

CSE 527

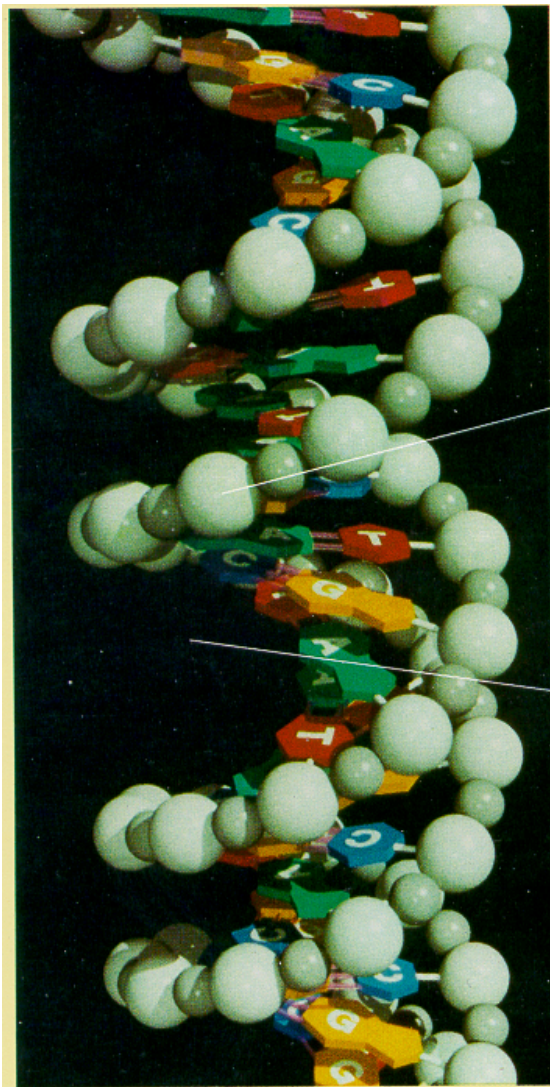
Autumn 2007

Lectures 8-9 (& part of 10)
Motifs: Representation & Discovery

DNA Binding Proteins

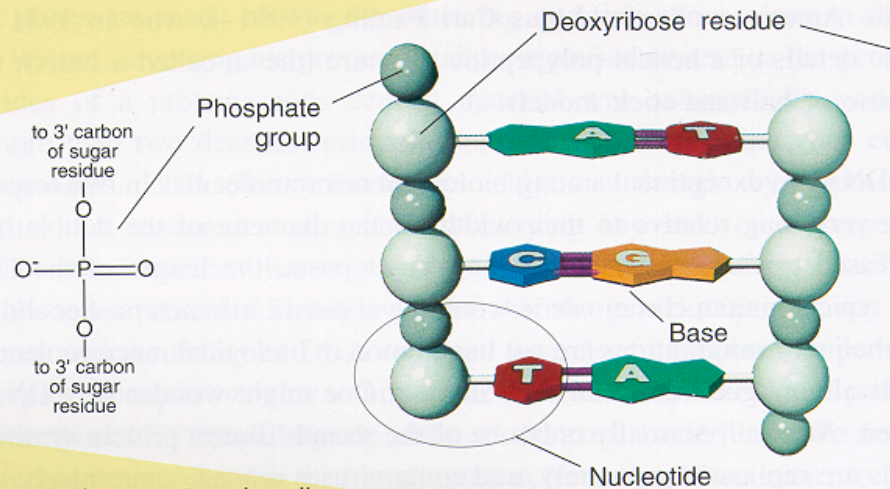
A variety of DNA binding proteins
("transcription factors"; a significant fraction,
perhaps 5-10%, of all human proteins)
modulate transcription of protein coding
genes

The Double Helix



(a) Computer-generated Image of DNA (by Mel Prueitt)

(b) Uncoiled DNA Fragment

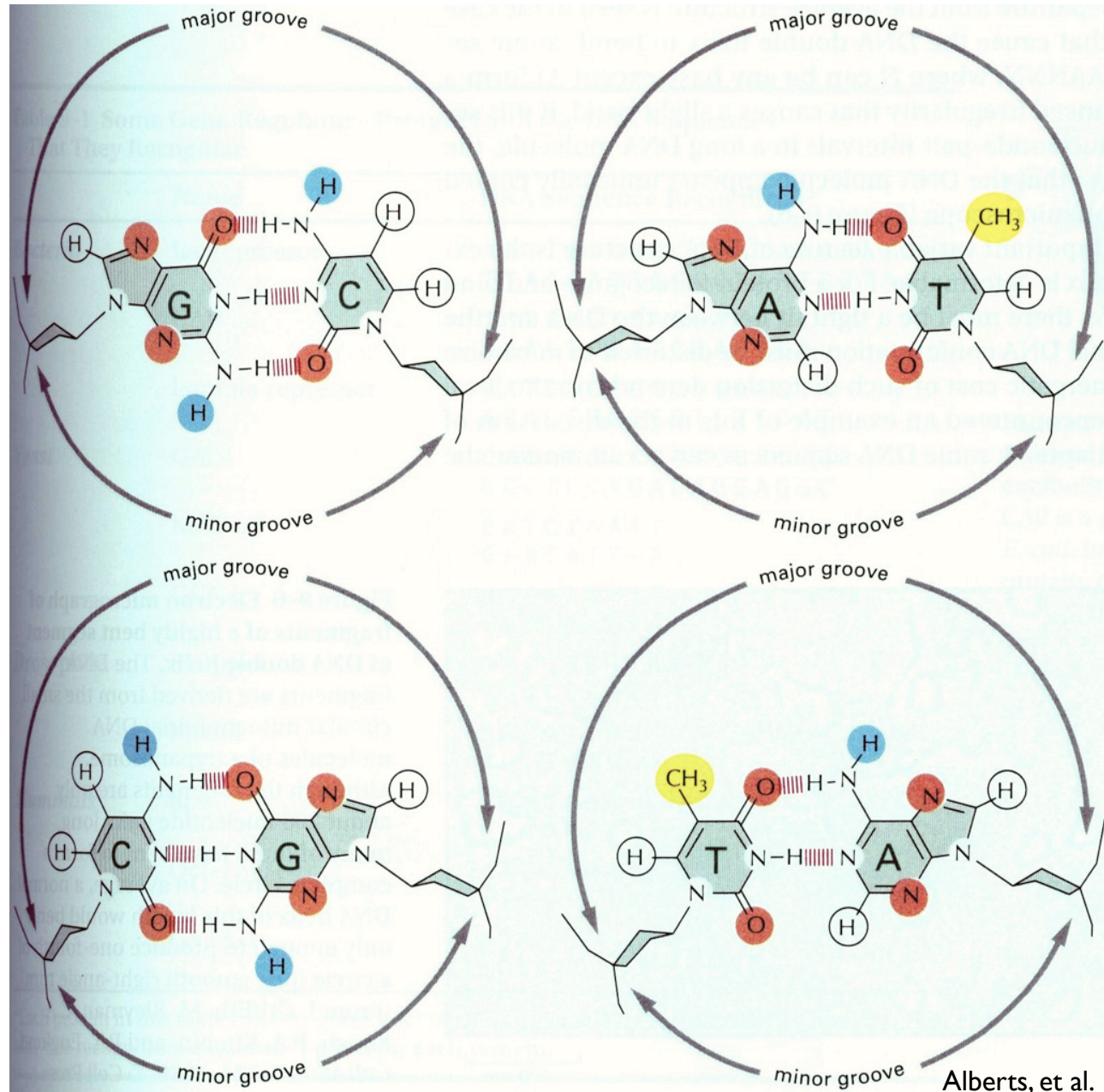


As shown, the two strands coil about each other in a fashion such that all the bases project inward toward the helix axis. The two strands are held together by hydrogen bonds (pink rods) linking each base projecting from one backbone to its so-called complementary base projecting from the other backbone. The base A always bonds to T (A and T are comple-

Shown in (b) is an uncoiled fragment of (a) three complementary base pair chemist's viewpoint, each strand a polymer made up of four re-called deoxyribonucleotides

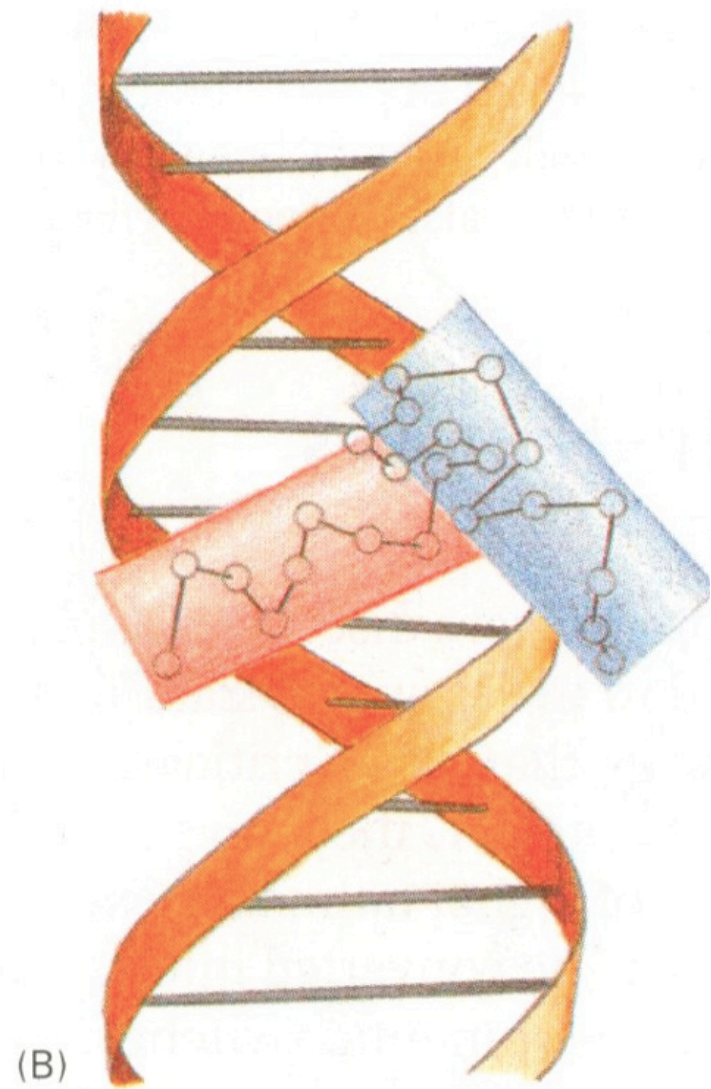
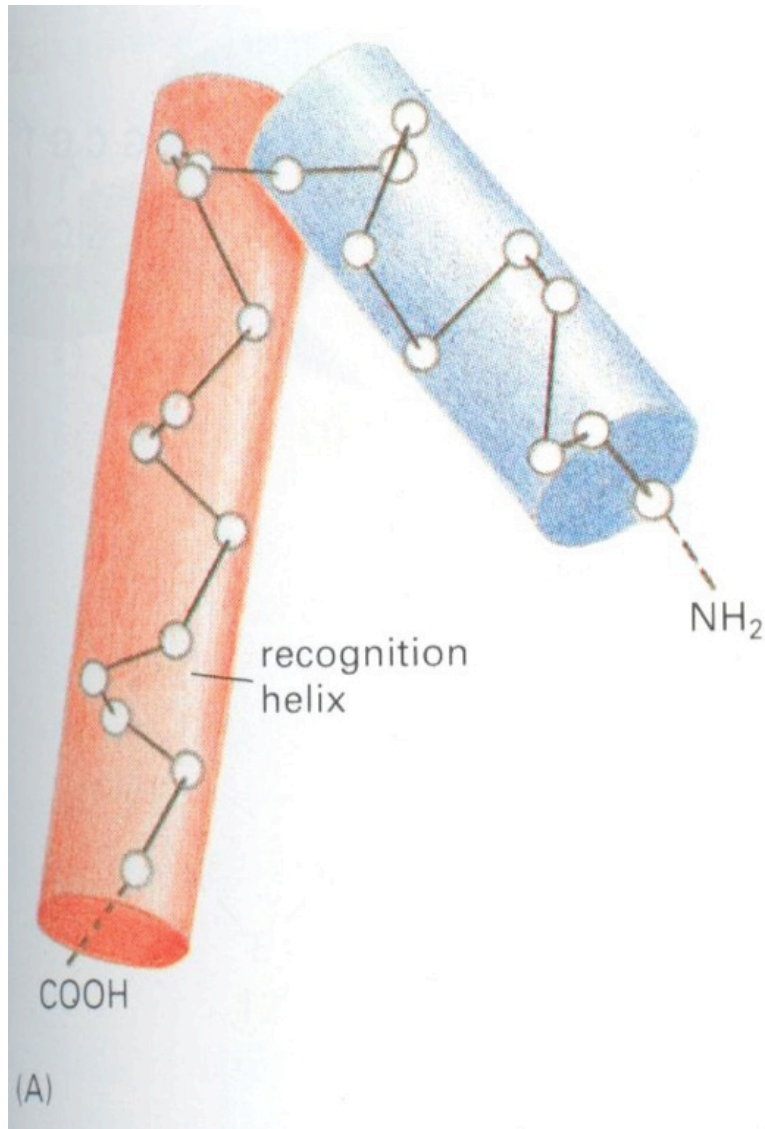
In the groove

