\frac{\alpha_{H}}{\theta}-\frac{\alpha_{T}}{1-\theta}=0 \quad \hat{\theta}_{M L E}=\frac{\alpha_{H}}{\alpha_{H}+\alpha_{T}}
\end{aligned}
$$

## But, how many flips do I need?

$$
\hat{\theta}_{M L E}=\frac{\alpha_{H}}{\alpha_{H}+\alpha_{T}}
$$

3 heads and 2 tails.
$\theta=3 / 5$, I can prove it!
What if I flipped 30 heads and 20 tails?
Same answer, I can prove it!
What's better?
Umm... The more the merrier???

## A bound

## (from Hoeffding's inequality)

For $N=\alpha_{H}+\alpha_{T}$, and $\quad \hat{\theta}_{M L E}=\frac{\alpha_{H}}{\alpha_{H}+\alpha_{T}}$
Let $\theta^{*}$ be the true parameter, for any $\varepsilon>0$ :

$$
P\left(\left|\hat{\theta}-\theta^{*}\right| \geq \epsilon\right) \leq 2 e^{-2 N \epsilon^{2}}
$$



## What if I have prior beliefs?

Wait, I know that the thumbtack is "close" to 50-50. What can you do for me now?

Rather than estimating a single $\theta$, we obtain a distribution over possible values of $\theta$

In the beginning


After observations


## How to use Prior

Use Bayes rule:


- Or equivalently: $P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta) P(\theta)$
- Also, for uniform priors:
$\rightarrow$ reduces to MLE objective

$$
P(\theta) \propto 1 \quad P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta)
$$

## Beta prior distribution - $P(\theta)$

$$
P(\theta)=\frac{\theta^{\beta_{H}-1}(1-\theta)^{\beta_{T}-1}}{B\left(\beta_{H}, \beta_{T}\right)} \sim \operatorname{Beta}\left(\beta_{H}, \beta_{T}\right)
$$






Likelihood function:

$$
P(\mathcal{D} \mid \theta)=\theta^{\alpha_{H}}(1-\theta)^{\alpha_{T}}
$$ Posterior:

$$
P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta) P(\theta)
$$

$P(\theta \mid \mathcal{D}) \propto \theta^{\alpha_{H}}(1-\theta)^{\alpha_{T}} \theta^{\beta_{H}-1}(1-\theta)^{\beta_{T}-1}$

$$
=\theta^{\alpha_{H}+\beta_{H}-1}(1-\theta)^{\alpha_{T}+\beta_{t}+1}
$$

$$
=\operatorname{Beta}\left(\alpha_{H}+\beta_{H}, \alpha_{T}+\beta_{T}\right)
$$

## MAP for Beta distribution



$$
P(\theta \mid \mathcal{D})=\frac{\theta^{\beta_{H}+\alpha_{H}-1}(1-\theta)^{\beta_{T}+\alpha_{T}-1}}{B\left(\beta_{H}+\alpha_{H}, \beta_{T}+\alpha_{T}\right)} \sim \operatorname{Beta}\left(\beta_{H}+\alpha_{H}, \beta_{T}+\alpha_{T}\right)
$$

MAP: use most likely parameter:

$$
\hat{\theta}=\arg \max _{\theta} P(\theta \mid \mathcal{D})=\frac{\alpha_{H}+\beta_{H}-1}{\alpha_{H}+\beta_{H}+\alpha_{T}+\beta_{T}-2}
$$

## What about continuous variables?

## We like Gaussians because

Affine transformation (multiplying by scalar and adding a constant) are Gaussian

- $X \sim N\left(\mu, \sigma^{2}\right)$
- $\mathrm{Y}=\mathrm{aX}+\mathrm{b} \rightarrow \mathrm{Y} \sim N\left(\mathrm{a} \mu+\mathrm{b}, \mathrm{a}^{2} \sigma^{2}\right)$

Sum of Gaussians is Gaussian

- $X \sim N\left(\mu_{x}, \sigma^{2}{ }_{x}\right)$
- $\left.Y \sim N\left(\mu_{Y}, \sigma^{2}\right)^{2}\right)$
- $Z=X+Y \rightarrow Z \sim N\left(\mu_{X}+\mu_{Y}, \sigma_{X}^{2}+\sigma_{Y}^{2}\right)$


Easy to differentiate

## Learning a Gaussian

- Collect a bunch of data
-Hopefully, i.i.d. samples
- e.g., exam scores

| $x_{i}$ | Exam |
| :--- | :--- |
| $i=$ | Score |
| 0 | 85 |
| 1 | 95 |
| 2 | 100 |
| 3 | 12 |
| $\ldots$ | $\ldots$ |
| 99 | 89 |

- Learn parameters
-Mean: $\mu$
-Variance: $\sigma$

$$
P(x \mid \mu, \sigma)=\frac{1}{\sigma \sqrt{2 \pi}} e^{\frac{-(x-\mu)^{2}}{2 \sigma^{2}}}
$$

MLE for Gaussian: $\quad P(x \mid \mu, \sigma)=\frac{1}{\sigma \sqrt{2 \pi}} e^{\frac{-(x-\mu)^{2}}{2 \sigma^{2}}}$
Prob. of i.i.d. samples $D=\left\{x_{1}, \ldots, x_{N}\right\}$ :

$$
\begin{aligned}
& P(\mathcal{D} \mid \mu, \sigma)=\left(\frac{1}{\sigma \sqrt{2 \pi}}\right)^{N} \prod_{i=1}^{N} e^{\frac{-\left(x_{i}-\mu\right)^{2}}{2 \sigma^{2}}} \\
& \quad \mu_{M L E}, \sigma_{M L E}=\arg \max _{\mu, \sigma} P(\mathcal{D} \mid \mu, \sigma)
\end{aligned}
$$

