Bayesian Learning

Instructor: Jesse Davis



- Homework 1 is due today
- Homework 2 is out
- Slides for this lecture are online
- We'll review some of homework 1 next class
 - Techniques for efficient implementation of collaborative filtering
 - Common mistakes made on rest of HW



- Probability overview
- Naïve Bayes
- Bayesian learning
- Bayesian networks

Random Variables

- A random variable is a number (or value) determined by chance
 - More formally it is drawn from a probability distribution
- Types of random variables
 - Continuous
 - Binary
 - Discrete

Why Random Variables

- Our goal is to predict a target variable
- We are not given the true function
- We are given observations
 - Number of times a dice lands on 4
 - Can estimate the probability of this event
 - We don't know where the dice will land
 - Can only guess what is likely to happen

- Bernoulli RV takes two values: 0 and 1
 Prob(1) = p and P(0) = 1 p
 P(x) = $\begin{cases} p^{x}(1-p)^{x}, & \text{if } x = 0 \text{ or } 1\\ 0, & \text{otherwise} \end{cases}$
- The performance of one trial with fixed probability of success (p) is a Bernoulli trial

- Like Bernoulli distribution, two values: 0 or 1 and probability P(1)=p and P(0)=1-p
- What is the probability of k successes, P(k), in a series of n independent trials? (n>=k)
- P(k) is a binomial random variable: P(x) = $\begin{bmatrix} n \\ k \end{bmatrix} p^k (1-p)^{n-k}$, where $\begin{bmatrix} n \\ k \end{bmatrix} = \frac{n!}{k!(n-k)!}$
- Bernoulli distribution is a special case of the binomial distribution (i.e., n=1)

Multinomial Distribution

- Generalizes binomial distribution to multiple outputs (classes)
- N independent trials
 - r possible outcomes
 - Each outcome c_r has $P(c_r) = p_r$
 - $\Sigma P(c_r) = 1$
- Multinomial RV: Probability that in n trials, the frequency of the r classes is (n₁,...,n_r)

$$P(x) = \begin{bmatrix} n \\ n_1 \dots n_r \end{bmatrix} p_1^{n_1 *} \dots * p_r^{n_r}, \text{ where } \begin{bmatrix} n \\ n_1 \dots n_r \end{bmatrix} = \frac{n!}{n_1! * \dots * n_r!}$$

Axioms of Probability Theory

Just three are enough to build entire theory!

1. All probabilities between 0 and 1

 $0 \leq P(A) \leq 1$

- 2. P(true) = 1 and P(false) = 0
- 3. Probability of disjunction of events is: $P(A \lor B) = P(A) + P(B) - P(A \land B)$



Conditional Probability

 P(A | B) is the probability of A given B
 Assumes that B is the only info known.
 Defined as P(A | B) = P(A \wedge B) P(B)





- A and B are independent iff:
 - $\bullet P(A \mid B) = P(A)$

$$\bullet P(B \mid A) = P(B)$$

These two constraints are logically equivalent

Therefore if A and B are independent

$$P(A | B) = \frac{P(A \land B)}{P(B)} = P(A)$$
$$P(A \land B) = P(A)P(B)$$





Independence is powerful, but rarely holds

Conditional Independence



But: A&B are *made* independent by ¬C





Bayes rule is:
$$P(A | B) = \frac{P(B | A) P(A)}{P(B)}$$

Proof: $P(A \wedge B)$ $P(A \mid B)$ P(B)Defn of cond. prob P(A ^ B) P(B | A) =P(A)Rearrange line 2 $P(A \land B) = P(B \mid A) P(A)$ P(B | A) P(A)Sub in prev result P(A | B) P(B)15 Use to Compute Diagnostic Probability from Causal Probability

- For example, let M be meningitis, S be stiff neck
 P(M) = 0.0001
 - P(S) = 0.1
 - P(S|M) = 0.8
- P(M | S) = 0.8 x 0.0001 / 0.1 = 0.0008
- Probability is very low!



