

**CSE 312**

# **Foundations of Computing II**

**Lecture 6: Bayesian Inference, Chain Rule,  
Independence**

## Review Conditional & Total Probabilities

- **Conditional Probability**

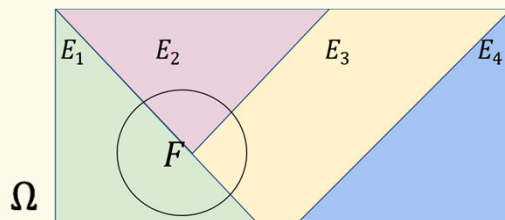
$$P(B|A) = \frac{P(A \cap B)}{P(A)}$$

- **Bayes Theorem**

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)} \quad \text{if } P(A) \neq 0, P(B) \neq 0$$

- **Law of Total Probability**

$E_1, \dots, E_n$  partition  $\Omega$




$$P(F) = \sum_{i=1}^n P(F \cap E_i) = \sum_{i=1}^n P(F|E_i)P(E_i)$$

## Conditional Probability Defines a Probability Space

The probability conditioned on  $\mathcal{A}$  follows the same properties as (unconditional) probability.

**Example.**  $P(\mathcal{B}^c|\mathcal{A}) = 1 - P(\mathcal{B}|\mathcal{A})$

**Formally.**  $(\Omega, P)$  is a probability space and  $P(\mathcal{A}) > 0$

  $(\mathcal{A}, P(\cdot | \mathcal{A}))$  is a probability space

# Agenda

- Bayes Theorem + Law of Total Probability ◀
- Chain Rule
- Independence
- Infinite process and Von Neumann's trick
- Conditional independence

## Example – Zika Testing

Suppose we know the following Zika stats

- A test is 98% effective at detecting Zika (“true positive”)  $P(T|Z)$
- However, the test may yield a “false positive” 1% of the time  $P(T|Z^c)$
- 0.5% of the US population has Zika.  $P(Z)$

What is the probability you have Zika (event  $Z$ ) if you test positive (event  $T$ )?.

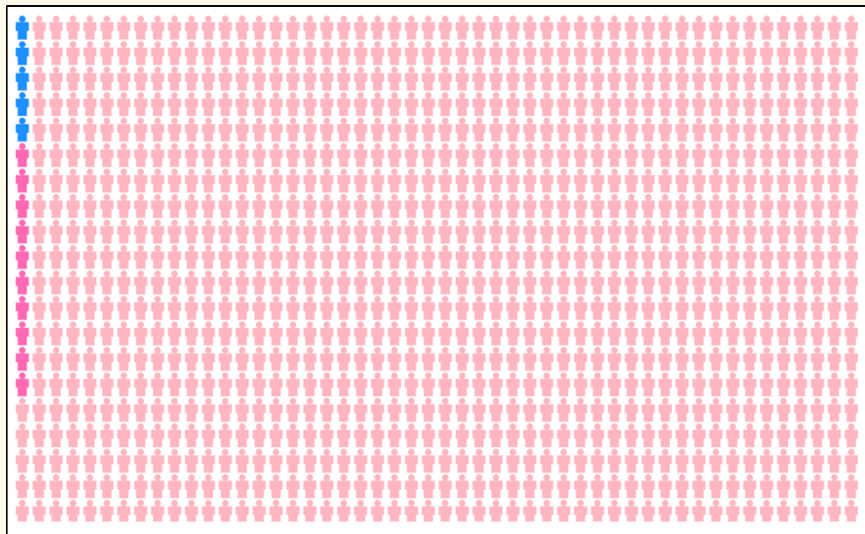
# Example – Zika Testing

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500 have Zika (0.5%)  
99,500 do not

What is the probability you have Zika (event  $Z$ ) if you test positive (event  $T$ )?



Suppose we had 100,000 people:

- 490 have Zika and test positive
- 10 have Zika and test negative
- 995 do not have Zika and test positive
- 98,505 do not have Zika and test negative