- Log-likelihood of data:

$$
\begin{aligned}
\ln P(\mathcal{D} \mid \mu, \sigma) & =\ln \left[\left(\frac{1}{\sigma \sqrt{2 \pi}}\right)^{N} \prod_{i=1}^{N} e^{\frac{-\left(x_{i}-\mu\right)^{2}}{2 \sigma^{2}}}\right] \\
& =-N \ln \sigma \sqrt{2 \pi}-\sum_{i=1}^{N} \frac{\left(x_{i}-\mu\right)^{2}}{2 \sigma^{2}}
\end{aligned}
$$

## MLE for mean of a Gaussian

What's MLE for mean?

$$
\begin{aligned}
& \frac{d}{d \mu} \ln P(\mathcal{D} \mid \mu, \sigma)=\frac{d}{d \mu}\left[-N \ln \sigma \sqrt{2 \pi}-\sum_{i=1}^{N} \frac{\left(x_{i}-\mu\right)^{2}}{2 \sigma^{2}}\right] \\
&=\frac{d}{d \mu}[-N \ln \sigma \sqrt{2 \pi}]-\sum_{i=1}^{N} \frac{d}{d \mu}\left[\frac{\left(x_{i}-\mu\right)^{2}}{2 \sigma^{2}}\right] \\
&=-\sum_{i=1}^{N} \frac{\left(x_{i}-\mu\right)}{\sigma^{2}}=0 \\
&=-\sum_{i=1}^{N} x_{i}+N \mu=0 \\
& \quad \widehat{\mu}_{M L E}=\frac{1}{N} \sum_{i=1}^{N} x_{i}
\end{aligned}
$$

## MLE for variance

Again, set derivative to zero:

$$
\begin{aligned}
\frac{d}{d \sigma} \ln P(\mathcal{D} \mid \mu, \sigma) & =\frac{d}{d \sigma}\left[-N \ln \sigma \sqrt{2 \pi}-\sum_{i=1}^{N} \frac{\left(x_{i}-\mu\right)^{2}}{2 \sigma^{2}}\right] \\
& =\frac{d}{d \sigma}[-N \ln \sigma \sqrt{2 \pi}]-\sum_{i=1}^{N} \frac{d}{d \sigma}\left[\frac{\left(x_{i}-\mu\right)^{2}}{2 \sigma^{2}}\right] \\
& =-\frac{N}{\sigma}+\sum_{i=1}^{N} \frac{\left(x_{i}-\mu\right)^{2}}{\sigma^{3}}=0
\end{aligned}
$$

$$
\widehat{\sigma}_{M L E}^{2}=\frac{1}{N} \sum_{i=1}^{N}\left(x_{i}-\widehat{\mu}\right)^{2}
$$

## Learning Gaussian parameters

MLE:

$$
\begin{aligned}
\widehat{\mu}_{M L E} & =\frac{1}{N} \sum_{i=1}^{N} x_{i} \\
\widehat{\sigma}_{M L E}^{2} & =\frac{1}{N} \sum_{i=1}^{N}\left(x_{i}-\widehat{\mu}\right)^{2}
\end{aligned}
$$

## Fitting a Gaussian to Skin samples

$$
\widehat{\mu}_{M L E}=\frac{1}{N} \sum_{i=1}^{N} x_{i}
$$

## Skin detection results



Figure 25.3. The figure shows a variety of images toget her with the output of the skin detector of Jones and Rehg applied to the image. Pixels marked black are skin pixels, and white are background. Notice that this process is relatively effective, and could certainly be used to focus attention on, say, faces and hands. Figure from "Statistionl color models with application to skin defection," M.J. Jones and J. Rehg, Proc. Computer Vision and Pattern Recognition, 1999 (C) 1999, IEEE

## Supervised Learning: find $f$

Given: Training set $\left\{\left(x_{i}, y_{i}\right) \mid i=1 \ldots n\right\}$
Find: A good approximation to $f: X \rightarrow Y$

What is x ?
What is $y$ ?

## Simple Example: Digit Recognition

Input: images / pixel grids
Output: a digit 0-9
Setup:

- Get a large collection of example images, each labeled with a digit
- Note: someone has to hand label all this data!
- Want to learn to predict labels of new, future digit images

Features: ?

Screw You, I want to use Pixels :D


0 2

## Lets take a probabilistic approach!!!

Can we directly estimate the data distribution $\mathrm{P}(\mathrm{X}, \mathrm{Y})$ ?
How do we represent these?
How many parameters?

- Prior, P(Y):
- Suppose $Y$ is composed of $k$ classes
- Likelihood, $\mathrm{P}(\mathrm{X} \mid \mathrm{Y})$ :
- Suppose $\mathbf{X}$ is composed of $n$ binary features


## Conditional Independence

$X$ is conditionally independent of $Y$ given $Z$, if the probability distribution for X is independent of the value of $Y$, given the value of $Z$

> e.g.,
> $(\forall i, j, k) P(X=i \mid Y=j, Z=k)=P(X=i \mid Z=k)$

Equivalent to:
$P($ Thunder $\mid$ Rain, Lightning $)=P($ Thunder $\mid$ Lightning $)$

$$
P(X, Y \mid Z)=P(X \mid Z) P(Y \mid Z)
$$

## Naïve Bayes

Naïve Bayes assumption:

- Features are independent given class:

$$
\begin{aligned}
P\left(X_{1}, X_{2} \mid Y\right) & =P\left(X_{1} \mid X_{2}, Y\right) P\left(X_{2} \mid Y\right) \\
& =P\left(X_{1} \mid Y\right) P\left(X_{2} \mid Y\right)
\end{aligned}
$$

- More generally:

$$
P\left(X_{1} \ldots X_{n} \mid Y\right)=\prod_{i} P\left(X_{i} \mid Y\right)
$$

## The Naïve Bayes Classifier

## Given:

- Prior P(Y)
- $n$ conditionally independent features $\mathbf{X}$ given the class $Y$
- For each $X_{i}$, we have likelihood $P\left(X_{i} \mid Y\right)$