- Probability overview
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Naïve Bayes: Motivation

- We will see many draws of X₁,...,X_n and the response (class) Y
- We want to estimate the most likely value of Y for the input, that is, P(Y| X₁,...,X_n)
- What difficulty arises?
 - Exponentially many settings for X₁,...,X_n
 - Next case probably has not been seen

One Approach: Assume Conditional Independence

By Bayes Rule (with normalization):

- $P(Y|X_1,...,X_n) = aP(X_1,...,X_n|Y)P(Y)$
- Normalization: Compute above for each value of Y, then normalize so sum to 1
- Recall Conditional independence: $P(X_1,...,X_n|Y) = P(X_1|Y)...P(X_n|Y)$

•
$$P(Y|X_1,...,X_n) = aP(X_1|Y)...P(X_n|Y)P(Y)$$

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- Assumes (naïvely) that all features are conditionally independent given the class
 - $P(A \land B \mid Class) = P(A \mid Class) * P(B \mid Class)$
 - Avoids estimating P(A ^ B), etc.
- Surprisingly, though the assumption is often violated naïve Bayes works well in practice
 - Bag of words for text, spam filtering, etc.

Naïve Bayes in Practice

 Empirically, estimates <u>relative</u> probabilities more reliably than absolute ones:

 $\frac{P(Pos | Features)}{P(Neg | Features)} = \frac{P(Features | Pos) * P(Pos)}{P(Features | Neg) * P(Neg)}$

Better than

P(Pos | Features) = P(Features | Pos) * P(Pos)

 Naïve Bayes tends to push probability estimates towards either 0 or 1

Technical Detail: Underflow

Assume we have 100 features

- We multiple 100 numbers in [0,1]
- If values are small, we are likely to `underflow' the min positive/float value
- Solution: $\prod \text{ probs} = e^{\sum \log(\text{prob})}$
- Sum log's of prob's
- Subtract logs since log $\frac{P(+)}{P(-)} = \log P(+) \log P(-)$



$Odds = \frac{P(F_1 \mid Pos)^* ...^* P(F_n \mid Pos) * P(Pos)}{P(F_1 \mid Neg)^* ...^* P(F_n \mid Neg) * P(Neg)}$

 $log(Odds) = [\Sigma log{ P(F_i | Pos) / P(F_i | Neg)}] + log(P(Pos) / P(Neg))$

Notice if a feature value is more likely in a pos, the log is pos and if more likely in neg, the log is neg (0 if tie)

Naïve Bayes Example

Color	Shape	Size	Category
red	•	big	+
blue	Δ	small	+
red		small	+
red	Δ	big	
blue	•	small	
red	Δ	small	?

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• For the new example (red, Δ , small) $\frac{P(+|F's)}{P(-|F's)} = \frac{P(red|+)*P(\Delta|+)*P(small|+)*P(+)}{P(-|F's)} = \frac{P(red|-)*P(\Delta|-)*P(small|-)*P(-)}{P(red|-)*P(\Delta|-)*P(small|-)*P(-)} = \frac{2/3*1/3*2/3*3/5}{1/2*1/2*1/2*2/5} = 1.77$

So most likely a POS example

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Dealing with Zeroes (and Small Samples)

- If we never see something (eg, in the train set), should we assume its probability is zero?
- If we only see 3 pos ex's and 2 are red, do we really think

P(red | pos) = 2/3 ?

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M-estimates (Eq 6.22 in Mitchell; Eq 7 in draft chapter)

- Imagine we had *m* hypothetical pos ex's
- Assume p is prob these examples are red
- Improved estimate:

$$P(red | pos) = 2 + p * m$$

3 + m

(In general, **red** is some feature value and 2 and 3 are actual counts in the training set)

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Laplace Smoothing

Special case of *m* estimates

- Let *m* = #colors, *p* = 1/m
- Ie, assume one hypothetical pos ex of each color
- Implementation trick

. . .