Different patterns of potential H bonds at edges of different base pairs, accessible esp. in major groove

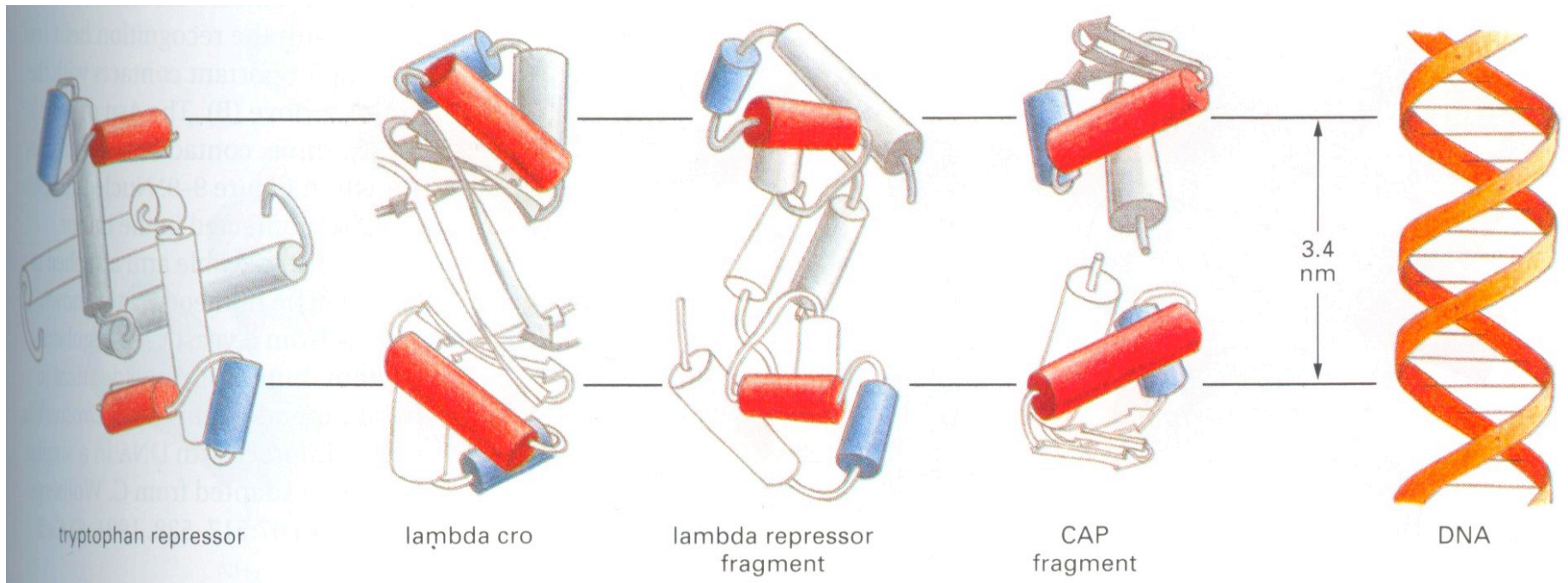


Alberts, et al.

Helix-Turn-Helix DNA Binding Motif



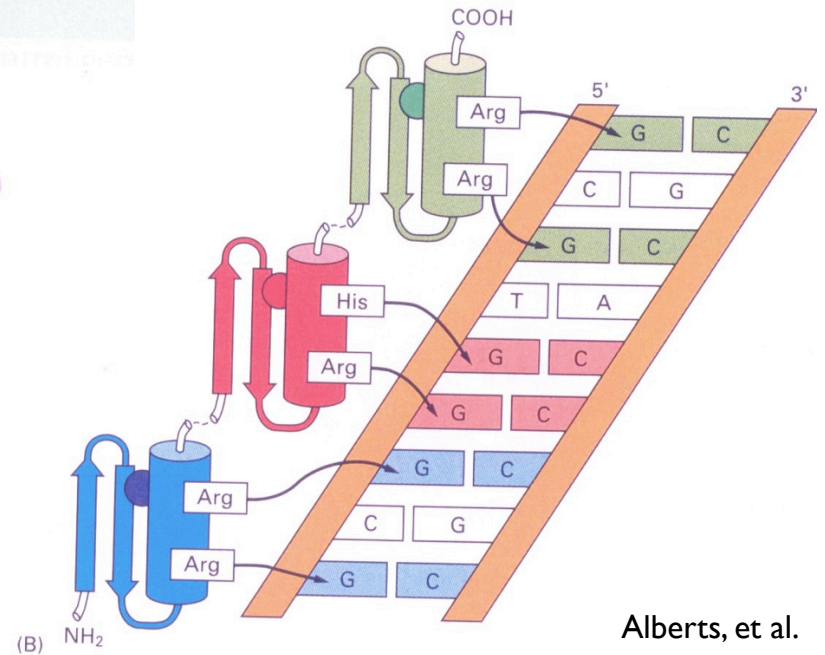
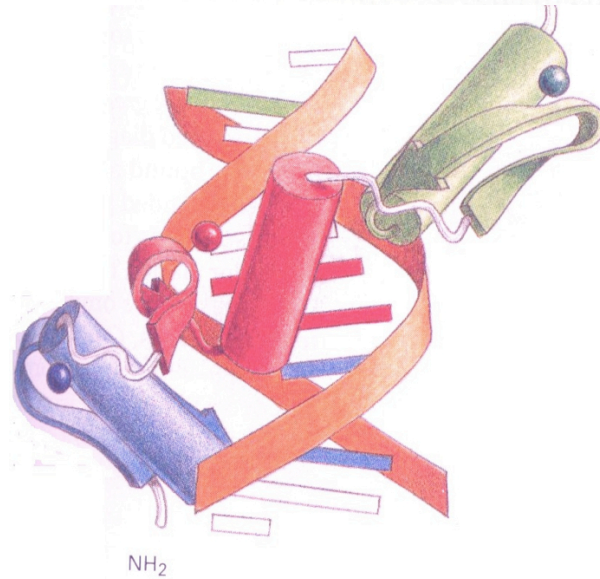
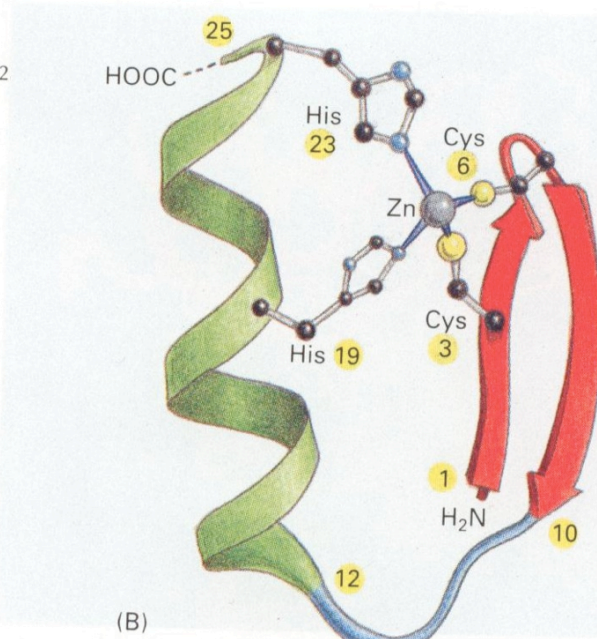
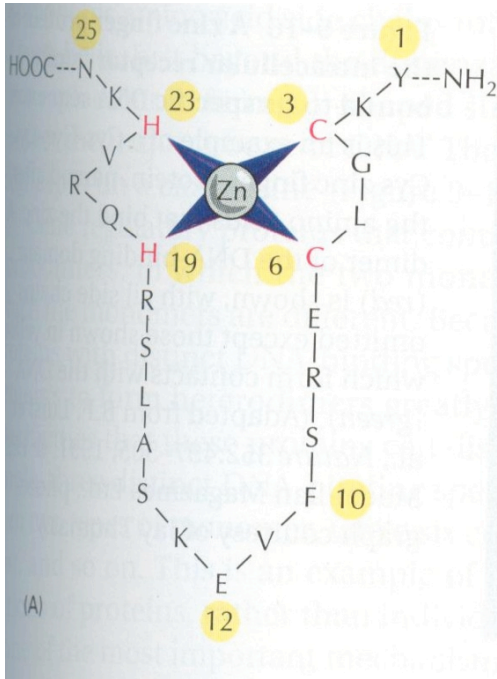
H-T-H Dimers



Alberts, et al.

