# CSEP 573: Artificial Intelligence 

## Hidden Markov Models

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Many slides over the course adapted from either Dan Klein, Stuart Russell, Andrew Moore, Ali Farhadi, or Dan Weld

## Outline

- Probabilistic sequence models (and inference)
- Markov Chains
- Hidden Markov Models
- Particle Filters


## Ghostbusters, Revisited

- Let's say we have two distributions:
- Prior distribution over ghost location: P(G)
- Let's say this is uniform
- Sensor reading model: $P(R \mid G)$
- Given: we know what our sensors do
- $\mathrm{R}=$ reading color measured at $(1,1)$
- E.g. $P(R=$ yellow $\mid G=(1,1))=0.1$
- We can calculate the posterior distribution $\mathrm{P}(\mathrm{G} \mid \mathrm{r})$ over ghost locations given a reading using Bayes' rule:

$$
P(g \mid r) \propto P(r \mid g) P(g)
$$

| 0.11 | 0.11 | 0.11 |
| :--- | :--- | :--- |
| 0.11 | 0.11 | 0.11 |
| 0.11 | 0.11 | 0.11 |


| 0.17 | 0.10 | 0.10 |
| :--- | :--- | :--- |
| 0.09 | 0.17 | 0.10 |
| $<0.01$ | 0.09 | 0.17 |

## Markov Models (Markov Chains)

- A Markov model is:
- a MDP with no actions (and no rewards)
- a chain-structured Bayesian Network (BN)

- A Markov model includes:
- Random variables $X_{t}$ for all time steps $t$ (the state)
- Parameters: called transition probabilities or dynamics, specify how the state evolves over time (also, initial probs)

$$
P\left(X_{1}\right) \quad \text { and } \quad P\left(X_{t} \mid X_{t-1}\right)
$$

## Markov Models (Markov Chains)



- A Markov model defines
- a joint probability distribution:

$$
P\left(X_{1}, \ldots, X_{n}\right)=P\left(X_{1}\right) \prod_{t=2}^{N} P\left(X_{t} \mid X_{t-1}\right)
$$

- One common inference problem:
- Compute marginals $P\left(X_{t}\right)$ for all time steps $t$


## Conditional Independence



- Basic conditional independence:
- Past and future independent of the present
- Each time step only depends on the previous
- This is called the (first order) Markov property
- Note that the chain is just a (growing) BN
- As we will see later, we can use generic BN reasoning on it if we truncate the chain at a fixed length


## Example: Markov Chain

- Weather:
- States: $X=$ \{rain, sun $\}$
- Transitions:

- Initial distribution: 1.0 sun
- What's the probability distribution after one step?

$$
\begin{aligned}
P\left(X_{2}=\text { sun }\right)=\quad & P\left(X_{2}=\operatorname{sun} \mid X_{1}=\text { sun }\right) P\left(X_{1}=\text { sun }\right)+ \\
& P\left(X_{2}=\operatorname{sun} \mid X_{1}=\text { rain }\right) P\left(X_{1}=\text { rain }\right) \\
& 0.9 \cdot 1.0+0.1 \cdot 0.0=0.9
\end{aligned}
$$

## Markov Chain Inference

- Question: probability of being in state $x$ at time t?
- Slow answer:
- Enumerate all sequences of length $t$ which end in $s$
- Add up their probabilities

$$
\begin{gathered}
P\left(X_{t}=\text { sun }\right)=\sum_{x_{1} \ldots x_{t-1}} P\left(x_{1}, \ldots x_{t-1}, \text { sun }\right) \\
P\left(X_{1}=\operatorname{sun}\right) P\left(X_{2}=\operatorname{sun} \mid X_{1}=\operatorname{sun}\right) P\left(X_{3}=\operatorname{sun} \mid X_{2}=\operatorname{sun}\right) P\left(X_{4}=\operatorname{sun} \mid X_{3}=\text { sun }\right) \\
P\left(X_{1}=\text { sun }\right) P\left(X_{2}=\text { sain } \mid X_{1}=\operatorname{sun}\right) P\left(X_{3}=\operatorname{sun} \mid X_{2}=\text { rain }\right) P\left(X_{4}=\text { sun } \mid X_{3}=\text { sun }\right)
\end{gathered}
$$