98% of those with Zika

2% of those with Zika

1% of those without Zika

$$\frac{490}{490 + 995} \approx 0.33$$

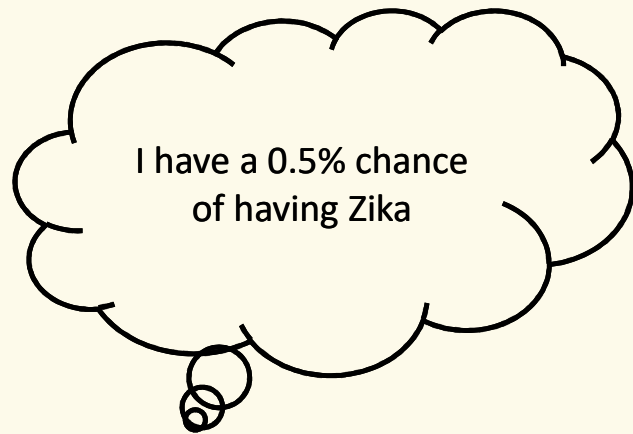
Demo

# Philosophy – Updating Beliefs

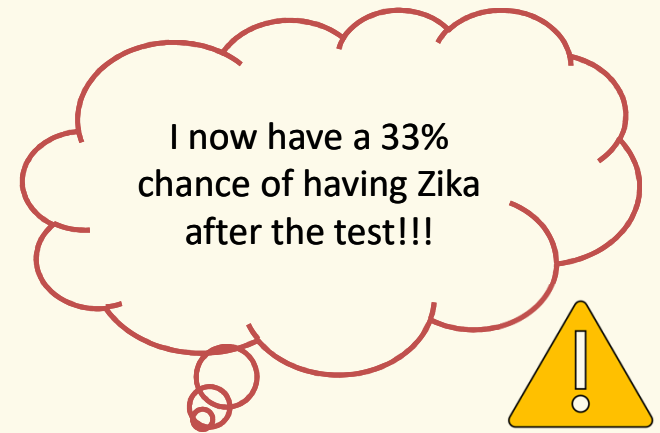
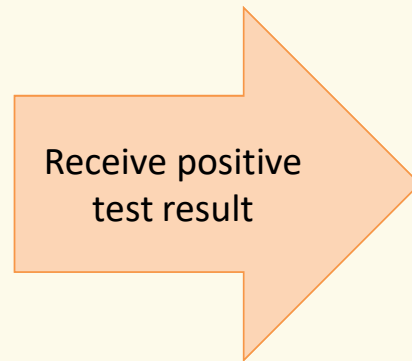
While it's not 98% that you have the disease, your beliefs changed **drastically**

$Z$  = you have Zika

$T$  = you test positive for Zika



**Prior:**  $P(Z)$



**Posterior:**  $P(Z|T)$

## Example – Zika Testing

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- 0.5% of the US population has Zika.  $P(Z)$

What is the probability you test negative (event  $T^c$ ) if you have Zika (event  $Z$ )?

$$P(T^c|Z) = 1 - P(T|Z) = 2\%$$



## Example – Zika Testing

Suppose we know the following Zika stats

- A test is 98% effective at detecting Zika (“true positive”)  $P(T|Z)$
- However, the test may yield a “false positive” 1% of the time  $P(T|Z^c)$
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What is the probability you test negative (event  $T^c$ ) if you have Zika (event  $Z$ )?

$$P(T^c|Z) = 1 - P(T|Z) = 2\%$$

What is the probability you have Zika (event  $Z$ ) if you test negative (event  $T^c$ )?

$$\text{By Bayes Rule, } P(Z|T^c) = \frac{P(T^c|Z)P(Z)}{P(T^c)}$$

By the Law of Total Probability,  $P(T^c) = P(T^c|Z)P(Z) + P(T^c|Z^c)P(Z^c)$

$$= \frac{2}{100} \cdot \frac{5}{1000} + \left(1 - \frac{1}{100}\right) \cdot \frac{995}{1000} = \frac{10}{100000} + \frac{98505}{100000}$$

$$\text{So, } P(Z|T^c) = \frac{10}{10+98505} \approx 0.01\%$$

## Bayes Theorem with Law of Total Probability

**Bayes Theorem with LTP:** Let  $E_1, E_2, \dots, E_n$  be a partition of the sample space, and  $F$  an event. Then,

$$P(E_1|F) = \frac{P(F|E_1)P(E_1)}{P(F)} = \frac{P(F|E_1)P(E_1)}{\sum_{i=1}^n P(F|E_i)P(E_i)}$$

**Simple Partition:** In particular, if  $E$  is an event with non-zero probability, then

$$P(E|F) = \frac{P(F|E)P(E)}{P(F|E)P(E) + P(F|E^C)P(E^C)}$$

## Bayes Theorem with Law of Total Probability

**Bayes Theorem with LTP:** Let  $E_1, E_2, \dots, E_n$  be a partition of the sample space, and  $F$  and event. Then,

$$P(E_1|F) = \frac{P(F|E_1)P(E_1)}{P(F)} = \frac{P(F|E_1)P(E_1)}{\sum_{i=1}^n P(F|E_i)P(E_i)}$$

We just used this implicitly on the negative Zika test example with  $E = Z$  and  $F = T^c$

**Simple Partition:** In particular, for conditional probability, then

$$P(E|F) = \frac{P(F|E)P(E)}{P(F|E)P(E) + P(F|E^c)P(E^c)}$$

## Our First Machine Learning Task: Spam Filtering


Subject: "FREE \$\$\$ CLICK HERE"

What is the probability this email is spam, given the subject contains "FREE"?

