## Lecture \#2: Advanced hashing and concentration bounds

- Bloom filters
- Cuckoo hashing
- Load balancing
- Tail bounds

Idea: For the sake of efficiency, sometime we allow our data structure to make mistakes

## Bloom filter: Hash table that has only false positives

(may report that a key is present when it is not, but always reports a key that is present)
Very simple and fast
Example: $\quad$ Google Chrome uses a Bloom filter to maintain its list of potentially malicious web sites.

- Most queried keys are not in the table
- If a key is in the table, can check against a slower (errorless) hash table

Many applications in networking (see survey by Broder and Mitzenmacher)

Data structure: Universe $\mathcal{U}$. Parameters $k, M \geq 1$
Maintain an array $A$ of $M$ bits; initially $A[0]=A[1]=\cdots=A[M-1]=0$ Choose $k$ hash functions $h_{1}, h_{2}, \ldots, h_{k}: \mathcal{U} \rightarrow[M]$
(assume completely random functions for sake of analysis)

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(assume completely random functions for sake of analysis)
To add a key $x \in \mathcal{U}$ to the dictionary $S \subseteq \mathcal{U}$, set bits

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A\left[h_{1}(x)\right]:=1, A\left[h_{2}(x)\right]:=1, \ldots, A\left[h_{k}(x)\right]:=1
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| 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |

To answer a query: $q \in S$ ?
Check whether $A\left[h_{i}(x)\right]=1$ for all $i=1,2, \ldots, k$ If yes, answer Yes. If no, answer No.


No false negatives: Clearly if $x \in S$, we return Yes.
But there is some chance that other keys have caused the bits in positions $h_{1}(x), \ldots, h_{k}(x)$ to be set even if $x \notin S$.


## Bloom filters

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## Heuristic analysis:

Let us assume that $|S|=n$.
Compute $\mathbb{P}[A[\ell]=0]$ for some location $\ell \in[M]$ :

$$
p(k, N)=\left(1-\frac{1}{M}\right)^{k N} \approx e^{-\frac{k N}{M}}
$$ for $M$ large enough.)

| 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
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If each location in $A$ is 0 with probability $p(k, N)$, then a false positive for $x \notin S$ should happen with probability at most

$$
(1-p(k, N))^{k} \approx\left(1-e^{-\frac{k N}{M}}\right)^{k}
$$

(Here we use the approximation $\left(1-\frac{1}{M}\right)^{M} \approx e^{-1}$ for $M$ large enough.)

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| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |



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But the actual fraction of $0^{\prime} s$ in the hash table is a random variable $X_{k, N}$ with expectation

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\mathbb{E}\left[X_{k, N}\right]=p(k, N)
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To get the analysis right, we need a concentration bound: Want to say that $X_{k, N}$ is close to its expected value with high probability. [We will return to this in the $2^{\text {nd }}$ half of the lecture]

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If the heuristic analysis is correct, it gives nice estimates:
For instance, if $M=8 N$, then choosing the optimal value of $k=7$ gives
false positive rate about $2 \%$.

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Cuckoo hashing is a hash scheme with worst-case constant lookup time. The name derives from the behavior of some species of cuckoo, where the cuckoo chick pushes the other eggs or young out of the nest when it hatches; analogously, inserting a new key into a cuckoo hashing table may push an older key to a different location in the table.

Idea: Simple hashing without errors
Lookups are worst case $O(1)$ time Deletions are worst case $O$ (1) time Insertions are expected $O(1)$ time Insertion time is $O(1)$ with good probability [will require a concentration bound]

Data structure: Two tables $A_{1}$ and $A_{2}$ both of size $M=O(N)$
Two hash functions $h_{1}, h_{2}: U \rightarrow[M]$ (will assume hash functions are fully random)

When an element $x \in S$ is inserted, if either $A_{1}\left[h_{1}(x)\right]$ or $A_{2}\left[h_{2}(x)\right]$ is empty, store $x$ there.


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## Bump:

If both locations are occupied, then place $x$ in $A_{1}\left[h_{1}(x)\right]$ and bump the current occupant.

Whenever an element $z$ is bumped from $A_{i}\left[h_{i}(z)\right]$, attempt to store it in the other location $A_{j}\left[h_{j}(z)\right]$ (here $(i, j)=(1,2)$ or $(2,1)$ )


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Abort: After $6 \log N$ consecutive bumps, stop the process and build a fresh hash table using new random hash functions $h_{1}, h_{2}$.


Alternately (as in the picture), we can use a single table with $2 M$ entries and two hash functions $h_{1}, h_{2}: U \rightarrow[2 M]$ (with the same "bumping" algorithm)

Arrows represent the alternate location for each key.
If we insert an item at the location of $A$, it will get bumped, thereby bumping $B$, and then we are done.

Cycles are possible (where the insertion process never completes). What's an example?


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## Theorem:

Expected time to perform an insert operation is $O(1)$ if $M \geq 4 N$.


