## **Lecture 8: Linear Regression**

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Su-In Lee, CSE & GS suinlee@uw.edu

#### Goals

- Develop basic concepts of linear regression from a probabilistic framework
- Estimating parameters and hypothesis testing with linear models
- Linear regression in R

# Regression

- Technique used for the modeling and analysis of numerical data
- Exploits the relationship between two or more variables so that we can gain information about one of them through knowing values of the other
- Regression can be used for prediction, estimation, hypothesis testing, and modeling causal relationships

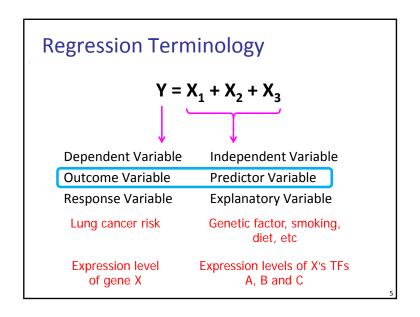
## Why Linear Regression?

 Suppose we want to model the outcome variable Y in terms of three predictors, X<sub>1</sub>, X<sub>2</sub>, X<sub>3</sub>

$$Y = f(X_1, X_2, X_3)$$

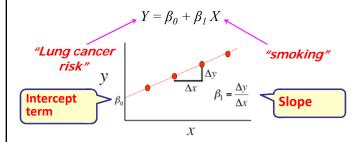
- Typically will not have enough data to try and directly estimate f
- Therefore, we usually have to assume that it has some restricted form, such as linear

$$Y = X_1 + X_2 + X_3$$





 Much of mathematics is devoted to studying variables that are deterministically related to one another



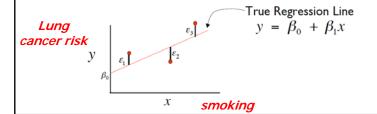
 But we're interested in understanding the relationship between variables related in a nondeterministic fashion

#### A Linear Probabilistic Model

■ **Definition:** There exists parameters  $\beta_0$ ,  $\beta_1$  and  $\sigma^2$ , such that for any fixed value of the predictor variable X, the outcome variable Y is related to X through the model equation

$$Y = \beta_0 + \beta_1 X + \varepsilon,$$

where  $\epsilon$  is a RV assumed to be  $N(0,\,\sigma^2)$ 



## **Implications**

■ The **expected value of Y** is a linear function of X, but for fixed value x, the variable Y differs from its expected value by a *random amount* 

Variables & Symbols: How is x different from X?

**Capital letter** *X*: a random variable **Lower case letter** *x*: corresponding values

(i.e. the real numbers the RV X map into)

For example,

*X*: Genotype of a certain locus

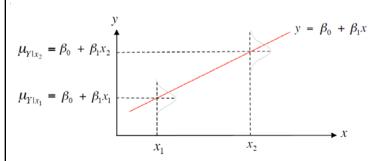
x: 0, 1 or 2 (meaning AA, AG and GG, respectively)

#### **Implications**

- The **expected value of** *Y* is a linear function of *X*, but for fixed value *x*, the variable *Y* differs from its expected value by a *random amount*
- Formally, let  $x^*$  denote a particular value of the predictor variable X, then our linear probabilistic model says:

$$E(Y | x^*) = \mu_{Y|x^*} = \text{mean value of } Y \text{ when } X \text{ is } x^*$$
  
 $V(Y | x^*) = \sigma_{Y|x^*}^2 = \text{variance of } Y \text{ when } X \text{ is } x^*$ 

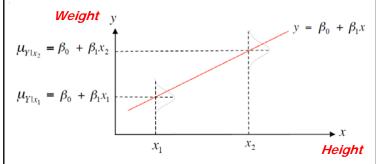
## **Graphical Interpretation**



 $E(Y \mid x^*) = \mu_{Y \mid x^*} = \text{mean value of } Y \text{ when } X \text{ is } x^*$  $V(Y \mid x^*) = \sigma_{Y \mid x^*}^2 = \text{variance of } Y \text{ when } X \text{ is } x^*$ 