Some useful stats:

- 10% of ham (i.e., not spam) emails contain the word "FREE" in the subject.
- 70% of spam emails contain the word "FREE" in the subject.
- 80% of emails you receive are spam.

# Agenda

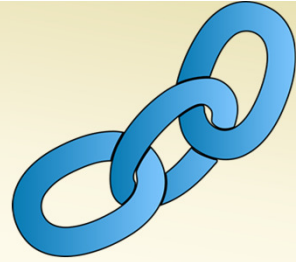
- Bayes Theorem + Law of Total Probability
- Chain Rule 
- Independence
- Infinite process and Von Neumann's trick
- Conditional independence

## Chain Rule

$$P(B|A) = \frac{P(A \cap B)}{P(A)}$$

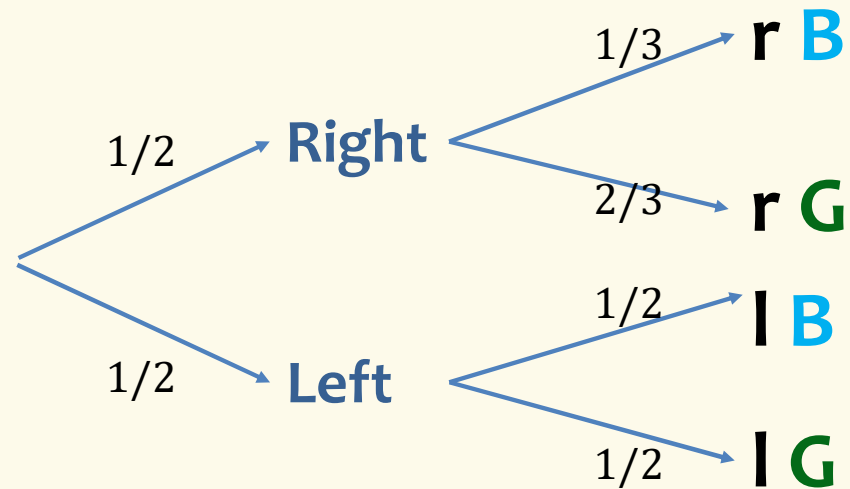


$$P(A)P(B|A) = P(A \cap B)$$



Often probability space  $(\Omega, \mathbb{P})$  is given **implicitly** via sequential process

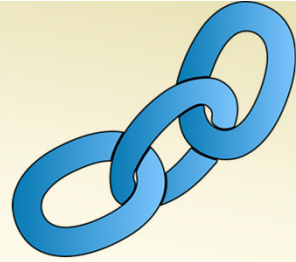
Recall from last time:



$$P(\mathbf{B}) = P(\text{Left}) \times P(\mathbf{B}|\text{Left}) + P(\text{Right}) \times P(\mathbf{B}|\text{Right})$$

What if we have more than two (e.g.,  $n$ ) steps?

## Chain Rule



$$P(B|A) = \frac{P(A \cap B)}{P(A)} \quad \longrightarrow \quad P(A)P(B|A) = P(A \cap B)$$

**Theorem. (Chain Rule)** For events  $A_1, A_2, \dots, A_n$ ,

$$P(A_1 \cap \dots \cap A_n) = P(A_1) \cdot P(A_2|A_1) \cdot P(A_3|A_1 \cap A_2)$$

$$\dots P(A_n|A_1 \cap A_2 \cap \dots \cap A_{n-1})$$

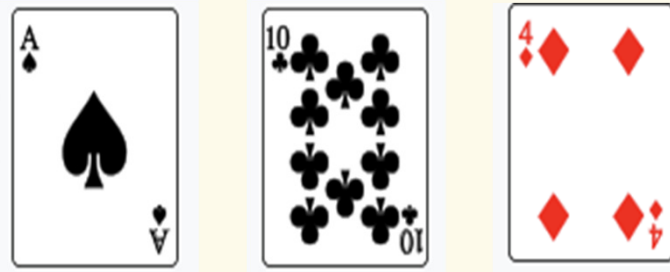
An easy way to remember: We have  $n$  tasks and we can do them **sequentially**, conditioning on the outcome of previous tasks



## Chain Rule Example

Shuffle a standard 52-card deck and draw the top 3 cards.  
(uniform probability space)

What is  $P(\text{Ace of Spades, 10 of Clubs, 4 of Diamonds}) = P(A \cap B \cap C)$ ?