## Mini-Forward Algorithm

- Question: What's $P(X)$ on some day t?
- We don't need to enumerate every sequence!


$$
P\left(x_{t}\right)=\sum_{x_{t-1}} P\left(x_{t} \mid x_{t-1}\right) P\left(x_{t-1}\right)
$$

$$
P\left(x_{1}\right)=\text { known }
$$



Forward simulation

## Proof that $P\left(x_{t}\right)=\sum_{x_{t-1}} P\left(x_{t} \mid x_{t-1}\right) P\left(x_{t-1}\right)$

$$
\begin{aligned}
P\left(x_{t}\right) & =\sum_{x_{1}, \ldots, x_{t-1}} P\left(x_{1}, \ldots, x_{t}\right) \\
& =\sum_{x_{1}, \ldots, x_{t-1}} P\left(x_{1}\right) \prod_{i=1}^{t} P\left(x_{i} \mid x_{i-1}\right) \\
& =\sum_{x_{t-1}} P\left(x_{t} \mid x_{t-1}\right) \sum_{x_{1}, \ldots, x_{t-2}} P\left(x_{1}\right) \prod_{i=1}^{t-1} P\left(x_{i} \mid x_{i-1}\right) \\
& =\sum_{x_{t-1}} P\left(x_{t} \mid x_{t-1}\right) \sum_{x_{1}, \ldots, x_{t-2}} P\left(x_{1}, \ldots, x_{t-1}\right) \\
& =\sum_{x_{t-1}} P\left(x_{t} \mid x_{t-1}\right) P\left(x_{t-1}\right)
\end{aligned}
$$

## Example

- From initial observation of sun

$$
\left. \begin{array}{l}
0.5 \\
0.5
\end{array}\right\rangle
$$

- From initial observation of rain

$$
\left. \begin{array}{l}
0.5 \\
0.5
\end{array}\right\rangle
$$

## Stationary Distributions

- If we simulate the chain long enough:
- What happens?
- Uncertainty accumulates
- Eventually, we have no idea what the state is!
- Stationary distributions:
- For most chains, the distribution we end up in is independent of the initial distribution
- Called the stationary distribution of the chain
- Usually, can only predict a short time out


## Pac-man Markov Chain

Pac-man knows the ghost's initial position, but gets no observations!



## Web Link Analysis

- PageRank over a web graph
- Each web page is a state
- Initial distribution: uniform over pages
- Transitions:
- With prob. c, uniform jump to a random page (dotted lines, not all shown)
- With prob. 1-c, follow a random
 outlink (solid lines)
- Stationary distribution
- Will spend more time on highly reachable pages
- E.g. many ways to get to the Acrobat Reader download page
- Somewhat robust to link spam
- Google 1.0 returned the set of pages containing all your keywords in decreasing rank, now all search engines use link analysis along with many other factors (rank actually getting less important over time)


## Hidden Markov Models

- Markov chains not so useful for most agents
- Eventually you don't know anything anymore
- Need observations to update your beliefs
- Hidden Markov models (HMMs)
- Underlying Markov chain over states S
- You observe outputs (effects) at each time step
- POMDPs without actions (or rewards).
- As a Bayes' net:



## Example



- An HMM is defined by:
- Initial distribution: $P\left(X_{1}\right)$
- Transitions:
$P\left(X_{t} \mid X_{t-1}\right)$
- Emissions:
$P(E \mid X)$


## Hidden Markov Models



- Defines a joint probability distribution:

$$
P\left(X_{1}, \ldots, X_{n}, E_{1}, \ldots, E_{n}\right)=
$$

$$
P\left(X_{1: n}, E_{1: n}\right)=
$$

$$
P\left(X_{1}\right) P\left(E_{1} \mid X_{1}\right) \prod_{t=2}^{N} P\left(X_{t} \mid X_{t-1}\right) P\left(E_{t} \mid X_{t}\right)
$$