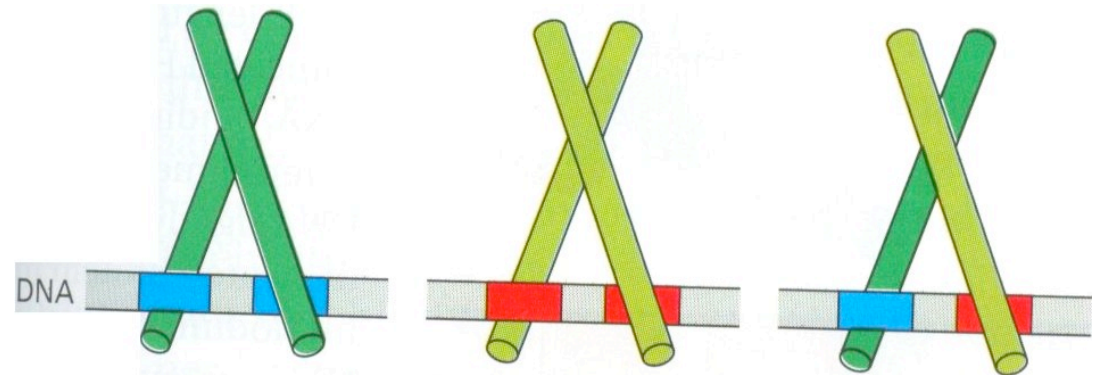
Bind 2 DNA patches, ~ 1 turn apart
Increases both specificity and affinity

Zinc Finger Motif



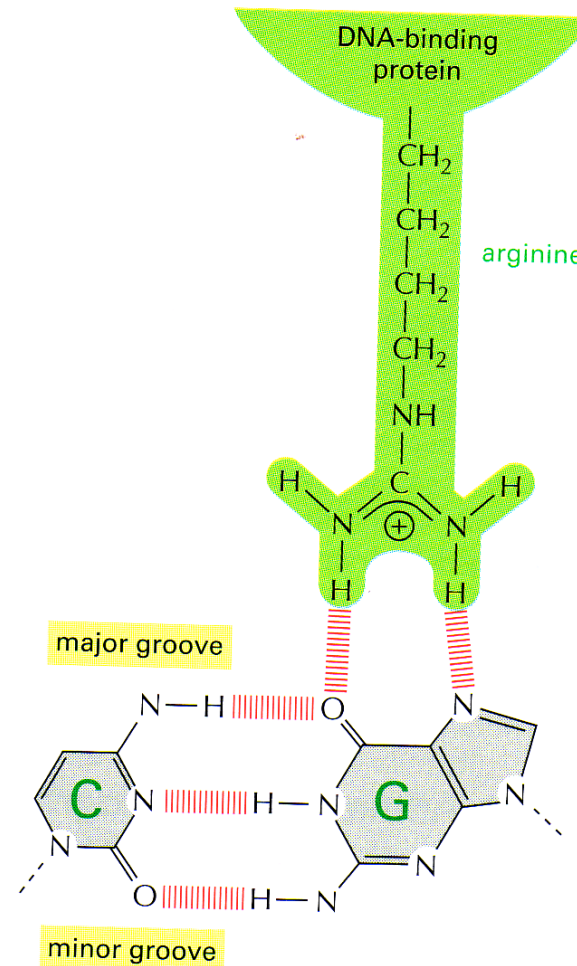
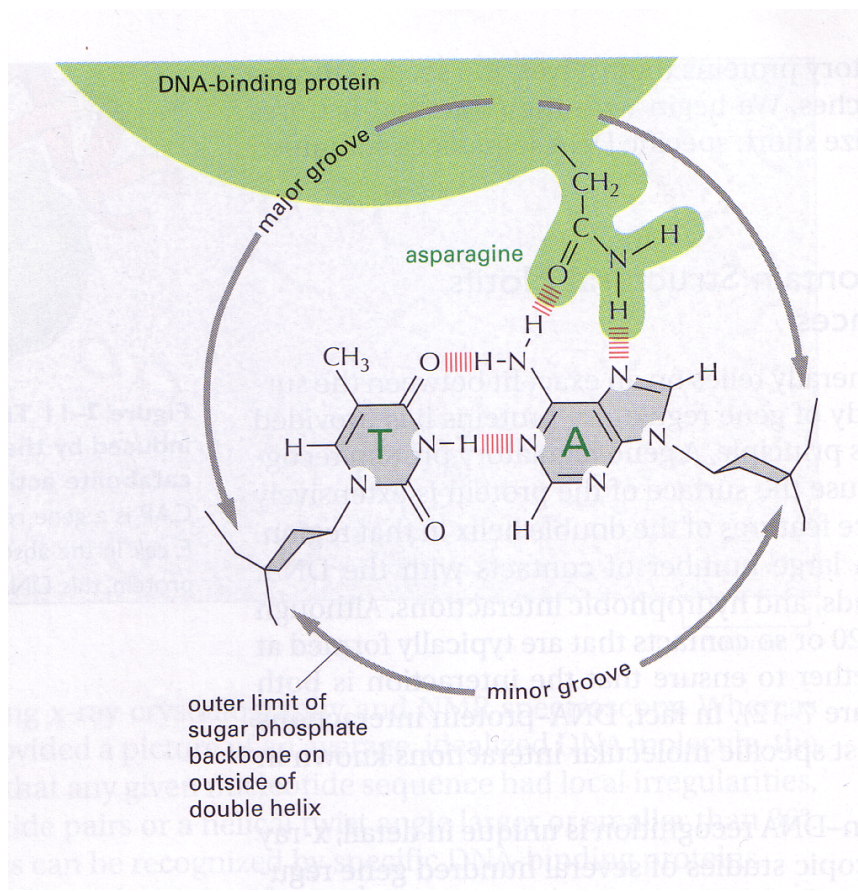
Alberts, et al.

Leucine Zipper Motif

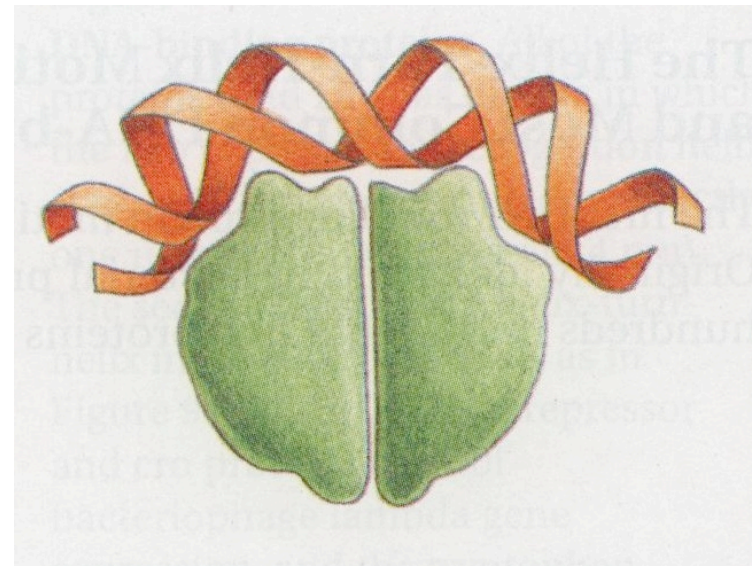
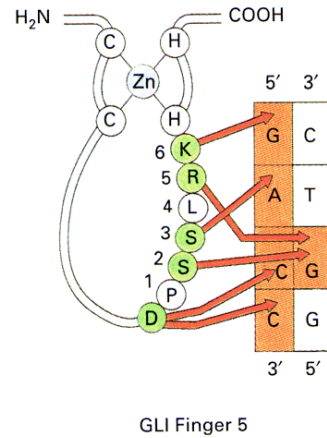
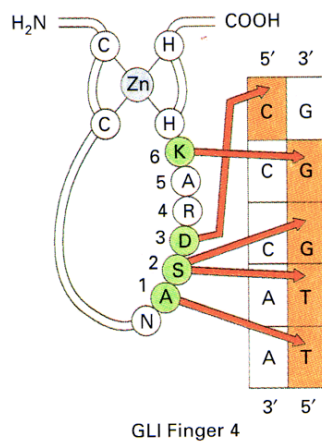
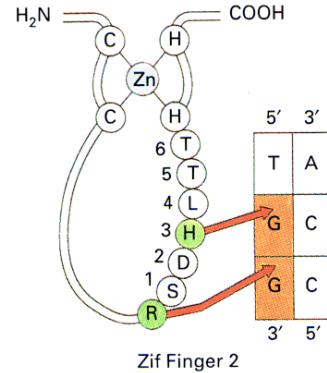
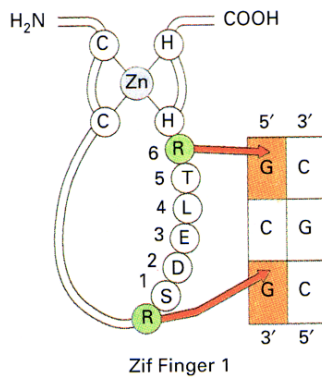
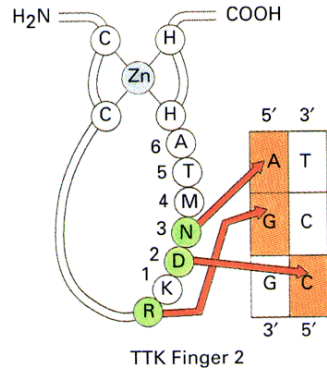
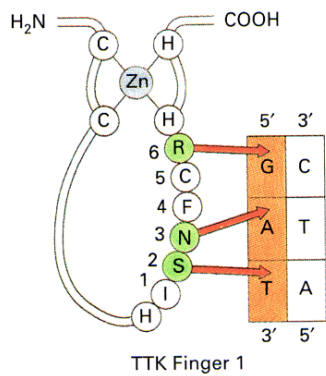


Homo-/hetero-dimers
and combinatorial
control

Some Protein/DNA interactions well-understood



But the overall DNA binding “code” still defies prediction

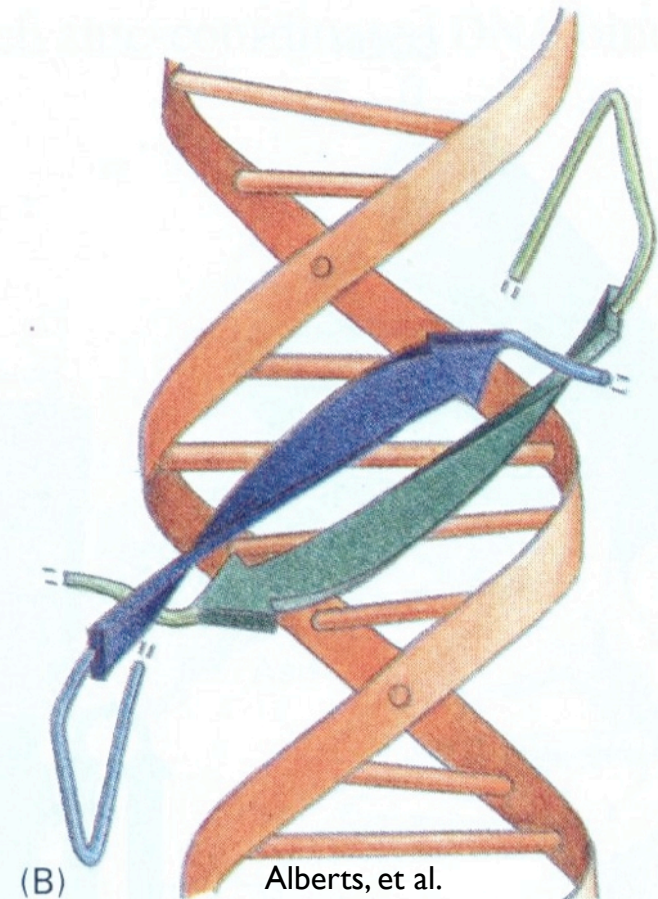
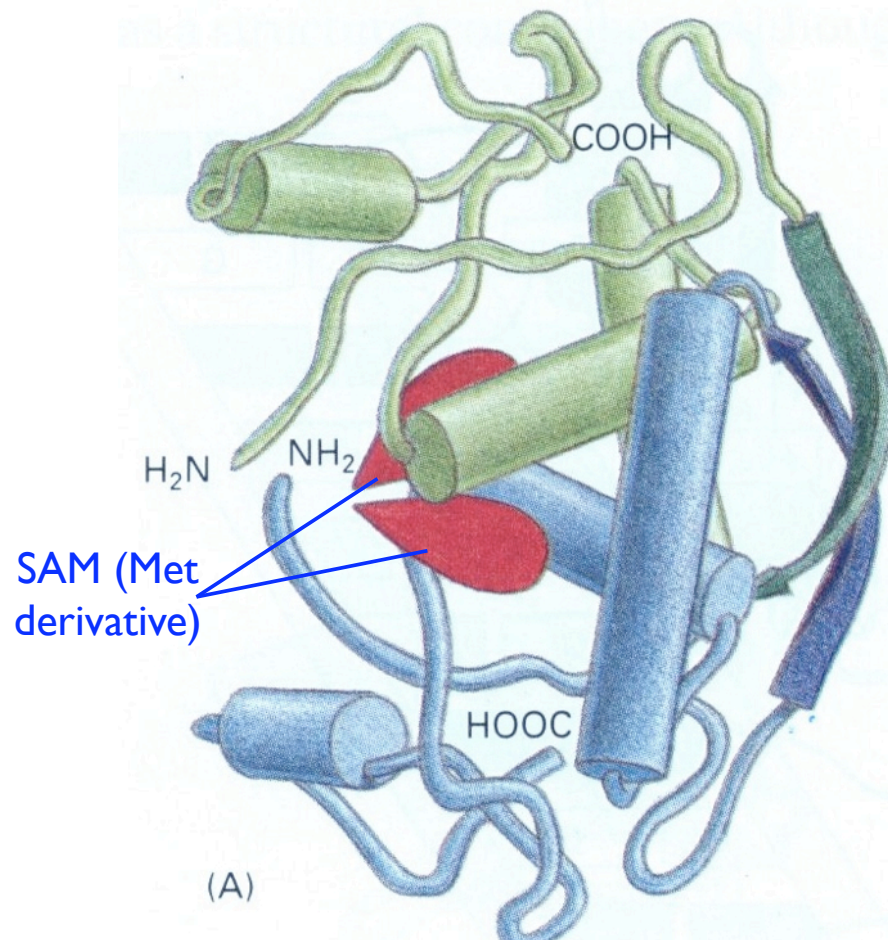