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Pretty good... but only $25 \%$ memory utilization.
Can actually get about 50\% memory utilization.
Experimentally, with 3 hash functions instead of 2, can get $\approx 90 \%$ utilization, but it is an open question to provide tight analyses for $d$ hash functions when $d \geq 3$.


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Claim: The max-loaded server has $<8 \log N / \log \log N$ jobs with probability at least $1-1 / N$
Proof: Probability that a fixed server $i \in\{1,2, \ldots, N\}$ gets at least $k$ jobs is at most

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\binom{N}{k}\left(\frac{1}{N}\right)^{k} \leq \frac{N^{k}}{k!} \cdot \frac{1}{N^{k}} \leq \frac{1}{k!} \leq k^{-\frac{k}{2}}
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If we choose $k=\frac{8 \log N}{\log \log N}$ this is at most $1 / N^{2}$
Explanation: $k^{\frac{k}{2}} \geq(\sqrt{\log N})^{\frac{4 \log N}{)^{\log \log N}}} \geq 2^{2 \log N}=N^{2}$

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This is an example of a concentration bound.
Let $X_{i}$ be the number of jobs assigned to the $i$ th server.
By linearity of expectation, $\mathbb{E}\left[X_{i}\right]=\sum_{j=1}^{N} \mathbb{P}[$ job $j \rightarrow$ server $i]=N \cdot(1 / N)=1$.

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This is a common analysis technique:
If a random variable (like $X_{i}$ ) depends in a "smooth" way on the outcome of many independent events, then it is likely not too far from its expectation.

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If a random variable (like $X_{i}$ ) depends in a "smooth" way on the outcome of many independent events, then it is likely not too far from its expectation.
"Smooth" in this case means that the outcome of any decision (where to put job $j$ ) does not affect the value of $X_{i}$ by too much (only by 1 ).

Is it concentrated? [why or why not?]
\#1: Choose a uniformly random vector $X \in \mathbb{R}^{n}$ with $\|X\|=\sqrt{X_{1}^{2}+X_{2}^{2}+\cdots+X_{n}^{2}}=1$
What is $\mathbb{E}\left[X_{1}^{2}\right]$ ?
What is the typical value of the maximum: max $\left(\left|X_{1}\right|,\left|X_{2}\right|, \ldots,\left|X_{n}\right|\right)$ ?
\#2 Rich get richer: Suppose we have $N$ people. Everyone starts with 1 dollar.
We assign $N^{2}$ more dollars in rounds.

$i$ th round: If person $j$ already has $n_{j}$ dollars, we give them the $i$ th dollar with probability

$$
\frac{n_{j}}{i-1}
$$

i.e., with probability to the proportional the amount of money they already have. Let $X_{i}$ be the amount of money person $i$ ends up with.

What is the typical value of $X_{1}$ ? Is $X_{1}$ concentrated?
What is the typical value of $\max \left(X_{1}, X_{2}, \ldots, X_{n}\right)$ ? Is it concentrated?


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The most basic concentration bound is Markov's inequality. It requires knowing only the expected value:

If $X$ is a non-negative random variable, then for any $\lambda \geq 1$,

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## Example:

If your expected revenue is $\$ 10,000$, then the probability to make $\$ 1,000$ is at most $1 / 10$.

Markov's inequality: If $X$ is a non-negative random variable, then for any $\lambda \geq 1$,

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A permutation is an invertible mapping $\pi:\{1,2, \ldots, n\} \rightarrow\{1,2, \ldots, n\}$
A number $j$ is called a fixed point of $\pi$ if $\pi(j)=j$.
Exercise: Prove that if $\pi$ is a uniformly random permutation, then

$$
\mathbb{P}[\pi \text { has more than } k \text { fixed points }] \leq \frac{1}{k}
$$

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Proof: Apply Markov's inequality to the random variable $Y=(X-\mathbb{E} X)^{2}$

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Let $L_{i}$ be the load of bin $i$.
Let $X_{i j}$ be the indicator random variable such that $X_{i j}=1 \leftrightarrow i$ th bin gets the $j$ th ball.
Note that $\mathbb{E}\left[X_{i j}\right]=1 / N$ for each $i=1, \ldots, N$.

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So write: $\quad \operatorname{var}\left(L_{i}\right)=\mathbb{E}\left[\left(X_{i 1}+\cdots+X_{i N}\right)^{2}\right]-1$
We have $\mathbb{E}\left[X_{i j}^{2}\right]=\mathbb{E}\left[X_{i j}\right]=1 / N$ and $\mathbb{E}\left[X_{i j} X_{i k}\right]=\mathbb{P}[h(j)=h(k)=i] \leq 1 / N^{2}$ using the 2 -universal property, so

$$
\operatorname{var}\left(L_{i}\right) \leq N \cdot\left(\frac{1}{N}\right)+\frac{N(N-1)}{N^{2}}-1=1-\frac{1}{N} \leq 1
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Apply Chebyshev's inequality to $L_{i}$, yielding

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Thus $\mathbb{P}\left[\left|L_{i}-1\right| \geq \sqrt{2 N}\right] \leq \frac{1}{2 N}$, so a union bound yields

$$
\mathbb{P}\left[\max \left(L_{1}, \ldots, L_{N}\right) \geq \sqrt{2 N}+1\right] \leq \frac{1}{2}
$$

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Suppose we choose $n$ independent random voters and ask them whether they prefer candidate \#1 over candidate \#2.