$A$ : Ace of Spades First  
 $B$ : 10 of Clubs Second  
 $C$ : 4 of Diamonds Third

$$P(A) \cdot P(B|A) \cdot P(C|A \cap B)$$

$$\frac{1}{52} \cdot \frac{1}{51} \cdot \frac{1}{50}$$

# Agenda

- Bayes Theorem + Law of Total Probability
- Chain Rule
- Independence ◀
- Infinite process and Von Neumann's trick
- Conditional independence

# Independence

**Definition.** Two events  $A$  and  $B$  are (statistically) **independent** if

$$P(A \cap B) = P(A) \cdot P(B).$$

Equivalent formulations:

- If  $P(A) \neq 0$ , equivalent to  $P(B|A) = P(B)$
- If  $P(B) \neq 0$ , equivalent to  $P(A|B) = P(A)$

“The probability that  $B$  occurs after observing  $A$ ” – Posterior  
= “The probability that  $B$  occurs” – Prior

## Independence - Example

Assume we toss two fair coins

*“first coin is heads”*

$$A = \{HH, HT\}$$

*“second coin is heads”*

$$B = \{HH, TH\}$$

$$P(A) = 2 \times \frac{1}{4} = \frac{1}{2}$$

$$P(B) = 2 \times \frac{1}{4} = \frac{1}{2}$$

$$P(A \cap B) = P(\{HH\}) = \frac{1}{4} = P(A) \cdot P(B)$$

## Example – Independence

Toss a coin 3 times. Each of 8 outcomes equally likely.

- $A = \{\text{at most one } T\} = \{HHH, HHT, HTH, THH\}$
- $B = \{\text{at most 2 } H\text{'s}\} = \{HHH\}^c$

Independent?

$$P(A \cap B) \stackrel{?}{=} P(A) \cdot P(B)$$

$$\frac{3}{8} \neq \frac{1}{2} \cdot \frac{7}{8}$$

Poll:

A. Yes, independent

B. No

[pollev/paulbeame028](https://polllev.com/paulbeame028)

## Multiple Events – Mutual Independence

**Definition.** Events  $A_1, \dots, A_n$  are **mutually independent** if for every non-empty subset  $I \subseteq \{1, \dots, n\}$ , we have

$$P\left(\bigcap_{i \in I} A_i\right) = \prod_{i \in I} P(A_i).$$

## Example – Network Communication

Each link works with the probability given, **independently**

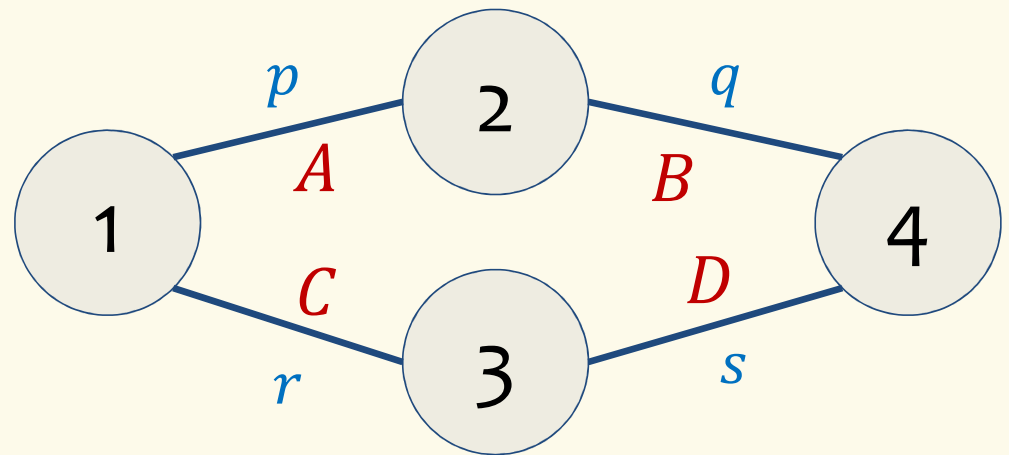
i.e., mutually independent events  $A, B, C, D$  with

$$P(A) = p$$

$$P(B) = q$$

$$P(C) = r$$

$$P(D) = s$$



## Example – Network Communication

If each link works with the probability given, **independently**:  
What's the probability that nodes 1 and 4 can communicate?