Bacterial Met Repressor

a beta-sheet DNA binding domain

Negative feedback loop:

high Met level \Rightarrow repress Met synthesis genes



DNA binding site summary

- complex “code”
- short patches (6-8 bp)
- often near each other (1 turn = 10 bp)
- often reverse-complements
- not perfect matches

Sequence Motifs

Last few slides described *structural* motifs in proteins

Equally interesting are the *DNA sequence* motifs to which these proteins bind - e.g. , one leucine zipper dimer might bind (with varying affinities) to dozens or hundreds of similar sequences

E. coli Promoters

“**TATA Box**” ~ 10bp upstream of transcription start

How to define it?

Consensus is TATAAT

BUT all differ from it

Allow k mismatches?

Equally weighted?

Wildcards like R, Y? ($\{A, G\}$, $\{C, T\}$, resp.)

TACGAT

TAAAAT

TATACT

GATAAT

TATGAT

TATGTT

E. coli Promoters

- “**TATA Box**” - consensus TATAAT ~ 10bp upstream of transcription start
- *Not exact*: of 168 studied (mid 80's)
 - nearly all had 2/3 of TAx_zyT
 - 80-90% had all 3
 - 50% agreed in each of x,y,z
 - **no** perfect match
- Other common features at -35, etc.

TATA Box Frequencies

pos base	1	2	3	4	5	6
A	2	95	26	59	51	1
C	9	2	14	13	20	3
G	10	1	16	15	13	0
T	79	3	44	13	17	96

TATA Scores

pos base	1	2	3	4	5	6
A	-36	19	1	12	10	-46
C	-15	-36	-8	-9	-3	-31
G	-13	-46	-6	-7	-9	-46 ^(?)
T	17	-31	8	-9	-6	19

Scanning for TATA

A	-38	19	1	12	10	-48	= -93
C	-15	-38	-8	-10	-3	-32	
G	-13	-48	-6	-7	-10	-48	
T	17	-32	8	-9	-6	19	

A C T A T A A T C G

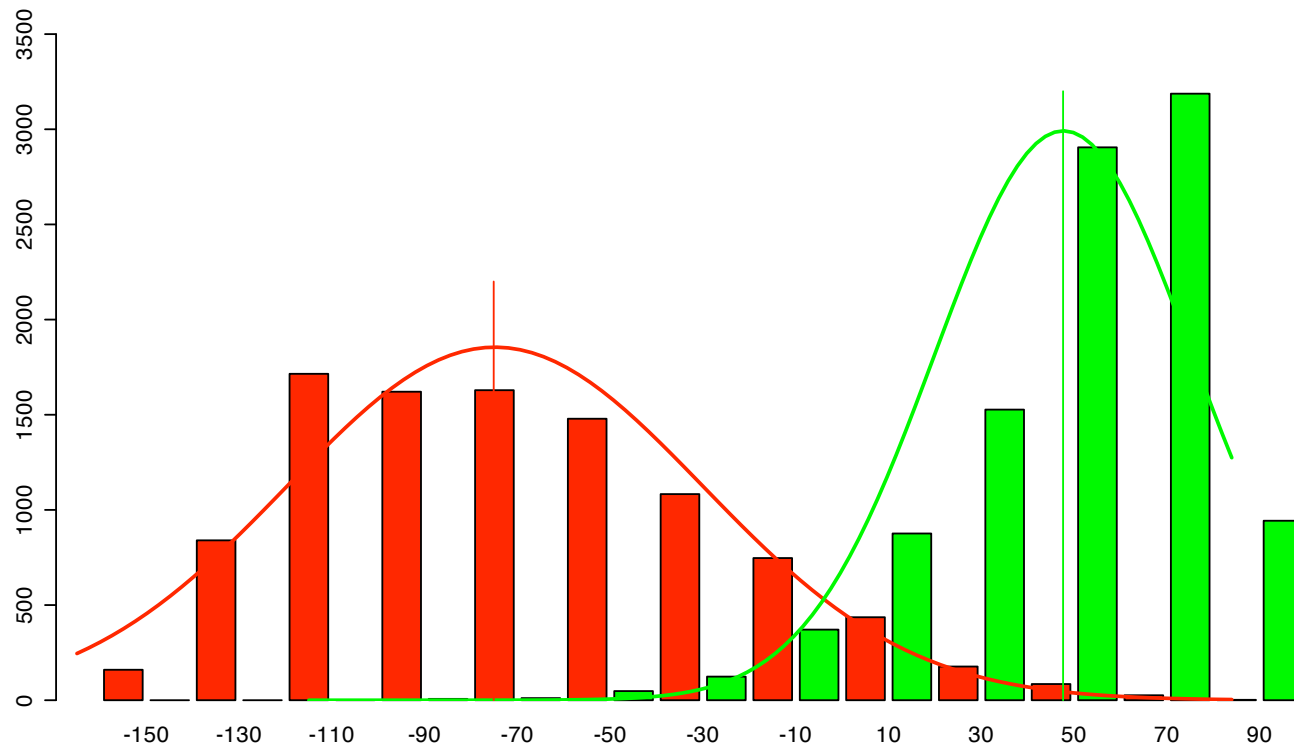
A	-38	19	1	12	10	-48	= 85
C	-15	-38	-8	-10	-3	-32	
G	-13	-48	-6	-7	-10	-48	
T	17	-32	8	-9	-6	19	

A C T A T A A T C G

A	-38	19	1	12	10	-48	= -95
C	-15	-38	-8	-10	-3	-32	
G	-13	-48	-6	-7	-10	-48	
T	17	-32	8	-9	-6	19	

A C T A T A A T C G

Score Distribution (Simulated)



Weight Matrices: Statistics

Assume:

$f_{b,i}$ = frequency of base b in position i in *TATA*

f_b = frequency of base b in all sequences

Log likelihood ratio, given $S = B_1B_2\dots B_6$:

$$\log \left(\frac{P(S|\text{"tata"})}{P(S|\text{"non-tata"})} \right) = \log \frac{\prod_{i=1}^6 f_{B_i,i}}{\prod_{i=1}^6 f_{B_i}} = \sum_{i=1}^6 \log \frac{f_{B_i,i}}{f_{B_i}}$$

Assumes independence

Neyman-Pearson

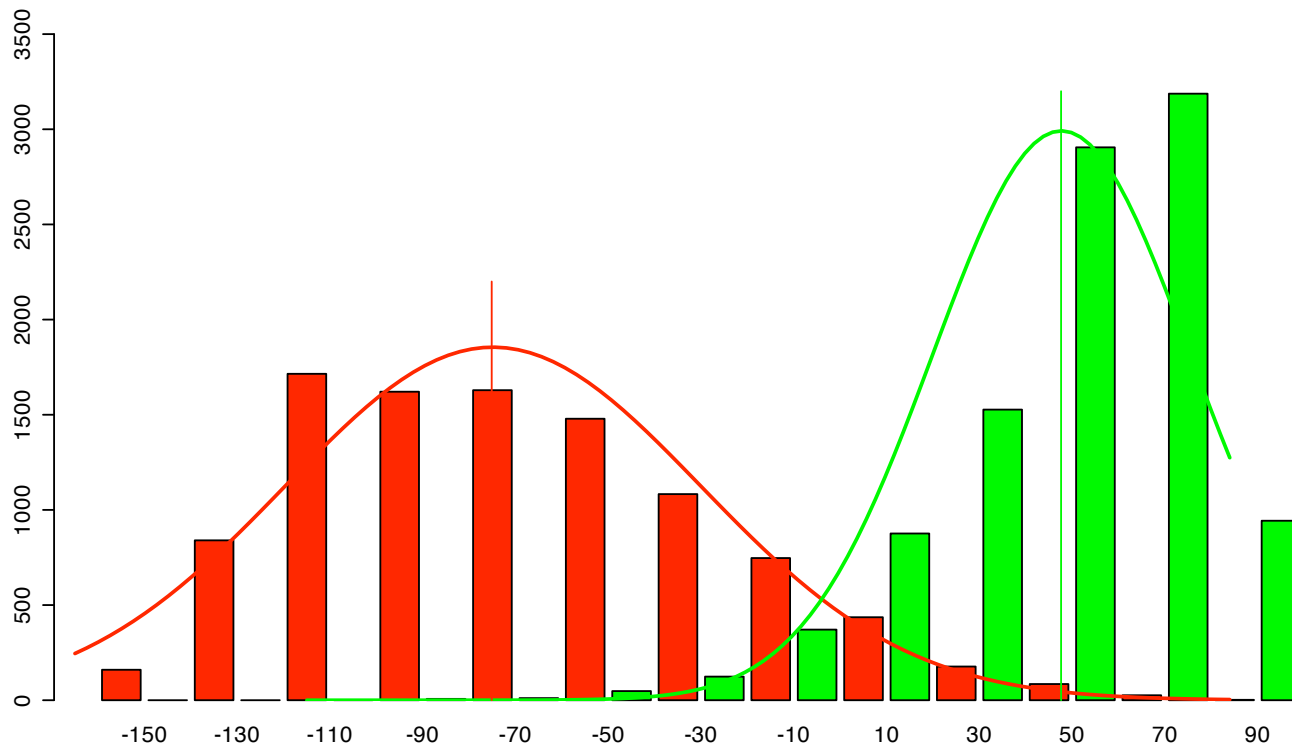
Given a sample x_1, x_2, \dots, x_n , from a distribution f ($\dots|\Theta$) with parameter Θ , want to test hypothesis $\Theta = \theta_1$ vs $\Theta = \theta_2$.