We see outcomes $X_{1}, X_{2}, \ldots, X_{n} \in\{0,1\}$.
Let $p$ be the actual percentage of the population the prefers candidate $\# 1$ and let $\hat{p}=\left(X_{1}+\cdots+X_{n}\right) / n$ denote the empirical mean.

## Exercise:

Prove that if we want $|p-\hat{p}| \leq \epsilon$ to hold with $99 \%$ probability, then we need only sample $n=O\left(1 / \epsilon^{2}\right)$ voters.

Hoeffding's inequality:
Let $X_{1}, \ldots, X_{n}$ be a sequence of independent random variables where, for each $1 \leq i \leq n$, we have $a_{i} \leq X_{i} \leq b_{i}$. Let $X=\left(X_{1}+\cdots+X_{n}\right) / n$. Then:

$$
\mathbb{P}[|X-\mathbb{E} X| \geq \lambda] \leq 2 e^{-\frac{2 \lambda^{2} n^{2}}{\sum_{i=1}^{n}\left(a_{i}-b_{i}\right)^{2}}}
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Setting $a_{i}=0, b_{i}=1$, and $\lambda=\epsilon$ in Hoeffding's inequality gives

$$
\mathbb{P}[|\hat{p}-p| \geq \epsilon] \leq 2 e^{-2 \epsilon^{2} n}
$$

so we only need $n \leq O\left(\frac{\log \left(\frac{1}{\delta}\right)}{\epsilon^{2}}\right)$ samples.

## Chernoff bound (multiplicative):

Let $X_{1}, \ldots, X_{n}$ be a sequence of independent $\{0,1\}$-valued random variables.
Let $p_{i}=\mathbb{E}\left[X_{i}\right], \quad X=X_{1}+X_{2}+\cdots+X_{n}, \quad \mu=\mathbb{E}[X]$. Then for every $\beta \geq 1$ :

$$
\mathbb{P}[X \geq \beta \mu] \leq\left(\frac{e^{\beta-1}}{\beta^{\beta}}\right)^{\mu} \quad \mathbb{P}[X \leq \mu / \beta] \leq\left(\frac{e^{\frac{1}{\beta}-1}}{\beta^{\beta}}\right)^{\mu}
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## Reproduce balls in bins:

$N$ balls thrown randomly into $N$ bins.
$X_{i}=1$ if ith ball ends up in first bin and $X_{i}=0$ otherwise.
Then $X=\#$ of balls in first bin. As we calculated earlier, $\mathbb{E}[X]=1$
For $\beta \approx \frac{\log N}{\log \log N^{\prime}}$, the Chernoff bound gives $\mathbb{P}[X \geq \beta] \leq 1 / N^{2}$

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## Heuristic analysis:

If each location in $A$ is 0 with probability $p(k, N)$, then a false positive for $x \notin S$ should happen with probability at most

$$
(1-p(k, N))^{k} \approx\left(1-e^{-\frac{k N}{M}}\right)^{k}
$$

But the actual fraction of $0^{\prime} s$ in the hash table is a random variable $X_{k, N}$ with expectation

$$
\mathbb{E}\left[X_{k, N}\right]=p(k, N)
$$

To get the analysis right, we need a concentration bound: Want to say that $X_{k, N}$ is close to its expected value with high probability.

Let's analyze!

We have an array with $M$ bits and to hash an element $x \in \mathcal{U}$, we set the bits in positions $h_{1}(x), h_{2}(x), \ldots, h_{k}(x)$ to 1.

| 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |



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Consider the $N$ elements to hash:

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Note that $x_{1}, \ldots, x_{N}$ are any set of keys. The randomness here is all in the choice of the hash functions $h_{1}, \ldots, h_{k}$.

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Such a sequence of random variables is called a martingale

Suppose that $\left\{X_{0}, X_{1}, \ldots, X_{N}\right\}$ is a martingale such that for some constants $\left\{c_{j}\right\}$, $\left|X_{j+1}-X_{j}\right| \leq c_{j}$ for all $j=0,1, \ldots, N-1$. Then for any $\lambda>0$,

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\mathbb{P}\left[\left|X_{N}-X_{0}\right| \geq \lambda\right] \leq 2 \exp \left(-\frac{\lambda^{2}}{2\left(c_{1}^{2}+\cdots+c_{N}^{2}\right)}\right)
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For our problem: $c_{1}=c_{2}=\cdots=c_{N}=k$
So the probability that the \# of 0's differs from its expectation by more than $\lambda$ is at most

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So the deviation is $\approx k \sqrt{N}$ and is tightly concentrated in this window.

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So the deviation is $\approx k \sqrt{N}$ and is tightly concentrated in this window.
Improve the error probability to $2 \exp \left(-\lambda^{2} / 2 k N\right)$ using a different martingale.