- Start all counters at 1 instead of 0
 - Eg, initialize count(pos, feature(i), value(i, j)) = 1
 - count(pos, color, red), count(neg, color, red), count(pos, color, blue),



Node *i* stores P(F_i | POS) and P(F_i | NEG)

How Does Naïve Bayes Partition Feature Space?



Homework: Spam Filtering

Task:

From: Branded anti-ED Pills <otubu9068@telesp.net.br> To: andrey.kolobov@gmail.com Date: Fri, Apr 2, 2010 at 7:23 PM Subject: Hot Sale, andrey.kolobov! 77% off on top goods Emen Mailed-by: telesp.net.br

Why aren't you on our site, andrey.kolobov? We have 77% off today!!

P(E|C)

- Ham or Spam? P(C) P(C|E) = P(E|C)P(C) / P(E) $C \leftarrow argmax \{ P(E|C)P(C) \}$
- $C \leftarrow argmax \{ P(E|C)P(C) \}$ C in {h, m}

... with Naïve Bayes



 $C \leftarrow \operatorname{argmax} \{ P(E|C)P(C) \} = \operatorname{argmax} \{ P(C) \prod P(W | C) \}$ C in {h, m} $C \text{ in } \{h, m\}$ W in E

Estimating Parameters

- Given:
 - Set of training spam emails S
 - Set of training ham emails H
- Probabilities:

P(w)

$$c) = (1 + (\#_c w)) / (\sum_{w' \text{ in } V} \#_c w'))$$

• P(c) = |c| / (|S| + |H|)

To avoid getting
P(w|c) = 0 due to data sparsity

Naïve Bayes Summary

- Fast, simple algorithm
- Effective in practice [good baseline comparison]
- Gives estimates of confidence in class label
- Makes simplifying assumptions
- Extensions to come...


- Probability overview
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$P(H|C_1) = 0.1$ $P(H|C_2) = 0.5$ $P(H|C_3) = 0.9$

Which coin will I use? $P(C_1) = 1/3$ $P(C_2) = 1/3$ $P(C_3) = 1/3$

Prior: Probability of a hypothesis before we make any observations



$P(H|C_1) = 0.1$ $P(H|C_2) = 0.5$ $P(H|C_3) = 0.9$

Which coin will I use? $P(C_1) = 1/3$ $P(C_2) = 1/3$ $P(C_3) = 1/3$

Uniform Prior: All hypothesis are equally likely before we make any observations



Which coin <u>did</u> I use? $P(C_1|H) = 0.066 P(C_2|H) = 0.333 P(C_3|H) = 0.6$

Posterior: Probability of a hypothesis given data



Terminology

- Prior: Probability of a hypothesis before we see any data
- Uniform prior: A prior that makes all hypothesis equally likely
- Posterior: Probability of hypothesis after we saw some data
- Likelihood: Probability of the data given the hypothesis

Which coin did I use? $P(C_1|HT) = ? P(C_2|HT) = ? P(C_3|HT) = ?$

 $P(C_1|HT) = \alpha P(HT|C_1)P(C_1) = \alpha P(H|C_1)P(T|C_1)P(C_1)$



Which coin <u>did</u> I use? $P(C_1|HT) = 0.21P(C_2|HT) = 0.58P(C_3|HT) = 0.21$

 $P(C_1|HT) = \alpha P(HT|C_1)P(C_1) = \alpha P(H|C_1)P(T|C_1)P(C_1)$





Which coin <u>did</u> I use? $P(C_1|HT) = 0.21P(C_2|HT) = 0.58P(C_3|HT) = 0.21$





What is the probability of heads after two experiments?



Your Estimate?