Decision rule:

$$
\begin{aligned}
y^{*}=h_{N B}(\mathbf{x}) & =\arg \max _{y} P(y) P\left(x_{1}, \ldots, x_{n} \mid y\right) \\
& =\arg \max _{y} P(y) \prod_{i} P\left(x_{i} \mid y\right)
\end{aligned}
$$

## A Digit Recognizer

Input: pixel grids


Output: a digit 0-9

## Naïve Bayes for Digits (Binary Inputs)

Simple version:

- One feature $\mathrm{F}_{\mathrm{ij}}$ for each grid position <i,j>
- Possible feature values are on / off, based on whether intensity is more or less than 0.5 in underlying image
- Each input maps to a feature vector, e.g.

$$
1 \rightarrow\left\langle F_{0,0}=0 \quad F_{0,1}=0 \quad F_{0,2}=1 \quad F_{0,3}=1 \quad F_{0,4}=0 \ldots F_{15,15}=0\right\rangle
$$

- Here: lots of features, each is binary valued

Naïve Bayes model:

$$
P\left(Y \mid F_{0,0} \ldots F_{15,15}\right) \propto P(Y) \prod_{i, j} P\left(F_{i, j} \mid Y\right)
$$

Are the features independent given class?
What do we need to learn?

## Example Distributions



## MLE for the parameters of NB

Given dataset

- Count( $A=a, B=b)$ number of examples where $A=a$ and $B=b$
MLE for discrete NB, simply:
- Prior:

$$
P(Y=y)=\frac{\operatorname{Count}(Y=y)}{\sum_{y^{\prime}} \operatorname{Count}\left(Y=y^{\prime}\right)}
$$

- Likelihood:

$$
P\left(X_{i}=x \mid Y=y\right)=\frac{\operatorname{Count}\left(X_{i}=x, Y=y\right)}{\sum_{x^{\prime}} \operatorname{Count}\left(X_{i}=x^{\prime}, Y=y\right)}
$$

## Violating the NB assumption

Usually, features are not conditionally independent:

$$
P\left(X_{1} \ldots X_{n} \mid Y\right) \neq \prod_{i} P\left(X_{i} \mid Y\right)
$$

- NB often performs well, even when assumption is violated
- [Domingos \& Pazzani '96] discuss some conditions for good performance


## Smoothing

$P($ features, $C=2)$

$$
P(C=2)=0.1
$$

$$
P(\mathrm{on} \mid C=2)=0.8
$$

$$
P(\mathrm{on} \mid C=3)=0.8
$$

$$
P(\mathrm{on} \mid C=2)=0.1
$$

$$
P(o f f \mid C=2)=0.1
$$

$$
P(\mathrm{on} \mid C=2)=0.01
$$



$$
P(o n \mid C=3)=0.9
$$

$$
P(\text { off } \mid C=3)=0.7
$$

$$
P(\text { on } \mid C=3)=0.0
$$

## 2 wins!!

Does this happen in vision?

## NB \& Bag of words model



## What about real Features? What if we have continuous $X_{i}$ ?

Eg., character recognition: $X_{i}$ is $\mathrm{i}^{\text {th }}$ pixel

Gaussian Naïve Bayes (GNB):

$$
P\left(X_{i}=x \mid Y=y_{k}\right)=\frac{1}{\sigma_{i k} \sqrt{2 \pi}} e^{\frac{-\left(x-\mu_{i k}\right)}{2 \sigma_{i k}^{2}}}
$$

Sometimes assume variance is independent of $Y$ (i.e., $\sigma_{\mathrm{i}}$ ), or independent of $X_{i}$ (i.e., $\sigma_{k}$ ) or both (i.e., os)

## Estimating Parameters

Maximum likelihood estimates:
Mean:

$$
\hat{\mu}_{i k}=\frac{1}{\sum_{j} \delta\left(Y^{j}=y_{k}\right)} \sum_{j} X_{i}^{j} \delta\left(Y^{j}=y_{k}\right)
$$

Variance:

$$
\begin{gathered}
\delta(x)=1 \text { if } x \text { true } \\
\text { else } 0
\end{gathered}
$$

$$
\widehat{\sigma}_{i k}^{2}=\frac{1}{\sum_{j} \delta\left(Y^{j}=y_{k}\right)-1} \sum_{j}\left(X_{i}^{j}-\widehat{\mu}_{i k}\right)^{2} \delta\left(Y^{j}=y_{k}\right)
$$

## another probabilistic approach!!!

Naïve Bayes: directly estimate the data distribution $P(X, Y)$ !

- challenging due to size of distribution!
- make Naïve Bayes assumption: only need $P\left(X_{i} \mid Y\right)$ !

But wait, we classify according to:

- $\max _{\mathrm{Y}} \mathrm{P}(\mathrm{Y} \mid \mathrm{X})$

Why not learn $\mathrm{P}(\mathrm{Y} \mid \mathrm{X})$ directly?

## Discriminative vs. generative

- Generative model
(The artist)

- Discriminative model
(The lousy painter)

- Classification function



## Logistic Regression

Logistic function (Sigmoid):

## Learn $\mathrm{P}(\mathrm{Y} \mid \mathrm{X})$ directly!

- Assume a particular functional form
- Sigmoid applied to a linear function of the data:


$$
\begin{aligned}
& P(Y=1 \mid X)=\frac{1}{1+\exp \left(w_{0}+\sum_{i=1}^{n} w_{i} X_{i}\right)} \\
& P(Y=0 \mid X)=\frac{\exp \left(w_{0}+\sum_{i=1}^{n} w_{i} X_{i}\right)}{1+\exp \left(w_{0}+\sum_{i=1}^{n} w_{i} X_{i}\right)}
\end{aligned}
$$

## Logistic Regression: decision boundary

$$
P(Y=1 \mid X)=\frac{1}{1+\exp \left(w_{0}+\sum_{i=1}^{n} w_{i} X_{i}\right)} \quad P(Y=0 \mid X)=\frac{\exp \left(w_{0}+\sum_{i=1}^{n} w_{i} X_{i}\right)}{1+\exp \left(w_{0}+\sum_{i=1}^{n} w_{i} X_{i}\right)}
$$

- Prediction: Output the Y with highest $\mathrm{P}(\mathrm{Y} \mid \mathrm{X})$
- For binary Y , output $\mathrm{Y}=0$ if

$$
\begin{aligned}
& 1<\frac{P(Y=0 \mid X)}{P(Y=1 \mid X)} \\
& 1<\exp \left(w_{0}+\sum_{i=1}^{n} w_{i} X_{i}\right) \\
& 0<w_{0}+\sum_{i=1}^{n} w_{i} X_{i}
\end{aligned}
$$



A Linear Classifier!