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# **Graphical Interpretation**



- Say that *X* = height and *Y* = weight
- Then  $\mu_{Y|x=60}$  is the average weight for all individuals 60 inches tall in the population

## One More Example

 Suppose the relationship between the predictor variable height (X) and outcome variable weight (Y) is described by a simple linear regression model with true regression line

$$Y = 7.5 + 0.5 X$$
,  $\varepsilon \sim N(0, \sigma^2)$  and  $\sigma = 3$ 

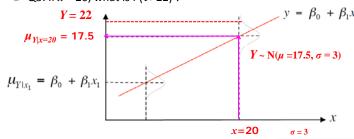
- Q1: What is the interpretation of β<sub>I</sub> = 0.5?
   The expected change in weight (Y) associated with a 1-unit increase in height (X)
- Q2: If x = 20, what is the expected value of Y?

 $\mu_{Y|x=20}$  = 7.5 + 0.5 (20) = 17.5

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## One More Example

• Q3: If x = 20, what is P(Y>22)?



• Given  $Y \sim N(\mu = 17.5, \sigma = 3)$ ,

$$P(Y > 22 \mid x = 20) = 1 - \phi(\frac{22 - 17.5}{3}) = 1 - \phi(1.5) = 0.067$$

where  $\phi$  means the CDF of Normal dist. N(0,1)

## **Estimating Model Parameters**

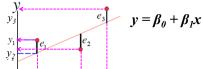
- Where are the parameters  $\beta_0$  and  $\beta_1$  from?
- **Predicted**, or fitted, values are values of y predicted by plugging  $x_1, x_2, \dots, x_n$  into the  $\hat{y}_2 = \beta_0 + \beta_1 x_2$ estimated regression line:  $y = \beta_0 + \beta_1 x$

 $\hat{y}_1 = \beta_0 + \beta_1 x_1$  $\hat{y}_3 = \beta_0 + \beta_1 x_3$ 

• Residuals are the deviations of observed (red dots) and predicted values (red line)

$$e_1 = y_1 - \hat{y}_1 e_2 = y_2 - \hat{y}_2$$

 $e_3 = y_3 - \hat{y}_3$ 



#### Residuals Are Useful!

• The error sum of squares (SSE) can tell us how well the line fits to the data

SSE = 
$$\sum_{i=1}^{n} (e_i)^2 = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

$$\hat{y}_1 = \beta_0 + \beta_1 x_1$$

$$\hat{y}_2 = \beta_0 + \beta_1 x_2$$

$$\hat{y}_3 = \beta_0 + \beta_1 x_3$$

$$\hat{y}_1$$

$$\hat{y}_1$$

$$\hat{y}_2$$

$$\hat{y}_3$$

$$\hat{y}_4$$

$$\hat{y}_1$$

$$\hat{y}_1$$

$$\hat{y}_1$$

$$\hat{y}_2$$

$$\hat{y}_3$$

$$\hat{y}_4$$

$$\hat{y}_1$$

$$\hat{y}_1$$

$$\hat{y}_2$$

$$\hat{y}_3$$

$$\hat{y}_4$$

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$$\hat{y}_3$$

$$\hat{y}_4$$

$$\hat{$$

Y > 22

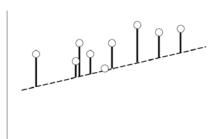
 $\mu = 17.5$ 

- Least squares
  - Find  $\beta_0$  and  $\beta_1$  that minimizes SSE

$$f(\beta_0, \beta_1) = \sum_{i=1}^{n} [y_i - (\beta_0 + \beta_1 x_i)]^2$$

• Denote the solutions by  $\hat{\beta}_0$  and  $\hat{\beta}_1$ 

**Least Squares** 



- Least squares
  - Find  $\beta_0$  and  $\beta_1$  that minimizes SSE

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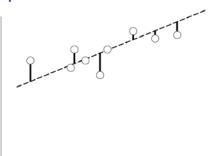
## **Least Squares**