Might as well look at *likelihood ratio*:

$$\frac{f(x_1, x_2, \dots, x_n | \theta_1)}{f(x_1, x_2, \dots, x_n | \theta_2)} > \tau$$

(or *log likelihood ratio*)

Score Distribution (Simulated)



What's best WMM?

Given 20 sequences s_1, s_2, \dots, s_k of length 8, assumed to be generated at random according to a WMM defined by $8 \times (4-1)$ parameters θ , what's the best θ ?

E.g., what's MLE for θ given data s_1, s_2, \dots, s_k ?

Answer: count frequencies per position.

Weight Matrices: Chemistry

Experiments show ~80% correlation of log likelihood weight matrix scores to measured binding energy of RNA polymerase to variations on TATAAT consensus
[Stormo & Fields]

Another WMM example

8 Sequences:

ATG
ATG
ATG
ATG
ATG
GTG
GTG
TTG

Freq.	Col 1	Col 2	Col3
A	.625	0	0
C	0	0	0
G	.250	0	1
T	.125	1	0

LLR	Col 1	Col 2	Col 3
A	1.32	$-\infty$	$-\infty$
C	$-\infty$	$-\infty$	$-\infty$
G	0	$-\infty$	2.00
T	-1.00	2.00	$-\infty$

Log-Likelihood Ratio:

$$\log_2 \frac{f_{x_i,i}}{f_{x_i}}, f_{x_i} = \frac{1}{4}$$

Non-uniform Background

- *E. coli* - DNA approximately 25% A, C, G, T
- *M. jannaschi* - 68% A-T, 32% G-C

LLR from previous example, assuming

$$f_A = f_T = 3/8$$

$$f_C = f_G = 1/8$$

LLR	Col 1	Col 2	Col 3
A	.74	$-\infty$	$-\infty$
C	$-\infty$	$-\infty$	$-\infty$
G	1.00	$-\infty$	3.00
T	-1.58	1.42	$-\infty$

e.g., G in col 3 is 8 x more likely via WMM than background, so (\log_2) score = 3 (bits).

Relative Entropy

AKA Kullback-Liebler Distance/Divergence,
AKA Information Content

Given distributions P, Q

$$H(P||Q) = \sum_{x \in \Omega} P(x) \log \frac{P(x)}{Q(x)} \geq 0$$

Notes:

Let $P(x) \log \frac{P(x)}{Q(x)} = 0$ if $P(x) = 0$ [since $\lim_{y \rightarrow 0} y \log y = 0$]

Undefined if $0 = Q(x) < P(x)$

WMM: How “Informative”?

Mean score of site vs bkg?

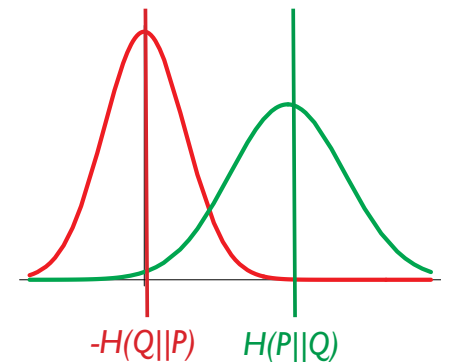
For any fixed length sequence x , let

$P(x)$ = Prob. of x according to WMM

$Q(x)$ = Prob. of x according to background

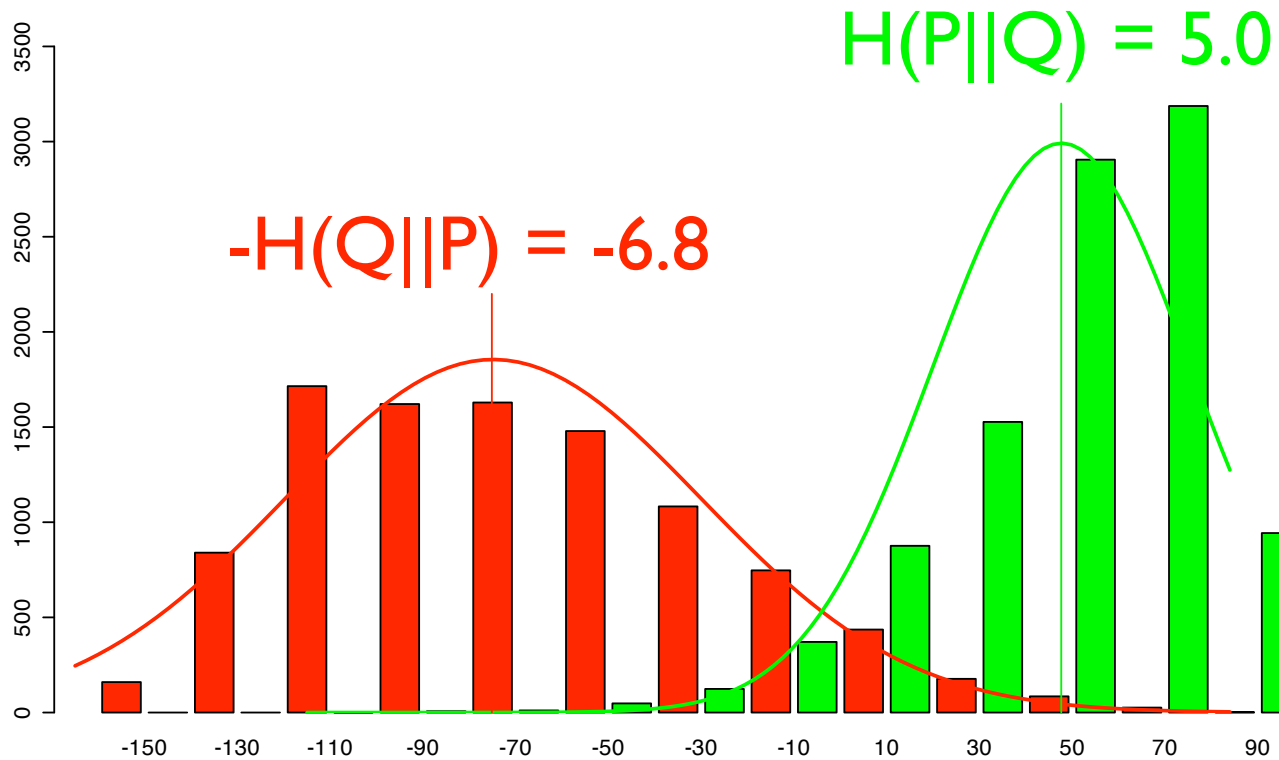
Relative Entropy:

$$H(P||Q) = \sum_{x \in \Omega} P(x) \log_2 \frac{P(x)}{Q(x)}$$



$H(P||Q)$ is *expected log likelihood score* of a sequence randomly chosen from **WMM**;
 $-H(Q||P)$ is expected score of **Background**

WMM Scores vs Relative Entropy



For WMM, you can show (based on the assumption of independence between columns), that :

$$H(P||Q) = \sum_i H(P_i||Q_i)$$

where P_i and Q_i are the WMM/background distributions for column i .

WMM Example, cont.

Freq.	Col 1	Col 2	Col 3
A	.625	0	0
C	0	0	0
G	.250	0	1
T	.125	1	0

Uniform

LLR	Col 1	Col 2	Col 3	
A	1.32	$-\infty$	$-\infty$	
C	$-\infty$	$-\infty$	$-\infty$	
G	0	$-\infty$	2.00	
T	-1.00	2.00	$-\infty$	
RelEnt	.70	2.00	2.00	4.70

Non-uniform

LLR	Col 1	Col 2	Col 3	
A	.74	$-\infty$	$-\infty$	
C	$-\infty$	$-\infty$	$-\infty$	
G	1.00	$-\infty$	3.00	
T	-1.58	1.42	$-\infty$	
RelEnt	.51	1.42	3.00	4.93

Pseudocounts

Are the $-\infty$'s a problem?

Certain that a given residue *never* occurs in a given position? Then $-\infty$ just right

Else, it may be a small-sample artifact

Typical fix: add a *pseudocount* to each observed count—small constant (e.g., .5, 1)

Sounds *ad hoc*; there is a Bayesian justification

WMM Summary

- Weight Matrix Model (aka Position Specific Scoring Matrix, PSSM, “possum”, 0th order Markov models)
- Simple statistical model assuming independence between adjacent positions
- To build: count (+ pseudocount) letter frequency per position, log likelihood ratio to background
- To scan: add LLRs per position, compare to threshold
- Generalizations to higher order models (i.e., letter frequency per position, conditional on neighbor) also possible, with enough training data

How-to Questions

Given aligned motif instances, build model?

- Frequency counts (above, maybe with pseudocounts)

Given a model, find (probable) instances

- Scanning, as above

Given unaligned strings thought to contain a motif, find it? (e.g., upstream regions for co-expressed genes from a microarray experiment)

- Hard... rest of lecture.

Motif Discovery

Unfortunately, finding a site of max relative entropy in a set of unaligned sequences is NP-hard [Akutsu]

Motif Discovery: 4 example approaches

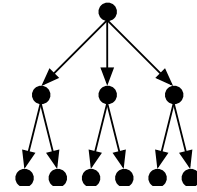
Brute Force

Greedy search

Expectation Maximization

Gibbs sampler

Brute Force



Input:

- Seqs s_1, s_2, \dots, s_k (len $\sim n$, say), each w/ one instance of an unknown length l motif

Algorithm:

- create singleton set with each length l subsequence of each s_1, s_2, \dots, s_k ($\sim nk$ sets)
- for each set, add each possible length l subsequence not already present ($\sim n^2 k(k-1)$ sets)
- repeat until all have k sequences ($\sim n^k k!$ sets)
- compute relative entropy of each; pick best

problem:
astronomically sloooow

Greedy Best-First Approach

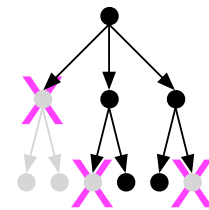
[Hertz & Stormo]

Input:

- Sequence s_1, s_2, \dots, s_k ; motif length l ; “breadth” d

Algorithm:

- create singleton set with each length l subsequence of each s_1, s_2, \dots, s_k
- for each set, add each possible length l subsequence not already present
- compute relative entropy of each
- discard all but d best
- repeat until all have k sequences



usual “greedy” problems

Expectation Maximization

[MEME, Bailey & Elkan, 1995]

Input (as above):

Sequence s_1, s_2, \dots, s_k ; motif length l ; background model; again assume one instance per sequence (variants possible)

Algorithm: EM

Visible data: the sequences

Hidden data: where's the motif

$$Y_{i,j} = \begin{cases} 1 & \text{if motif in sequence } i \text{ begins at position } j \\ 0 & \text{otherwise} \end{cases}$$

Parameters θ : The WMM

MEME Outline

Typical EM algorithm:

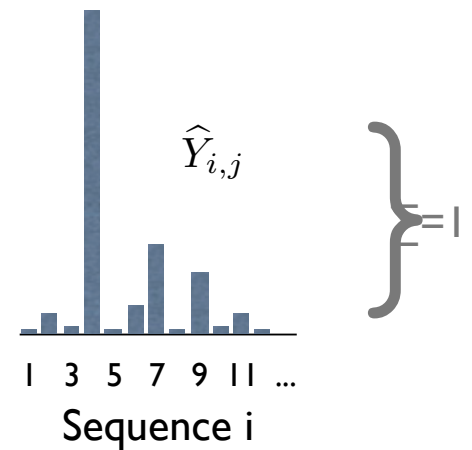
- Parameters θ^t at t^{th} iteration, used to estimate where the motif instances are (the hidden variables)
- Use those estimates to re-estimate the parameters θ to maximize likelihood of observed data, giving θ^{t+1}
- Repeat

Key: given a few good matches to best motif, expect to pick out more

Expectation Step

(where are the motif instances?)

$$\begin{aligned}
 \hat{Y}_{i,j} &= E(Y_{i,j} \mid s_i, \theta^t) \xrightarrow{\text{E} = 0 \cdot P(0) + 1 \cdot P(1)} \\
 &= P(Y_{i,j} = 1 \mid s_i, \theta^t) \xrightarrow{\text{Bayes}} \\
 &= P(s_i \mid Y_{i,j} = 1, \theta^t) \frac{P(Y_{i,j}=1|\theta^t)}{P(s_i|\theta^t)} \\
 &= cP(s_i \mid Y_{i,j} = 1, \theta^t) \\
 &= c' \prod_{k=1}^l P(s_{i,j+k-1} \mid \theta^t)
 \end{aligned}$$



where c' is chosen so that $\sum_j \hat{Y}_{i,j} = 1$.