## Loss functions / Learning Objectives: Likelihood v. Conditional Likelihood

Generative (Naïve Bayes) Loss function:
Data likelihood

$$
\begin{aligned}
\ln P(\mathcal{D} \mid \mathbf{w}) & =\sum_{j=1}^{N} \ln P\left(\mathbf{x}^{j}, y^{j} \mid \mathbf{w}\right) \\
& =\sum_{j=1}^{N} \ln P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right)+\sum_{j=1}^{N} \ln P\left(\mathbf{x}^{j} \mid \mathbf{w}\right)
\end{aligned}
$$

But, discriminative (logistic regression) loss function:
Conditional Data Likelihood

$$
\ln P\left(\mathcal{D}_{Y} \mid \mathcal{D}_{\mathbf{X}}, \mathbf{w}\right)=\sum_{j=1}^{N} \ln P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right)
$$

- Doesn't waste effort learning $\mathrm{P}(\mathrm{X})$ - focuses on $\mathrm{P}(\mathrm{Y} \mid \mathrm{X})$ all that matters for classification
- Discriminative models cannot compute $\mathrm{P}\left(\mathbf{x}^{j} \mid \mathbf{w}\right)$ !


## Conditional Log Likelihood

$$
\begin{array}{ll} 
& P(Y=0 \mid \mathbf{X}, \mathbf{w})=\frac{1}{1+\exp \left(w_{0}+\sum_{i} w_{i} X_{i}\right)} \\
l(\mathbf{w}) \equiv \sum_{j} \ln P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right) & P(Y=1 \mid \mathbf{X}, \mathbf{w})=\frac{\exp \left(w_{0}+\sum_{i} w_{i} X_{i}\right)}{1+\exp \left(w_{0}+\sum_{i} w_{i} X_{i}\right)} \\
&
\end{array}
$$

$l(\mathbf{w})=\sum_{j} y^{j} \ln P\left(y^{j}=1 \mid \mathbf{x}^{j}, \mathbf{w}\right)+\left(1-y^{j}\right) \ln P\left(y^{j}=0 \mid \mathbf{x}^{j}, \mathbf{w}\right)$
remaining steps: substitute definitions, expand logs, and simplify
$=\sum_{j} y^{j} \ln \frac{e^{w_{0}+\sum_{i} w_{i} X_{i}}}{1+e^{w_{0}+\sum_{i} w_{i} X_{i}}}+\left(1-y^{j}\right) \ln \frac{1}{1+e^{w_{0}+\sum_{i} w_{i} X_{i}}}$

## Logistic Regression Parameter Estimation: Maximize Conditional Log Likelihood

$$
\begin{aligned}
l(\mathbf{w}) & \equiv \ln \prod_{j} P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right) \\
& =\sum_{j} y^{j}\left(w_{0}+\sum_{i}^{n} w_{i} x_{i}^{j}\right)-\ln \left(1+\exp \left(w_{0}+\sum_{i}^{n} w_{i} x_{i}^{j}\right)\right)
\end{aligned}
$$

Good news: $/(\mathbf{w})$ is concave function of $\mathbf{w}$
$\rightarrow$ no locally optimal solutions!
Bad news: no closed-form solution to maximize /(w)
Good news: concave functions "easy" to optimize

## Optimizing concave function Gradient ascent

Conditional likelihood for Logistic Regression is concave !


Gradient: $\quad \nabla_{\mathbf{w}} l(\mathbf{w})=\left[\frac{\partial l(\mathbf{w})}{\partial w_{0}}, \ldots, \frac{\partial l(\mathbf{w})}{\partial w_{n}}\right]^{\prime}$

Update rule:

$$
\Delta \mathbf{w}=\eta \nabla_{\mathbf{w}} l(\mathbf{w})
$$

$$
w_{i}^{(t+1)} \leftarrow w_{i}^{(t)}+\eta \frac{\partial l(\mathbf{w})}{\partial w_{i}}
$$

Gradient ascent is simplest of optimization approaches

- e.g., Conjugate gradient ascent much better


## Maximize Conditional Log Likelihood: Gradient

 ascent$$
P(Y=1 \mid X, W)=\frac{\exp \left(w_{0}+\sum_{i} w_{i} X_{i}\right)}{1+\exp \left(w_{0}+\sum_{i} w_{i} X_{i}\right)}
$$

$$
l(\mathbf{w})=\sum_{j} y^{j}\left(w_{0}+\sum_{i}^{n} w_{i} x_{i}^{j}\right)-\ln \left(1+\exp \left(w_{0}+\sum_{i}^{n} w_{i} x_{i}^{j}\right)\right)
$$

$$
\frac{\partial l(w)}{\partial w_{i}}=\sum_{j}\left[\frac{\partial}{\partial w} y^{j}\left(w_{0}+\sum_{i} w_{i} x_{i}^{j}\right)-\frac{\partial}{\partial w} \ln \left(1+\exp \left(w_{0}+\sum_{i} w_{i} x_{i}^{j}\right)\right)\right]
$$

$$
\begin{aligned}
& =\sum_{j}\left[y^{j} x_{i}^{j}-\frac{x_{i}^{j} \exp \left(w_{0}+\sum_{i} w_{i} x_{i}^{j}\right)}{1+\exp \left(w_{0}+\sum_{i} w_{i} x_{i}^{j}\right)}\right] \\
& =\sum_{j} x_{i}^{j}\left[y^{j}-\frac{\exp \left(w_{0}+\sum_{i} w_{i} x_{i}^{j}\right)}{1+\exp \left(w_{0}+\sum_{i} w_{i} x_{i}^{j}\right)}\right]
\end{aligned}
$$

$$
\frac{\partial l(w)}{\partial w_{i}}=\sum_{j} x_{i}^{j}\left(y^{j}-P\left(Y^{j}=1 \mid x^{j}, w\right)\right)
$$

