# Lecture 12

CSE 527: Computational Biology November 13, 2001 Notes: Scott Votaw

## **The Motif-Finding Problem**

Given: *n* strings  $x_1, x_2, ..., x_n$ , each with length *m* each string has an instance of a motif of length  $l \ll m$ 

Example:

$x_{l}$	$\Rightarrow$	А	С	Т	G	С	Т	А	T	A	A	T	С	T	G	Т	Т	А	G	С
$x_2$	$\Rightarrow$	Т	А	G	С	T	A	T	A	A	T	G	G	С	Т	Т	А	Т	G	А
<i>x</i> <sub>3</sub>	$\Rightarrow$	G	А	G	А	А	Т	А	Т	G	G	С	С	С	T	A	T	A	A	T
$x_n$	$\Rightarrow$	Т	T	A	С	A	$\boldsymbol{A}$	T	Т	G	Т	А	G	G	G	Т	Α	А	А	G

The simplest motif model would simply designate an exact sequence of characters to match (e.g. "TATAAT"). A measure of the match at any location k in the string would simply be the number of mismatches.

A slightly more advanced model of a motif of length l=6 (so-called "TATAAT" box) can be represented by a probability table. This is better than the previous model since the motif does not have to be an exact sequence. (This is a 0<sup>th</sup> order Markov model since all positions are independent)

	1	2	3	4	5	6
A	5	85	2	80	82	1
С	3	5	3	10	8	3
G	2	6	3	12	5	2
Т	90	4	92	8	5	94

Hypothetical probability model for a TATAAT box

In order to give an estimate of the likelihood of a model, there must be a comparison model. The "background" model can assume equal distribution, or represent the actual distribution of each character (e.g. 42% GC, 58% AT).

	1	2	3	4	5	6
Α	29	29	29	29	29	29
С	21	21	21	21	21	21
G	21	21	21	21	21	21
Т	29	29	29	29	29	29

Background model representing overall base pair frequencies

Each model ( $M_{test}$  and  $M_{background}$ ) has a probability of generating *every* possible string. For the 0<sup>th</sup> order Markov models given above, the probability of a particular sequence occurring is simply the product of the probabilities for each character. This is true since the individual characters are independent in this type of model.

 $p_{\text{test}}(\text{TATAAT}) = .90 * .85 * .92 * .80 * .82 * .94$  $p_{\text{background}}(\text{TATAAT}) = .29 * .29 * .29 * .29 * .29 * .29$ 

The natural way to compare the two models is to generate a ratio of the probabilities. This ratio is called the natural score.

natural score = 
$$\frac{p(\text{test model})}{p(\text{background model})}$$

Multiplying small numbers can be inaccurate computationally, so alternatively we can sum the logs.

$$\log \frac{\prod p_i}{\prod q_i} = \sum_{i=1}^l \log \frac{p_i}{q_i}$$

Defining  $\theta = \log \frac{p_i}{q_i}$ , finding a good motif is equivalent to finding a good  $\theta$ . We wish to maximize the value of  $\theta$  so that the relative probability of our test model is greatest with respect

#### Using E.M. to solve the Motif Finding Problem

Given: *n* strings  $x_1, x_2, ..., x_n$ , each with length *m* 

Define:  $y_{ik} = 1$  if motif starts at position k in string i = 0 otherwise

$$\theta = \log \frac{p_i}{q_i}$$

to the background model.

We need:

1)  $p(x_i | \theta, y_{ik} = 1)$ 2) a way to find  $\theta'$  maximizing above given some data

#### The E.M. (Expectation Maximization) algorithm

**Given:**  $\theta^{t}$  (an initial estimate of  $\theta$ )

### **E** Step: estimates $E_{\theta^{t}}(y_{ik})$

(Calculate the expectation that the motif exists at location k in each string i)

$$E_{\theta^{t}}(y_{ik}) = E(y_{ik} = 1 | x_{i}, \theta^{t})$$
  
= 1 ·  $p(y_{ik} = 1 | x_{i}, \theta^{t}) + 0 · p(...)$   
=  $\frac{p(x_{i} | y_{ik} = 1, \theta^{t}) \cdot p(y_{ik} = 1 | \theta^{t})}{p(x_{i} | \theta^{t})}$   
(by Bayes' rule)

Trick 1: The denominator is independent of  $\underline{y_{ik}}$ , so it cancels out in the ratio Trick 2: The  $p(y_{ik} = 1 | \theta^t)$  is the prior belief, and can often be assumed a constant that can be factored out

#### M Step: maximize $\theta$ given data

$$Q(\theta \mid \theta^{t}) = E_{\theta^{t}} (\log (x, y \mid \theta))$$
  
= E ( log  $\prod_{i=1}^{n} p(x_{ij}, y_i \mid \theta)$ )  
(Trick: only one  $y_i$  is 1, the rest are zero)

$$= E \left(\log \prod_{i=1}^{n} \prod_{k=1}^{m} p(x_{ij} \mid \theta, y_{ik} = 1)\right)^{y_{ik}}$$
  
$$= E \left(\sum_{i=1}^{n} \sum_{k=1}^{m} y_{ik} \log(p(x_i \mid \theta, y_k = 1))\right)$$
  
$$= \sum_{i=1}^{n} \sum_{k=1}^{m} E(y_{ik}) \log(p(x_i \mid \theta, y_k = 1))$$
  
(since  $p(x_i \mid \theta, y_k = 1)$  is independent of expectation)

Goal: find  $\theta$  maximizing Q function given data

### How do you start the E.M. algorithm? (i.e. how to determine initial $\theta^{t}$ )

Setting  $\theta$ 's equal to the background is a bad idea since there will not be any progress on the first iteration.

Other possible starts:

- Start with random values and repeat many times
- Use prior knowledge
- MEME (Bailey and Elkan, San Diego Super Computer Center)
  - 1. Try all length l substrings of  $x_i$ 's
    - 2. Do 2 iterations of E.M. for each seed
    - 3. Pick best few and do full iteration

## **Model Selection**

General problem: Given 2 models  $M_1$  and  $M_2$ , which is better?

Given:  $M_1(\theta_1)$ ,  $M_2(\theta_2)$ , and observed data D

What's  $p(M_1 | D)$ ?

by Bayes rule:  $p(M_1 | D) = \frac{p(D | M_1) \cdot p(M_1)}{p(D | M_1)p(M_1) \cdot p(D | M_2)p(M_2)}$ 

What's  $p(M_1)$  and  $p(M_2)$ ?

These are the a priori probabilities of each model being correct.

- could assume equal probability of both models (i.e., 50/50)
- could calculate some other a priori probability using knowledge

e.g., data is a distribution of 8 points

$$\begin{split} M_1 &= 1 \text{ distribution with} \\ 1 \ \mu \text{ and } \sigma^2 \\ M_2 &= 2 \text{ populations with} \\ & \text{ different } \mu \text{ and } \sigma^2 \end{split}$$



We could use likelihood directly, but more complicated models fit better (e.g.  $M_3 = 8$  populations with mean  $\mu$  and  $\sigma^2 = 0$ ). Intuition: penalize extra degrees of freedom. This can be done using a BIC score.

## **BIC score (Bayesian Information Criterion)**

BIC = likelihood - penalty for degrees of freedom =  $2 \log[p(x | \hat{\theta})] - d \log n$ 

(where d = number of free parameters and n = number of data points)