Maximization Step

(what is the motif?)

Find θ maximizing expected value:

$$\begin{aligned} Q(\theta | \theta^t) &= E_{Y \sim \theta^t} [\log P(s, Y | \theta)] \\ &= E_{Y \sim \theta^t} [\log \prod_{i=1}^k P(s_i, Y_i | \theta)] \\ &= E_{Y \sim \theta^t} [\sum_{i=1}^k \log P(s_i, Y_i | \theta)] \\ &= E_{Y \sim \theta^t} [\sum_{i=1}^k \sum_{j=1}^{|s_i|} Y_{i,j} \log P(s_i, Y_{i,j} = 1 | \theta)] \\ &= E_{Y \sim \theta^t} [\sum_{i=1}^k \sum_{j=1}^{|s_i|} Y_{i,j} \log(P(s_i | Y_{i,j} = 1, \theta) P(Y_{i,j} = 1 | \theta))] \\ &= \sum_{i=1}^k \sum_{j=1}^{|s_i|} E_{Y \sim \theta^t} [Y_{i,j}] \log P(s_i | Y_{i,j} = 1, \theta) + C \\ &= \sum_{i=1}^k \sum_{j=1}^{|s_i|} \hat{Y}_{i,j} \log P(s_i | Y_{i,j} = 1, \theta) + C \end{aligned}$$

M-Step (cont.)

$$Q(\theta | \theta^t) = \sum_{i=1}^k \sum_{j=1}^{|s_i|} \hat{Y}_{i,j} \log P(s_i | Y_{i,j} = 1, \theta) + C$$

Exercise: Show this is maximized by “counting” letter frequencies over all possible motif instances, with counts weighted by $\hat{Y}_{i,j}$, again the “obvious” thing.

s_1 : A**CGG**ATT...

...
 s_k : GC...T**CGG**AC

$\hat{Y}_{1,1}$	ACGG
$\hat{Y}_{1,2}$	CGGA
$\hat{Y}_{1,3}$	GGAT
\vdots	\vdots
$\hat{Y}_{k,l-1}$	CGGA
$\hat{Y}_{k,l}$	GGAC

Initialization

1. Try every motif-length substring, and use as initial θ a WMM with, say 80% of weight on that sequence, rest uniform
2. Run a few iterations of each
3. Run best few to convergence
(Having a supercomputer helps)

Another Motif Discovery Approach The Gibbs Sampler

Lawrence, *et al.* “Detecting Subtle Sequence Signals: A Gibbs Sampling Strategy for Multiple Sequence Alignment,” *Science* 1993

Sigma-37	223	IIDLTYIQNK	SQKETGDILGISQMHVSR	LQRKAVKKLR	240	A25944
SpoIIIC	94	RFGLDLKKEK	TQREIAKELGISRSYVSR	IEKRALMKMF	111	A28627
NahR	22	VVFNQLLVDR	RVSITAENLGLTQPAVSN	ALKRLRTSLQ	39	A32837
Antennapedia	326	FHFNRYLTRR	RRIEIAHALCLTERQIKI	WFQNRMMKWK	343	A23450
NtrC (Brady.)	449	LTAALAAATRG	NQIRAADLLGLNRNTLRK	KIRDLDIQVY	466	B26499
DicA	22	IRYRRKNLKH	TQRS LAKALKISHVSVSQ	WERGDSEPTG	39	B24328 (BVECDA)
MerD	5		MNAY TVSRLALDAGVSVHIVRD	YLLRGLLRPV	22	C29010
Fis	73	LDMVMQYTRG	NQTR AALMMGINRGTLRK	KLKKYGMN	90	A32142 (DNECF5)
MAT a1	99	FRRKQSLNSK	EKEEVAKKCGITPLQVRV	WFINKRMR5K	116	A90983 (JEBY1)
Lambda cII	25	SALLNKIAML	GTEKTAEAVGV DKSQISR	WKR DWIPKFS	42	A03579 (QCBP2L)
Crp (CAP)	169	THPDGMQIKI	TRQEIGQIVGCSRET VGR	ILKMLE DQNL	186	A03553 (QRECC)
Lambda Cro	15	ITLKDYAMRF	GQTKTAKDLGVYQSAINK	AIHAGR KIFL	32	A03577 (RCBPL)
P22 Cro	12	YKKDVIDHFG	TQRAVAKALGISDA AVSQ	WKÉVIPEKDA	29	A25867 (RGBP22)
AraC	196	ISDHLADSNF	DIASVAQH VCLSPSRLSH	LFRQQLGISV	213	A03554 (RGECA)
Fnr	196	FSPREFRLTM	TRGDIGNYLGLTVETISR	LLGRFQKSGM	213	A03552 (RGECE)
HtpR	252	ARWLDEDNKS	TLQELADRYGVSA ERVRQ	LEKNAMKKLR	269	A00700 (RGECH)
NtrC (K.a.)	444	LTTALRHTQG	HKQEAARLLGWGRNTLTR	KLKELGME	461	A03564 (RGKBCP)
Cytr	11	MKAKKQETAA	TMKDVALKAKVSTATVSR	ALMNPDKVSQ	28	A24963 (RPECCT)
DeoR	23	LQELKRSDKL	HLKDAAALLGVSEMTIRR	DLNNHSAPVV	40	A24076 (RPECDO)
GalR	3		MA TIKDVARLAGVSVATVSR	VINNSPKASE	20	A03559 (RPECG)
LacI	5		MKPV TLYDVAEYAGVSYQTVSR	VVNQASHVSA	22	A03558 (RPECL)
TetR	26	LLNEVGIEGL	TTRKLAQKLGVEQPTLYW	HVKNKRALLD	43	A03576 (RPECTN)
TrpR	67	IVEELLRGEM	SQRELKNE LGAGIATITR	GSNSLKAAPV	84	A03568 (RPECW)
NifA	495	LIAALEKAGW	VQAKAARLLGMTPRQVAY	RIQIMDITMP	512	S02513
SpoIIG	205	RFGLVGEEEK	TQKDVADMMGISQSYISR	LEKRIIKRLR	222	S07337
Pin	160	QAGRLIAAGT	PRQKVAI IYDVG VSTLYK	TFPAGDK	177	S07958
PurR	3		MA TIKDVAKRANVSTTTVSH	VINKTRFVAE	20	S08477
EbgR	3		MA TLKDIAIEAGVSLATVSR	VLNDDPTLNV	20	S09205
LexA	27	DHISQTGMPP	TRAEIAQRLGFRSPNAAE	EHLKALARKG	44	S11945
P22 cI	25	SSILNRIAIR	GQRKVADALGINESQISR	WKGDFIPKMG	42	B25867 (Z1BPC2)

***** ***

B	Position in site																	
	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18
Arg	94	222	265	137	9	9	137	137	9	9	9	52	222	94	94	9	265	606
Lys	9	133	442	380	9	71	380	194	9	133	9	9	71	9	9	9	71	256
Glu	53	9	96	401	9	9	140	140	9	9	9	53	140	140	9	9	9	53
Asp	67	9	9	473	9	9	299	125	9	67	9	67	67	9	9	9	9	67
Gln	9	600	224	9	9	9	224	9	9	9	9	9	278	63	278	9	9	170
His	240	9	9	9	9	9	125	125	9	9	9	9	125	125	125	9	9	240
Asn	168	9	9	9	9	9	168	89	9	89	9	248	9	168	89	9	89	89
Ser	117	9	117	117	9	9	9	9	9	9	9	819	63	387	63	9	819	9
Gly	151	9	56	9	9	151	9	9	9	1141	9	151	9	56	9	9	56	9
Ala	9	9	112	43	181	901	43	181	215	9	43	9	43	181	112	43	78	9
Thr	915	130	130	9	251	9	9	9	9	9	9	311	130	70	855	9	130	9
Pro	76	9	9	9	9	9	9	9	9	9	9	9	210	210	9	9	9	9
Cys	9	9	9	9	9	9	9	9	295	581	295	9	9	9	9	9	9	9
Val	58	107	9	9	500	9	9	9	156	9	598	9	205	58	9	746	9	58
Leu	9	121	9	9	149	9	93	149	458	9	149	9	37	37	9	177	9	9
Ile	9	166	114	61	323	9	114	166	9	9	427	9	61	9	61	427	9	61
Met	9	104	9	9	9	9	9	198	198	9	104	9	9	198	9	9	9	9
Tyr	9	9	136	9	9	9	9	262	262	9	9	136	136	9	262	9	262	136
Phe	9	9	9	9	9	9	9	9	9	9	108	9	9	9	9	9	9	9
Trp	9	9	9	9	9	9	9	9	9	9	366	9	9	9	9	9	9	366

Some History

Geman & Geman, IEEE PAMI 1984

Hastings, Biometrika, 1970

Metropolis, Rosenbluth, Rosenbluth, Teller, & Teller, "Equations of State Calculations by Fast Computing Machines," J. Chem. Phys. 1953

Josiah Williard Gibbs, 1839-1903, American physicist, a pioneer of thermodynamics

How to Average

An old problem:

n random variables:

$$x_1, x_2, \dots, x_k$$

Joint distribution (p.d.f.):

$$P(x_1, x_2, \dots, x_k)$$

Some function:

$$f(x_1, x_2, \dots, x_k)$$

Want Expected Value:

$$E(f(x_1, x_2, \dots, x_k))$$

How to Average

$$E(f(x_1, x_2, \dots, x_k)) = \int_{x_1} \int_{x_2} \cdots \int_{x_k} f(x_1, x_2, \dots, x_k) \cdot P(x_1, x_2, \dots, x_k) dx_1 dx_2 \cdots dx_k$$

Approach 1: direct integration

(rarely solvable analytically, esp. in high dim)

Approach 2: numerical integration

(often difficult, e.g., unstable, esp. in high dim)

Approach 3: Monte Carlo integration

sample $\vec{x}^{(1)}, \vec{x}^{(2)}, \dots, \vec{x}^{(n)} \sim P(\vec{x})$ and average:

$$E(f(\vec{x})) \approx \frac{1}{n} \sum_{i=1}^n f(\vec{x}^{(i)})$$

Markov Chain Monte Carlo (MCMC)