## Gradient ascent for LR

Gradient ascent algorithm: (learning rate $\eta>0$ ) do:

$$
w_{0}^{(t+1)} \leftarrow w_{0}^{(t)}+\eta \sum_{j}\left[y^{j}-\widehat{P}\left(Y^{j}=1 \mid \mathbf{x}^{j}, \mathbf{w}\right)\right]
$$

For $i=1 . . . n:(i t e r a t e ~ o v e r ~ w e i g h t s) ~$
until "change" < $\varepsilon$

$$
w_{i}^{(t+1)} \leftarrow w_{i}^{(t)}+\eta \sum_{j} x_{i}^{j}\left[y^{j}-\widehat{P}\left(Y^{j}=1 \mid \mathbf{x}^{j}, \mathbf{w}\right)\right]
$$

Loop over training examples!

## Large parameters...

$$
\frac{1}{1+e^{-a x}}
$$





Maximum likelihood solution: prefers higher weights

- higher likelihood of (properly classified) examples close to decision boundary
- larger influence of corresponding features on decision
- can cause overfitting!!!

Regularization: penalize high weights

- again, more on this later in the quarter


## How about MAP?

$$
p(\mathbf{w} \mid Y, \mathbf{X}) \propto P(Y \mid \mathbf{X}, \mathbf{w}) p(\mathbf{w})
$$

One common approach is to define priors on w

- Normal distribution, zero mean, identity covariance
Often called Regularization

$$
p(\mathrm{w})=\prod_{i} \frac{1}{\kappa \sqrt{2 \pi}} e^{\frac{-w_{i}^{2}}{2 \kappa^{2}}}
$$

- Helps avoid very large weights and overfitting

MAP estimate:

$$
\mathbf{w}^{*}=\arg \max _{\mathbf{w}} \ln \left[p(\mathbf{w}) \prod_{j=1}^{N} P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right)\right]
$$

## $\mathrm{M}(\mathrm{C}) \mathrm{AP}$ as Regularization

$$
\mathbf{w}^{*}=\arg \max _{\mathbf{w}} \ln \left[p(\mathbf{w}) \prod_{j=1}^{N} P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right)\right] \quad p(\mathbf{w})=\prod_{i} \frac{1}{\kappa \sqrt{2 \pi}} e^{\frac{-w_{i}^{2}}{2 \kappa^{2}}}
$$

Add $\log p(w)$ to objective:

$$
\ln p(w) \propto-\frac{\lambda}{2} \sum_{i} w_{i}^{2} \quad \frac{\partial \ln p(w)}{\partial w_{i}}=-\lambda w_{i}
$$

- Quadratic penalty: drives weights towards zero
- Adds a negative linear term to the gradients


## MLE vs. MAP

Maximum conditional likelihood estimate

$$
\begin{aligned}
& \mathbf{w}^{*}=\arg \max _{\mathbf{w}} \ln \left[\prod_{j=1}^{N} P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right)\right] \\
& w_{i}^{(t+1)} \leftarrow w_{i}^{(t)}+\eta \sum_{j} x_{i}^{j}\left[y^{j}-\widehat{P}\left(Y^{j}=1 \mid \mathbf{x}^{j}, \mathbf{w}\right)\right]
\end{aligned}
$$

Maximum conditional a posteriori estimate

$$
\begin{aligned}
& \mathbf{w}^{*}=\arg \max _{\mathbf{w}} \ln \left[p(\mathbf{w}) \prod_{j=1}^{N} P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right)\right] \\
& w_{i}^{(t+1)} \leftarrow w_{i}^{(t)}+\eta\left\{-\lambda w_{i}^{(t)}+\sum_{j} x_{i}^{j}\left[y^{j}-\widehat{P}\left(Y^{j}=1 \mid \mathbf{x}^{j}, \mathbf{w}\right)\right]\right\}
\end{aligned}
$$

## Logistic regression v. Naïve Bayes

Consider learning $\mathrm{f}: \mathrm{X} \rightarrow \mathrm{Y}$, where

- $X$ is a vector of real-valued features, $<X_{1} \ldots X_{n}>$
- Y is boolean

Could use a Gaussian Naïve Bayes classifier

- assume all $X_{i}$ are conditionally independent given $Y$
- model $\mathrm{P}\left(\mathrm{X}_{\mathrm{i}} \mid \mathrm{Y}=\mathrm{y}_{\mathrm{k}}\right)$ as Gaussian $\mathrm{N}\left(\mu_{\mathrm{i} k}, \sigma_{\mathrm{i}}\right)$
- model $P(Y)$ as Bernoulli $(\theta, 1-\theta)$

What does that imply about the form of $\mathrm{P}(\mathrm{Y} \mid \mathrm{X})$ ?

$$
P\left(Y=1 \mid X=<X_{1}, \ldots X_{n}>\right)=\frac{1}{1+\exp \left(w_{0}+\sum_{i} w_{i} X_{i}\right)}
$$

## Derive form for $\mathrm{P}(\mathrm{Y} \mid \mathrm{X})$ for continuous $\mathrm{X}_{\mathrm{i}}$

$$
P(Y=1 \mid X)=\frac{P(Y=1) P(X \mid Y=1)}{P(Y=1) P(X \mid Y=1)+P(Y=0) P(X \mid Y=0)}
$$

$$
\begin{aligned}
& =\frac{1}{1+\frac{P(Y=0) P(X \mid Y=0)}{P(Y=1) P(X \mid Y=1)}} \\
& =\frac{1}{1+\exp \left(\ln \frac{P(Y=0) P(X \mid Y=0)}{P(Y=1) P(X \mid Y=1)}\right)}
\end{aligned}
$$

up to now, all arithmetic
$=\frac{1}{1+\exp \left(\left(\ln \frac{1-\theta}{\theta}\right)+\sum_{i} \ln \frac{P\left(X_{i} \mid Y=0\right)}{P\left(X_{i} \mid Y=1\right)}\right)}$
Can we solve for $w_{i}$ ?
Looks like a setting for $\mathrm{w}_{0}$ ?