Maximum Likelihood Estimate: The best hypothesis that fits observed data assuming uniform prior

Most likely coin:

 C_2

Best estimate for P(H)

$$P(H|C_2) = 0.5$$



Using Prior Knowledge

- Should we always use a Uniform Prior ?
- Background knowledge:
 - Heads => we have take-home midterm
 - Jesse likes take-homes...
 - => Jesse is more likely to use a coin biased in his favor





We can encode it in the prior:



 $P(C_3) = 0.70$







 $P(H|C_3) = 0.9$ $P(H|C_1) = 0.1$ $P(H|C_2) = 0.5$

Experiment 1: Heads Which coin did I use? $P(C_1|H) = ?$ $P(C_2|H) = ?$ $P(C_3|H) = ?$

 $P(C_1|H) = \alpha P(H|C_1)P(C_1)$



Experiment 1: Heads

Which coin did I use?

 $P(C_1|H) = 0.006 P(C_2|H) = 0.165 P(C_3|H) = 0.829$

Compare with ML posterior after Exp 1: $P(C_1|H) = 0.066 P(C_2|H) = 0.333 P(C_3|H) = 0.600$





Which coin did l use? $P(C_1|HT) = ? P(C_2|HT) = ? P(C_3|HT) = ?$

 $P(C_1|HT) = \alpha P(HT|C_1)P(C_1) = \alpha P(H|C_1)P(T|C_1)P(C_1)$





Which coin did I use? $P(C_1|HT) = 0.035 P(C_2|HT) = 0.481 P(C_3|HT) = 0.485$

 $P(C_1|HT) = \alpha P(HT|C_1)P(C_1) = \alpha P(H|C_1)P(T|C_1)P(C_1)$





Which coin did I use? $P(C_1|HT) = 0.035 P(C_2|HT) = 0.481 P(C_3|HT) = 0.485$





What is the probability of heads after two experiments?

Most likely coin:

 C_3

Best estimate for P(H)

 $P(H|C_3) = 0.9$



Your Estimate?

Maximum A Posteriori (MAP) Estimate: The best hypothesis that fits observed data assuming a non-uniform prior

Most likely coin:



Best estimate for P(H) P(H|C₃) = 0.9





$P(C_1|HT)=0.035$ $P(C_2|HT)=0.481$ $P(C_3|HT)=0.485$





$P(C_1|HT) = 0.035 P(C_2|HT) = 0.481 P(C_3|HT) = 0.485$

C₂ and C₃ are almost equally likely



$P(H|C_1) = 0.1$ $P(H|C_2) = 0.5$ $P(H|C_3) = 0.9$



Recall: $P(H) = \sum_{i=1}^{n} P(H|C_i)P(C_i) = 0.680$ i=1

A Better Estimate

Bayesian Estimate

Bayesian Estimate: Minimizes prediction error, given data and (generally) assuming a non-uniform prior

$$P(H) = \sum_{i=1}^{3} P(H|C_i) P(C_i) = 0.680$$

 $P(C_1|HT)=0.035$ $P(C_2|HT)=0.481$ $P(C_3|HT)=0.485$



Comparison After more Experiments

- Seen: HTHHHHHHHH
- Maximum likelihood:
 - P(H) = 0.5
 - After 10 experiments: P(H) = 0.9
- Maximum a posteriori:
 - P(H) = 0.9
 - After 10 experiments: P(H) = 0.9
- Bayesian:
 - P(H) = 0.68
 - After 10 experiments: P(H) = 0.9

Comparison

- ML:
 - Easy to compute
- MAP:
 - Easy to compute
 - Incorporates prior knowledge
- Bayesian:
 - Minimizes error -> great with little data
 - Potentially very difficult to compute

Brute-Force MAP Hypothesis Learner

1. For each hypothesis h in H, calculate the posterior probability

$$P(h|D) = \frac{P(D|h)P(h)}{P(D)}$$

2. Output the hypothesis h_{MAP} with the highest posterior probability

$$h_{MAP} = \operatorname*{argmax}_{h \in H} P(h|D)$$

Relation to Concept Learning

Let $D = \langle c(x_1), \dots, c(x_m) \rangle$ (examples' classes) Choose P(D|h)

- P(D|h) = 1 if h consistent with D
- P(D|h) = 0 otherwise

Choose P(h) to be uniform distribution

•
$$P(h) = \frac{1}{|H|}$$
 for all h in H

Then

$$P(h|D) = \begin{cases} rac{1}{|VS_{H,D}|} & ext{if } h ext{ is consistent with } D \\ 0 & ext{otherwise} \end{cases}$$

Most Probable Classification of New Instances

So far we've sought the most probable hypothesis given the data D (i.e., h_{MAP})

Given new instance x, what is its most probable classification? Not $h_{MAP}(x)$!