- *Independent* sampling also often hard, but *not required* for expectation

- MCMC $\vec{X}_{t+1} \sim P(\vec{X}_{t+1} | \vec{X}_t)$ w/ stationary dist = P

- Simplest & most common: Gibbs Sampling

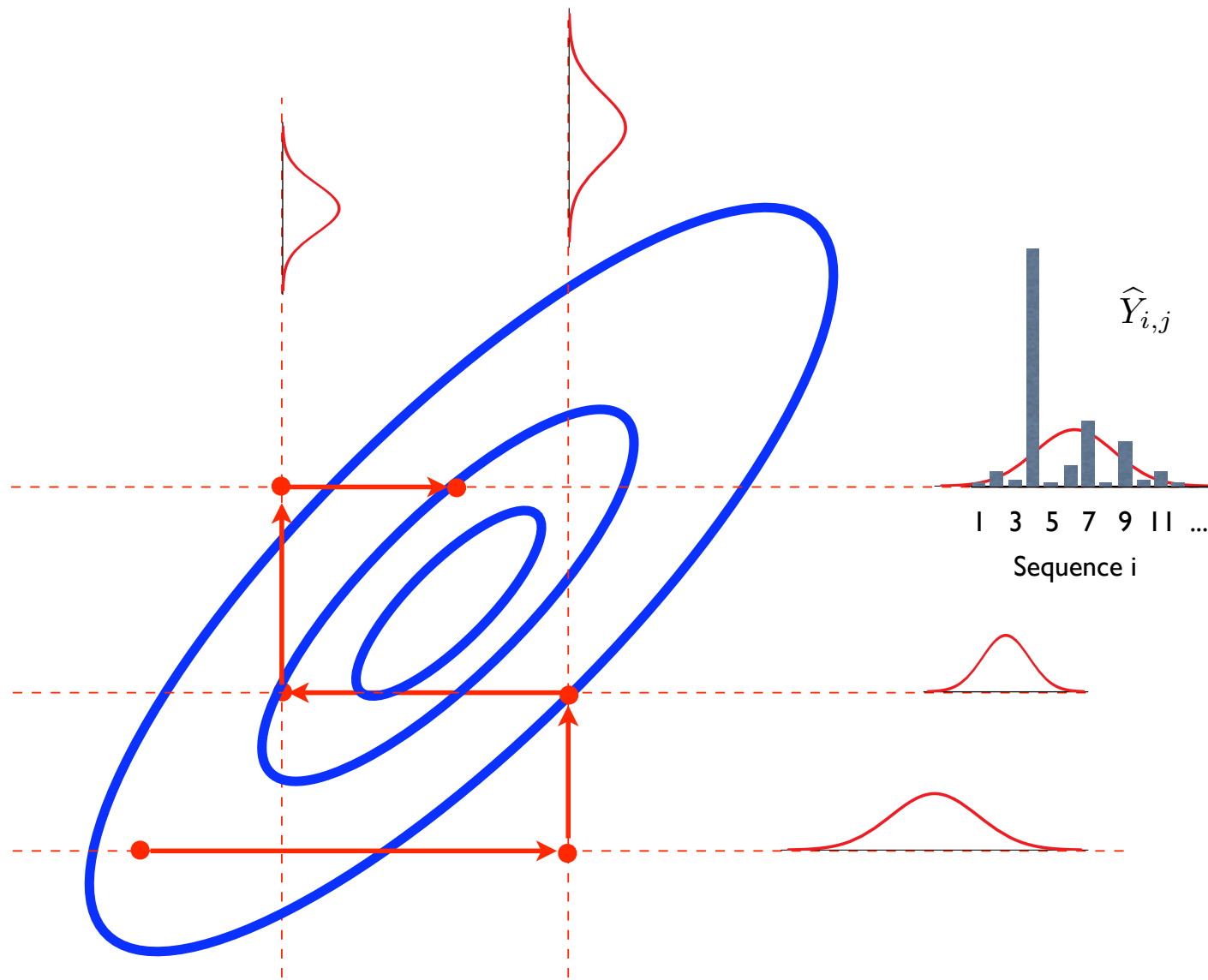
$$P(x_i | x_1, x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_k)$$

- Algorithm

for $t = 1$ to ∞

for $i = 1$ to k do :

$$x_{t+1,i} \sim P(x_{t+1,i} | \underbrace{x_{t+1,1}, x_{t+1,2}, \dots, x_{t+1,i-1}}_{t+1}, \underbrace{x_{t,i+1}, \dots, x_{t,k}}_t)$$



- Input: again assume sequences s_1, s_2, \dots, s_k with one length w motif per sequence
- Motif model: WMM
- Parameters: Where are the motifs?
for $1 \leq i \leq k$, have $1 \leq x_i \leq |s_i| - w + 1$
- “Full conditional”: to calc

$$P(x_i = j \mid x_1, x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_k)$$
 build WMM from motifs in all sequences except i , then calc prob that motif in i^{th} seq occurs at j by usual “scanning” alg.

Overall Gibbs Alg

Randomly initialize x_i 's

for $t = 1$ to ∞

for $i = 1$ to k

discard motif instance from s_i ;

recalc WMM from rest

for $j = 1 \dots |s_i| - w + 1$

calculate prob that i^{th} motif is at j :

→ $P(x_i = j \mid x_1, x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_k)$

pick new x_i according to that distribution

Similar to
MEME, but it
would
average over,
rather than
sample from

Issues

Burnin - how long must we run the chain to reach stationarity?

Mixing - how long a post-burnin sample must we take to get a good sample of the stationary distribution? (Recall that individual samples are not independent, and may not “move” freely through the sample space. Also, many isolated modes.)

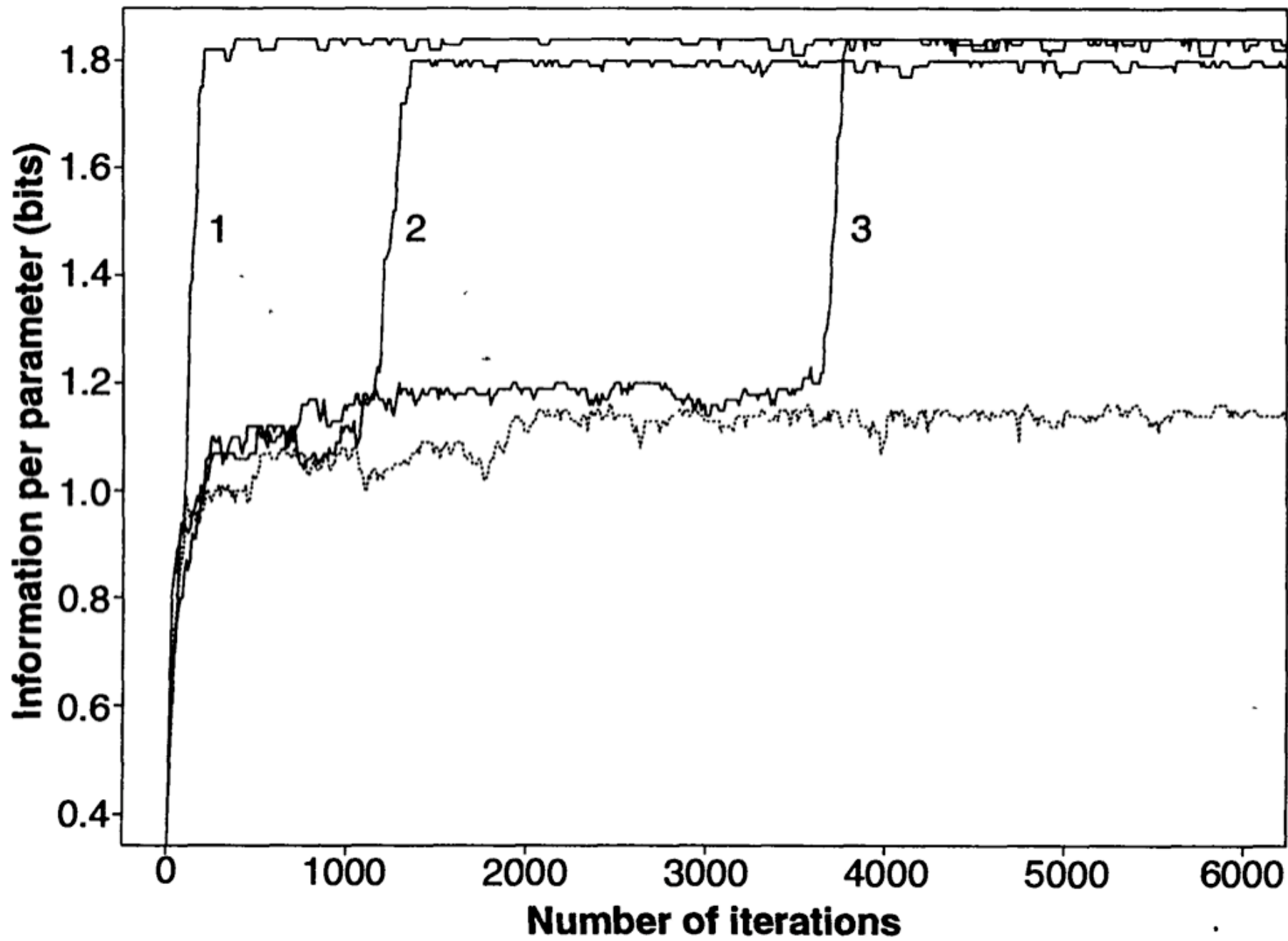
Variants & Extensions

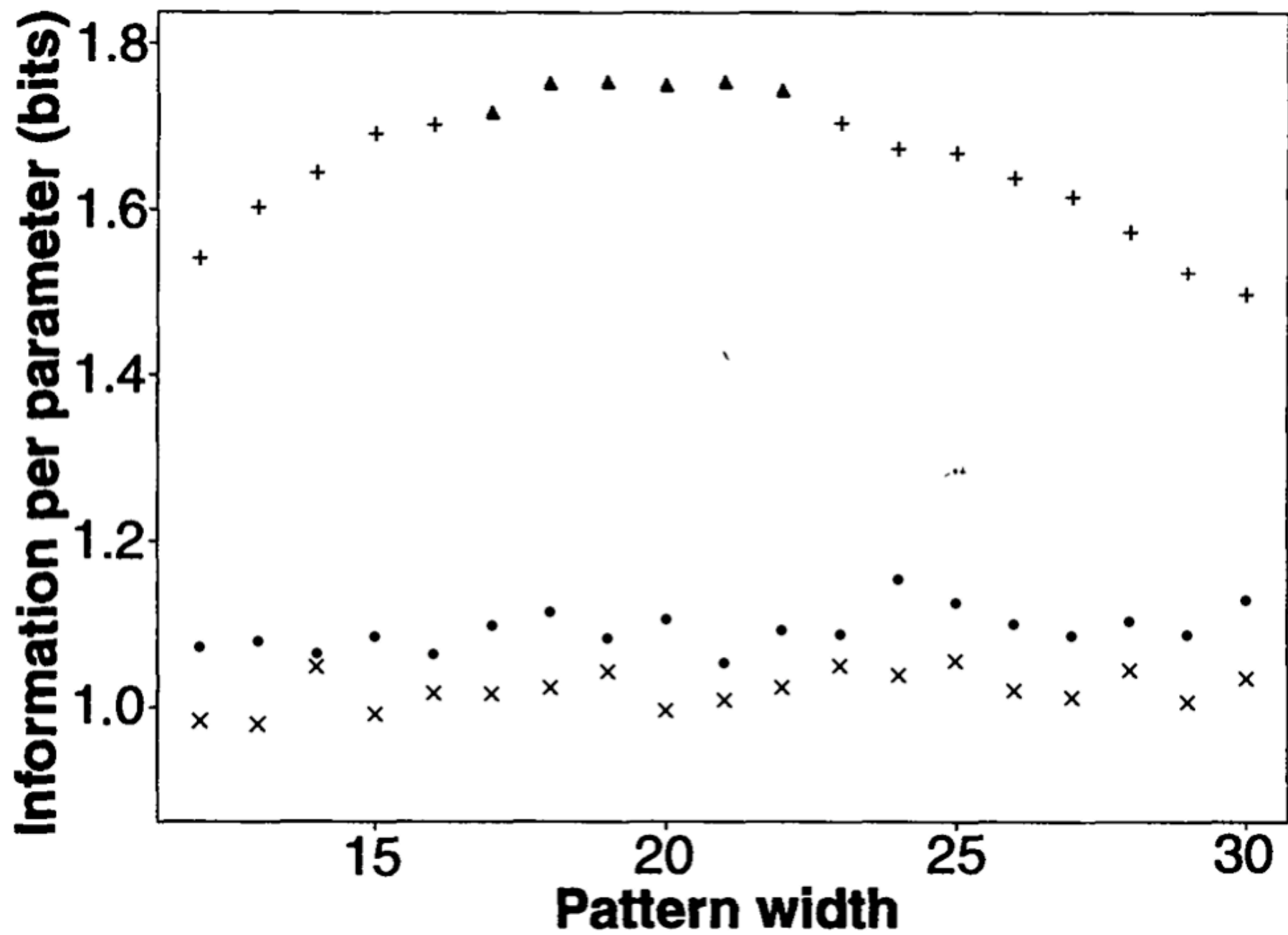
“Phase Shift” - may settle on suboptimal solution that overlaps part of motif.