## Ghostbusters HMM

- $P\left(X_{1}\right)=$ uniform
- $P\left(X^{\prime} \mid X\right)=$ usually move clockwise, but sometimes move in a random direction or stay in place
- $P(E \mid X)=$ same sensor model as before: red means close, green means far away.

| $1 / 9$ | $1 / 9$ | $1 / 9$ |
| :--- | :--- | :--- |
| $1 / 9$ | $1 / 9$ | $1 / 9$ |
| $1 / 9$ | $1 / 9$ | $1 / 9$ |
| $P\left(X_{1}\right)$ |  |  |



| $1 / 6$ | $1 / 6$ | $1 / 2$ |
| :---: | :---: | :---: |
| 0 | $1 / 6$ | 0 |
| 0 | 0 | 0 |
|  | $P\left(X^{\prime} \mid X=<1,2>\right)$ |  |

## HMM Computations

- Given
- joint $P\left(X_{1: n}, E_{1: n}\right)$
- evidence $E_{1: n}=e_{1: n}$
- Inference problems include:
- Filtering, find $P\left(X_{t} \mid e_{1: t}\right)$ for all $t$
- Smoothing, find $P\left(X_{t} \mid e_{1: n}\right)$ for all $t$
- Most probable explanation, find

$$
x_{1: n}=\operatorname{argmax}_{x_{1: n}} P\left(x_{1: n} \mid e_{1: n}\right)
$$

## Real HMM Examples

- Speech recognition HMMs:
- Observations are acoustic signals (continuous valued)
- States are specific positions in specific words (so, tens of thousands)
- Machine translation HMMs:
- Observations are words (tens of thousands)
- States are translation options
- Robot tracking:
- Observations are range readings (continuous)
- States are positions on a map (continuous)


## Filtering / Monitoring

- Filtering, or monitoring, is the task of tracking the distribution $B(X)$ (the belief state) over time
- We start with $B(X)$ in an initial setting, usually uniform
- As time passes, or we get observations, we update $B(X)$
- The Kalman filter was invented in the 60's and first implemented as a method of trajectory estimation for the Apollo program


## Example: Robot Localization

Example from
Michael Pfeiffer

$t=0$
Sensor model: never more than 1 mistake Motion model: may not execute action with small prob.

## Example: Robot Localization



Prob

$t=1$

## Example: Robot Localization



Prob

$t=2$

## Example: Robot Localization



Prob

$t=3$

## Example: Robot Localization



Prob

$t=4$

## Example: Robot Localization



## Inference: Simple Cases



$$
P\left(X_{1} \mid e_{1}\right)
$$

$$
\begin{aligned}
P\left(x_{1} \mid e_{1}\right) & =P\left(x_{1}, e_{1}\right) / P\left(e_{1}\right) \\
& \propto_{X_{1}} P\left(x_{1}, e_{1}\right) \\
& =P\left(x_{1}\right) P\left(e_{1} \mid x_{1}\right)
\end{aligned}
$$


$P\left(X_{2}\right)$

$$
\begin{aligned}
P\left(x_{2}\right) & =\sum_{x_{1}} P\left(x_{1}, x_{2}\right) \\
& =\sum_{x_{1}} P\left(x_{1}\right) P\left(x_{2} \mid x_{1}\right)
\end{aligned}
$$

## Online Belief Updates

- Every time step, we start with current $\mathrm{P}(\mathrm{X} \mid$ evidence $)$
- We update for time:

$$
P\left(x_{t} \mid e_{1: t-1}\right)=\sum_{x_{t-1}} P\left(x_{t-1} \mid e_{1: t-1}\right) \cdot P\left(x_{t} \mid x_{t-1}\right)
$$

- We update for evidence:

$$
P\left(x_{t} \mid e_{1: t}\right) \propto_{X} P\left(x_{t} \mid e_{1: t-1}\right) \cdot P\left(e_{t} \mid x_{t}\right)
$$