- Yes, but only in Gaussian case


## Ratio of class-conditional probabilities

$$
\begin{aligned}
& \ln \frac{P\left(X_{i} \mid Y=0\right)}{P\left(X_{i} \mid Y=1\right)} \\
= & \ln \left[\begin{array}{ll}
\left.\frac{\frac{1}{\sigma_{i} \sqrt{2 \pi}} e^{-\frac{\left(x_{i}-\mu_{i 0}\right)^{2}}{2 \sigma_{i}^{2}}}}{\frac{1}{\sigma_{i} \sqrt{2 \pi}} e^{-\frac{\left(x_{i}-\mu_{i 1}\right)^{2}}{2 \sigma_{i}^{2}}}}\right] & P\left(X_{i}=x \mid Y=y_{k}\right)=\frac{1}{\sigma_{i} \sqrt{2 \pi}} e^{\frac{-\left(x-\mu_{i k}\right)^{2}}{2 \sigma_{i}^{2}}} \\
= & -\frac{\left(x_{i}-\mu_{i 0}\right)^{2}}{2 \sigma_{i}^{2}}+\frac{\left(x_{i}-\mu_{i 1}\right)^{2}}{2 \sigma_{i}^{2}} \\
\cdots & \text { Linear function! } \\
= & \begin{array}{l}
\text { Coefficients } \\
\text { expressed with } \\
\text { original Gaussian }
\end{array} \\
\sigma_{i}^{2} & \text { parameters! }
\end{array}\right. \\
2 \sigma_{i}+\frac{\mu_{i 0}^{2}+\mu_{i 1}^{2}}{2 \sigma_{i}^{2}} &
\end{aligned}
$$

## Derive form for $\mathrm{P}(\mathrm{Y} \mid \mathrm{X})$ for continuous $\mathrm{X}_{\mathrm{i}}$

$$
\begin{aligned}
& P(Y=1 \mid X)=\frac{P(Y=1) P(X \mid Y=1)}{P(Y=1) P(X \mid Y=1)+P(Y=0) P(X \mid Y=0)} \\
&=\frac{1}{1+\exp \left(\left(\ln \frac{1-\theta}{\theta}\right)+\sum_{i} \ln \frac{P\left(X_{i} \mid Y=0\right)}{P\left(X_{i} \mid Y=1\right)}\right)} \\
& P(Y=1 \mid X)=\frac{\sum_{i}\left(\frac{\mu_{i 0}-\mu_{i 1}}{\sigma_{i}^{2}} X_{i}+\frac{\mu_{i 1}^{2}-\mu_{i 0}^{2}}{2 \sigma_{i}^{2}}\right)}{1+\exp \left(w o+\sum_{i=1}^{n} w_{i} X_{i}\right)} \\
& w_{0}=\ln \frac{1-\theta}{\theta}+\frac{\mu_{i 0}^{2}+\mu_{i 1}^{2}}{2 \sigma_{i}^{2}} \quad w_{i}=\frac{\mu_{i 0}+\mu_{i 1}}{\sigma_{i}^{2}}
\end{aligned}
$$

## Gaussian Naïve Bayes vs. Logistic Regression

Set of Gaussian Naïve Bayes parameters
(feature variance
independent of class label)


## Set of Logistic <br> Regression parameters

Representation equivalence

- But only in a special case!!! (GNB with class-independent variances)
But what's the difference???
LR makes no assumptions about $\mathrm{P}(\mathrm{X} \mid \mathrm{Y})$ in learning!!!
Loss function!!!
- Optimize different functions ! Obtain different solutions


## Naïve Bayes vs. Logistic Regression

Consider $Y$ boolean, $X_{i}$ continuous, $X=<X_{1} \ldots X_{n}>$

Number of parameters:
Naïve Bayes: $4 \mathrm{n}+1$
Logistic Regression: n+1

Estimation method:
Naïve Bayes parameter estimates are uncoupled
Logistic Regression parameter estimates are coupled

## Naïve Bayes vs. Logistic Regression

[ Ng \& Jordan, 2002]
Generative vs. Discriminative classifiers
Asymptotic comparison
(\# training examples $\rightarrow$ infinity)

- when model correct
- GNB (with class independent variances) and LR produce identical classifiers
- when model incorrect
- LR is less biased - does not assume conditional independence
» therefore LR expected to outperform GNB


## Naïve Bayes vs. Logistic Regression

[ Ng \& Jordan, 2002]
Generative vs. Discriminative classifiers Non-asymptotic analysis

- convergence rate of parameter estimates, ( $\mathrm{n}=$ \# of attributes in X)
- Size of training data to get close to infinite data solution
- Naïve Bayes needs O(log n) samples
- Logistic Regression needs O(n) samples
- GNB converges more quickly to its (perhaps less helpful) asymptotic estimates


## What you should know about Logistic Regression (LR)

Gaussian Naïve Bayes with class-independent variances representationally equivalent to LR

- Solution differs because of objective (loss) function

In general, NB and LR make different assumptions

- NB: Features independent given class ! assumption on $\mathrm{P}(\mathbf{X} \mid \mathrm{Y})$
- LR: Functional form of $\mathrm{P}(\mathrm{Y} \mid \mathbf{X})$, no assumption on $\mathrm{P}(\mathbf{X} \mid \mathrm{Y})$

LR is a linear classifier

- decision rule is a hyperplane

LR optimized by conditional likelihood

- no closed-form solution
- concave ! global optimum with gradient ascent
- Maximum conditional a posteriori corresponds to regularization

Convergence rates

- GNB (usually) needs less data
- LR (usually) gets to better solutions in the limit


## Decision Boundary



## Voting (Ensemble Methods)

Instead of learning a single classifier, learn many weak classifiers that are good at different parts of the data
Output class: (Weighted) vote of each classifier

- Classifiers that are most "sure" will vote with more conviction
- Classifiers will be most "sure" about a particular part of the space
- On average, do better than single classifier!