Periodically try moving all motif instances a few spaces left or right.

Algorithmic adjustment of pattern width:
Periodically add/remove flanking positions to maximize (roughly) average relative entropy per position

Multiple patterns per string





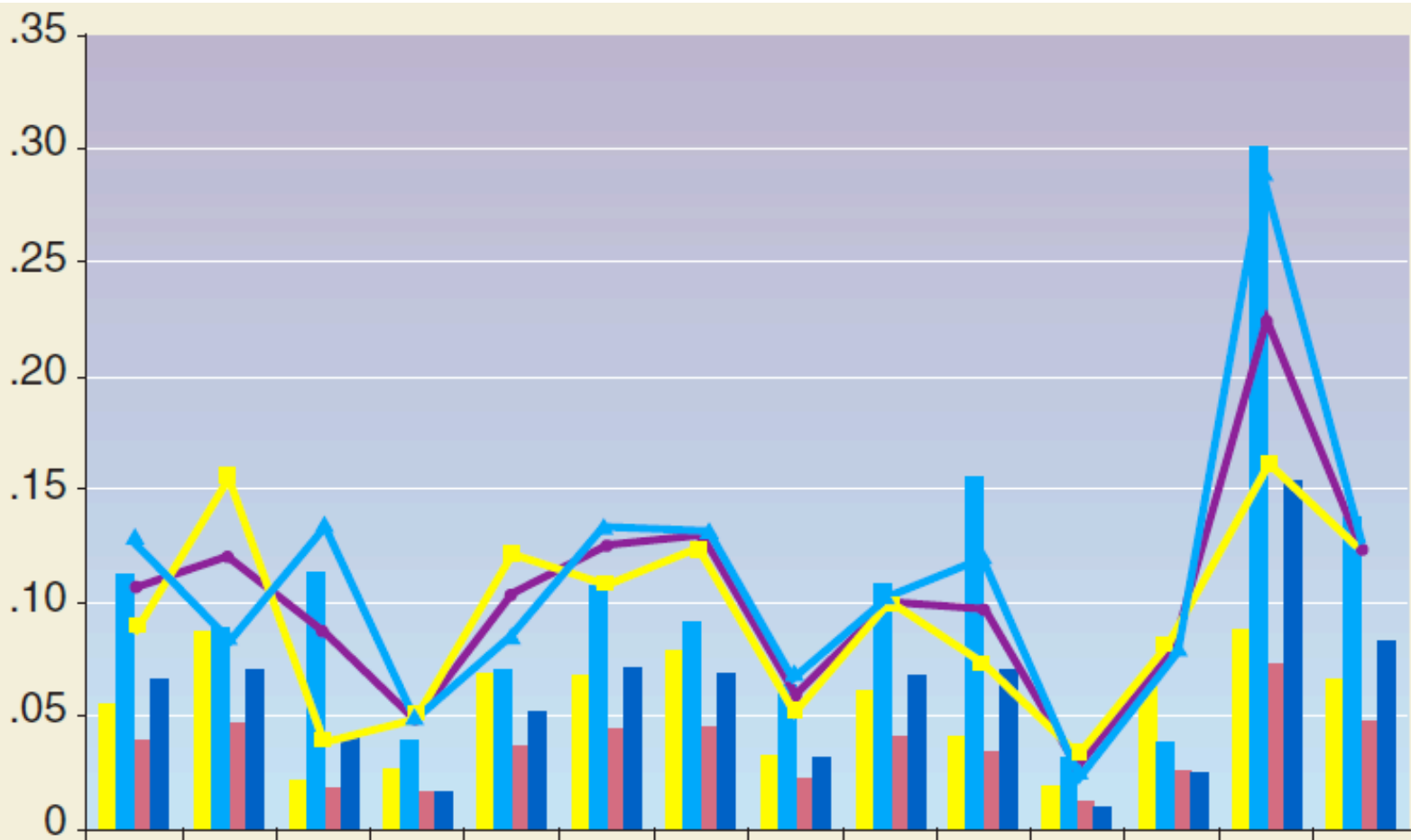
NATURE BIOTECHNOLOGY VOLUME 23 NUMBER 1 JANUARY 2005

Assessing computational tools for the discovery of transcription factor binding sites

Martin Tompa^{1,2}, Nan Li¹, Timothy L Bailey³, George M Church⁴, Bart De Moor⁵, Eleazar Eskin⁶, Alexander V Favorov^{7,8}, Martin C Frith⁹, Yutao Fu⁹, W James Kent¹⁰, Vsevolod J Makeev^{7,8}, Andrei A Mironov^{7,11}, William Stafford Noble^{1,2}, Giulio Pavesi¹², Graziano Pesole¹³, Mireille Régnier¹⁴, Nicolas Simonis¹⁵, Saurabh Sinha¹⁶, Gert Thijs⁵, Jacques van Helden¹⁵, Mathias Vandenbogaert¹⁴, Zhiping Weng⁹, Christopher Workman¹⁷, Chun Ye¹⁸ & Zhou Zhu⁴

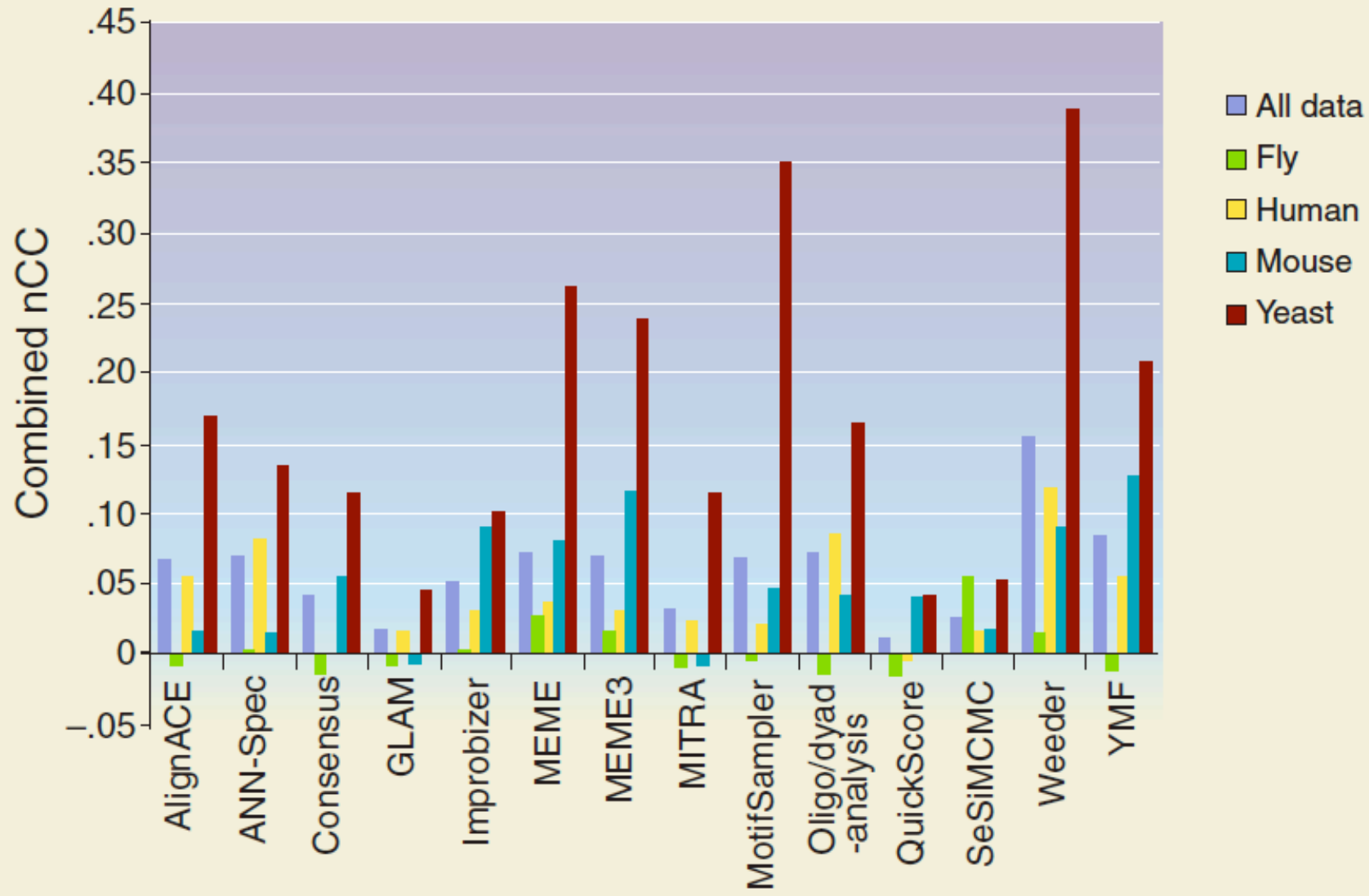
Methodology

- 13 tools
- Real 'motifs' (Transfac)
- 56 data sets (human, mouse, fly, yeast)
- 'Real', 'generic', 'Markov'
- Expert users, top prediction only

a

\$ Greed
 * Gibbs
 ^ EM

■ nSn ■ nPPV ■ nPC ■ nCC ■ sSn ▲ sPPV ● sASP

b

Lessons

- Evaluation is hard (esp. when “truth” is unknown)
- Accuracy low
 - partly reflects defects in evaluation methodology (e.g. ≤ 1 prediction per data set; results better in synth data)
 - partly reflects difficult task, limited knowledge (e.g. yeast $>$ others)
- No clear winner re methods or models

Motif Discovery Summary

- Important problem: a key to understanding gene regulation
- Hard problem: short, degenerate signals amidst much noise
- *Many* variants have been tried, for representation, search, and discovery. We looked at only a few:
 - Weight matrix models for representation & search
 - greedy, MEME and Gibbs for discovery
- Still much room for improvement. *Comparative genomics*, i.e. cross-species comparison is very promising