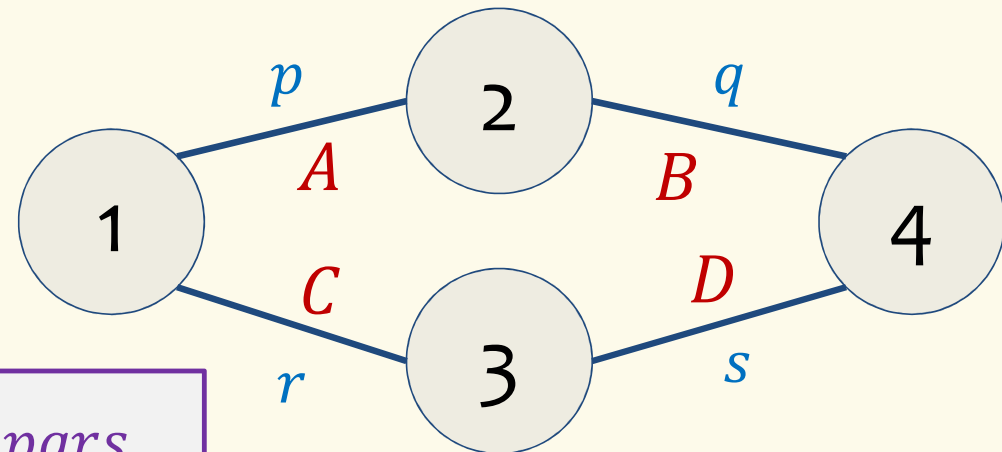
$$\begin{aligned}P(1-4 \text{ connected}) &= P((A \cap B) \cup (C \cap D)) \\ &= P(A \cap B) + P(C \cap D) - P(A \cap B \cap C \cap D)\end{aligned}$$

$$P(A \cap B) = P(A) \cdot P(B) = pq$$

$$P(C \cap D) = P(C) \cdot P(D) = rs$$

$$P(A \cap B \cap C \cap D)$$

$$= P(A) \cdot P(B) \cdot P(C) \cdot P(D) = pqrs$$



$$P(1-4 \text{ connected}) = pq + rs - pqrs$$



## Independence as an assumption

- People often assume it **without justification**

- Example: A skydiver has two chutes

$A$ : event that the main chute doesn't open       $P(A) = 0.02$

$B$ : event that the back-up doesn't open       $P(B) = 0.1$

- What is the chance that at least one opens assuming independence?

Assuming independence doesn't justify the assumption!

Both chutes could fail because of the same rare event e.g., freezing rain.

## Independence – Another Look

**Definition.** Two events  $A$  and  $B$  are (statistically) **independent** if

$$P(A \cap B) = P(A) \cdot P(B).$$

**“Equivalently.”**  $P(A|B) = P(A)$ .

It is important to understand that independence is a property of probabilities of outcomes, not of the root cause generating these events.

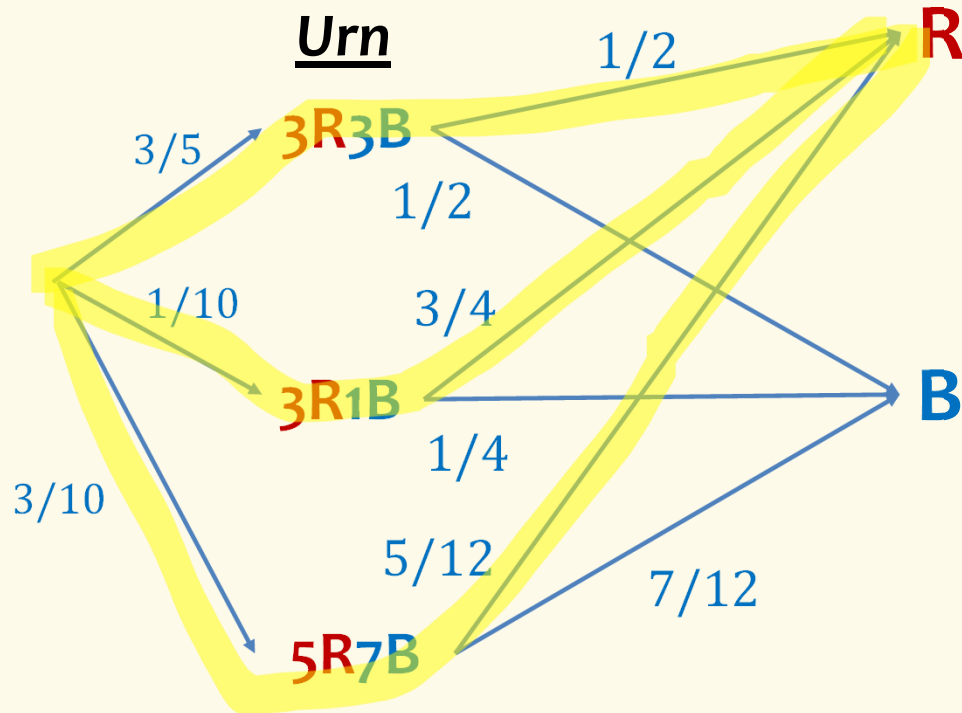
*Events generated independently → their probabilities satisfy independence*

*← Not necessarily*

This can be counterintuitive!