But how???

- force classifiers to learn about different parts of the input space? different subsets of the data?
- weigh the votes of different classifiers?


## BAGGing $=\underline{\text { Bootstrap AGGregation }}$

## (Breiman, 1996)

- for $\mathrm{i}=1,2, \ldots, \mathrm{~K}$ :
$-T_{i} \leftarrow$ randomly select $M$ training instances with replacement
$-\mathrm{h}_{\mathrm{i}} \leftarrow \operatorname{learn}\left(\mathrm{T}_{\mathrm{i}}\right) \quad$ [ID3, $N B, k N N$, neural net, ...]
- Now combine the $T_{i}$ together with uniform voting ( $w_{i}=1 / K$ for all $i$ )


## Bagging Example



## Decision Boundary



## 100 bagged trees


shades of blue/red indicate strength of vote for particular classification

## Fighting the bias-variance tradeoff

Simple (a.k.a. weak) learners are good

- e.g., naïve Bayes, logistic regression, decision stumps (or shallow decision trees)
- Low variance, don't usually overfit

Simple (a.k.a. weak) learners are bad

- High bias, can't solve hard learning problems

Can we make weak learners always good???

- No!!!
- But often yes...


## Boosting

Idea: given a weak learner, run it multiple times on
(reweighted) training data, then let learned classifiers vote

On each iteration $t$ :

- weight each training example by how incorrectly it was classified
- Learn a hypothesis - $h_{t}$
- A strength for this hypothesis $-\alpha_{t}$

Final classifier:

$$
h(x)=\operatorname{sign}\left(\sum_{i} \alpha_{i} h_{i}(x)\right)
$$

## Practically useful

Theoretically interesting

```
                            Iools Help
* \ 人 (x) http://www1.cs.columbie.edu/freund/daboost/
```



First, generate a data-set by clicking on the left and right buttons in the main window of the applet. Then, press "split" to split the data into training and test set