- Least squares
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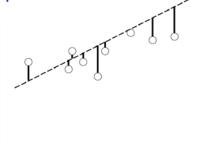
# **Least Squares**



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**Least Squares** 



- Least squares
  - Find  $\beta_0$  and  $\beta_1$  that minimizes SSE

$$f(\beta_0, \beta_1) = \sum_{i=1}^{n} [y_i - (\beta_0 + \beta_1 x_i)]^2$$

#### Coefficient of Determination

 Important statistic referred to as the coefficient of determination (R<sup>2</sup>):

$$R^{2} = 1 - \frac{\text{SSE}}{\text{SST}}$$

$$\text{SSE} = \sum_{i=1}^{n} (e_{i})^{2} = \sum_{i=1}^{n} (y_{i} - \hat{y}_{i})^{2}$$

$$\text{Error Sum Squares}$$

$$\text{SST} = \sum_{i=1}^{n} (y_{i} - \overline{y})^{2}$$

$$\text{when } \beta_{0} = \text{avg}(y)$$

$$\text{and } \beta_{1} = 0$$

$$y = \beta_{0} + \beta_{1}x$$

$$y = \text{average } y$$

## Multiple Linear Regression

 Extension of the simple linear regression model to two or more independent variables

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_n x_n + \varepsilon$$

Expression = Baseline + Age + Tissue + Sex + Error

Partial Regression Coefficients:

 $\beta_i$  = effect on the outcome variable when increasing the  $i^{th}$  predictor variable by 1 unit, **holding all other predictors** constant

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#### **Categorical Independent Variables**

- Qualitative variables are easily incorporated in regression framework through dummy variables
- Simple example: sex can be coded as 0/1
- What if my categorical variable contains three levels:

$$X_i = \begin{cases} 0 & \text{if AA} \\ 1 & \text{if AG} \\ 2 & \text{if GG} \end{cases}$$

NO!

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## **Categorical Independent Variables**

- Previous coding would result in collinearity
- Solution is to set up a series of dummy variable. In general for k levels you need (k-1) dummy variables

$$X_1 = \begin{cases} 1 \text{ if AA} \\ 0 \text{ otherwise} \end{cases} \qquad X_2 = \begin{cases} 1 \text{ if AG} \\ 0 \text{ otherwise} \end{cases}$$

$$X_{i} = \begin{cases} 0 & \text{if AA} \\ 0 & \text{if AG} \\ 0 & \text{if AG} \\ 0 & \text{if AG} \end{cases} AG = \begin{cases} X_{1} & X_{2} \\ 0 & \text{AA} & 1 & 0 \\ 0 & \text{AG} & 0 & 1 \\ 0 & \text{AG} & 0 & 0 \end{cases}$$

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## Hypothesis Testing: Model Utility Test

■ The first thing we want to know after fitting a model is whether any of the predictor variables (X's) are significantly related to the outcome variable (Y):

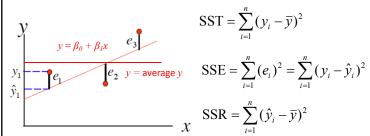
$$\mathbf{H}_0: \, \boldsymbol{\beta}_1 = \boldsymbol{\beta}_2 = \dots = \boldsymbol{\beta}_k = \mathbf{0}$$

 $H_A$ : At least one  $\beta_i \neq 0$ 

- Let's frame this in our ANOVE framework
- In ANOVA, we partitioned total variance (SST) into two components:
  - SSE (unexplained variation)
  - SSR (variation explained by linear model)

## **Model Utility Test**

- Partition total variance (SST) into two components:
  - SSE (unexplained variation)
  - SSR (variation explained by linear model)
- Let's consider n (=3) data points and k (=1) predictor model



$$SST = \sum_{i=1}^{n} (y_i - \overline{y})$$

SSE = 
$$\sum_{i=1}^{n} (e_i)^2 = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

$$SSR = \sum_{i=1}^{n} (\hat{y}_i - \overline{y})^2$$

#### ANOVA Formulation of Model Utility Test

- Partition total variance (SST) into two components:
  - SSE (unexplained variation)
  - SSR (variation explained by linear model)