## Sequential Process

## Ball drawn



**Setting:** An urn contains:

- 3 **red** and 3 **blue** balls w/ probability  $3/5$
- 3 **red** and 1 **blue** balls w/ probability  $1/10$
- 5 **red** and 7 **blue** balls w/ probability  $3/10$

We draw a ball at random from the urn.

$$P(\mathbf{R}) = \frac{3}{5} \times \frac{1}{2} + \frac{1}{10} \times \frac{3}{4} + \frac{3}{10} \times \frac{5}{12} = \frac{1}{2}$$

$$P(\mathbf{3R3B}) \times P(\mathbf{R} \mid \mathbf{3R3B})$$

Are **R** and **3R3B** independent?

**Independent!**  $P(\mathbf{R}) = P(\mathbf{R} \mid \mathbf{3R3B})$



# Agenda

- Bayes Theorem + Law of Total Probability
- Chain Rule
- Independence
- **Infinite process and Von Neumann's trick**
- Conditional independence

Often probability space  $(\Omega, P)$  is given **implicitly** via sequential process

- *Experiment proceeds in  $n$  sequential steps, each step follows some **local rules** defined by the chain rule and independence*
- *Natural extension: Allows for easy definition of experiments where  $|\Omega| = \infty$*

## Fun: Von Neumann's Trick with a biased coin

- How to use a biased coin to get a fair coin flip:
  - Suppose that you have a biased coin:
    - $P(H) = p$      $P(T) = 1 - p$

1. Flip coin twice: If you get  $HH$  or  $TT$  go to step 1
2. If you got  $HT$  output  $H$ ; if you got  $TH$  output  $T$ .

Why is it fair?  $P(H) = P(HT) = p(1 - p) = (1 - p)p = P(TH) = P(T)$

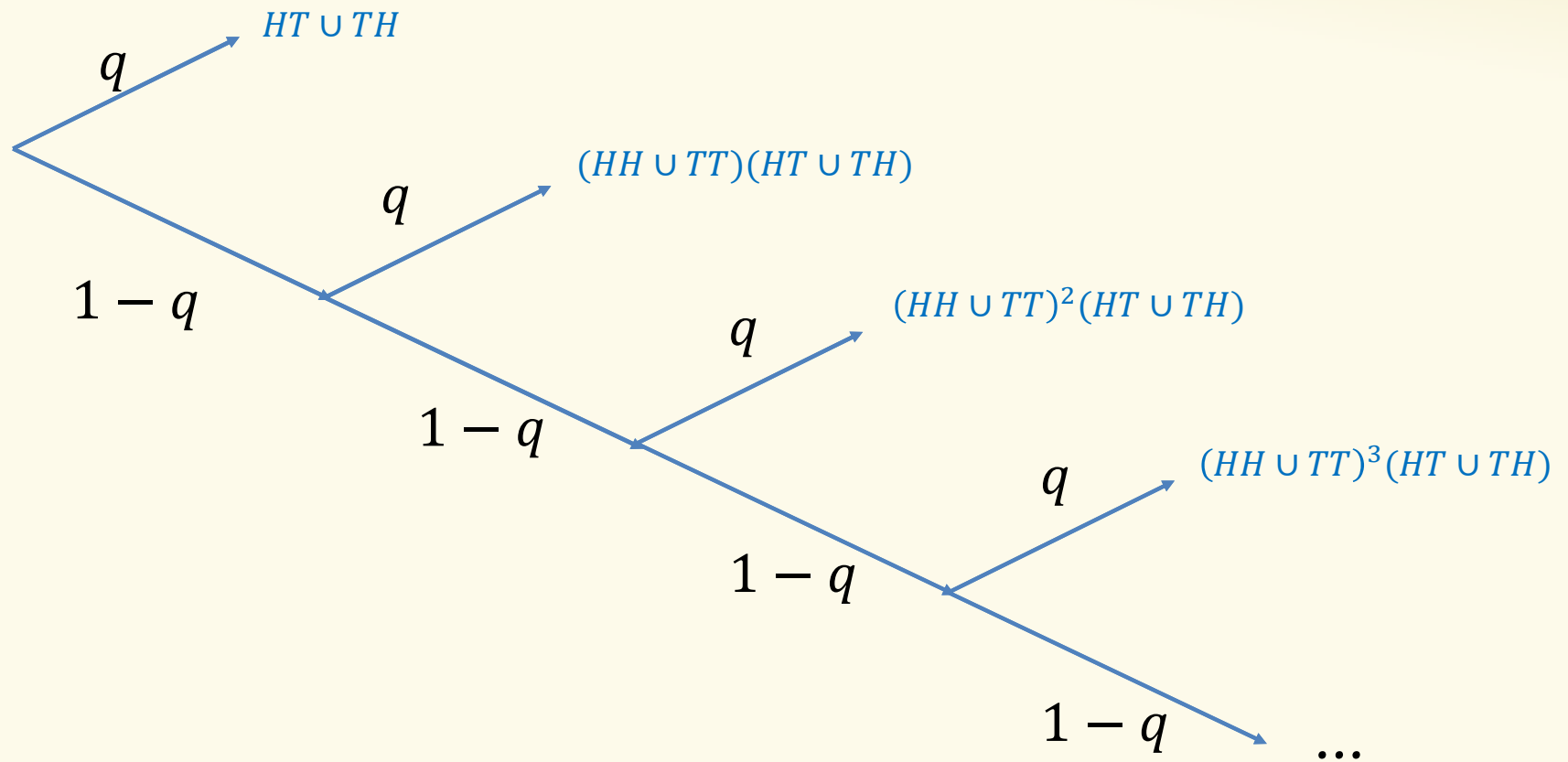
Drawback: You may never get to step 2.