Consider:

• Three possible hypotheses:

$$P(h_1|D) = .4, \ P(h_2|D) = .3, \ P(h_3|D) = .3$$

• Given new instance x,

$$h_1(x) = +, \ h_2(x) = -, \ h_3(x) = -$$

• What's most probable classification of x?

Bayes Optimal Classifier

Bayes optimal classification:

$$\arg\max_{v_j\in V}\sum_{h_i\in H}P(v_j|h_i)P(h_i|D)$$

Example:

$$P(h_1|D) = .4, \quad P(-|h_1) = 0, \quad P(+|h_1) = 1$$

$$P(h_2|D) = .3, \quad P(-|h_2) = 1, \quad P(+|h_2) = 0$$

$$P(h_3|D) = .3, \quad P(-|h_3) = 1, \quad P(+|h_3) = 0$$

therefore

$$\sum_{h_i \in H} P(+|h_i)P(h_i|D) = .4$$
$$\sum_{h_i \in H} P(-|h_i)P(h_i|D) = .6$$

and

$$\arg\max_{v_j \in V} \sum_{h_i \in H} P(v_j|h_i) P(h_i|D) = -$$

Gibbs Classifier

Bayes optimal classifier is hopelessly inefficient

Gibbs algorithm:

- 1. Choose one hypothesis at random, according to P(h|D)
- 2. Use this to classify new instance

Surprising fact: Assume target concepts are drawn at random from H according to priors on H. Then

 $E[error_{Gibbs}] \leq 2 \times E[error_{BayesOptimal}]$



- Probability overview
- Naïve Bayes
- Bayesian learning
- Bayesian networks
 - Representation
 - Inference
 - Parameter learning
 - Structure learning

- In general, a joint distribution P over variables (X₁,...,X_n) requires exponential space
- A Bayesian network is a graphical representation of the conditional independence relations in P
 - Usually quite compact
 - Requires fewer parameters than the full joint distribution
 - Can yield more efficient inference and belief updates

Bayesian Network

Formally, a Bayesian network is

- A directed, acyclic graph
- Each node is a random variable
- Each node X has a conditional probability distribution
 P(X | Parents(X))
- Intuitively, an arc from X to Y means that X and Y are related


Terminology

- If X and its parents are discrete, we represent
 P(X|Parents(X)) by
 - A conditional probability table (CPT)
 - It specifies the probability of each value of X, given all possible settings for the variables in *Parents(X)*.
 - Number of parameters *locally* exponential in *|Parents(X)|*
- A conditioning case is a row in this CPT: A setting of values for the parent nodes

Bayesian Network Semantics

 A Bayesian network completely specifies a full joint distribution over variables X₁,...,X_n

•
$$P(x_1,...,x_n) = \prod_{i=1}^{n} P(x_i | Parents(x_i))$$

 Here P(x₁,...,x_n) represents a specific setting for all variables (i.e., P(X₁ = x₁,..., X_n = x_n))

Conditional Indepencies

- A node X is conditionally independent of its predecessors given its parents
- Markov Blanket of X_i consists of:
 - Parents of X_i
 - Children of X_i
 - Other parents of X_i's children
- X is conditionally independent of all nodes in the network given its Markov Blanket

































- Probability overview
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- The graphical independence representation yields efficient inference schemes
- Generally, we want to compute
 - P(X) or
 - P(X | E), where E is (conjunctive) evidence
- Computations organized by network topology
- Two well-known exact algorithms:
 - Variable elimination
 - Junction trees