## Passage of Time

- Assume we have current belief $\mathrm{P}(\mathrm{X} \mid$ evidence to date)

$$
B\left(X_{t}\right)=P\left(X_{t} \mid e_{1: t}\right)
$$

- Then, after one time step passes:


$$
P\left(X_{t+1} \mid e_{1: t}\right)=\sum_{x_{t}} P\left(X_{t+1} \mid x_{t}\right) P\left(x_{t} \mid e_{1: t}\right)
$$

- Or, compactly:

$$
B^{\prime}\left(X^{\prime}\right)=\sum_{x} P\left(X^{\prime} \mid x\right) B(x)
$$

- Basic idea: beliefs get "pushed" through the transitions
- With the "B" notation, we have to be careful about what time step $t$ the belief is about, and what evidence it includes


## Example: Passage of Time

- As time passes, uncertainty "accumulates"

| $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ |
| $<0.01$ | $<0.01$ | 1.00 | $<0.01$ | $<0.01$ | $<0.01$ |
| $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ |
| $<$ |  |  |  |  |  |


| $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $<0.01$ | $<0.01$ | 0.06 | $<0.01$ | $<0.01$ | $<0.01$ |
| $<0.01$ | 0.76 | 0.06 | 0.06 | $<0.01$ | $<0.01$ |
| $<0.01$ | $<0.01$ | 0.06 | $<0.01$ | $<0.01$ | $<0.01$ |

$$
\begin{array}{cc}
\mathrm{T}=1 & \mathrm{~T}=2 \\
B^{\prime}\left(X^{\prime}\right)= & \sum_{x} P\left(X^{\prime} \mid x\right) B(x)
\end{array}
$$

| 0.05 | 0.01 | 0.05 | $<0.01$ | $<0.01$ | $<0.01$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 0.02 | 0.14 | 0.11 | 0.35 | $<0.01$ | $<0.01$ |
| 0.07 | 0.03 | 0.05 | $<0.01$ | 0.03 | $<0.01$ |
| 0.03 | 0.03 | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ |

$$
\mathrm{T}=5
$$

Transition model: ghosts usually go clockwise

## Observation

- Assume we have current belief $\mathrm{P}(\mathrm{X} \mid$ previous evidence $)$ :

$$
B^{\prime}\left(X_{t+1}\right)=P\left(X_{t+1} \mid e_{1: t}\right)
$$

- Then:

$$
P\left(X_{t+1} \mid e_{1: t+1}\right) \propto P\left(e_{t+1} \mid X_{t+1}\right) P\left(X_{t+1} \mid e_{1: t}\right)
$$

- Or:

$$
B\left(X_{t+1}\right) \propto P(e \mid X) B^{\prime}\left(X_{t+1}\right)
$$

- Basic idea: beliefs reweighted by likelihood of evidence
- Unlike passage of time, we have to renormalize


## Example: Observation

- As we get observations, beliefs get reweighted, uncertainty "decreases"

| 0.05 | 0.01 | 0.05 | $<0.01$ | $<0.01$ | $<0.01$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 0.02 | 0.14 | 0.11 | 0.35 | $<0.01$ | $<0.01$ |
| 0.07 | 0.03 | 0.05 | $<0.01$ | 0.03 | $<0.01$ |
| 0.03 | 0.03 | $<0.01$ | $<0.01$ | $<0.01$ | $<0.01$ |

Before observation


After observation

$$
B(X) \propto P(e \mid X) B^{\prime}(X)
$$

## The Forward Algorithm

- We to know: $B_{t}(X)=P\left(X_{t} \mid e_{1: t}\right)$
- We can derive the following updates

$$
\begin{aligned}
P\left(x_{t} \mid e_{1: t}\right) & \propto{ }_{X} P\left(x_{t}, e_{1: t}\right) \\
& =\sum_{x_{t-1}} P\left(x_{t-1}, x_{t}, e_{1: t}\right) \\
& =\sum_{x_{t-1}} P\left(x_{t-1}, e_{1: t-1}\right) P\left(x_{t} \mid x_{t-1}\right) P\left(e_{t} \mid x_{t}\right) \\
& =P\left(e_{t} \mid x_{t}\right) \sum_{x_{t-1}} P\left(x_{t} \mid x_{t-1}\right) P\left(x_{t-1}, e_{1: t-1}\right)
\end{aligned}
$$