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```
                            Iools Help
* \ 人 (x) h http://www1.cs.columbi.edu/freund/daboost/
```


time $=3$

First, generate a data-set by clicking on the left and right buttons in the main window of the applet. Then, press "split" to split the data into training and test sets

## Applet adaboost started



First, generate a data-set by clicking on the left and right buttons in the main window of the applet. Then, press "split" to split the data into training and test sets

## Learning from weighted data

## Consider a weighted dataset

- $\mathrm{D}(\mathrm{i})$ - weight of $i$ th training example ( $\left.\mathbf{x}^{i}, \mathrm{y}^{i}\right)$
- Interpretations:
- ith training example counts as if it occurred D (i) times
- If I were to "resample" data, I would get more samples of "heavier" data points
Now, always do weighted calculations:
- e.g., MLE for Naïve Bayes, redefine $\operatorname{Count}(Y=y)$ to be weighted count:

$$
\operatorname{Count}(Y=y)=\sum_{j=1}^{n} D(j) \delta\left(Y^{j}=y\right)
$$

- setting $D(j)=1$ (or any constant value!), for all $j$, will recreates unweighted case

Given: $\left(x_{1}, y_{1}\right), \ldots,\left(x_{m}, y_{m}\right)$ where $x_{i} \in X, y_{i} \in Y=\{-1,+1\}$
Initialize $D_{1}(i)=1 / \mathrm{m}$.
For $t=1, \ldots, T$ :
How? Many possibilities. Will see one shortly!

- Train base learner using distribution $D_{t}$.
- Get base classifier $h_{t} \cdot X \rightarrow \mathbb{R}$.
- Choose $\alpha_{t} \in \mathbb{R}$.
- Update:

$$
D_{t+1}(i)=\frac{D_{t}(i) \exp \left(-\alpha_{t} y_{i} h_{t}\left(x_{i}\right)\right)}{Z_{t}}
$$

where $Z_{t}$ is a normalization factor

$$
Z_{t}=\sum_{i=1}^{m} D_{t}(i) \exp \left(-\alpha_{t} y_{i} h_{t}\left(x_{i}\right)\right)
$$

Output the final classifier:

$$
H(x)=\operatorname{sign}\left(\sum_{t=1}^{T} \alpha_{t} h_{t}(x)\right) . \quad \begin{aligned}
& \text { "base" or "weak" classifier } \\
& \text { outputs. }
\end{aligned}
$$

Figure 1: The boosting algorithm AdaBoost.

Given: $\left(x_{1}, y_{1}\right), \ldots,\left(x_{m}, y_{m}\right)$

$$
\begin{array}{r}
\epsilon_{t}=P_{i \sim D_{t}(i)}\left[h_{t}\left(\mathbf{x}^{i}\right) \neq y^{i}\right] \\
\epsilon_{t}=\sum_{i=1}^{m} D_{t}(i) \delta\left(h_{t}\left(x_{i}\right) \neq y_{i}\right)
\end{array}
$$

Initialize $D_{1}(i)=1 / m$.
For $t=1, \ldots, T$ :

- Train base learner using distribution $D_{t}$.
- Get base classifier $h_{t}: X \rightarrow \mathbb{R}$.
- Choose $\alpha_{t} \in \mathbb{R}$.

$$
\alpha_{t}=\frac{1}{2} \ln \left(\frac{1-\epsilon_{t}}{\epsilon_{t}}\right)
$$

- Update:

$$
D_{t+1}(i)=\frac{D_{t}(i) \exp \left(-\alpha_{t} y_{i} h_{t}\left(x_{i}\right)\right)}{Z_{t}}
$$

where $Z_{t}$ is a normalization factor (chosen so that $D_{t+1}$ will be a distribution).

Output the final classifier:

$$
H(x)=\operatorname{sign}\left(\sum_{t=1}^{T} \alpha_{t} h_{t}(x)\right) .
$$

Figure 1: The boosting algorithm AdaBoost.

## What $\alpha_{t}$ to choose for hypothesis $h_{t}$ ?

[Schapire, 1989]
Idea: choose $\alpha_{t}$ to minimize a bound on training error!

$$
\frac{1}{m} \sum_{i=1}^{m} \delta\left(H\left(x_{i}\right) \neq y_{i}\right) \leq \frac{1}{m} \sum_{i=1}^{m} \exp \left(-y_{i} f\left(x_{i}\right)\right)
$$

Where $f(x)=\sum_{t} \alpha_{t} h_{t}(x) ; H(x)=\operatorname{sign}(f(x))$


## What $\alpha_{t}$ to choose for hypothesis $h_{t}$ ?

[Schapire, 1989]
Idea: choose $\alpha_{t}$ to minimize a bound on training error!

$$
\begin{aligned}
& \frac{1}{m} \sum_{i=1}^{m} \delta\left(H\left(x_{i}\right) \neq y_{i}\right) \leq \frac{1}{m} \sum_{i} \exp \left(-y_{i}\right. \\
& \text { here } \\
& f(x)=\sum_{t} \alpha_{t} h_{t}(x) ; H(x)=\operatorname{sign}(f(x))
\end{aligned}
$$

And

$$
Z_{t}=\sum_{i=1}^{m} D_{t}(i) \exp \left(-\alpha_{t} y_{i} h_{t}\left(x_{i}\right)\right)
$$

This equality isn't obvious! Can be shown with algebra (telescoping sums)!

If we minimize $\prod_{t} Z_{t}$, we minimize our training error!!!
We can tighten this bound greedily, by choosing $\alpha_{t}$ and $h_{t}$ on each iteration to minimize $Z_{t}$.
$h_{t}$ is estimated as a black box, but can we solve for $\alpha_{t}$ ?

Summary: choose $\alpha_{t}$ to minimize error bound
[Schapire, 1989]
We can squeeze this bound by choosing $\alpha_{t}$ on each iteration to minimize $Z_{t}$.

$$
\begin{aligned}
& Z_{t}=\sum_{i=1}^{m} D_{t}(i) \exp \left(-\alpha_{t} y_{i} h_{t}\left(x_{i}\right)\right) \\
& \epsilon_{t}=\sum_{i=1}^{m} D_{t}(i) \delta\left(h_{t}\left(x_{i}\right) \neq y_{i}\right)
\end{aligned}
$$

For boolean Y : differentiate, set equal to 0 , there is a closed form solution! [Freund \& Schapire '97]:

$$
\alpha_{t}=\frac{1}{2} \ln \left(\frac{1-\epsilon_{t}}{\epsilon_{t}}\right)
$$

## Strong, weak classifiers

If each classifier is (at least slightly) better than random: $\varepsilon_{t}<0.5$
Another bound on error:

$$
\frac{1}{m} \sum_{i=1}^{m} \delta\left(H\left(x_{i}\right) \neq y_{i}\right) \leq \prod_{t} Z_{t} \leq \exp \left(-2 \sum_{t=1}^{T}\left(1 / 2-\epsilon_{t}\right)^{2}\right)
$$

What does this imply about the training error?

- Will get there exponentially fast!

Is it hard to achieve better than random training error?

## Boosting results - Digit recognition

[Schapire, 1989]


## Boosting:

- Seems to be robust to overfitting
- Test error can decrease even after training error is zero!!!


## Boosting generalization error bound

[Freund \& Schapire, 1996]
error $_{\text {true }}(H) \leq$ error $_{\text {train }}(H)+\tilde{\mathcal{O}}$
Constants:
$T$ : number of boosting rounds

- Higher $\mathrm{T} \rightarrow$ Looser bound, what does this imply?
d: VC dimension of weak learner, measures complexity of classifier
- Higher $d \rightarrow$ bigger hypothesis space $\rightarrow$ looser bound
$m$ : number of training examples
- more data $\rightarrow$ tighter bound


## Boosting generalization error bound

[Freund \& Schapire, 1996]

Constants:

- Theory does not match practice:
- Robust to overfitting
- Test set error decreases even after training error is zero
- Need better analysis tools
- we'll come back to this later in the quarter
- more data $\rightarrow$ tignter bound


## Logistic Regression as Minimizing Loss

Logistic regression assumes:

$$
P(Y=1 \mid X)=\frac{1}{1+\exp (f(x))} \quad f(x)=w_{0}+\sum_{i} w_{i} h_{i}(x)
$$

And tries to maximize data likelihood, for $Y=\{-1,+1\}$ :

$$
\begin{aligned}
P\left(y_{i} \mid \mathbf{x}_{i}\right)=\frac{1}{1+e^{-y_{i} f\left(\mathbf{x}_{i}\right)}} \ln P\left(\mathcal{D}_{Y} \mid \mathcal{D}_{\mathbf{X}}, \mathbf{w}\right) & =\sum_{j=1}^{N} \ln P\left(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}\right) \\
= & -\sum_{i=1}^{m} \ln \left(1+\exp \left(-y_{i} f\left(x_{i}\right)\right)\right)
\end{aligned}
$$

Equivalent to minimizing log loss:

$$
\sum_{i=1}^{m} \ln \left(1+\exp \left(-y_{i} f\left(x_{i}\right)\right)\right)
$$

## Boosting and Logistic Regression

Logistic regression equivalent to minimizing log loss:

Both smooth approximations of $0 / 1$ loss!