- A factor is a function from set of variables to a specific value: CPTS are factors
 - E.g.: p(A | E,B) is a function of A,E,B
- VE works by eliminating all variables in turn until there is a factor with only query variable

Joint Distributions & CPDs Vs. Potentials

CPT for P(B | A) *b* ¬*b*



Represent probability distributions
1. For CPT, specific setting of parents, values of child must sum to 1
2. For joint all entries sum to 1

2. For joint, all entries sum to 1

Potential



Potentials occur when we temporarily forget meaning associated with table1. Must be non-negative2. Doesn't have to sum to 1Arise when incorporating evidence



















	a		¬ <i>a</i>	
	С	٦C	С	٦C
b	.02	.04	.10	.20
¬b	.06	.10	.24	.40

Marginalize/sum out a variable





Normalize a potential







Variable Elimination Procedure

- The initial potentials are the CPTS in the BN
- Repeat until only query variable remains:
 - Choose a variable to eliminate
 - Multiply all potentials that contain the variable
 - If no evidence for the variable then sum the variable out and replace original potential by the new result
 - Else, remove variable based on evidence
- Normalize the remaining potential to get the final distribution over the query variable

P(A,B,C,D,E,F) = P(A) P(B|A) P(C|A) P(D|B) P(E|C) P(F|D,E)





Before eliminating A, multiple all potentials involving A



Now, eliminate B, multiple all potentials involving B



Next, eliminate C, multiple all potentials involving C



Next, eliminate D, multiple all potentials involving D



Next, eliminate E



Notes on Variable Elimination

- Each operation is a simple multiplication of factors and summing out a variable
- Complexity determined by size of largest factor
 - E.g., in example 3 variables (not 6)
 - Linear in number of variables, exponential in largest factor
 - Elimination ordering greatly impacts factor size
 - Optimal elimination ordering: NP-hard



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Parameter Estimate for Bayesian Networks



Ε	B	R	A	J	M
Т	F	Т	Т	F	Т
F	F	F	F	F	Т
F	Т	F	Т	Т	Т
F	F	F	Т	Т	Т
F	Т	F	F	F	F
•••					

We have:

- Bayes Net structure and observations
- We need: Bayes Net parameters


Now compute either MAP or Bayesian estimate



- The following are two common priors
- Binary variable Beta
 - Posterior distribution is binomial
 - Easy to compute posterior
- Discrete variable Dirichlet
 - Posterior distribution is multinomial
 - Easy to compute posterior

One Prior: Beta Distribution

$$eta(x) = rac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} x^{a-1}(1-x)^{b-1},$$
a,b

 $0 \leq x \leq 1 ext{ and } a, b > 0$ Here $\Gamma(y) = \int_0^\infty x^{y-1} e^{-x} dx$

For any positive integer y, $\Gamma(y) = (y-1)!$



Figure 3. Beta distributions with a = 4 and b = 8 (solid line) and with a = 8 and b = 4 (dashed line). To get a higher peak (and stronger skew), use a and b that sum to a higher value.

- Example: Flip coin with Beta distribution as prior p [prob(heads)]
 - Parameterized by two positive numbers: a and b
 - Mode of distribution (E[p]) is a /(a+b)
 - Specify our prior belief for p = a / (a+b)
 - Specify confidence with initial values of a and b
- Updating our prior belief based on data by
 - Increment a for each head outcome
 - Increment b for each tail outcome
- Posterior is a binomial distribution!



Parameter Estimate for Bayesian Networks



$$P(A|E,B) = ?$$

 $P(A|E,\neg B) = ?$
 $P(A|\neg E,B) = ?$
 $P(A|\neg E,\neg B) = ?$





Parameter Estimate for Bayesian Networks





P(A|E,B) = ? Prior ... $P(A|E,\neg B) = ?$ $P(A|\neg E,B) = Beta(2,3) + data = (3,4)$ $P(A|\neg E,\neg B) = ?$ General EM Framework: Handling Missing Values

 Given: Data with missing values, space of possible models, initial model

Repeat until no change greater than threshold:

- Expectation (E) Step: Compute expectation over missing values, given model.
- Maximization (M) Step: Replace current model with model that maximizes probability of data.