- To get $B_{t}(X)$, compute each entry and normalize


## Example: Run the Filter



- An HMM is defined by:
- Initial distribution: $P\left(X_{1}\right)$
- Transitions:
$P\left(X_{t} \mid X_{t-1}\right)$
- Emissions:
$P(E \mid X)$


## Example HMM

$\begin{array}{ll}\text { True } & 0.500 \\ \text { False } & 0.500\end{array}$


## Example HMM



## Example Pac-man



SCORE:
0

## Summary: Filtering

- Filtering is the inference process of finding a distribution over $X_{T}$ given $e_{1}$ through $e_{T}: P\left(X_{T} \mid e_{1: t}\right)$
- We first compute $\mathrm{P}\left(\mathrm{X}_{1} \mid \mathrm{e}_{1}\right): \quad P\left(x_{1} \mid e_{1}\right) \propto P\left(x_{1}\right) \cdot P\left(e_{1} \mid x_{1}\right)$
- For each t from 2 to $T$, we have $P\left(X_{t-1} \mid e_{1: t-1}\right)$
- Elapse time: compute $P\left(X_{t} \mid \mathrm{e}_{1: \mathrm{t}-1}\right)$

$$
P\left(x_{t} \mid e_{1: t-1}\right)=\sum_{x_{t-1}} P\left(x_{t-1} \mid e_{1: t-1}\right) \cdot P\left(x_{t} \mid x_{t-1}\right)
$$

- Observe: compute $P\left(X_{t} \mid e_{1:-1-1}, e_{t}\right)=P\left(X_{t} \mid e_{1: t}\right)$

$$
P\left(x_{t} \mid e_{1: t}\right) \propto P\left(x_{t} \mid e_{1: t-1}\right) \cdot P\left(e_{t} \mid x_{t}\right)
$$

## Filtering Complexity

- Problem size:
- |X| states, |E| observations, T time steps
- Each Time Step
- Computation: $\mathrm{O}\left(|\mathrm{X}|^{2}\right)$
- Space: O(|X|)
- Total
- Computation: $\mathrm{O}\left(|X|^{2} \mathrm{~T}\right)$
- Space: O(|X|)


## Recap: Reasoning Over Time

- Stationary Markov models


$$
P\left(X_{1}\right) \quad P\left(X \mid X_{-1}\right)
$$

$$
P(E \mid X)
$$

- Hidden Markov models


| X | E | P |
| :---: | :---: | :---: |
| rain | umbrella | 0.9 |
| rain | no umbrella | 0.1 |
| sun | umbrella | 0.2 |
| sun | no umbrella | 0.8 |

## Particle Filtering

- Sometimes $|X|$ is too big to use exact inference
- $|X|$ may be too big to even store $B(X)$
- E.g. X is continuous
- $|X|^{2}$ may be too big to do updates
- Solution: approximate inference
- Track samples of X, not all values
- Samples are called particles
- Time per step is linear in the number of samples
- But: number needed may be large
- In memory: list of particles, not states
- This is how robot localization
 works in practice


## Representation: Particles

- Our representation of $P(X)$ is now a list of N particles (samples)
- Generally, N << |X|
- Storing map from $X$ to counts would defeat the point

- $P(x)$ approximated by number of particles with value $x$
- So, many $x$ will have $P(x)=0$ !
- More particles, more accuracy
- For now, all particles have a weight of 1

Particles:
$(3,3)$
$(3,3)$
$(3,2)$
$(2,1)$

## Particle Filtering: Elapse Time

- Each particle is moved by sampling its next position from the transition model

$$
x^{\prime}=\operatorname{sample}\left(P\left(X^{\prime} \mid x\right)\right)
$$

- This is like prior sampling - samples' frequencies reflect the transition probs
- Here, most samples move clockwise, but some move in another direction or stay in place
- This captures the passage of time
- If we have enough samples, close to the exact values before and after (consistent)