Source of Variation	df	Sum of Squares	MS	F
Regression	k	$SSR = \sum (\hat{y}_i - \overline{y})^2$	$\frac{SSR}{k}$	$\frac{MS_R}{MS_E}$
Error	n-(k+1)	$SSE = \sum (y_i - \hat{y}_i)^2$	<u>SSE</u> n-(k+1)	
Total	n-1	$SST = \sum (y_i - \overline{y})^2$		

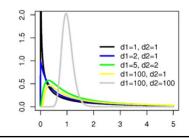
# data points – (# parameters in the model) Rejection Region :  $F_{\alpha,k,n-(k+1)}$ 

#### **ANOVA Formulation of Model Utility Test**

F-test statistic

$$F = \frac{MS_R}{MS_E} = \frac{SSR/k}{SSE/[n - (k+1)]} = \frac{R^2}{1 - R^2} \cdot \frac{n - (k+1)}{k}$$

Rejection Region :  $F_{\alpha k n-(k+1)}$ 



- Pick the distribution function. based on k and n-(k+1).
- Choose the critical value based on  $\alpha$  (F<sub>a,k,n-(k+1)</sub>)
  - Say that  $\alpha$  =0.05
  - Prob(F>F<sub>a,k,n-(k+1)</sub>) = 0.05

F Test For Subsets of Independent Variables

- A powerful tool in multiple regression analysis is the ability to compare two models
- For instance say we want to compare

Full Model:  $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \beta_4 x_4 + \varepsilon$ 

Reduced Model:  $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \varepsilon$ 

Again, another example of ANOVA

 $SSE_p = error sum of squares for$ reduced model with l predictors

 $F = \frac{(SSE_R - SSE_F)/(k-l)}{SSE_F/[n-(k+1)]}$ 

 $SSE_{E} = error sum of squares for$ full model with k predictors

## **Example of Model Comparison**

• We have a quantitative trait and want to test the effects at two markers, M1 and M2.

Full Model: Trait = Mean + M1 + M2 + (M1 × X2) +  $\varepsilon$ Reduced Model: Trait = Mean + M1 + M2 +  $\varepsilon$ 

$$F = \frac{(SSE_R - SSE_F)/(k - l)}{SSE_F/[n - (k + 1)]} = \frac{(SSE_R - SSE_F)/(3 - 2)}{SSE_F/[100 - (3 + 1)]}$$
$$= \frac{(SSE_R - SSE_F)}{SSE_F/96}$$

Rejection Region :  $F_{\alpha,1.96}$ 

# How To Do In R

- You can fit a least-squares regression using the function
  - mm <- Isfit(x,y)</p>
- The coefficients of the fit are then given by
  - mm\$coef
- The residuals are
  - mmŚresiduals
- And to print out the tests for zero slope just do
  - Is.print (mm)

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#### **Input Data**

- http://www.cs.washington.edu/homes/suinlee/geno me560/data/cats.txt
- Data on fluctuating proportions of marked cells in marrow from heterozygous Safari cats
- Proportions of cells of one cell type in samples from cats (taken in our department many years ago).
   Column 1 is the ID number of the particular cat. You will want to plot the data from one cat.
  - For example cat 40004 is rows 1:17, 40005a is 18:31, 40005b is 32:47, 40006 is 48:65, 40665 is 66:83 and so on.

**Input Data** 

- 2<sup>nd</sup> column: Time, in weeks from the start of monitoring, that the measurement from marrow is recorded.
- 3<sup>rd</sup> column: Percent of domestic-type progenitor cells observed in a sample of cells at that time.
- 4<sup>th</sup> column: Sample size at that time, i.e. the number of progenitor cells analyzed.

40005a 34 13 56 40005a 37 17 65

40004 11 33 72

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