"Soft" EM vs. "Hard" EM

- Soft EM: Expectation is a probability distribution
- Hard EM: Expectation is "all or nothing," assign most likely/probable value
- Advantage of hard EM is computational efficiency when expectation is over state consisting of values for multiple variables

- For each data point with missing values
 - Compute the probability of each possible completion of that data point
 - Replace the original data point with all completions, weighted by probabilities
- Computing the probability of each completion (expectation) is just answering query over missing variables given others

- Use the completed data set to update our Beta/Dirichlet distributions
 - Same as if complete data set
 - Note: Counts may be fractional now
- Update CPTs based on new Beta/Dirichlet distribution
 - Same as if complete data set

Subtlety for Parameter Learning

- Overcounting based on number of iterations required to converge to settings for the missing values
- After each E step, reset all Beta/Dirichlet distributions before repeating M step.

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Lecture #5, Slide 121

EM for Parameter Learning



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Data



P(A=0) * P(B=0) * P(C=0 | A=0,B=0) *P(D=0 | C=0) *P(E=0 | C =0) = ? P(A=0) * P(B=0) * P(C=1 | A=0,B=0) *P(D=0 | C=1) *P(E=0 | C =1) = ?



Data



P(A=0) * P(B=0) * P(C=0 | A=0,B=0) *P(D=0 | C=0) *P(E=0 | C=0) = .41472 P(A=0) * P(B=0) * P(C=1 | A=0,B=0) *P(D=0 | C=1) *P(E=0 | C=1) = .00288

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Data



$$P(C=0) = .41472$$

.4176
 $P(C=0) = .99$

$$P(C=1) = .00288$$

.4176
 $P(C=1) = .01$

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Lecture #5, Slide 129



- Only local optimum
- Deterministic: Uniform priors can cause issues
 - See next slide
 - Use randomness to overcome this problem

What will EM do here?



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Data



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Lecture #5, Slide 131



- Probability overview
- Naïve Bayes
- Bayesian learning
- Bayesian networks
 - Representation
 - Inference
 - Parameter learning
 - Structure learning

Learning the Structure of a Bayesian Network

- Search through the space of possible structures
- For each structure, learn parameters
- Pick the one that fits observed data the best
 - Problem: Will get a fully connected structure?
 - Solution: Add a penalty term
- Problem?
 - Exponential number of networks!
 - Exhaustive search infeasible
- What now?

Structure Learning as Search

Local search

- Start with some network structure
- Try to make a change: Add, delete or reverse an edge
- See if the new structure is better
- What should the initial state be
 - Uniform prior over random networks?
 - Based on prior knowledge
 - Empty network?
- How do we evaluate networks?

Structure Search Example





Bayesian Information Criteion (BIC)

- P(D | BN) penalty
- Penalty = ½ (# parameters) Log (# data points)
- MAP score
 - P(BN | D) = P(D | BN) P(BN)
 - P(BN) must decay exponential with # of parameters for this to work well
- Note: We use log P(D | BN)





Models limited set of dependencies Guaranteed to find best structure Runs in polynomial time

Tree-Augmented Naïve Bayes

- Each feature has at most one parent in addition to the class attribute
- For every pair of features, compute the conditional mutual information

 $I_{cm}(x;y|c) = \Sigma_{x,y,c} P(x,y,c) \log [p(x,y|c)/[p(x|c)*p(y|c)]]$

- Add arcs between all pairs of features, weighted by this value
- Compute the maximum weight spanning tree, and direct arcs from the root
- Compute parameters as already seen



- Proposition rule induction
- First-order rule induction
- Read Mitchell Chapter 10



- Homework 2 is now available
- Naïve Bayes: Reasonable, simple baseline
- Different ways to incorporate prior beliefs
- Bayesian networks are an efficient way to represent joint distributions
 - Representation
 - Inference
 - Learning