## Particle Filtering: Observe

- Slightly trickier:
- Don't do rejection sampling (why not?)
- We don't sample the observation, we fix it
- This is similar to likelihood weighting, so we downweight our samples based on the evidence

$$
\begin{aligned}
w(x) & =P(e \mid x) \\
B(X) & \propto P(e \mid X) B^{\prime}(X)
\end{aligned}
$$

- Note that, as before, the probabilities don't sum to one, since most have been downweighted (in fact they sum to an approximation of $\mathrm{P}(\mathrm{e})$ )



## Particle Filtering: Resample

- Rather than tracking weighted samples, we resample
- N times, we choose from our weighted sample distribution (i.e. draw with replacement)
- This is equivalent to renormalizing the distribution
- Now the update is complete for this time step, continue with the next one

Old Particles:
$(3,3) w=0.1$
$(2,1) w=0.9$
$(2,1) w=0.9$
$(3,1) w=0.4$
$(3,2) w=0.3$
$(2,2) w=0.4$
$(1,1) w=0.4$
$(3,1) w=0.4$
$(2,1) w=0.9$
$(3,2) w=0.3$

New Particles:
$(2,1) w=1$
$(2,1) w=1$
$(2,1) w=1$
$(3,2) w=1$
$(2,2) w=1$
$(2,1) w=1$
$(1,1) w=1$
$(3,1) w=1$
$(2,1) w=1$
$(1,1) w=1$


## Particle Filtering Summary

- Represent current belief $P(X \mid$ evidence to date) as set of $n$ samples (actual assignments $X=x$ )
- For each new observation e:

1. Sample transition, once for each current particle x

$$
x^{\prime}=\operatorname{sample}\left(P\left(X^{\prime} \mid x\right)\right)
$$

2. For each new sample x', compute importance weights for the new evidence e:

$$
w\left(x^{\prime}\right)=P\left(e \mid x^{\prime}\right)
$$

3. Finally, normalize the importance weights and resample N new particles

## Robot Localization

- In robot localization:
- We know the map, but not the robot's position
- Observations may be vectors of range finder readings
- State space and readings are typically continuous (works basically like a very fine grid) and so we cannot store $B(X)$
- Particle filtering is a main technique



## Robot Localization



## Which Algorithm?

## Exact filter, uniform initial beliefs



## Which Algorithm?

Particle filter, uniform initial beliefs, 300 particles


SCORE: 0

## Which Algorithm?

Particle filter, uniform initial beliefs, 25 particles


SCORE: 0

## P4: Ghostbusters

- Plot: Pacman's grandfather, Grandpac, learned to hunt ghosts for sport.
- He was blinded by his power, but could hear the ghosts' banging and clanging.
- Transition Model: All ghosts move randomly, but are sometimes biased
- Emission Model: Pacman knows a "noisy" distance to each ghost

Noisy distance prob
True distance = 815131211

10
9

## Dynamic Bayes Nets (DBNs)

- We want to track multiple variables over time, using multiple sources of evidence
- Idea: Repeat a fixed Bayes net structure at each time
- Variables from time $t$ can condition on those from $t-1$

- Discrete valued dynamic Bayes nets are also HMMs


## DBN Particle Filters

- A particle is a complete sample for a time step
- Initialize: Generate prior samples for the $t=1$ Bayes net
- Example particle: $\mathbf{G}_{1}{ }^{\mathbf{a}}=(3,3) \mathbf{G}_{1}{ }^{\mathbf{b}}=(5,3)$
- Elapse time: Sample a successor for each particle
- Example successor: $\mathbf{G}_{2}{ }^{\mathbf{a}}=(2,3) \mathbf{G}_{\mathbf{2}}{ }^{\mathbf{b}}=(6,3)$
- Observe: Weight each entire sample by the likelihood of the evidence conditioned on the sample
- Likelihood: $\mathrm{P}\left(\mathbf{E}_{1}{ }^{\mathrm{a}} \mid \mathbf{G}_{1}{ }^{\mathbf{a}}\right)^{*} \mathrm{P}\left(\mathbf{E}_{1}{ }^{\mathrm{b}} \mid \mathbf{G}^{1}{ }^{\mathbf{b}}\right)$
- Resample: Select prior samples (tuples of values) in proportion to their likelihood