## The sample space for Von Neumann's trick

- For each round of Von Neumann's trick we flipped the biased coin twice.
  - If  $HT$  or  $TH$  appears, the experiment ends:
    - Total probability each round:  $2p(1 - p)$  call this  $q$
  - If  $HH$  or  $TT$  appears, the experiment continues:
    - Total probability each round:  $p^2 + (1 - p)^2$  this is  $1 - q$
- Probability that flipping ends in round  $t$  is  $(1 - q)^{t-1} \cdot q$ 
  - Conditioned on ending in round  $t$ ,  $P(H) = P(T) = 1/2$



## Sequential Process – Example



## The sample space for Von Neumann's trick

More precisely, the sample space contains the successful outcomes:

$$\bigcup_{t=1}^{\infty} (HH \cup TT)^{t-1} (HT \cup TH)$$

which together have probability  $\sum_{t=1}^{\infty} (1-q)^{t-1} q$  for  $q = 2p(1-p)$

as well as all of the failing outcomes in  $(HH \cup TT)^{\infty}$ .

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Observe that  $q \neq 0$  iff  $0 < p < 1$ . We have two cases:

- If  $q \neq 0$  then  $\sum_{t=1}^{\infty} (1-q)^{t-1} = 1/q$  so successful outcomes account for total probability 1.
- If  $q = 0$  then either:
  - $p = 1$  and  $(HH)^{\infty}$  has probability 1.
  - $p = 0$  and  $(TT)^{\infty}$  has probability 1.

## Agenda

- Bayes Theorem + Law of Total Probability
- Chain Rule
- Independence
- Infinite process and Von Neumann's trick
- **Conditional independence** ◀

## Conditional Independence

**Definition.** Two events  $A$  and  $B$  are **independent** conditioned on  $C$  if  $P(C) \neq 0$  and  $P(A \cap B | C) = P(A | C) \cdot P(B | C)$ .

- If  $P(A \cap C) \neq 0$ , equivalent to  $P(B|A \cap C) = P(B | C)$
- If  $P(B \cap C) \neq 0$ , equivalent to  $P(A|B \cap C) = P(A | C)$

**Plain Independence.** Two events  $A$  and  $B$  are **independent** if

$$P(A \cap B) = P(A) \cdot P(B).$$

- If  $P(A) \neq 0$ , equivalent to  $P(B|A) = P(B)$
- If  $P(B) \neq 0$ , equivalent to  $P(A|B) = P(A)$

## Example – Throwing Dice

Suppose that Coin 1 has probability of heads 0.3  
and Coin 2 has probability of head 0.9.

We choose one coin randomly with equal probability and flip that coin 3 times independently. What is the probability we get all heads?

$$P(HHH) = P(HHH | C_1) \cdot P(C_1) + P(HHH | C_2) \cdot P(C_2) \quad \text{Law of Total Probability (LTP)}$$

$$= P(H|C_1)^3 P(C_1) + P(H | C_2)^3 P(C_2) \quad \text{Conditional Independence}$$

$$= 0.3^3 \cdot 0.5 + 0.9^3 \cdot 0.5 = 0.378$$

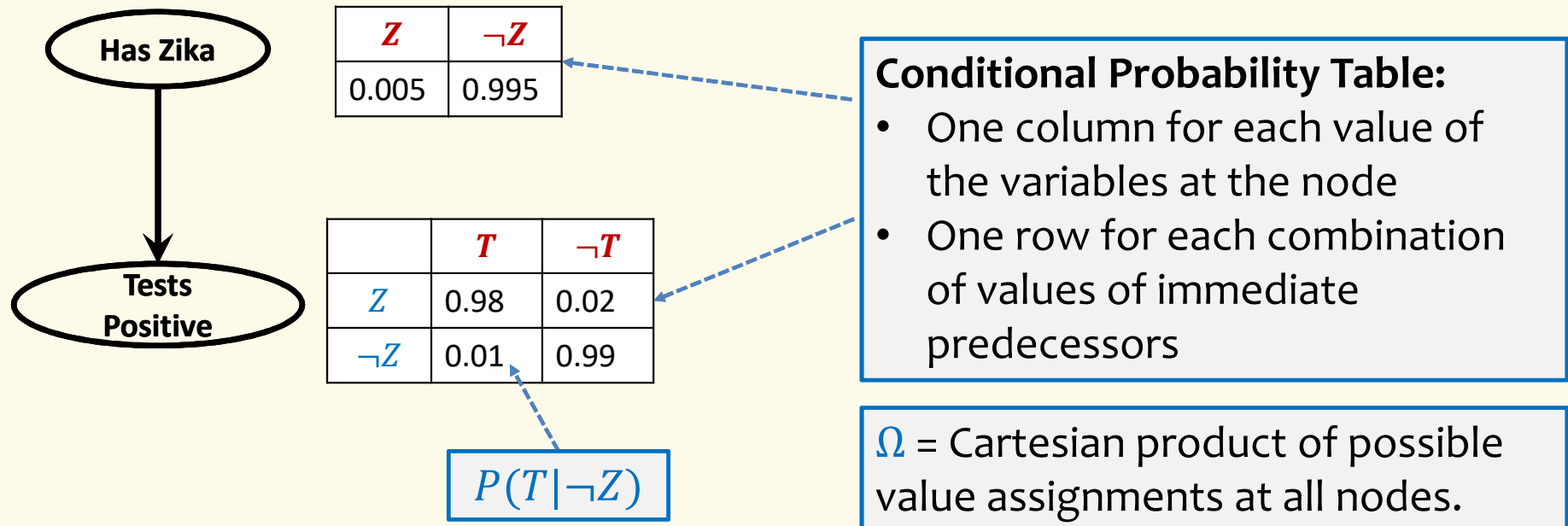
$C_i$  = coin  $i$  was selected

## Conditional independence and Bayesian inference in practice: Graphical models

- The sample space  $\Omega$  is often the Cartesian product of possibilities of many different variables
- We often can understand the probability distribution  $P$  on  $\Omega$  based on **local properties** that involve a few of these variables at a time
- We can represent this via a directed acyclic graph augmented with probability tables (called a **Bayes net**) in which each node represents one or more variables...

## Graphical Models/Bayes Nets

- Bayes net for the Zika testing probability space  $(\Omega, P)$







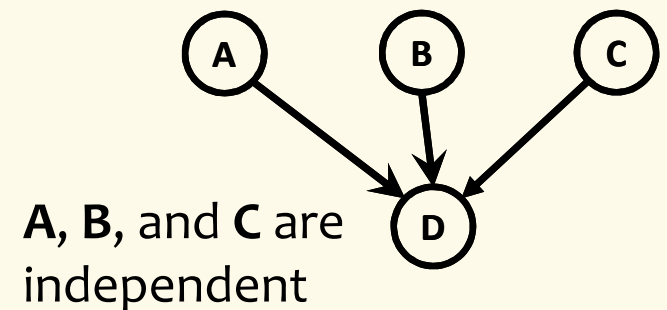
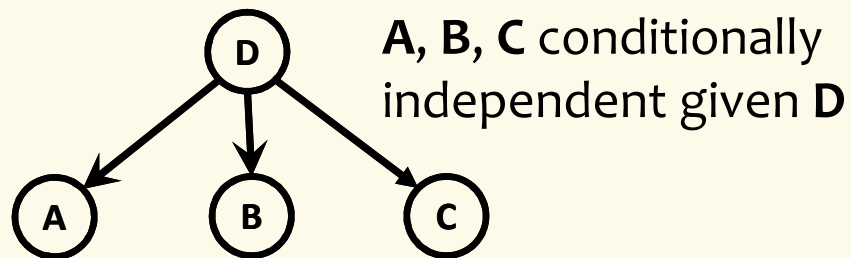
## Graphical Models/Bayes Nets

### Bayes Net assumption/requirement

- The only dependence between variables is given by paths in the Bayes Net graph:

- if only edges are 

then **A** and **C** are *conditionally independent* given the value of **B**



Defines a unique global probability space  $(\Omega, P)$

